

CS680, Spring 2020, Assignment 10

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Exercise 1

$$\begin{aligned}
 & \max_{\forall j, \|\mathbf{z}_j\|_2 \leq \lambda} \|(X + Z)\mathbf{w} - \mathbf{y}\|_2 \\
 & \max_{\forall j, \|\mathbf{z}_j\|_2 \leq \lambda} \|X\mathbf{w} - \mathbf{y} + (\mathbf{z}_1, \mathbf{z}_2, \dots, \mathbf{z}_d)\mathbf{w}\|_2 \\
 & \leq \|X\mathbf{w} - \mathbf{y}\|_2 + \max_{\forall j, \|\mathbf{z}_j\|_2 \leq \lambda} \left\| \sum_{i=1}^d \mathbf{z}_i w_i \right\|_2 \\
 & \leq \|X\mathbf{w} - \mathbf{y}\|_2 + \sum_{i=1}^d |\mathbf{w}_i| \lambda \\
 & \|X\mathbf{w} - \mathbf{y}\|_2 + \lambda \|\mathbf{w}\|_1
 \end{aligned} \tag{1}$$

Therefore, $\min_{\mathbf{w} \in \mathbb{R}^d} \max_{\forall j, \|\mathbf{z}_j\|_2 \leq \lambda} \|(X + Z)\mathbf{w} - \mathbf{y}\|_2$ is equivalent to $\min_{\mathbf{w} \in \mathbb{R}^d} \|X\mathbf{w} - \mathbf{y}\|_2 + \lambda \|\mathbf{w}\|_1$

Exercise 2

1. Re-write the summation as vector and matrix multiplication.

$$\begin{aligned}
 & \sum_{i=1}^n \mathbb{E} \left[(y_i - \mathbf{w}^T \tilde{\mathbf{x}}_i)^2 \right] \\
 & \mathbb{E} \left[(\mathbf{y} - \tilde{X}\mathbf{w})^2 \right] \\
 & \mathbb{E} \left[\mathbf{y}^T \mathbf{y} - 2\mathbf{w}^T \tilde{X}^T \mathbf{y} + \mathbf{w}^T \tilde{X}^T \tilde{X} \mathbf{w} \right] \\
 & \mathbb{E} \left[\mathbf{y}^T \mathbf{y} - 2\mathbf{w}^T (X + \epsilon)^T \mathbf{y} + \mathbf{w}^T (X + \epsilon)^T (X + \epsilon) \mathbf{w} \right] \\
 & \mathbb{E} \left[\mathbf{y}^T \mathbf{y} - 2\mathbf{w}^T X^T \mathbf{y} - 2\mathbf{w}^T \epsilon^T \mathbf{y} + \mathbf{w}^T X^T X \mathbf{w} + \mathbf{w}^T \epsilon^T X \mathbf{w} + \mathbf{w}^T X^T \epsilon \mathbf{w} + \mathbf{w}^T \epsilon^T \epsilon \mathbf{w} \right] \\
 & \mathbb{E} \left[\mathbf{y}^T \mathbf{y} - 2\mathbf{w}^T X^T \mathbf{y} + \mathbf{w}^T X^T X \mathbf{w} \right] + \mathbb{E} \left[-2\mathbf{w}^T \epsilon^T \mathbf{y} + \mathbf{w}^T \epsilon^T X \mathbf{w} + \mathbf{w}^T X^T \epsilon \mathbf{w} \right] + \mathbb{E} \left[\mathbf{w}^T \epsilon^T \epsilon \mathbf{w} \right] \\
 & (\mathbf{y} - X\mathbf{w})^2 + 0 + \mathbf{w}^T \mathbb{E} \left[\epsilon^T \epsilon \right] \mathbf{w} \\
 & \sum_{i=1}^n (y_i - \mathbf{w}^T \mathbf{x}_i)^2 + \mathbf{w}^T \sum_{i=1}^n \mathbb{E} \left[\epsilon_i^2 \right] \mathbf{w} \\
 & \sum_{i=1}^n (y_i - \mathbf{w}^T \mathbf{x}_i)^2 + \mathbf{w}^T \sum_{i=1}^n \left\{ \mathbb{E} \left[\epsilon_i^2 \right] + \text{Var} \left[\epsilon_i \right] \right\} \mathbf{w} \\
 & \sum_{i=1}^n (y_i - \mathbf{w}^T \mathbf{x}_i)^2 + \mathbf{w}^T \sum_{i=1}^n \{0_{d \times d} + \lambda \mathbb{I}_{d \times d}\} \mathbf{w} \\
 & \sum_{i=1}^n (y_i - \mathbf{w}^T \mathbf{x}_i)^2 + n\lambda \|\mathbf{w}\|_2^2
 \end{aligned} \tag{2}$$

