

Department of Economics

Working Paper Series

Second-order approximation of dynamic models without the use of tensors

09-004	Paul Gomme Concordia University Paul Klein University of Western Ontario
--------	---



UNIVERSITÉ
Concordia
UNIVERSITY

Department of Economics

Working Paper Series

Second-order approximation of dynamic models without the use of tensors

09-004 | Paul Gomme
Concordia University
Paul Klein
University of Western Ontario



UNIVERSITÉ

Concordia

UNIVERSITY

Second-order approximation of dynamic models without the use of tensors*

Paul Gomme
Concordia University

Paul Klein
University of Western Ontario

First draft: May 17, 2005
This version: March 25, 2009

Abstract

Several approaches to finding the second-order approximation to a dynamic model have been proposed recently. This paper differs from the existing literature in that it makes use of the [Magnus and Neudecker \(1999\)](#) definition of the Hessian matrix. The key result is a linear system of equations that characterizes the second-order coefficients. No use is made of multi-dimensional arrays or tensors, a practical implication of which is that it is much easier to transcribe the mathematical representation of the solution into usable computer code. Matlab code is available from <http://www.ssc.uwo.ca/economics/faculty/klein>; Fortran 90 code is available from <http://alcor.concordia.ca/~pgomme/>

Keywords: Solving dynamic models; second-order approximation
JEL classification: E0; C63

*We thank Audra Bowlus, Elizabeth Caucutt, Martin Gervais, Lance Lochner and Igor Livshits.

1 Introduction

There are a number of methods for solving for first-order approximations of dynamic models, including: [Blanchard and Kahn \(1980\)](#); [King *et al.* \(2002\)](#); [Klein \(2000\)](#); [Uhlig \(1999\)](#); [Sims \(2001\)](#); [Christiano \(2002\)](#). Each approach has distinct advantages and disadvantages, but they all deliver essentially the same solutions. This paper contributes to the growing literature on finding *second-order* approximations to dynamic models; this literature includes [Schmitt-Grohé and Uribe \(2004\)](#); [Kim *et al.* \(2005\)](#); [Lombardo and Sutherland \(2007\)](#). Like first-order solutions, the papers describing second-order solutions have distinct advantages and disadvantages. A distinct disadvantage of [Schmitt-Grohé and Uribe \(2004\)](#) and [Kim *et al.* \(2005\)](#) is their use of tensor notation (multidimensional arrays). Like [Lombardo and Sutherland \(2007\)](#), our paper uses “standard” matrix algebra making the presentation of the solution method more transparent; coding the solution method is likewise more straightforward.¹ Unlike previous work in the literature, we use the [Magnus and Neudecker \(1999\)](#) definition of the Hessian matrix; associated with this definition of the Hessian matrix, there is a chain rule. We use this chain rule to solve for the matrices characterizing the second-order approximation to the decision rules.

At this stage, one might well ask why a second-order solution method is needed in the first place. The answer lies in the intersection between speed and accuracy. First-order solution methods are incredibly fast compared with more accurate solution methods like [Coleman \(1990\)](#) and the methods described in [Judd \(1998\)](#). For many purposes, first-order solution methods are more than adequate; for example, the usual set of second moments generated by business cycle theorists are virtually identical across all solution methods. There are a number of applications for which first-order solutions are inadequate, including optimal policy

¹Granted, [Schmitt-Grohé and Uribe \(2004\)](#) and [Kim *et al.* \(2005\)](#) have made their code publicly available, partially obviating the need for others to actually code their solution methods. However, their code is written in `Matlab` which may be prohibitively expensive to some researchers, including students. Further, compiled languages like `Fortran` and `C` or `C++` are considerably faster than interpreted languages like `Matlab` or `Gauss`.

exercises that require accurate solutions to the ratio of marginal utilities, and in addressing asset pricing. For such applications, second-order solution methods provide more accurate solutions at a marginal time cost.

The paper is organized as follows. Section 2 lays the groundwork by establishing the mathematical foundations for what we do. Section 3 discusses a simple asset-pricing application of our approach and compares our solution to an almost-exact solution method. Section 4 gives a practical guide, suitable for anyone who is impatient with the mathematical details and wants to get started with computing quadratic approximations as quickly as possible. Section 5 gives a detailed example of our solution method, applied to a home production model to show how to incorporate correlated shock innovations. Section 6 concludes.

2 Theory

This section describes the various building blocks necessary to define and characterize the second-order accurate approximation around the non-stochastic steady state to the solution of a dynamic model.

2.1 The model

In general, the equilibrium conditions (Euler equations, constraints and market clearing conditions) of a wide variety of dynamic economic models can be expressed as:

$$\mathbb{E}_t[f(x_{t+1}, y_{t+1}, x_t, y_t)] = 0 \quad (1)$$

where f maps $\mathbb{R}^{2n_x+2n_y}$ into $\mathbb{R}^{n_x+n_y}$. x_t is a vector of state variables and y_t as a vector of non-state variables (including control variables). x_t is of length n_x while y_t is of length n_y . Following [Schmitt-Grohé and Uribe \(2004\)](#), we introduce σ , a variable that scales the variance in order to define the second-order approximation around the non-stochastic steady state.

