
COM 525000 – Statistical Learning

Lecture 5 – Resampling Methods

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Resampling Methods

- **Resampling** refers to the **repeated drawing** of samples from a training set and **refitting** a model of interest on each set of samples.
 - ➔ Used to obtain additional information about the fitted model.
 - ➔ E.g., for estimating the test error (for model assessment and selection), or to estimate the variability of coefficient estimates.
- Two common approaches:
 - **Cross-validation**
 - **Bootstrap**

Training versus Test Error (1/2)

- The **training error** is the average error between the fitted and true responses of data points in the training set. That is,

$$\overline{\text{err}} = \frac{1}{n} \sum_{i=1}^n L(y_i, \hat{f}(x_i; \mathcal{D})) \left(\stackrel{\text{e.g.}}{=} \frac{1}{n} \sum_{i=1}^n (y_i - \hat{f}(x_i; \mathcal{D}))^2 \right)$$

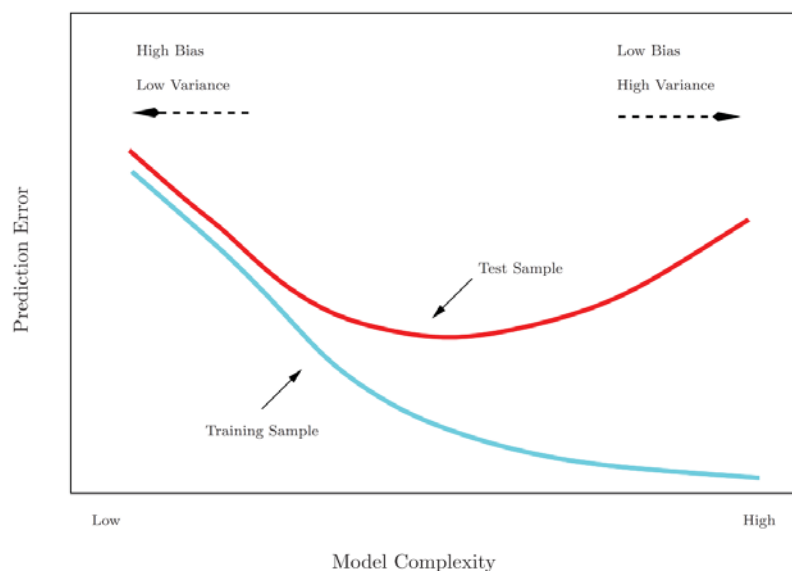
where $\mathcal{D} = \{(x_1, y_1), \dots, (x_n, y_n)\}$ is the available data.

- The **test error** (or, generalization error) is the average error of predicting the response on a *new observation*, i.e.,

$$\text{Err} = E[L(Y, \hat{f}(X; \mathcal{D}))] \left(\stackrel{\text{e.g.}}{=} E[(Y - \hat{f}(X; \mathcal{D}))^2] \right)$$

where the expectation is taken over X , Y and \mathcal{D} .

Training versus Test Error (2/2)



- Test set is often unavailable and, thus, the actual test error is often unknown. (➔ It must be estimated!)
 - Key Idea:** Hold out a subset of the available data for testing later on, and train on the remaining subset.
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The Validation Set Approach

- **The Validation Set Approach:**

- Split into a training set $\mathcal{D}_{\text{train}}$ and a validation set \mathcal{D}_{val} .

