A Flexible Bayesian Framework to Study Viral Trait Evolution

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5 June 2019





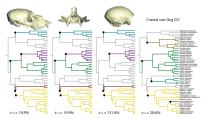


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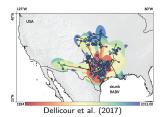
³ Department of Biomathematics, David Geffen School of Medicine at UCLA, USA

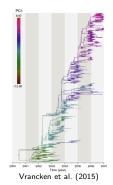
⁴ Department of Human Genetics, David Geffen School of Medicine at UCLA, USA

Phylogenetic Comparative Methods



Aristide et al. (2018)





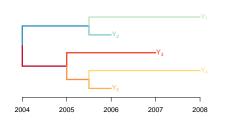
- Various time scales: Myr decade.
- Various traits: morpho, geo, viral.

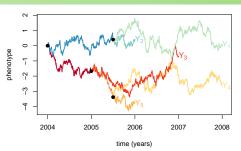
Question: Trait dynamics for an evolving organism?

Outline

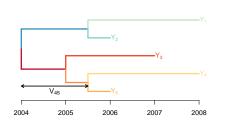
- Models of Trait Evolution
- 2 Efficient Bayesian Inference
- 3 HIV Virulence Heritability Study

BM on a Tree





BM on a Tree





EDS: $dX_t = \sigma dB_t$

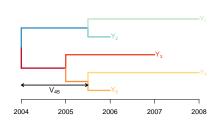
Variance: \mathbb{C} ov $[Y_4; Y_5] = \sigma^2 \times V_{45}$

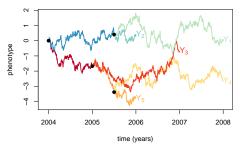
shared evolution time

Expectation: $\mathbb{E}[Y_i] = \mu$

ancestral root value

BM on a Tree

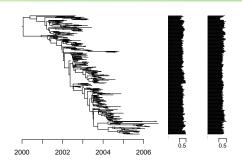




Distribution: Normal

$$\mathbf{Y} \sim \mathcal{N}(\mu \mathbf{1}_n, \sigma^2 \mathbf{V})$$

Multivariate BM

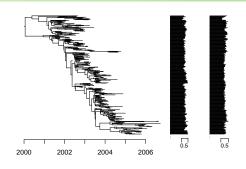


Data: Vectors of p traits

$$\mathbf{Y}_i^T = (Y_{i1}, \dots, Y_{ip})$$

Tree: Influenza H3N2 (Lemey et al., 2014)

Multivariate BM



Data: Vectors of *p* traits

$$\mathbf{Y}_{i}^{T}=\left(Y_{i1},\ldots,Y_{ip}\right)$$

EDS:
$$d \mathbf{X}_t = \mathbf{\Sigma} d \mathbf{B}_t$$

$$\mathbf{R} = \mathbf{\Sigma}^T \mathbf{\Sigma}$$

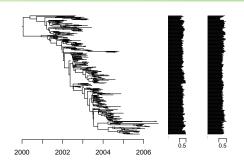
Variance:
$$\mathbb{C}$$
ov $[Y_{ik}; Y_{jl}] = R_{kl} \times V_{ij}$

shared evolution time

Expectation:
$$\mathbb{E}\left[\mathbf{Y}_{\cdot k}\right] = \mu_k$$

Tree: Influenza H3N2 (Lemey et al., 2014)

Multivariate BM



Data: Vectors of *p* traits

$$\mathbf{Y}_{i}^{T}=\left(Y_{i1},\ldots,Y_{ip}\right)$$

Distribution: Matrix Normal

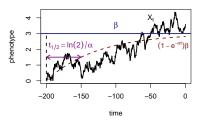
$$oldsymbol{\mathsf{Y}} \sim \mathcal{M} \mathcal{N}(\mathbf{1}_n oldsymbol{\mu}^T, oldsymbol{\mathsf{V}}, oldsymbol{\mathsf{R}})$$

$$\mathbb{V}$$
ar [$\mathsf{vec}(\mathbf{Y})$] = $\mathbf{R} \otimes \mathbf{V}$

Tree: Influenza H3N2 (Lemey et al., 2014)

Ornstein-Uhlenbeck Modeling

(Hansen, 1997)



$$dX_t = \alpha[\beta - X_t] dt + \sigma dB_t$$

Deterministic part:

- β : optimum value.
- α : selection strength (attenuation).