Eq. (1), together with a stability condition, define (exact) solution functions g and h whose roles are defined via

$$y_t = g(x_t, \sigma) \quad (2)$$

and

$$x_{t+1} = h(x_t, \sigma) + \sigma \varepsilon_{t+1}. \quad (3)$$

where $\{\varepsilon_t\}$ is an exogenous, i.i.d. sequence of random variable with zero mean and variance matrix Σ . Notice that this specification allows for arbitrary cross-correlations between the shocks. In [Schmitt-Grohé and Uribe \(2004\)](#), correlations between the shocks can be handled by premultiplying the shocks by a matrix η . If Σ is positive definite, then η can easily be obtained by computing the Cholesky decomposition of Σ . However, if Σ is merely positive *semidefinite*, then our approach is more straightforward.

Our approximation is computed around the non-stochastic state at which $x_t = x_{t+1} = \bar{x}$ and $y_t = y_{t+1} = \bar{y}$, defined via

$$f(\bar{x}, \bar{y}, \bar{x}, \bar{y}) = 0$$

and $\sigma = 0$.

For future reference, it is convenient to define

$$z(x, \sigma; \tilde{g}, \tilde{h}) := \mathbb{E} \left[f \left(\tilde{h}(x, \sigma) + \sigma \varepsilon, \tilde{g} \left(\tilde{h}(x, \sigma) + \sigma \varepsilon \right); x, \tilde{g}(x, \sigma) \right) \right] \quad (4)$$

for arbitrary functions \tilde{g} and \tilde{h} , and where ε has an expected value of 0 and variance Σ . Notice that

$$z(x, \sigma; g, h) \equiv 0.$$

2.2 Some Preliminaries

This subsection develops notation, definitions, and results that will prove useful later.

As stated in [Magnus and Neudecker \(1999\)](#), the second-order Taylor expansion of a twice differentiable function $f : \mathbb{R}^n \rightarrow \mathbb{R}^m$ is given by

$$f(x) \approx f(x_0) + [Df(x_0)](x - x_0) + \frac{1}{2} (I_m \otimes (x - x_0)') [Hf(x_0)](x - x_0) \quad (5)$$

where the matrix of first-order derivatives of f is defined by

$$Df(x) = \frac{\partial f(x)}{\partial x'} = \begin{bmatrix} \frac{\partial f^1(x)}{\partial x_1} & \frac{\partial f^1(x)}{\partial x_2} & \cdots & \frac{\partial f^1(x)}{\partial x_n} \\ \frac{\partial f^2(x)}{\partial x_1} & \ddots & & \vdots \\ \vdots & & \ddots & \vdots \\ \frac{\partial f^m(x)}{\partial x_1} & \cdots & \cdots & \frac{\partial f^m(x)}{\partial x_n} \end{bmatrix}$$

and the matrix of second-order derivatives of f is defined by

$$Hf(x) = \frac{\partial^2 f(x)}{\partial x \partial x'} = D \text{vec}((Df(x))').$$

In the above, $f^i(x)$ denotes the i^{th} equation; that is,

$$f(x) = \begin{bmatrix} f^1(x) \\ f^2(x) \\ \vdots \\ f^m(x) \end{bmatrix}$$

Notice that

$$Df(x) = \begin{bmatrix} Df^1(x) \\ Df^2(x) \\ \vdots \\ Df^m(x) \end{bmatrix}$$

and that

$$Hf(x) = \begin{bmatrix} Hf^1(x) \\ Hf^2(x) \\ \vdots \\ Hf^m(x) \end{bmatrix}.$$

Thus the Hessian $Hf(x)$ is of dimension $mn \times n$ and consists of m vertically concatenated symmetric $n \times n$ matrices. Whereas [Schmitt-Grohé and Uribe \(2004\)](#) and [Kim *et al.* \(2005\)](#) adopt tensor notation to represent the Hessian, here we stack the matrices of second-order derivatives so that standard matrix algebra can be used.

The main tool here is the chain rule for Hessian matrices – missing from [Schmitt-Grohé and Uribe \(2004\)](#); [Kim *et al.* \(2005\)](#); [Lombardo and Sutherland \(2007\)](#) – stated as Theorem 9 in chapter 6 of [Magnus and Neudecker \(1999\)](#).

Theorem 1 (Chain rule) *Let $f : \mathbb{R}^n \rightarrow \mathbb{R}^m$ and $g : \mathbb{R}^m \rightarrow \mathbb{R}^p$ and be twice differentiable and define*

$$h(x) = g(f(x)).$$

Then, letting $y = f(x)$,

$$Hh(x) = (I_p \otimes Df(x))'(Hg(y))Df(x) + (Dg(y) \otimes I_n)Hf(x).$$

Proof: See [Magnus and Neudecker \(1999\)](#). ■

2.3 Second-order Approximation of the Economic Model

Solving for the second-order approximate solutions proceeds via a “guess-and-verify” strategy. Given the representation of the second-order approximation in Eq. (5), the second-order Taylor series approximations of g and h can be represented as

$$\widehat{g}(x) = \sigma^2 k_y + Fx + \frac{1}{2}(I_{n_y} \otimes x')Ex \quad (6)$$

and

$$\widehat{h}(x) = \sigma^2 k_x + Px + \frac{1}{2}(I_{n_x} \otimes x')Gx \quad (7)$$

where it is assumed, without loss of generality, that all variables have been expressed as deviations from their steady state values; that is, $\bar{x} = 0$ and $\bar{y} = 0$. In Eq. (6), F is the gradient of g with respect to x , E is the corresponding Hessian, and k_y is the Hessian of g with respect to σ . Similarly, in Eq. (7) P is the gradient of h with respect to x , G the corresponding Hessian, and k_x the Hessian of h with respect to σ . Notice that there is no linear term in σ in Eqs. (6) and (7), nor are there cross terms involving x and σ ; these results were established in [Schmitt-Grohé and Uribe \(2004\)](#). The reason why these terms are missing is because the second-order approximation is around $\sigma = 0$.