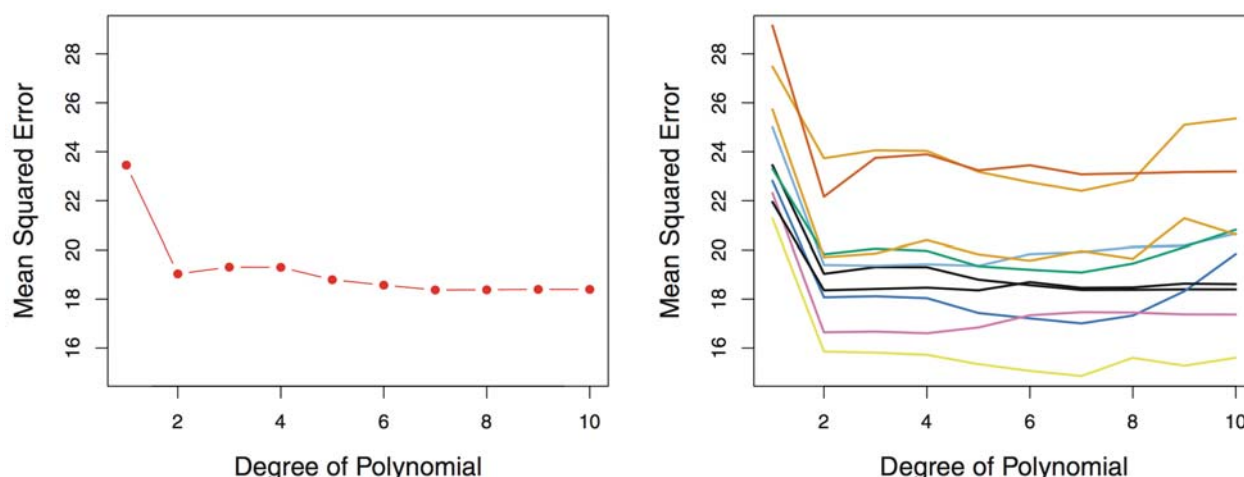


- Train (or fit) the model on the training set, and predict the responses for observations in the validation set.

➔ The validation set error rate provides an estimate of the test error rate.

Example: Auto Data Set

- X : horsepower, Y : mpg
- 392 observations=196 training set+196 validation set



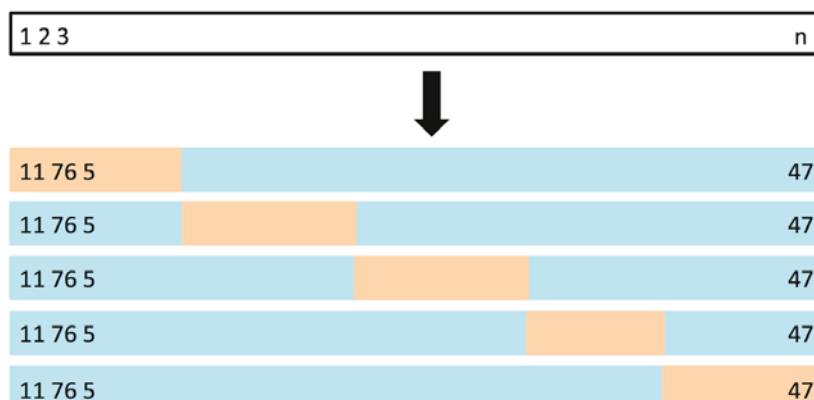
➔ The estimate of the test error is highly variable.

➔ The validation set error rate tend to *overestimate* the test error rate for the model fit on the entire data.

k-Fold Cross-Validation

- **k-Fold Cross-Validation:**

- Split the data into k groups (or folds) $\mathcal{D}_1, \dots, \mathcal{D}_k$ of size n_1, \dots, n_k (e.g., $n_1 = \dots = n_k = \frac{n}{k}$).



- Use the j th fold \mathcal{D}_j as the validation set and the remaining $k-1$ folds $\mathcal{D} \setminus \mathcal{D}_j$ as the training set.
- Repeat for $j=1, \dots, k$ to get MSEs $\text{MSE}_1, \dots, \text{MSE}_k$.

MSE of k-Fold CV

- The overall MSE is

$$\text{CV}_{(k)} = \sum_{j=1}^k \frac{n_j}{n} \text{MSE}_j \left(\stackrel{\text{e.g.}}{=} \frac{1}{k} \sum_{j=1}^k \text{MSE}_j, \text{ for } n_j = \frac{n}{k}, \forall j \right)$$

➔ Typical values of k are 5 and 10.