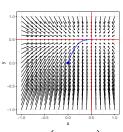
Stochastic part:

- X_t : trait value.
- $\sigma dB(t)$: Brownian fluctuations.

Multivariate OU Modeling

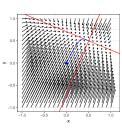
$$d \mathbf{X}_t = \mathbf{A}[\boldsymbol{\beta} - \mathbf{X}_t] d t + \mathbf{\Sigma} d \mathbf{B}_t$$

Diagonal



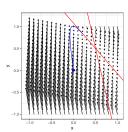
$$\mathbf{A} = \begin{pmatrix} 0.1 & 0 \\ 0 & 0.3 \end{pmatrix}$$

Symmetric



$$\mathbf{A} = \begin{pmatrix} 0.1 & 0.1 \\ 0.1 & 0.3 \end{pmatrix}$$

Diagonalizable in $\mathbb R$

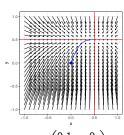


$$\mathbf{A} = \begin{pmatrix} -0.02 & -0.04 \\ 0.2 & 0.2 \end{pmatrix}$$

Multivariate OU Modeling

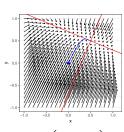
$$\mathrm{d}\, \mathbf{X}_t = \mathbf{A}[eta - \mathbf{X}_t]\,\mathrm{d}\, t + \mathbf{\Sigma}\,\mathrm{d}\, \mathbf{B}_t$$





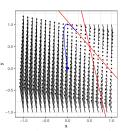
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$$\mathbf{A} = \begin{pmatrix} -0.02 & -0.04 \\ 0.2 & 0.2 \end{pmatrix}$$

Distribution on a Tree: Still Gaussian. No nice Kronecker product.

Goal:

$$p(\boldsymbol{\theta}, \mathcal{T}, \psi \mid \mathbf{Y}, \mathbf{S})$$

Goal:

$$p(\theta, \mathcal{T}, \psi \mid \mathbf{Y}, \mathbf{S}) \propto p(\mathbf{Y}, \mathbf{S} \mid \theta, \mathcal{T}, \psi) p(\theta, \mathcal{T}, \psi)$$

Goal:

$$egin{aligned}
ho\left(m{ heta}, \mathcal{T}, m{\psi} \mid m{ heta}, m{ heta}
ight) &\propto
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ight) \\ &\propto
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ight) \end{aligned}$$

Assumption: \mathbf{Y} and \mathbf{S} independent conditionally on \mathcal{T} .

Goal:

$$\begin{split} \rho\left(\boldsymbol{\theta}, \mathcal{T}, \boldsymbol{\psi} \mid \mathbf{Y}, \mathbf{S}\right) &\propto \rho\left(\mathbf{Y}, \mathbf{S} \mid \boldsymbol{\theta}, \mathcal{T}, \boldsymbol{\psi}\right) \rho\left(\boldsymbol{\theta}, \mathcal{T}, \boldsymbol{\psi}\right) \\ &\propto \rho\left(\mathbf{Y} \mid \boldsymbol{\theta}, \mathcal{T}\right) \rho\left(\mathbf{S} \mid \mathcal{T}, \boldsymbol{\psi}\right) \rho\left(\boldsymbol{\theta}, \mathcal{T}, \boldsymbol{\psi}\right) \\ &\propto \rho\left(\mathbf{Y} \mid \boldsymbol{\theta}, \mathcal{T}\right) \rho\left(\boldsymbol{\theta}\right) \rho\left(\mathbf{S} \mid \mathcal{T}, \boldsymbol{\psi}\right) \rho\left(\mathcal{T}, \boldsymbol{\psi}\right) \end{split}$$

Assumption: \mathbf{Y} and \mathbf{S} independent conditionally on \mathcal{T} .

This talk: \mathcal{T} fixed.