Based on the definition of z in Eq. (4), we characterize the second-order approximations \widehat{g} and \widehat{h} via

$$Dz(0, 0; \widehat{g}, \widehat{h}) = 0 \quad (8)$$

$$Hz(0, 0; \widehat{g}, \widehat{h}) = 0. \quad (9)$$

In particular, Eq. (8) characterizes the matrices F and P (the linear parts) while Eq. (9) characterizes the matrices E , G , k_x and k_y (the quadratic parts). While it may not be obvious, Eq. (8) does not involve any E , G , k_x or k_y , and so can be used to obtain F and P independent of the quadratic terms. In other words, there is a recursive structure to these equations that allows us to solve first for the linear parts, then for the quadratic parts.

2.4 First-order approximation

There are a variety of methods for finding F and P in terms of Df , starting in the economics literature with [Blanchard and Kahn \(1980\)](#). We follow [Klein \(2000\)](#) in applying a generalized Schur method. [King and Watson \(2002\)](#) show how to manipulate the system of log-linearized

equations so that a Schur method can be applied when the matrix A in Eq. (10) is singular. The generalized Schur method handles the problem of a singular A by providing a unified treatment of finite unstable and infinite generalized eigenvalues. A practical implication of the generalized Schur method is that the algebraic manipulations are kept to an absolute minimum. Blanchard and Kahn (1980) use the Jordan form which can suffer from problems related to numerical stability; small perturbations to the underlying matrices can lead to large changes in the Jordan form.

Following Klein (2000) one may proceed as follows, keeping in mind that we are after a non-explosive solution only. The linear approximation of the equilibrium conditions can be written as

$$A \begin{bmatrix} x_{t+1} \\ E_t y_{t+1} \end{bmatrix} = B \begin{bmatrix} x_t \\ y_t \end{bmatrix} + \begin{bmatrix} \epsilon_{t+1} \\ 0 \end{bmatrix} \quad (10)$$

where $A = [f_1 \ f_2]$ and $B = [f_3 \ f_4]$ and f_i denotes that part of the gradient associated with the i^{th} argument. That is, A corresponds to the first-order derivatives of f (the function characterizing the equilibrium of the model) with respect variables dated $t + 1$, and B the first-order derivatives with respect to variables dated t . Notice that the matrices A and B are both $(n_x + n_y) \times (n_x + n_y)$.

The key theorem required here is stated as Theorem 7.7.1 in Golub and van Loan (1996).

Theorem 2 (Generalized Schur Form) *Let A and B be $n \times n$ matrices. If there is a $z \in \mathbb{C}$ such that $|B - zA| \neq 0$, then there exist matrices Q, Z, T and S such that*

1. Q and Z are Hermitian, i.e. $Q^H Q = Q Q^H = I_n$ and similarly for Z , where H denotes the Hermitian transpose (transpose followed by complex conjugation or vice versa).
2. T and S are upper triangular (all entries below the main diagonal are zero).
3. $QA = SZ^H$ and $QB = TZ^H$.

4. There is no i such that $s_{ii} = t_{ii} = 0$.²

Moreover, the matrices Q , Z , S and T can be chosen in such a way as to make the diagonal entries s_{ii} and t_{ii} appear in any desired order.

Proof: See [Golub and van Loan \(1996\)](#). ■

We will choose the following ordering of the pairs s_{ii}, t_{ii} : the ones satisfying $|s_{ii}| > |t_{ii}|$ appear first. We will call these pairs the *stable generalized eigenvalues*.

Taking conditional expectations of Eq. (10) yields

$$A\mathbb{E}_t \begin{bmatrix} x_{t+1} \\ y_{t+1} \end{bmatrix} = B \begin{bmatrix} x_t \\ y_t \end{bmatrix}. \quad (11)$$

Premultiply by Q :

$$QA\mathbb{E}_t \begin{bmatrix} x_{t+1} \\ y_{t+1} \end{bmatrix} = QB \begin{bmatrix} x_t \\ y_t \end{bmatrix}. \quad (12)$$

From Theorem 2, $QA = SZ^H$ and $QB = TZ^H$ which means that the preceding equation can be rewritten as:

$$SZ^H\mathbb{E}_t \begin{bmatrix} x_{t+1} \\ y_{t+1} \end{bmatrix} = TZ^H \begin{bmatrix} x_t \\ y_t \end{bmatrix}, \quad (13)$$

or,

$$S\mathbb{E}_t \begin{bmatrix} s_{t+1} \\ u_{t+1} \end{bmatrix} = T \begin{bmatrix} s_t \\ u_t \end{bmatrix}, \quad (14)$$

where

$$\begin{bmatrix} s_t \\ u_t \end{bmatrix} := Z^H \begin{bmatrix} x_t \\ y_t \end{bmatrix} \quad (15)$$

and where s_t is the same length as x_t and u_t the same length as y_t .