Leave-One-Out Cross-Validation (LOOCV)

- **Leave-one-out cross-validation (LOOCV)** is a special case of k-fold CV where $\mathcal{D} = \{(x_1, y_1), \dots, (x_n, y_n)\}$ is split into n groups $\mathcal{D}_1 = \{(x_1, y_1)\}, \dots, \mathcal{D}_n = \{(x_n, y_n)\}$.



- ➔ Less bias (since larger training sets are used).
- ➔ Computationally expensive.

Special Case of LOOCV MSE

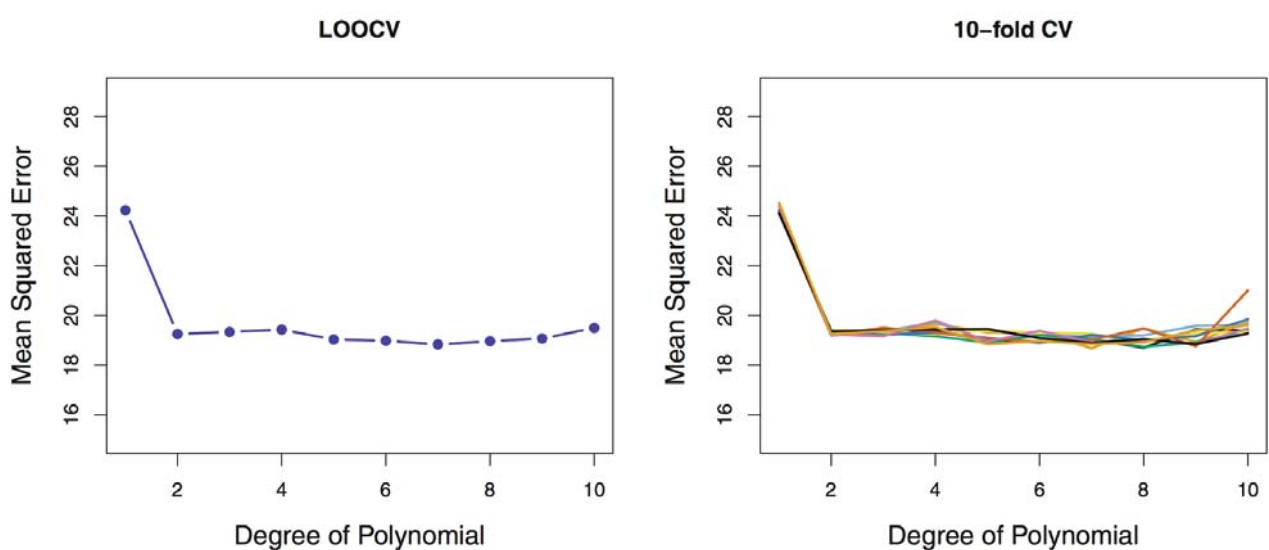
- **Special Case:** For least squares linear (or polynomial) regression, the cost of LOOCV can be computed as

$$CV_{(n)} = \frac{1}{n} \sum_{i=1}^n \left(\frac{y_i - \hat{y}_i}{1 - h_i} \right)^2$$

where $h_i \triangleq \{\mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T\}_{ii} = x_i^T (\mathbf{X}^T \mathbf{X})^{-1} x_i$.