BM: Gibbs with Conjugate Priors

Likelihood:

$$old Y | old R, old \mu \sim \mathcal{MN}(old 1_n old \mu^T, old V, old R)$$

Conjugate Priors:

$$\mathbf{R} \sim \mathcal{IW}(\mathbf{R}_0,
u)$$
 $oldsymbol{\mu} | \mathbf{R} \sim \mathcal{N}(oldsymbol{\mu}_0, \kappa_0^{-1} \mathbf{R})$

Gibbs:

$$\mathbf{R}|\mathbf{Y}, \boldsymbol{\mu} \sim \mathcal{IW}(\mathbf{R}_n, \nu_n)$$
 with $\mathbf{R}_n = f(\mathbf{Y}, \boldsymbol{\mu}, \mathbf{V})$

 \hookrightarrow Automatic sampling in the space of variance matrices.

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OU: No Gibbs

Likelihood:

$$Y|A,R,\mu \not\sim \mathcal{MN}$$

Conjugate Priors: ? \rightarrow No Gibbs

 \rightarrow Sample in constrained spaces (A, R)

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Conjugate Priors: ? \rightarrow No Gibbs

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Transformation:

$$f: egin{cases} \mathcal{C}_q o \mathbb{R}^q \ oldsymbol{ heta} \mapsto oldsymbol{
u} = f(oldsymbol{ heta}) & ; \qquad \pi_{oldsymbol{ heta}}(oldsymbol{ heta}) = \pi_{oldsymbol{
u}}(f(oldsymbol{ heta})) imes |J_f(oldsymbol{ heta})| \end{cases}$$

Metropolis - Hasting Iterate:

- Draw ν^* in $q(\nu \mid \nu^t)$.
- Set $\theta^{(t+1)} = \hat{\theta}^* = f^{-1}(\nu^*)$ with probability

$$r_t = \min \left\{ 1, \frac{p\left(\mathbf{Y} \mid \boldsymbol{\theta}^*\right)}{p\left(\mathbf{Y} \mid \boldsymbol{\theta}^t\right)} \frac{p\left(\boldsymbol{\theta}^*\right)}{p\left(\boldsymbol{\theta}^t\right)} \frac{q(\boldsymbol{\nu}^{(t)} | \boldsymbol{\nu}^*)}{q(\boldsymbol{\nu}^* | \boldsymbol{\nu}^{(t)})} \frac{|J_f(\boldsymbol{\theta}^{(t)})|}{|J_f(\boldsymbol{\theta}^*)|} \right\}.$$

Variance Matrix: LKJ

(Lewandowski, Kurowicka, and Joe, 2009)

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Decomposition Use correlation matrix **C**.

$$\mathbf{R} = \begin{pmatrix} \sigma_1 & & 0 \\ & \ddots & \\ 0 & & \sigma_p \end{pmatrix} \begin{pmatrix} 1 & & C_{kl} \\ & \ddots & \\ C_{kl} & & 1 \end{pmatrix} \begin{pmatrix} \sigma_1 & & 0 \\ & \ddots & \\ 0 & & \sigma_p \end{pmatrix}$$

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Cholesky
$$\mathbf{C} = \mathbf{W}^T \mathbf{W}$$
 with:

- $W_{1k}^2 + \cdots + W_{kk}^2 = 1$ (correlation)
- $W_{kk} > 0$ (identifiability)

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Attenuation Matrix:

Assumptions
$$\mathbf{A} = \mathbf{P} \mathbf{\Lambda} \mathbf{P}^{-1}$$
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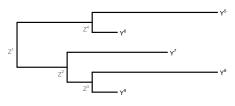
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BM, OU: Instance of a general Gaussian propagation model.

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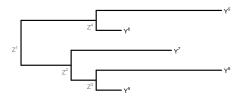
$$egin{aligned} oldsymbol{\mathsf{X}}^1 &\sim \mathcal{N}\left(oldsymbol{\mu}, oldsymbol{\Gamma}
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$$\textbf{X} = \begin{cases} \textbf{Z} & : \mathsf{latent} \ \mathsf{nodes} \\ \textbf{Y} & : \mathsf{tips} \end{cases}$$

root

tips and nodes

BM, OU: Instance of a general Gaussian propagation model.