²Here we denote the row i , column j element of any matrix M by m_{ij} .

From Theorem 2, the matrices S and T are upper triangular. Consequently, Eq. (14) can be written out as

$$\begin{bmatrix} S_{11} & S_{12} \\ 0 & S_{22} \end{bmatrix} \mathbb{E}_t \begin{bmatrix} s_{t+1} \\ u_{t+1} \end{bmatrix} = \begin{bmatrix} T_{11} & T_{12} \\ 0 & T_{22} \end{bmatrix} \begin{bmatrix} s_t \\ u_t \end{bmatrix}. \quad (16)$$

The last block of this last equation can be written out as

$$S_{22} \mathbb{E}_t[u_{t+1}] = T_{22} u_t. \quad (17)$$

If S_{22} and T_{22} constitute an unstable matrix pair (meaning that the generalized eigenvalue pairs of these matrices satisfy $|s_{ii}| < |t_{ii}|$), then any solution to Eq. (10) with bounded mean must satisfy $u_t = 0$ for all t . If S_{22} and T_{22} constitute a weakly unstable matrix pair (meaning that the generalized eigenvalue pairs of these matrices satisfy $|s_{ii}| \leq |t_{ii}|$), then any solution to Eq. (10) with bounded variance must satisfy $u_t = 0$ for all t , unless $\Sigma = 0$.

Given $u_t = 0$ for all t , the first block of Eq. (16) says

$$S_{11} \mathbb{E}_t[s_{t+1}] = T_{11} s_t. \quad (18)$$

If S_{11} and T_{11} constitute a stable matrix pair (meaning, as on page 8, that the generalized eigenvalue pairs of these matrices satisfy $|s_{ii}| > |t_{ii}|$), then S_{11} is invertible. Hence we may write

$$\mathbb{E}_t[s_{t+1}] = S_{11}^{-1} T_{11} s_t. \quad (19)$$

Rewrite Eq. (15) which defined (s_t, u_t) as

$$\begin{bmatrix} x_t \\ y_t \end{bmatrix} = Z \begin{bmatrix} s_t \\ u_t \end{bmatrix} = \begin{bmatrix} Z_{11} & Z_{12} \\ Z_{21} & Z_{22} \end{bmatrix} \begin{bmatrix} s_t \\ u_t \end{bmatrix} \quad (20)$$

Given the earlier result that $u_t = 0$ for all t , it follows that

$$x_t = Z_{11}s_t \quad (21)$$

$$y_t = Z_{21}s_t. \quad (22)$$

If Z_{11} is invertible, then

$$s_t = Z_{11}^{-1}x_t \quad (23)$$

$$y_t = \underbrace{Z_{21}Z_{11}^{-1}}_F x_t. \quad (24)$$

Substitute Eq. (23) into Eq. (19):

$$\mathbb{E}_t[Z_{11}^{-1}x_{t+1}] = S_{11}^{-1}T_{11}Z_{11}^{-1}x_t, \quad (25)$$

or,

$$\mathbb{E}_t[x_{t+1}] = Z_{11}S_{11}^{-1}T_{11}Z_{11}^{-1}x_t. \quad (26)$$

Finally, dropping the expectation operator yields:

$$x_{t+1} = \underbrace{Z_{11}S_{11}^{-1}T_{11}Z_{11}^{-1}}_P x_t + \varepsilon_{t+1}. \quad (27)$$

In summary, if there are exactly as many state variables as there are stable generalized eigenvalues and if Z_{11} is invertible, then any solution to Eq. (10) with bounded mean and variance has the representation given by Eq. (24) and Eq. (27) and any pair of stochastic processes $\{x_t\}$ and $\{y_t\}$ satisfying Eqs. (24) and (27) will solve Eq. (10) and have bounded mean and variance, regardless of x_0 .

2.5 Second-order approximation

Using Eq. (9) and Theorem 1 (the chain rule) we may derive the following conditions, where H denotes the Hessian and D the gradient of f , both evaluated at the non-stochastic steady

state.

$$(I_{n_x+n_y} \otimes M')HM + (D \otimes I_{n_x}) \begin{bmatrix} G \\ (I_{n_y} \otimes P')EP + (F \otimes I_{n_x})G \\ 0 \\ E \end{bmatrix} = 0 \quad (28)$$

$$\text{trm}((I_{n_x+n_y} \otimes N')HNS) + D \begin{bmatrix} 2k_x \\ \text{trm}((I_{n_y} \otimes \Sigma)E) + 2Fk_x + 2k_y \\ 0 \\ 2k_y \end{bmatrix} = 0 \quad (29)$$

where

$$M = \begin{bmatrix} P \\ FP \\ I_{n_x} \\ F \end{bmatrix}, \quad N = \begin{bmatrix} I_{n_x} \\ F \\ 0 \\ (n_x+n_y) \times n_x \end{bmatrix} \quad (30)$$

and where we define the matrix trace (trm) of an $nm \times n$ matrix

$$\begin{bmatrix} A'_1 & A'_2 & \dots & A'_m \end{bmatrix}'$$

as the $m \times 1$ vector

$$\begin{bmatrix} \text{tr}(A_1) & \text{tr}(A_2) & \dots & \text{tr}(A_m) \end{bmatrix}'.$$

