# Monte Carlo simulation of a QML estimation of an ARCH(1) model

Paulina Kurowska Bingling Wang

Ladislaus von Bortkiewicz Chair of Statistics Humboldt–Universität zu Berlin https://github.com/paulinakurowska/SFM<sub>p</sub>roject.git



#### **Outline**

- □ ARCH(1) model
- Quasi Maximum Likelihood estimation
- Monte Carlo Simulation
- simulation Results
- Furthur information



## Definition: ARCH(1) Model

$$\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2$$

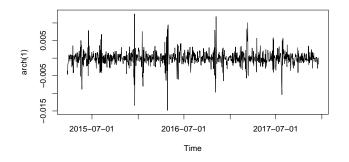
with  $\omega > 0, \alpha > 0$  and

- $ext{ } ext{ } ext$
- $\Box Var(\varepsilon_t | F_{t-1}) = \sigma_t^2 \text{ (semi-strong ARCH)}$
- $\stackrel{\square}{\longrightarrow} \mathcal{P}(\varepsilon_t^2|1,\varepsilon_{t-1},\varepsilon_{t-2},...,\varepsilon_{t-1}^2,\varepsilon_{t-2}^2,...) = \sigma_t^2 \text{(weak ARCH), where } \\ \mathcal{P} \text{ is the best linear projection}$



### Properties of ARCH(1) process

Necessary and sufficient condition for weak stationarity of a semi strong ARCH(1) process is  $\alpha < 1$ 



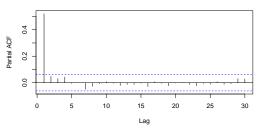


### Estimation of ARCH(1) Models

□ ARCH(1) process  $\varepsilon_t$  has an AR(1) representation:

$$\varepsilon_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \eta_t$$
$$\eta_t = \sigma_t^2 (Z_t^2 - 1)$$

#### Series sim10002



# Maximum Likelihood Estimation (MLE)

 $oxed{\Box}$  Assume pdf of  $\varepsilon_t | \mathcal{F}_{t-1}$  normal

$$p(\varepsilon_t|\mathcal{F}_{t-1}) = \frac{1}{\sqrt{2\pi}\sigma_t} exp\left\{-\frac{1}{2}\frac{\varepsilon_t^2}{\sigma_t^2}\right\}$$

□ Log-likelihood function  $I(\omega, \alpha)$ :

$$I(\omega, \alpha) = \sum_{t=2}^{n} I_t(\omega, \alpha) + logp_{\varepsilon}(\varepsilon_1)$$

where  $I(\omega, \alpha) = \log p(\varepsilon_t | \mathcal{F}_{t-1})$ 



## Maximum Likelihood Estimation (MLE)

First order conditions

$$\sum_{t=2}^{n} \frac{\partial I_{t}^{b}}{\partial \omega} = 0 \text{ and } \sum_{t=q+1}^{n} \frac{\partial I_{t}^{b}}{\partial \alpha} = 0$$

- Under the conditions
  - 1.  $E[Z_t|\mathcal{F}_{t-1}] = 0$  and  $E[Z_t^2|\mathcal{F}_{t-1}] = 1$
  - 2.  $E[log(\alpha Z_t^2)|\mathcal{F}_{t-1}] < 0$ (strong stationarity)

ML estimators are consistent



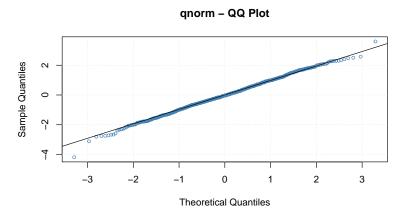
# Quasi Maximum likelihood Estimation (QMLE)

- ☑ If  $Z_t$  is not normally distributed,  $\hat{\theta}$  is consistent, and asymptotically normally distributed, but not efficient. In this case the method is interpreted as quasi ML (QML).

$$\sqrt{n}(\hat{\theta} - \theta) \xrightarrow{\mathcal{L}} N(0, \mathcal{H}^{-1}I\mathcal{H}^{-1})$$
 (1)



#### QQ Plot of standardized residuals





#### Requirements

- Control over the random variables and parameters (selection of pdf for random numbers)
- Reproducibility
- Efficience
- Documented steps
   Source: Elements of Computational Statistics, Gentle 2002



#### Monte Carlo Set up

R-Package: fGarch

Testing for different RNG: Marsenne Twister, Knuth-TAOCP-2002, Wichmann-Hill with seed=123.

We define the function called 'simulation' which takes as input size of the dataset and returns the string of:

- oxdot averaged value of estimated  $\hat{lpha}$
- oxdot standard deviation of estimated  $\hat{\alpha}$  from true parameter  $\alpha = 0.9$
- $oxed{oxed}$  number of  $\hat{lpha}$  which are larger or equal to 1 (evidence of the stationarity violation)

over k=1000 replications for datasets of size n=100,250,500,1000.



### Monte Carlo Set up continued

For building that simulation we use the function fitGarch() from fGarch package with specified conditional distribution to be calculated with the Quasi Maximum Likelihood Estimation, the simulated ARCH(1) models are derived with the help of the function garchSim().



### Monte Carlo Set up continued

QMLE assumes normal distribution and uses robust standard errors for inference. Bollerslev and Wooldridge (1992) proved that if the mean and the volatility equations are correctly specified, the QML estimates are consistent and asymptotically normally distributed. However, the estimates are not efficient.



#### **Results of Monte Carlo Simulation**

#### RNG: Marsenne Twister

n	$k^- 1 \sum_{j=1}^k \hat{\alpha_j}$	$\sqrt{k^-1\sum_{j=1}^k(\hat{\alpha_j}-\alpha)^2}$	$\#(\alpha_j \geq 1)$
100	0.824	0.209	34.5%
250	0.868	0.129	23.2%
500	0.880	0.095	17.3%
1000	0.890	0.073	7.4%



#### Results with other RNG

RNG: Knuth-TAOCP-2002

n	$k^- 1 \sum_{j=1}^k \hat{\alpha_j}$	$\sqrt{k^-1\sum_{j=1}^k(\hat{\alpha_j}-\alpha)^2}$	$\#(\alpha_j \geq 1)$
100	0.820	0.207	30.1%
250	0.869	0.128	24.1%
500	0.883	0.097	17.2%
1000	0.893	0.072	9.3%

RNG: Wichmann-Hill

_	n	$k^- 1 \sum_{j=1}^k \hat{\alpha_j}$	$\sqrt{k^-1\sum_{j=1}^k(\hat{\alpha_j}-\alpha)^2}$	$\#(\alpha_j \geq 1)$
	100	0.822	0.207	31.1%
	250	0.868	0.132	24.5%
	500	0.884	0.096	17.9%
	1000	0.896	0.070	9.7%
M	l estima	tion of an ARCH(1	1) model:	

QML estimation of an ARCH(1) model:

### For Further Reading

- Gentle, James E.

  Elements of Computational Statistics
- R-Package fGarch, Nov 2017
- Bollersev, T., Wooldridge, M.

  Quasi-maximum likelihood estimation and inference in dynamic models with time-varying covariances
- J.Franke , W. K. Haerdle , C. M. Hafner Statistics of Financial Markets: 2011

