

4. 6.5.4. Let  $X_1, X_2, \dots, X_n$  and  $Y_1, Y_2, \dots, Y_m$  be independent random samples from the two normal distributions  $N(0, \theta_1)$  and  $N(0, \theta_2)$ .

- (a) Find the likelihood ratio  $\Lambda$  for testing the composite hypothesis  $H_0 : \theta_1 = \theta_2$  against the composite alternative  $H_1 : \theta_1 \neq \theta_2$ .
- (b) This  $\Lambda$  is a function of what  $F$ -statistic that would actually be used in this test?

$$X \sim N(0, \theta_1) \quad Y \sim N(0, \theta_2)$$

$$\Lambda = \frac{L(\hat{\omega})}{L(\hat{\Omega})}$$

(a)  $\theta_1 = \theta_2 = \theta_0$  5%

$$L(\theta_0) = \frac{1}{(2\pi\theta_0)^{\frac{m+n}{2}}} e^{-\frac{1}{2\theta_0} \left( \sum_{i=1}^n X_i^2 + \sum_{j=1}^m Y_j^2 \right)}$$

$$\ell(\theta) = -\frac{m+n}{2} \ln(2\pi\theta) + \frac{1}{2\theta} \left( \sum_{i=1}^n X_i^2 + \sum_{j=1}^m Y_j^2 \right)$$

$$\frac{\partial}{\partial \theta} \ell(\theta) = -\frac{m+n}{2} \times \frac{1}{2\pi\theta} \times 2\pi + \frac{1}{2\theta^2} \left( \sum_{i=1}^n X_i^2 + \sum_{j=1}^m Y_j^2 \right) = 0 \quad \frac{m+n}{2\theta} = \frac{1}{2\theta^2} \left( \sum_{i=1}^n X_i^2 + \sum_{j=1}^m Y_j^2 \right)$$

$$\hat{\theta} = \frac{\sum_{i=1}^n X_i^2 + \sum_{j=1}^m Y_j^2}{m+n}$$

$\theta_1 \neq \theta_2$  5%

$$L(\theta_1, \theta_2) = (2\pi\theta_1)^{-\frac{n}{2}} e^{-\frac{1}{2\theta_1} \sum_{i=1}^n X_i^2} \cdot (2\pi\theta_2)^{-\frac{m}{2}} e^{-\frac{1}{2\theta_2} \sum_{j=1}^m Y_j^2}$$

$$\ell(\theta_1, \theta_2) = -\frac{n}{2} \ln(2\pi\theta_1) + \left( -\frac{1}{2\theta_1} \sum_{i=1}^n X_i^2 \right) + \left( -\frac{m}{2} \right) \ln(2\pi\theta_2) + \left( -\frac{1}{2\theta_2} \sum_{j=1}^m Y_j^2 \right)$$

$$\frac{\partial \ell(\theta_1, \theta_2)}{\partial \theta_1} = 0 \Rightarrow -\frac{n}{2\theta_1} + \frac{1}{2\theta_1^2} \sum_{i=1}^n X_i^2 = 0 \quad \hat{\theta}_1 = \frac{1}{n} \sum_{i=1}^n X_i^2$$

$$\frac{\partial \ell(\theta_1, \theta_2)}{\partial \theta_2} = 0 \Rightarrow -\frac{m}{2\theta_2} + \frac{1}{2\theta_2^2} \sum_{j=1}^m Y_j^2 = 0 \quad \hat{\theta}_2 = \frac{1}{m} \sum_{j=1}^m Y_j^2$$

$$\Lambda = \frac{L(\theta_0)}{L(\hat{\theta})} = \frac{(2\pi\theta_0)^{-\frac{m+n}{2}} e^{-\frac{1}{2\theta_0} \left( \sum_{i=1}^n X_i^2 + \sum_{j=1}^m Y_j^2 \right)}}{(2\pi\hat{\theta}_1)^{-\frac{n}{2}} e^{-\frac{1}{2\hat{\theta}_1} \sum_{i=1}^n X_i^2} (2\pi\hat{\theta}_2)^{-\frac{m}{2}} e^{-\frac{1}{2\hat{\theta}_2} \sum_{j=1}^m Y_j^2}} = \frac{\theta_0^{\frac{m+n}{2}} e^{-\frac{(m+n)}{2}}}{\theta_1^{\frac{n}{2}} \theta_2^{\frac{m}{2}} e^{-\frac{n}{2}} e^{-\frac{m}{2}}}$$

$$\Lambda = \frac{L(\theta_0)}{L(\hat{\theta})} =$$

$$\frac{\theta_1^{\frac{n}{2}} \theta_2^{\frac{m}{2}}}{\theta^{\frac{m+n}{2}}} \Rightarrow$$

$$\frac{\left(\frac{1}{n}\right)^{\frac{n}{2}} \left(\frac{1}{m}\right)^{\frac{m}{2}} \left(\sum_{i=1}^n x_i^2\right)^{\frac{n}{2}} \left(\sum_{j=1}^m y_j^2\right)^{\frac{m}{2}}}{\left(\frac{1}{m+n}\right)^{\frac{m+n}{2}} \left(\sum_{i=1}^n x_i^2 + \sum_{j=1}^m y_j^2\right)^{\frac{m+n}{2}}}$$

$$\Lambda = \frac{(m+n)^{\frac{m+n}{2}}}{n^{\frac{n}{2}} m^{\frac{m}{2}}} \frac{\left(\sum_{i=1}^n x_i^2\right)^{\frac{n}{2}} \left(\sum_{j=1}^m y_j^2\right)^{\frac{m}{2}}}{\left(\sum_{i=1}^n x_i^2 + \sum_{j=1}^m y_j^2\right)^{\frac{m+n}{2}}}$$

(b)  $F = \frac{\frac{1}{n} \sum_{i=1}^n x_i^2}{\frac{1}{m} \sum_{j=1}^m y_j^2}$  The relevant F statistic  $F = \frac{\frac{1}{n} \sum_{i=1}^n x_i^2}{\frac{1}{m} \sum_{j=1}^m y_j^2} \neq$