Using the partition $D = [f_1 \ f_2 \ f_3 \ f_4]$, we can summarize Eq. (28) via

$$A_1 + B_1G + B_2C_1EP + B_2C_2G + B_4E = 0$$

where $B_i = f_i \otimes I_{n_x}$, $C_1 = I_{n_y} \otimes P'$ and $C_2 = F \otimes I_{n_x}$. This equation is linear in $\text{vec}(E)$ and $\text{vec}(G)$. To see this, apply the vec operator to both sides. The result is

$$\text{vec}(A_1) + ((P' \otimes B_2C_1) + I_{n_x} \otimes B_4) \text{vec}(E) + (I_{n_x} \otimes (B_1 + B_2C_2)) \text{vec}(G) = 0.$$

Thus the linear system we have to solve is given by

$$\begin{bmatrix} (P' \otimes B_2 C_1) + I_{n_x} \otimes B_4 & I_{n_x} \otimes (B_1 + B_2 C_2) \end{bmatrix} \begin{bmatrix} \text{vec}(E) \\ \text{vec}(G) \end{bmatrix} = -\text{vec}(A_1). \quad (31)$$

Notice that the equation above involves neither k_x nor k_y which means that this equation gives a solution for (E, G) independently of (k_x, k_y) .

Finally, we use Eq. (29) to solve for k_x and k_y . Eq. (29) is linear in k_x and k_y . In order to write it in a way more amenable to solving, recall the partition

$$D = [f_1 \ f_2 \ f_3 \ f_4]$$

and rewrite Eq. (29) as

$$\text{trm}((I_{n_x+n_y} \otimes N') H N \Sigma) + 2f_1 k_x + f_2 \text{trm}((I_{n_y} \otimes \Sigma) E) + 2f_2 F k_x + 2f_2 k_y + 2f_4 k_y = 0$$

which in turn can be rewritten as

$$-2[f_1 + f_2 F \ f_2 + f_4] \begin{bmatrix} k_x \\ k_y \end{bmatrix} = f_2 \text{trm}((I_{n_y} \otimes \Sigma) E) + \text{trm}((I_{n_x+n_y} \otimes N') H N \Sigma). \quad (32)$$

2.6 Summary

Obtaining a second-order approximate solution to Eq. (1) involves the following steps:

1. Derive the gradient and Hessian matrices of the economic model, evaluated at steady state.
2. Cast the gradient in the form of Eq. (10).
3. Apply the generalized Schur decomposition to obtain the matrices S and T in Eq. (16).
4. Compute matrices F and P from Eqs. (24) and (27), respectively.
5. Compute matrices E and G from Eq. (31).

6. Solve for the vectors k_x and k_y via Eq. (32).

3 Application: A simple asset pricing model

Consider an endowment economy where a representative consumer maximizes

$$\mathbb{E} \left[\sum_{t=0}^{\infty} \beta^t \frac{c_t^{1-\gamma} - 1}{1-\gamma} \right]$$

and where the endowment satisfies the stochastic difference equation

$$\ln c_{t+1} = \rho \ln c_t + \varepsilon_{t+1}$$

where $\{\varepsilon_t\}$ is a white noise process with variance σ_ε^2 . Using annual U.S. consumption data from 1929-2005 and removing a geometric trend, we estimate $\rho = 0.953$ and $\sigma_\varepsilon = 0.0214$. Setting $\beta = 0.97$, we then compute bond prices and equity prices using a quadratic approximation and, for comparison, a discretization of the consumption process following [Tauchen and Hussey \(1991\)](#) with 100 Gauss-Hermite quadrature points.

Table 1: Returns in a simple asset pricing model

γ	Tauchen-Hussey		Linear		Quadratic	
	Equity	Bond	Equity	Bond	Equity	Bond
Mean Returns						
1	3.12%	3.07%	3.12%	3.09%	3.12%	3.07%
2	3.15%	3.00%	3.16%	3.11%	3.15%	3.00%
5	3.33%	2.52%	3.43%	3.11%	3.33%	2.53%
10	3.88%	0.81%	4.38%	3.17%	3.84%	0.82%
Standard Deviations of Returns						
1	2.23%	0.34%	2.23%	0.34%	2.23%	0.34%
2	3.61%	0.68%	3.60%	0.68%	3.61%	0.68%
5	7.77%	1.70%	7.75%	1.70%	7.77%	1.70%
10	15.06%	3.48%	14.82%	3.45%	14.79%	3.34%

We take the [Tauchen and Hussey \(1991\)](#) solutions as the “truth.” Section 3 shows that both the linear and quadratic approximations do quite well with respect to the standard deviations of returns. While the linear approximation does reasonably well with respect to average returns for low risk aversion (low γ), it does quite poorly when risk aversion is high. This problem is particularly noticeable with respect to average bond returns. In contrast, the quadratic approximation does quite well on mean returns for both equity and bonds and for a wider range of coefficients of relative risk aversion.

