

KU Leuven Summer School
Segment 5B
Information Flow in the IFR Inference Problem

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Consider a different/hypothetical pandemic (in code)

There are $K=10$ jurisdictions. Unbeknownst to us, here is the truth:

```
### infection fatality rate
### known to be constant across jurisdictions, but value unknown

ifr <- 0.025

### actual infection rates, again unknown

ir <- c(0.2, 0.1, 0.6, 0.5, 0.3, 0.35, 0.2, 0.1, 0.5, 0.45)

### actual extents of preferential testing

phi <- c(8, 2, 3, 3, 4, 7, 6, 2, 3, 7)
```

What does nature reveal to us?

```
### death rate
```

```
dr.obs <- ir*ifr
```

```
dr.obs
```

```
## [1] 0.00500 0.00250 0.01500 0.01250 0.00750 0.00875 0.00500
```

```
## [8] 0.00250 0.01250 0.01125
```

```
### test positivity rate
```

```
tp.obs <- 1-(1-ir)^phi
```

```
tp.obs
```

```
## [1] 0.832 0.190 0.936 0.875 0.760 0.951 0.738 0.190 0.875 0.985
```

```
### and ensure no cheating!
```

```
rm(ir, ifr, phi)
```

And what else do we have going for us?

We have a valid bound for all the preferential testing parameters

```
phi.int <- c(1,10)
```

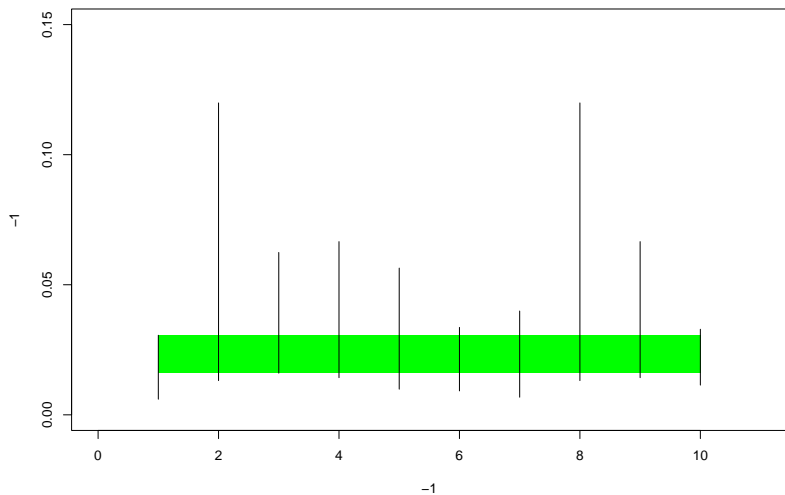
So what do individual jurisdictions tell us about the IFR

```
### e.g., jurisdiction 5 tells us that IFR must lie in  
dr.obs[5] / ( 1 - (1-tp.obs[5])^(1/phi.int) )
```

```
## [1] 0.00987 0.05641
```

```
### do for all jurisdictions  
ifr.bnd.lo <- dr.obs / ( 1 - (1-tp.obs)^(1/phi.int[1]) )  
ifr.bnd.hi <- dr.obs / ( 1 - (1-tp.obs)^(1/phi.int[2]) )
```

Visualize how well we have squeezed in on our target



Have learned that the IFR lies in

```
round(c(max(ifr.bnd.lo), min(ifr.bnd.hi)),3)
```

```
## [1] 0.016 0.031
```

But getting back to Bayes

The above proof-of-concept takes place in asymptotia, no explicit Bayes inference.

So now back to Bayes, but for demonstration, want/need a simpler sandbox than the full COVID example

Generative Model

prior $f(\mu_a)f(\sigma_a)f(\mu_b)f(\sigma_b)f(\omega)$

$$(a_1, \dots, a_k | \mu_a, \sigma_a^2) \stackrel{iid}{\sim} N(\mu_a, \sigma_a^2)$$

$$(b_1, \dots, b_k | \mu_b, \sigma_b^2) \stackrel{iid}{\sim} N(\mu_b, \sigma_b^2)$$

$$(c_1, \dots, c_k) \stackrel{iid}{\sim} U(0, \omega)$$

$$\log(D_k/P_k) \sim N(a_k + b_k, \nu_i^2)$$

$$\log(CC_k/T_k) \sim N(b_k + c_k, \kappa_i^2)$$

Code this up

```
genmod.string <- "  
model {  
  
  mn.a ~ dnorm(0, 0.01)  
  sd.a ~ dbeta(1,19) ### limited ifr variation  
  prc.a <- pow(sd.a,-2)  
  
  mn.b ~ dnorm(0, 0.01)  
  sd.b ~ dunif(0,1) ### not very limited ir variation  
  prc.b <- pow(sd.b,-2)  
  
  for (i in 1:k) {  
    a[i] ~ dnorm(mn.a, prc.a)  
    b[i] ~ dnorm(mn.b, prc.b)  
    c[i] ~ dunif(0,2) ### effect of preferential testing  
  
    log.dth.rate[i] ~ dnorm(a[i] + b[i], prc.dth[i])  
    log.tps.rate[i] ~ dnorm(b[i] + c[i], prc.tps[i])  
  }  
}"
```

