Monte Carlo simulation of a QML estimation of an ARCH(1) model

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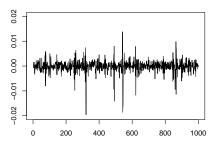
Outline

- 1. ARCH(1) model
- 2. Maximum Likelihood estimation
- 3. Quasi Maximum Likelihood estimation
- 4. Monte Carlo Set up
- 5. Monte Carlo Simulation
- 6. Simulation Results
- 7. Further information



Simulated ARCH(1) process

Simulated ARCH(1) process with 1000 observations, $\alpha < 1$





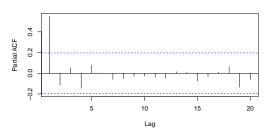


Estimation of ARCH(1) Models

ARCH(1) process ε_t has an AR(1) representation:

$$\varepsilon_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \eta_t, \eta_t = \sigma_t^2 (Z_t^2 - 1)$$

Series sim1002







Maximum Likelihood Estimation (MLE)

oxdot Assume pdf of $\varepsilon_t | \mathcal{F}_{t-1}$ normal

$$p(\varepsilon_t | \mathcal{F}_{t-1}) = \frac{1}{\sqrt{2\pi}\sigma_t} \exp\left\{-\frac{\varepsilon_t^2}{2\sigma_t^2}\right\}$$

□ Log-likelihood function $I(\omega, \alpha)$:

$$I(\omega,\alpha) = \sum_{t=2}^{n} I_t(\omega,\alpha) + \log p(\varepsilon_1)$$

where $I(\omega, \alpha) = \log p(\varepsilon_t | \mathcal{F}_{t-1})$

moralpark

Maximum Likelihood Estimation (MLE)

First order conditions

$$\sum_{t=2}^{n} \frac{\partial I_{t}^{b}}{\partial \omega} = 0$$

$$\sum_{t=a+1}^{n} \frac{\partial I_{t}^{b}}{\partial \alpha} = 0$$

$$\sum_{t=q+1}^{n} \frac{\partial I_t^b}{\partial \alpha} = 0$$

- Under the conditions
 - 1. $E[Z_t | \mathcal{F}_{t-1}] = 0$ and $E[Z_t^2 | \mathcal{F}_{t-1}] = 1$
 - 2. $E[log(\alpha Z_t^2)|\mathcal{F}_{t-1}] < 0$ (strong stationarity)

ML estimators are consistent



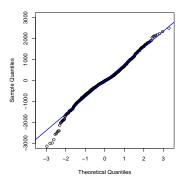
Quasi Maximum likelihood Estimation (QMLE)

- □ Quasi ML (QML): In the case Z_t is not normally distributed, $\hat{\theta}$ is consistent, and asymptotically normally distributed.

$$\sqrt{n}(\hat{\theta} - \theta) \xrightarrow{\mathcal{L}} N(0, J^{-1}IJ^{-1}) \tag{1}$$



QQ Plot of standardized residuals







Requirements

- Control over the random variables and parameters (selection of pdf for random numbers)
- Reproducibility
- Efficiency
- Documented steps



Monte Carlo Set up

- □ R-Package: fGarch
- Testing for different RNG: Mersenne Twister, Knuth-TAOCP-2002, Wichmann-Hill with seed=123.
- Custom function called 'simulation' which takes as input size of the dataset and returns the string of:
 - 1. averaged value of estimated $\hat{\alpha}$
 - 2. standard deviation of estimated $\hat{\alpha}$ from true parameter $\alpha = 0.9$
 - 3. number of $\hat{\alpha}$ which are larger or equal to 1 (evidence of the stationarity violation)

over k = 1000 replications for datasets of size n = 100, 250, 500, 1000.



Monte Carlo Set up continued

Functions used:

- fitGarch() with specified conditional distribution to be calculated with the Quasi Maximum Likelihood Estimation;



Results of Monte Carlo Simulation

n	$k^- 1 \sum_{j=1}^k \hat{\alpha_j}$	$\sqrt{k^- 1 \sum_{j=1}^k (\hat{\alpha_j} - \alpha)^2}$	$\#(\alpha_j \geq 1)$
100	0.824	0.209	34.5%
250	0.868	0.130	23.2%
500	0.880	0.095	17.3%
1000	0.890	0.073	7.4%

Table 1: RNG: Mersenne Twister

SFMqmle



Results with other RNG

n	$k^- 1 \sum_{j=1}^k \hat{\alpha_j}$	$\sqrt{k^-1\sum_{j=1}^k(\hat{\alpha_j}-\alpha)^2}$	$\#(\alpha_j \geq 1)$
100	0.820	0.207	30.1%
250	0.869	0.128	24.1%
500	0.883	0.097	17.2%
1000	0.890	0.072	9.3%

Table 2: RNG: Knuth-TAOCP-2002

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Results with other RNG

n	$k^{-1}\sum_{i=1}^{k}\hat{\alpha_i}$	$\sqrt{k^-1\sum_{j=1}^k(\hat{\alpha_j}-\alpha)^2}$	$\#(\alpha_j \geq 1)$
100	0.822	0.207	31.1%
250	0.868	0.132	24.5%
500	0.884	0.096	17.9%
1000	0.896	0.070	9.7%

Table 3: RNG: Wichmann-Hill

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For Further Reading

- James E Gentle.
 Elements of computational statistics. *QA276*, 4:G455.
- D Wuertz, Y Chalabi, and M Miklovic. fgarch: Rmetrics-autoregressive conditional heteroskedastic modelling, r package version 290.76, 2008
- Jürgen Franke, Wolfgang Karl Härdle, and Christian M Hafner. Statistics of financial markets Forth Edition, Springer-Verlag Berlin Heidelberg, 2015.