4 Practical guide

The purpose of this paper is mainly to make it easy for a user who wants to compute the quadratic approximation to a dynamic model. In this section, therefore, we briefly go through the steps required in computing the quadratic approximation, basing our discussion on a simple stochastic growth model.

A social planner wants to maximize

$$\mathbb{E} \left[\sum_{t=0}^{\infty} \beta^t \ln c_t \right]$$

subject to

$$c_t + k_{t+1} = z_t k_t^\theta + (1 - \delta)k_t$$

with $k_0 > 0$ given and the non-negativity constraints $k_{t+1} \geq 0$ and $c_t \geq 0$. In addition, the choices of c_t and k_{t+1} must not anticipate the information generated by the process z_t , which satisfies

$$\ln z_{t+1} = \rho \ln z_t + \varepsilon_{t+1}$$

where ε_t is an i.i.d. process with mean zero and variance σ^2 and where $0 < \rho < 1$.

The equilibrium conditions are given by the following equations.

$$\frac{1}{c_t} - \beta \theta E \left[\frac{\theta k_t^{\theta-1}}{c_{t+1}} \middle| z_t \right] = 0,$$

$$E [\ln z_{t+1} | z_t] - \rho \ln z_t = 0$$

and

$$c_t + k_{t+1} - z_t k_t^\theta - (1 - \delta) k_t = 0.$$

For convenience, let the state variables be given by logs, $x_t = [\ln z_t \ln k_t]'$, and similarly let the control variable be given by $y_t = \ln c_t$. Then the function f , which defines the equilibrium, should be defined as

$$f^1(x'_1, x'_2, y', x_1, x_2, y) = \frac{1}{\exp(y)} - \beta \left[\frac{1 + \theta \exp(x_2)^{\theta-1} - \delta}{\exp(y')} \right],$$

$$f^2(x'_1, x'_2, y', x_1, x_2, y) = \exp(y) + \exp(x'_2) - \exp(x_1) \exp(x_2)^\theta - (1 - \delta) \exp(x_2),$$

and

$$f^3(x'_1, x'_2, y', x_1, x_2, y) = x'_1 - \rho \ln x_1.$$

Now use these equations to find the deterministic steady state. That steady state (z, k, c) solves

$$c + \delta k = z k^\theta,$$

$$\beta [1 + z \theta k^{\theta-1} - \delta] = 1$$

and

$$\ln z = 0.$$

To form the gradient D , take the partial derivatives, evaluated at the steady state, of f^1 , f^2 and f^3 , all with respect to the six variables of which they are a function. We now have three (row) vectors of length six; concatenate them vertically (stack them on top of each other) to create a 3×6 matrix. Partition this matrix so as to fit the format of Eq. (10), i.e. define A as

the first three columns of the gradient and B as *minus* the last three columns.

The next step is to compute the linear part of the decision rule. To do this, compute the (ordered) generalized Schur decomposition (T, S) of the matrix pair (B, A) as defined in Eq. (16). Then compute matrices F and P from Eq. (24) and Eq. (27), respectively.

Next compute the Hessian matrix of second derivatives of each of the functions f^1 , f^2 and f^3 . Each matrix is 6×6 . Now vertically concatenate these matrices into an 18×6 matrix. This is our H matrix.

The final two pieces of information that our computer needs are the prediction error variance matrix of the state vector and the number of variables in that state vector. In this case, there are two state variables and their prediction error variance matrix is

$$\Sigma = \begin{bmatrix} \sigma_\varepsilon^2 & 0 \\ 0 & 0 \end{bmatrix}.$$

With all this information, we can compute matrices E and G from Eq. (31), and solve for the vectors k_x and k_y via Eq. (32).

Alternatively, we can just feed the gradient D , the Hessian H , the variance matrix Σ and the number of state variables into the Matlab function `solab2` available from <http://economics.uwo.ca/faculty/klein/personal/codes.htm> or a corresponding subroutine in Fortran.

Once we have computed the vectors and matrices k_x , k_y , F , P , E and G we can simulate solutions to the model by using Eqs. (6) and (7), keeping in mind that the variables in those equations are defined as deviations from the steady state.

5 Application

Of course, all the examples in Schmitt-Grohé and Uribe (2004) can be replicated using the method presented here. To show how to apply the method to a model with correlated error

innovations, consider a home production model along the lines of [Greenwood *et al.* \(1995\)](#).

To keep the presentation simple, consider the planner's version of the model:

$$\max E \sum_{t=0}^{\infty} \beta^t U(c_{mt}, c_{ht}, h_{mt}, h_{ht})$$

subject to:

$$\begin{aligned} c_{mt} + k_{m,t+1} + k_{h,t+1} &= F(k_{mt}, h_{mt}; z_{mt}) + (1 - \delta_m)k_{mt} + (1 - \delta_h)k_{ht} \\ c_{ht} &= H(k_{ht}, h_{ht}; z_{ht}). \end{aligned}$$

For now, assume that the shocks, z_{mt} and z_{ht} follow a time-homogeneous finite-state Markov chain.