LOOCV vs k-Fold CV

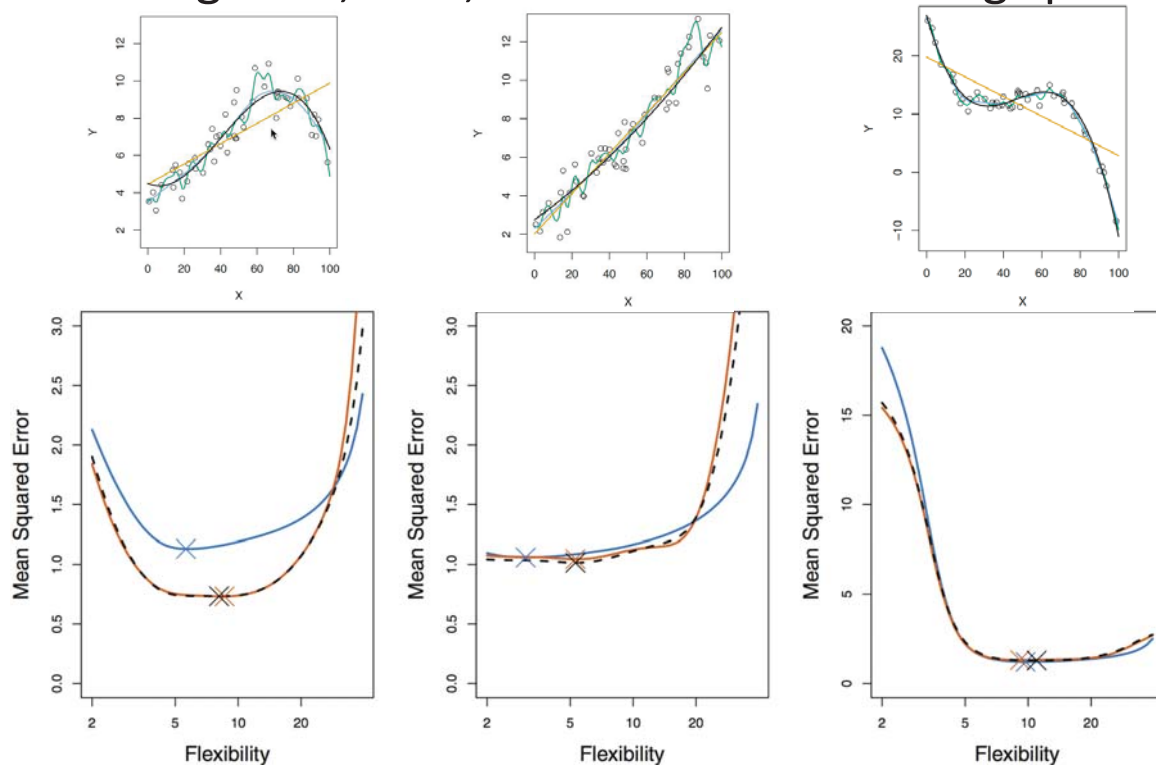
- Revisit Auto Data Set Example:



- ➔ LOOCV yields a deterministic test error estimate since there is only one way to split the data set.
- ➔ 10-fold CV exhibits some variability for random splits.

LOOCV vs k-Fold CV (Simulated Example)

- Recall Figs. 2.9, 2.10, and 2.11 on smoothing splines.



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Bias-Variance Tradeoff for k-Fold CV (1/2)

- CV methods are used to estimate the test error

$$E[(Y - \hat{f}(X; \mathcal{D}))^2]$$

where \mathcal{D} is the available data set.

- In k-fold CV (with $n_j = n/k, \forall j$), this is estimated by

$$\begin{aligned} CV_{(k)} &= \frac{1}{k} \sum_{j=1}^k \frac{1}{n/k} \sum_{i: (x_i, y_i) \in \mathcal{D}_j} (y_i - \hat{f}(x_i; \mathcal{D} \setminus \mathcal{D}_j))^2 \\ &= \frac{1}{n} \sum_{i=1}^n (y_i - \hat{f}(x_i; \mathcal{D} \setminus \mathcal{D}_{j^*(i)}))^2 \end{aligned}$$

where $j^*(i)$ is defined such that $(x_i, y_i) \in \mathcal{D}_{j^*(i)}$.

➔ Note that $E[CV_{(k)}] = E[(Y - \hat{f}(X; \mathcal{D}^{(n-n/k)}))^2]$.

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Bias-Variance Tradeoff for k-Fold CV (2/2)

- As k increases, the bias of the test error estimate

$$E[(Y - \hat{f}(X; \mathcal{D}^{(n-n/k)}))^2] - E[(Y - \hat{f}(X; \mathcal{D}))^2]$$

decreases, but the variance

$$E\left[\left\{\frac{1}{n} \sum_{i=1}^n \left[(Y_i - \hat{f}(X_i; \mathcal{D} \setminus \mathcal{D}_{j^*(i)}))^2 - E[(Y - \hat{f}(X; \mathcal{D}^{(n-n/k)}))^2]\right]\right\}^2\right]$$

increases.