First test dataset

```
k <- 30

set.seed(13)
a.true <- rnorm(k, log(0.04), log(1.005))
b.true <- rnorm(k, log(0.35), log(1.15))
c.true <- runif(k, 0,2)

prc.dth <- rep(100000*(0.02/0.98), k) ### ???
prc.tps <- rep(1000*(0.5/0.5), k)      ### ???

log.dth.rate <- rnorm(k, a.true+b.true, 1/sqrt(prc.dth))
log.tps.rate <- rnorm(k, b.true+c.true, 1/sqrt(prc.tps))
```

Take a look

```
head(cbind(log.dth.rate, sqrt(1/prc.dth),  
           log.tps.rate, sqrt(1/prc.tps)))
```

| ## | log.dth.rate | log.tps.rate |
|---------|--------------|----------------|
| ## [1,] | -4.29 0.0221 | -0.8597 0.0316 |
| ## [2,] | -4.11 0.0221 | -0.6240 0.0316 |
| ## [3,] | -4.32 0.0221 | 0.1447 0.0316 |
| ## [4,] | -4.33 0.0221 | 0.0265 0.0316 |
| ## [5,] | -4.29 0.0221 | -0.1526 0.0316 |
| ## [6,] | -4.35 0.0221 | -1.1286 0.0316 |

Turn the crank

```
### generative model, data go in
mod <- jags.model(
  textConnection(genmod.string),
  data=list(log.dth.rate=log.dth.rate, prc.dth=prc.dth,
            log.tps.rate=log.tps.rate, prc.tps=prc.tps, k=k),
  n.chains=5)

update(mod, 2500) # burn-in

### MC output comes out
opt.JAGS <- coda.samples(mod, n.iter=500000, thin=100,
  variable.names=c("mn.a", "a[1]", "b[1]", "c[1]"))
```

Get an answer

```
MCMCsummary(opt.JAGS)
```

| ## | | mean | sd | 2.5% | 50% | 97.5% | Rhat | n.eff |
|----|------|-------|--------|-------|--------|--------|------|-------|
| ## | a[1] | -3.17 | 0.0566 | -3.28 | -3.173 | -3.056 | 1 | 5869 |
| ## | b[1] | -1.12 | 0.0597 | -1.24 | -1.120 | -1.004 | 1 | 6424 |
| ## | c[1] | 0.26 | 0.0674 | 0.13 | 0.259 | 0.394 | 1 | 7616 |
| ## | mn.a | -3.17 | 0.0446 | -3.26 | -3.169 | -3.080 | 1 | 3769 |

Second test dataset

```
k <- 30

set.seed(13)
a.true <- rnorm(k, log(0.04), log(1.005))
b.true <- rnorm(k, log(0.35), log(1.05))
c.true <- runif(k, 0.4, 1.6)

prc.dth <- rep(100000*(0.02/0.98), k) ### ???
prc.tps <- rep(1000*(0.5/0.5), k)      ### ???

log.dth.rate <- rnorm(k, a.true+b.true, 1/sqrt(prc.dth))
log.tps.rate <- rnorm(k, b.true+c.true, 1/sqrt(prc.tps))
```

Turn the crank

```
### generative model, data go in
mod <- jags.model(
  textConnection(genmod.string),
  data=list(log.dth.rate=log.dth.rate, prc.dth=prc.dth,
            log.tps.rate=log.tps.rate, prc.tps=prc.tps, k=k),
  n.chains=5)

update(mod, 2500) # burn-in

### MC output comes out
opt.JAGS <- coda.samples(mod, n.iter=500000, thin=100,
  variable.names=c("mn.a", "a[1]", "b[1]", "c[1]"))
```


Get an answer

```
MCMCsummary(opt.JAGS)
```

| ## | | mean | sd | 2.5% | 50% | 97.5% | Rhat | n.eff |
|----|------|--------|-------|-------|--------|--------|------|-------|
| ## | a[1] | -3.138 | 0.231 | -3.57 | -3.110 | -2.789 | 1.02 | 90 |
| ## | b[1] | -1.150 | 0.231 | -1.50 | -1.178 | -0.719 | 1.02 | 90 |
| ## | c[1] | 0.608 | 0.233 | 0.17 | 0.636 | 0.969 | 1.02 | 89 |
| ## | mn.a | -3.132 | 0.230 | -3.56 | -3.103 | -2.786 | 1.02 | 93 |

Our pesky folk theorem rears its inconvenient head again!

Thoughts?