The relevant “Euler” equations are:

$$\begin{aligned} U_1(c_{mt}, c_{ht}, h_{mt}, h_{ht})F_2(k_{mt}, h_{mt}; z_{mt}) + U_3(c_{mt}, c_{ht}, h_{mt}, h_{ht}) &= 0 \\ U_3(c_{mt}, c_{ht}, h_{mt}, h_{ht})H_2(k_{ht}, h_{ht}; z_{ht}) + U_4(c_{mt}, c_{ht}, h_{mt}, h_{ht}) &= 0 \\ U_1(c_{mt}, c_{ht}, h_{mt}, h_{ht}) &= \beta E_t \{ U_1(c_{m,t+1}, c_{h,t+1}, h_{m,t+1}, h_{h,t+1}) [F_1(k_{m,t+1}, h_{m,t+1}; z_{m,t+1}) + 1 - \delta_m] \} \\ U_1(c_{mt}, c_{ht}, h_{mt}, h_{ht}) &= \beta E_t \{ U_2(c_{m,t+1}, c_{h,t+1}, h_{m,t+1}, h_{h,t+1}) H_1(k_{h,t+1}, h_{h,t+1}; z_{h,t+1}) \\ &\quad + U_1(c_{m,t+1}, c_{h,t+1}, h_{m,t+1}, h_{h,t+1})(1 - \delta_h) \} \end{aligned}$$

The equations characterizing the economic model are, then, the four Euler equations and the two constraints.

5.1 Functional Forms

Utility:

$$U(c_m, c_h, h_m, h_h) = \begin{cases} \frac{[C(c_m, c_h)(1 - h_m - h_h)^\omega]^{1-\gamma}}{1-\gamma} & \gamma \in (0, 1) \cup (1, \infty) \\ \ln C(c_m, c_h) + \omega(1 - h_m - h_h) & \gamma = 1 \end{cases}$$

Table 2: Parameter Values

β	γ	ω	ψ	ξ	α	δ_m	η	δ_h	ρ_m	ρ_h
0.99	1	0.55	0.5	0.35	0.3	0.018	0.4	0.015	0.95	0.95

Consumption aggregator:

$$C(c_m, c_h) = \begin{cases} [\psi c_m^\xi + (1 - \psi) c_h^\xi]^{1/\xi} & \xi \in (-\infty, 0) \cup (0, 1) \\ c_m^\psi c_h^{1-\psi} & \xi = 0 \end{cases}$$

Market production:

$$F(k_m, h_m; z_m) = z_m k_m^\alpha h_m^{1-\alpha}$$

Home production:

$$H(k_h, h_h; z_h) = z_h k_h^\alpha h_h^{1-\alpha}$$

The shock processes are:

$$\ln z_{mt} = \rho_m \ln z_{m,t-1} + \varepsilon_{mt}$$

$$\ln z_{ht} = \rho_h \ln z_{h,t-1} + \varepsilon_{ht}$$

where

$$\begin{bmatrix} \varepsilon_{mt} \\ \varepsilon_{ht} \end{bmatrix} \sim N(0, \tilde{\Sigma})$$

5.2 Parameter Values

Most of the parameter values are summarized in Table 2; the remainder are:

$$\Sigma = \begin{bmatrix} 0.00763^2 & \frac{2}{3}0.00763^2 & 0 & 0 \\ \frac{2}{3}0.00763^2 & 0.00763^2 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}$$

The standard deviation of the innovations to the shocks is equal at 0.00763, and their correlation is $2/3$. Notice that this matrix is *not* positive definite and so it does not have a unique Cholesky decomposition. Any attempt to compute the Cholesky decomposition will run into an error message.

5.3 Solution Matrices

The gradient and Hessian matrices are omitted owing to their size; the interested reader is directed to either of our web sites. The solution matrices are:

$$P = \begin{bmatrix} 0.9500 & 0.0000 & 0.0000 & 0.0000 \\ 0.0000 & 0.9500 & 0.0000 & 0.0000 \\ 1.2111 & -0.2034 & 0.2652 & 0.2939 \\ -0.9439 & 0.1851 & 0.6314 & 0.6996 \end{bmatrix}$$

$$F = \begin{bmatrix} 0.4880 & -0.2775 & 0.3325 & 0.2263 \\ -0.3298 & 1.1096 & -0.0470 & 0.5379 \\ 0.8286 & -0.0844 & 0.0731 & -0.3517 \\ -0.5497 & 0.1827 & -0.0784 & 0.2299 \end{bmatrix}$$

$$E = \begin{bmatrix} 2 \times 10^{-10} & 0. & 0. & 0. \\ 0. & 0. & 0. & 0. \\ 0. & 0. & 0. & 0. \\ 0. & 0. & 0. & 0. \\ 0. & 0. & 0. & 0. \\ 0. & 2 \times 10^{-10} & 0. & 0. \\ 0. & 0. & 0. & 0. \\ 0. & 0. & 0. & 0. \\ -2.5119 & 0.3341 & 0.4869 & 0.5223 \\ 0.3341 & -0.0129 & -0.0632 & -0.0697 \\ 0.4869 & -0.0632 & 0.0417 & -0.2378 \\ 0.5223 & -0.0697 & -0.2378 & 0.0312 \\ 0.1525 & 0.1001 & -0.0652 & -0.1130 \\ 0.1001 & -0.0592 & -0.0133 & -0.0141 \\ -0.0652 & -0.0133 & 0.3411 & -0.2981 \\ -0.1130 & -0.0141 & -0.2981 & 0.3710 \end{bmatrix}$$