CV on Classification Problems (1/2)

- Similarly, for classification problems, k-fold CV yields

$$CV_{(k)} = \frac{1}{k} \sum_{j=1}^k \text{Err}_j$$

where $\text{Err}_j = \frac{1}{|\mathcal{D}_j|} \sum_{i \in \mathcal{D}_j} L(y_i, \hat{y}_i^{(\mathcal{D} \setminus \mathcal{D}_j)})$.

- In classification problems, we may take

$$L(y, \hat{y}) = I(y \neq \hat{y})$$

(called the 0-1 loss) or

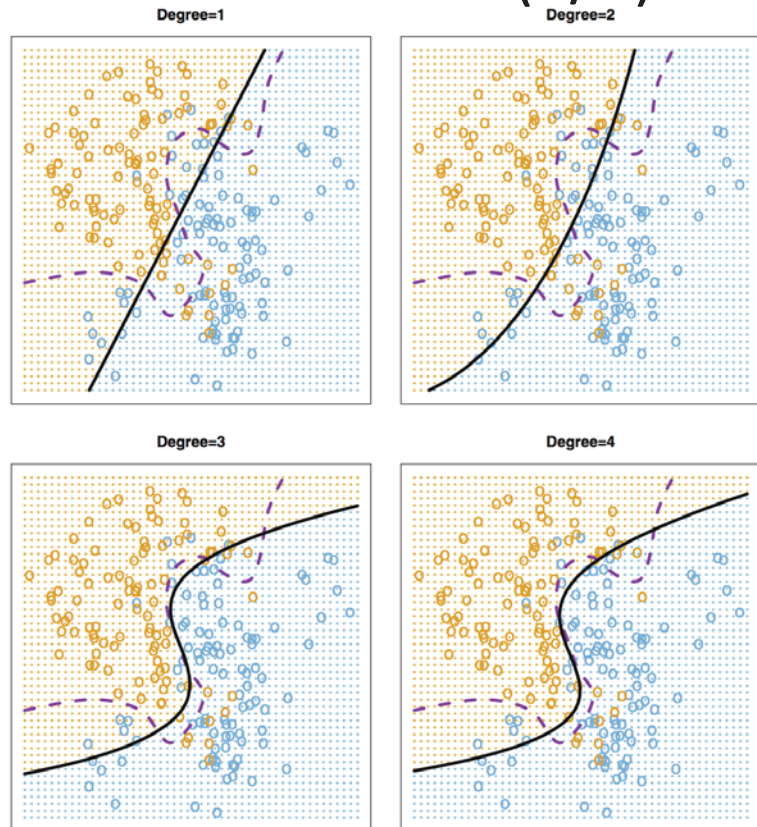
$$L(y, \hat{f}(x, \mathcal{D})) = -2 \log \hat{p}(x; \mathcal{D}) = -2 \log \hat{\text{Pr}}(y|x; \mathcal{D})$$

(called the log-likelihood loss, cross-entropy loss, or deviance).