Therefore, $\min_{\mathbf{w}} \sum_{i=1}^n \mathbb{E} \left[(y_i - \mathbf{w}^T \tilde{\mathbf{x}}_i)^2 \right]$ is equivalent to $\min_{\mathbf{w}} \sum_{i=1}^n (y_i - \mathbf{w}^T \mathbf{x}_i)^2 + \lambda \|\mathbf{w}\|_2^2$

2. Re-write the summation like Ex2.1.

$$\begin{aligned}
& \sum_{i=1}^n \mathbb{E} \left[(y_i - \mathbf{w}^T \tilde{\mathbf{x}}_i)^2 \right] \\
& \mathbb{E} \left[\mathbf{y}^T \mathbf{y} - 2\mathbf{w}^T \tilde{X}^T \mathbf{y} + \mathbf{w}^T \tilde{X}^T \tilde{X} \mathbf{w} \right] \\
& \mathbb{E} [\mathbf{y}^T \mathbf{y}] - 2\mathbf{w}^T \mathbb{E} [\tilde{X}]^T \mathbf{y} + \mathbf{w}^T \mathbb{E} [\tilde{X}^T \tilde{X}] \mathbf{w} \\
& \mathbf{y}^T \mathbf{y} - 2\mathbf{w}^T \mathbb{E} [\tilde{X}]^T \mathbf{y} + \mathbf{w}^T \mathbb{E} [\tilde{X}^T \tilde{X}] \mathbf{w}
\end{aligned} \tag{3}$$

For $\mathbb{E} [\tilde{X}]$,

$$\begin{aligned}
\mathbb{E} [\tilde{X}]_{ji} &= p \cdot \frac{x_{ji}}{p} = x_{ji} \\
\mathbb{E} [\tilde{X}] &= X
\end{aligned} \tag{4}$$

For $\mathbb{E} [\tilde{X}^T \tilde{X}]$

$$\begin{aligned}
& (\tilde{X}^T \tilde{X})_{ji} = \sum_{k=1}^n \tilde{x}_{kj} \tilde{x}_{ki} \\
\mathbb{E} [\tilde{X}^T \tilde{X}]_{ji} &= \begin{cases} \sum p \cdot \frac{x_{kj}}{p} \cdot \frac{x_{ki}}{p} = \frac{1}{p} (X^T X)_{ji} & j = i \\ \sum p \cdot \frac{x_{kj}}{p} \cdot \frac{x_{ki}}{p} = \frac{1}{p} (X^T X)_{ji} & j \neq i \end{cases} \\
\mathbb{E} [\tilde{X}^T \tilde{X}] &= X^T X + \left(\frac{1}{p} - 1 \right) \text{diag}(X^T X)
\end{aligned} \tag{5}$$

Back to Eq.(3)

$$\begin{aligned}
& \sum_{i=1}^n \mathbb{E} \left[(y_i - \mathbf{w}^T \tilde{\mathbf{x}}_i)^2 \right] \\
& \mathbf{y}^T \mathbf{y} - 2\mathbf{w}^T \mathbb{E} [\tilde{X}]^T \mathbf{y} + \mathbf{w}^T \mathbb{E} [\tilde{X}^T \tilde{X}] \mathbf{w} \\
& \mathbf{y}^T \mathbf{y} - 2\mathbf{w}^T X^T \mathbf{y} + \mathbf{w}^T X^T X \mathbf{w} - \mathbf{w}^T X^T X \mathbf{w} + \mathbf{w}^T \mathbb{E} [\tilde{X}^T \tilde{X}] \mathbf{w} \\
& \sum_{i=1}^n (y_i - \mathbf{w}^T \mathbf{x}_i)^2 + \left(\frac{1}{p} - 1 \right) \mathbf{w}^T \text{diag}(X^T X) \mathbf{w} \\
& \sum_{i=1}^n (y_i - \mathbf{w}^T \mathbf{x}_i)^2 + \left(\frac{1}{p} - 1 \right) \left\| \text{diag}(X^T X)^{\frac{1}{2}} \mathbf{w} \right\|_2^2
\end{aligned} \tag{6}$$

Therefore, $\min \sum_{i=1}^n \mathbb{E} \left[(y_i - \mathbf{w}^T \tilde{\mathbf{x}}_i)^2 \right]$ is equivalent to $\min \sum_{i=1}^n (y_i - \mathbf{w}^T \mathbf{x}_i)^2 + \left(\frac{1}{p} - 1 \right) \left\| \text{diag}(X^T X)^{\frac{1}{2}} \mathbf{w} \right\|_2^2$