$$G = \begin{bmatrix} 0.1149 & 0.0346 & -0.0445 & -0.0663 \\ 0.0346 & -0.0639 & 0.0175 & -0.0128 \\ -0.0445 & 0.0175 & 0.1566 & -0.1409 \\ -0.0663 & -0.0128 & -0.1409 & 0.178 \\ -0.1224 & 0.0318 & -0.0276 & 0.0469 \\ 0.0318 & -0.0310 & 0.0094 & -0.0127 \\ -0.0276 & 0.0094 & 0.0416 & -0.0329 \\ 0.0469 & -0.0127 & -0.0329 & 0.0220 \\ -0.7862 & 0.0648 & -0.0245 & 0.3494 \\ 0.0648 & 0.0023 & 0.0034 & -0.0279 \\ -0.0245 & 0.0034 & -0.1566 & 0.1607 \\ 0.3494 & -0.0279 & 0.1607 & -0.2893 \\ -0.2040 & 0.0530 & -0.0460 & 0.0782 \\ 0.0530 & -0.0517 & 0.0157 & -0.0212 \\ -0.0460 & 0.0157 & 0.0694 & -0.0549 \\ 0.0782 & -0.0212 & -0.0549 & 0.0366 \end{bmatrix}$$

$$k_x = 10^{-8} \times \begin{bmatrix} -1 \times 10^{-5} & -1 \times 10^{-5} & 4273.0 & -3535.6 \end{bmatrix}'$$

$$k_y = \begin{bmatrix} -2 \times 10^{-5} & -4 \times 10^{-6} & 1 \times 10^{-5} & -7 \times 10^{-6} \end{bmatrix}'$$

6 Conclusion

Our paper makes a number of contributions to the literature on second-order approximate solutions. First, since we use the [Magnus and Neudecker \(1999\)](#) definition of the Hessian matrix, we can also use their definition of the chain rule which is key to us being able to

characterize the solutions for the parameters of the second-order terms as systems of linear equations. Second, by eschewing the use of tensor notation, our solution method is arguably easier to code. Third, we show, both in the theory and in an example, how to incorporate correlated shock innovations.

References

- Blanchard, Oliver Jean and Charles M. Kahn (1980). "The Solution of Linear Difference Models Under Rational Expectations," *Econometrica*, 48 (5).
- Christiano, Lawrence J. (2002). "Solving Dynamic Equilibrium Models by a Method of Undetermined Coefficients," *Computational Economics*, 20.
- Coleman, Wilbur, II (1990). "Solving the Stochastic Growth Model by Policy-Function Iteration," *Journal of Business and Economic Statistics*, 8 (1): 27.
- Golub, Gene and Charles van Loan (1996). *Matrix Computations, Third Edition*, Baltimore and London: The Johns Hopkins University Press.
- Greenwood, Jeremy, Richard Rogerson and Randall Wright (1995). "Household Production in Real Business Cycle Theory," in Thomas Cooley, ed., "Frontiers of Business Cycle Research," Princeton, N.J.: Princeton University Press, pp. 157–174.
- Judd, Kenneth L. (1998). *Numerical Methods in Economics*, Cambridge: The MIT Press.
- Kim, Jinill, Sunghyun Kim, Ernst Schaumburg and Christopher Sims (2005). "Calculating and Using Second Order Accurate Solutions of Discrete Time Dynamic Equilibrium Models," Manuscript.
- King, Robert G., Charles I. Plosser and Sergio T. Rebelo (2002). "Production, Growth and Business Cycles: Technical Appendix," *Computational Economics*, 20 (1–2): 87–116.
- King, Robert G. and Mark Watson (2002). "System Reduction and Solution Algorithms for Singular Linear Difference Systems under Rational Expectations," *Computational Economics*, 20 (1–2): 57–86.
- Klein, Paul (2000). "Using the Generalized Schur Form to Solve a Multivariate Linear Rational Expectations Model," *Journal of Economic Dynamics and Control*, 24 (10): 1405–1423.
- Lombardo, Giovanni and Alan Sutherland (2007). "Computing Second-order-accurate Solutions for Rational Expectations Models Using Linear Solution Methods," *Journal of Economic Dynamics and Control*, 31 (2): 515–530.
- Magnus, Jan and Heinz Neudecker (1999). *Matrix Differential Calculus With Applications in Statistics and Econometrics*, John Wiley and Sons.
- Schmitt-Grohé, Stephanie and Martín Uribe (2004). "Solving dynamic general equilibrium models using a second-order approximation to the policy function," *Journal of Economic Dynamics and Control*, 28: 755–775.
- Sims, Christopher A. (2001). "Solving Linear Rational Expectations Models," *Computational Economics*, 20: 1–20.

- Tauchen, George and Robert Hussey (1991). "Quadrature-Based Methods for Obtaining Approximate Solutions to Nonlinear Asset Pricing Models," *Computational Economics*, 59 (2): 371–396.
- Uhlig, Harald (1999). "A Toolkit for Analyzing Nonlinear Dynamic Stochastic Models Easily," in Ramon Marimon and Andrew Scott, eds., "Computational Methods for the Study of Dynamic Economies," New York: Oxford University Press, pp. 30–61.