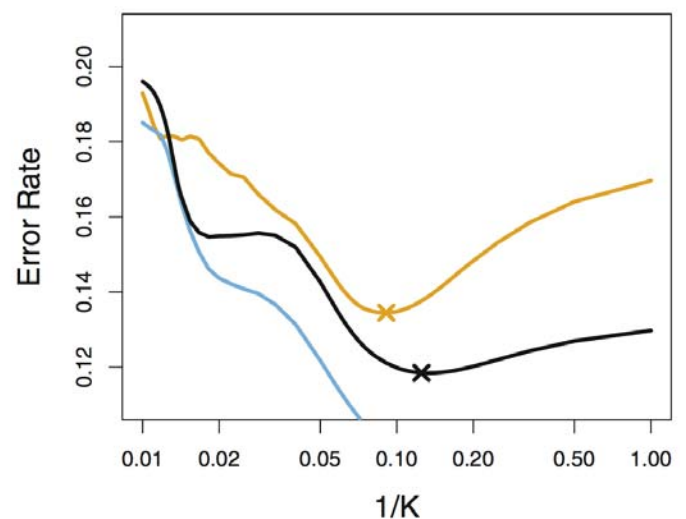
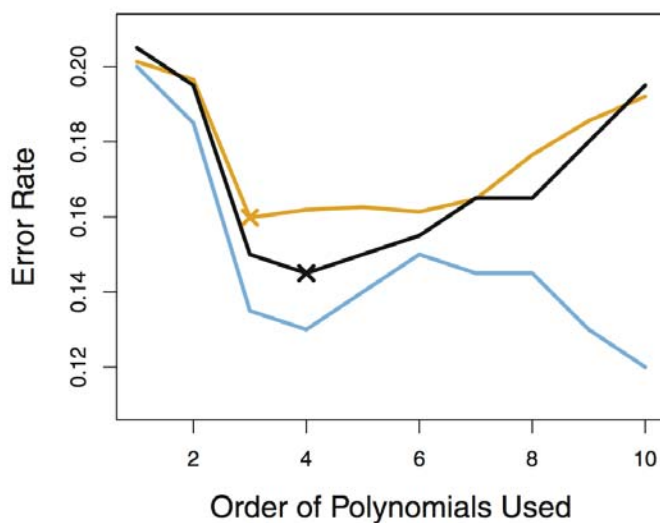
CV on Classification Problems (2/2)

Example:

- Polynomial logistic regression
- Test error rates are 0.201, 0.197, 0.160, 0.162 vs Bayes error rate 0.133.



CV for Model Selection



Bootstrap via A Toy Example (1/3)

- Bootstrapping is a tool in statistics (often for measuring accuracy) that involves *random sampling with replacement* of the available data set.

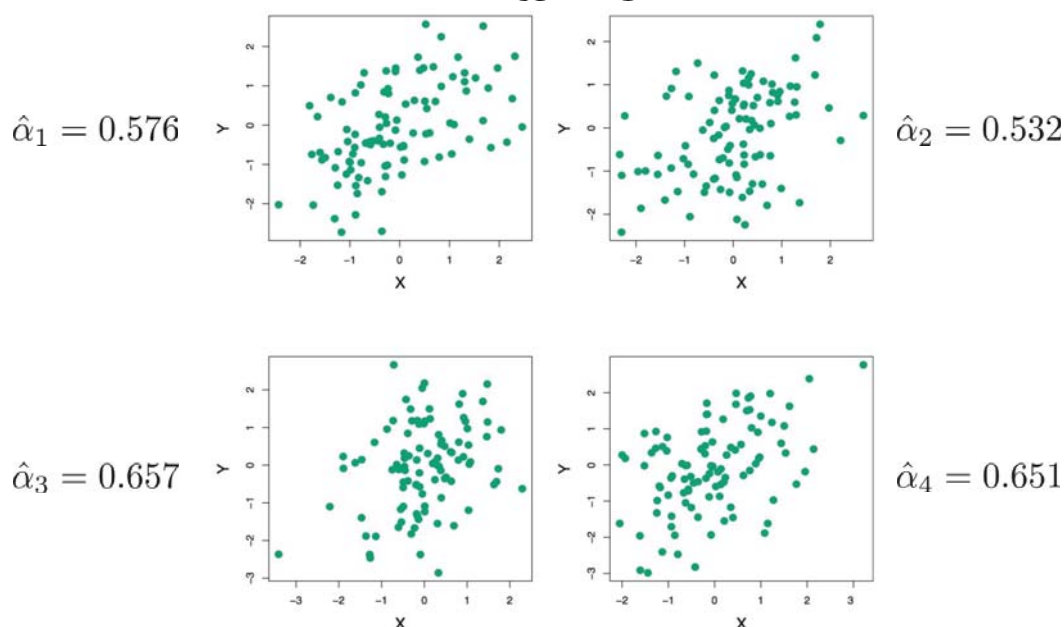
Toy Example:

- Investment of a fixed sum of money in two financial assets that yield returns of X and Y , respectively.
- By investing only a fraction α on X and the remaining $1 - \alpha$ on Y , the return is $\alpha X + (1 - \alpha)Y$.
- The choice of α that minimizes the variability is

$$\alpha = \arg \min_{\alpha \in [0,1]} \text{Var}(\alpha X + (1 - \alpha)Y) = \frac{\sigma_Y^2 - \sigma_{XY}}{\sigma_X^2 + \sigma_Y^2 - 2\sigma_{XY}}.$$

Bootstrap via A Toy Example (2/3)

- Simulate 100 realizations of the values of X and Y using parameters $\sigma_X^2 = 1$, $\sigma_Y^2 = 1.25$, $\sigma_{XY} = 0.5$, and use them to estimate σ_X^2 , σ_Y^2 , σ_{XY} and thus α .



Bootstrap via A Toy Example (3/3)

- By generating 1000 estimates, we can compute the mean $\bar{\alpha} = \frac{1}{1000} \sum_{r=1}^{1000} \hat{\alpha}_r$ and the standard error

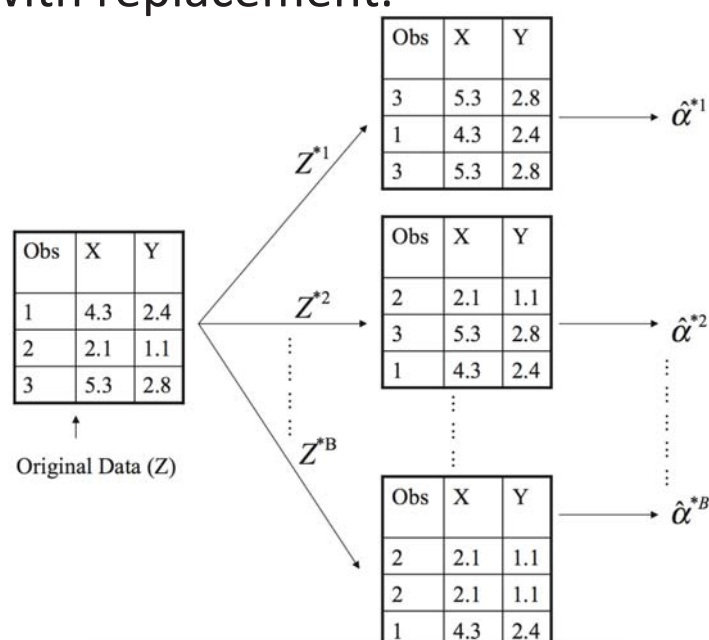
$$\widehat{\text{SE}}(\hat{\alpha}) = \sqrt{\frac{1}{1000 - 1} \sum_{r=1}^{1000} (\hat{\alpha}_r - \bar{\alpha})^2}.$$

- In practice, we cannot generate data at will to obtain these estimates.

➔ **Key idea:** Obtain distinct data sets by repeatedly sampling observations (with replacement) from the original data set. (This is called *bootstrapping*!)

Bootstrap Illustration

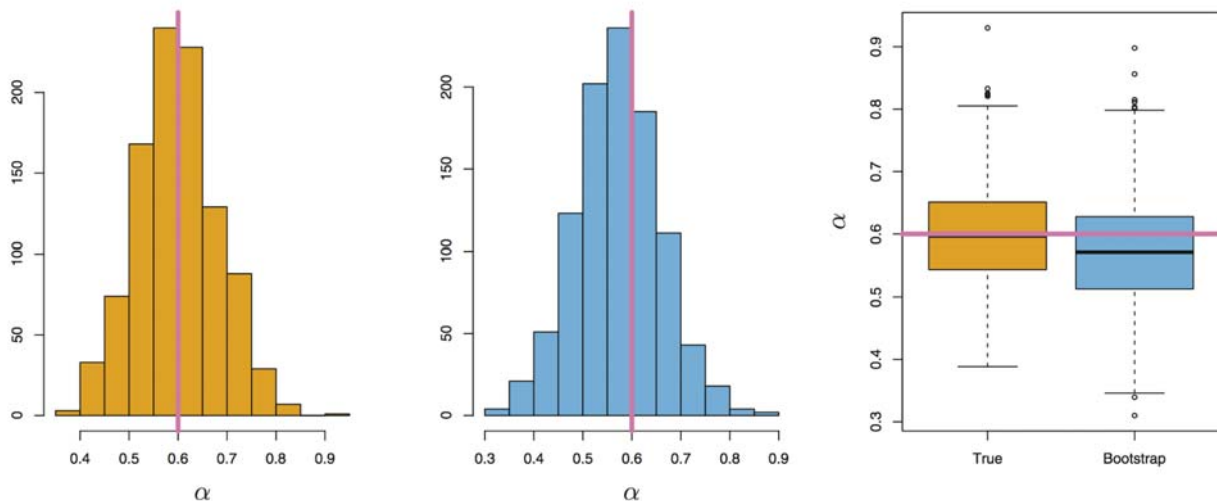
- Generate B bootstrap data sets $Z^{*1}, Z^{*2}, \dots, Z^{*B}$ by randomly sampling n observations from the original data set Z with replacement.



Bootstrap vs Simulated Data Sets

- The bootstrap estimates are $\hat{\alpha}^{*1}, \hat{\alpha}^{*2}, \dots, \hat{\alpha}^{*B}$ and the standard error is

$$\widehat{\text{SE}}_B(\hat{\alpha}) = \sqrt{\frac{1}{B-1} \sum_{r=1}^B \left(\hat{\alpha}^{*r} - \frac{1}{B} \sum_{r'=1}^B \hat{\alpha}^{*r'} \right)^2}.$$



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Bootstrap for Test Error Estimation (1/2)

Question: Can bootstrap be used to estimate test error?

- E.g., use the original data set \mathcal{D} for testing, and bootstrap data sets $\mathcal{D}^{*1}, \mathcal{D}^{*2}, \dots, \mathcal{D}^{*B}$ for training.
- The estimated test error is

$$\widehat{\text{Err}}_{\text{boot}} = \frac{1}{B} \frac{1}{n} \sum_{b=1}^B \sum_{i=1}^n L(y_i, \hat{f}(x_i; \mathcal{D}^{*b})).$$

➔ Problem: \mathcal{D}^{*b} contains the (x_i, y_i) with probability

Bootstrap for Test Error Estimation (2/2)

- Let $\mathcal{I}^{(-i)} = \{b : (x_i, y_i) \neq \mathcal{D}^{*b}\}$ be the set of indices of bootstrap data sets that do not include (x_i, y_i) . This yields the leave-one-out bootstrap with

$$\widehat{\text{Err}}^{(1)} = \frac{1}{n} \sum_{i=1}^n \sum_{b \in \mathcal{I}^{(-i)}} \frac{1}{|\mathcal{I}^{(-i)}|} L(y_i, \hat{f}(x_i; \mathcal{D}^{*b})).$$

➔ Problem: The number of distinct observations in each data set is only $0.632n$.

- Adopt a weighted estimate

$$\widehat{\text{Err}}^{(0.632+)} = \left(1 - \frac{0.632}{1 - 0.368\hat{R}}\right) \overline{\text{err}} + \frac{0.632}{1 - 0.368\hat{R}} \widehat{\text{Err}}^{(1)}$$

$$\text{where } \hat{R} = \frac{\widehat{\text{Err}}^{(1)} - \overline{\text{err}}}{\frac{1}{n^2} \sum_i \sum_{i'} L(y_i, \hat{f}(x_{i'}; \mathcal{D})) - \overline{\text{err}}}.$$