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Likelihood:
$$\log p(\mathbf{Y})$$

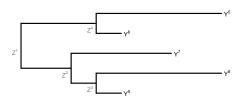
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tips and nodes

one post-order traversal

BM, OU: Instance of a general Gaussian propagation model.



$$\mathbf{X} = \begin{cases} \mathbf{Z} & \text{: latent nodes} \\ \mathbf{Y} & \text{: tips} \end{cases}$$

$$egin{align*} egin{align*} egin{align*}$$

root

tips and nodes

Likelihood: $\log p(\mathbf{Y})$

one post-order traversal

Gradient:
$$\frac{\partial}{\partial \phi_i} [\log p(\mathbf{Y})] = \mathbb{E} [\mathbf{F}(\mathbf{X}^j; \phi_j) \mid \mathbf{Y}]$$

one pre-order traversal

Implementation

(Suchard et al., 2018)



- MCMC for tree estimation
- Comprehensive set of tools:
 - Factor model
 - Marginal Likelihood
 - ..
- Developed in Java since 2002.
- This is BEAST 1.10.

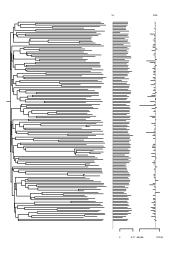
Don't ask about BEAST 2.

What's new:

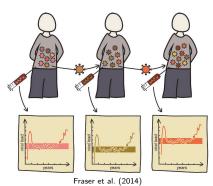
- Flexible OU models
- Efficient sampling of variance
- Efficient HMC (in progress)

HIV virulence heritability

(Alizon et al., 2010; Vrancken et al., 2015)

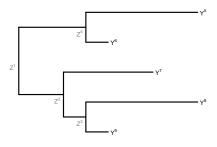


CD4: CD4+ T cells decline rate VL: Set point viral load



Questions: Is virulence "heritable"? What model of trait evolution?

Heritability

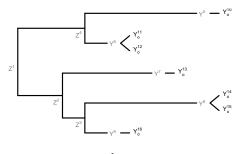


$$\begin{aligned} \mathbf{X}^{1} &\sim \mathcal{N}\left(\boldsymbol{\mu}, \boldsymbol{\Gamma}\right) \\ \mathbf{X}^{j} \;\middle|\; \mathbf{X}^{\mathsf{pa}(j)} &\sim \mathcal{N}\left(\mathbf{q}_{i} \mathbf{X}_{\mathsf{pa}(i)} + \mathbf{r}_{i}, \; \boldsymbol{\Sigma}_{i}\right) \end{aligned}$$

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root tips and nodes

Heritability



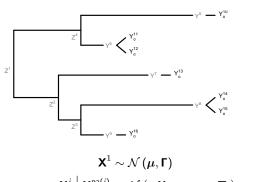
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root

latent tips and nodes observations

Heritability



$$\mathbf{X} = \begin{cases} \mathbf{Z} & \text{: latent nodes} \\ \mathbf{Y} & \text{: latent tips} \\ \mathbf{Y}_o & \text{: observed traits} \end{cases}$$

$$\mathbf{X}^{j} \mid \mathbf{X}^{\mathsf{pa}(j)} \sim \mathcal{N}\left(\mathbf{q}_{i}\mathbf{X}_{\mathsf{pa}(i)} + \mathbf{r}_{i}, \; \mathbf{\Sigma}_{i}\right)$$
 $\mathbf{Y}_{o}^{i} \mid \mathbf{Y}^{\mathsf{pa}(i)} \sim \mathcal{N}\left(\mathbf{Y}^{\mathsf{pa}(i)}, \mathbf{S}\right)$

latent tips and nodes

$$\mathbf{Y}_o^i \mid \mathbf{Y}^{\mathsf{pa}(i)} \sim \mathcal{N}\left(\mathbf{Y}^{\mathsf{pa}(i)}, \mathbf{S}\right)$$

observations

root

"Heritability":
$$h_k^2 = \frac{V(\mathbf{Y}_{\cdot k})}{V(\mathbf{Y}_{o \cdot k})} \approx \frac{\sigma_k^2 \tilde{t}}{\sigma_k^2 \tilde{t} + s_k}$$

Results

Model:

VL OU (stabilizing selection)

CD4 BM (no selection)

The two traits can be correlated.

Heritability:

VL
$$h^2 = 17\%$$
 [0.007, 82.5]% (95% CI)
CD4 $h^2 = 0.02\%$ [0.0024, 0.16]%

 \hookrightarrow "Consistent" with previous estimates.

Conclusion and Perspectives

A general framework for trait evolution with dated tips.

Main Features:

- Flexible models and implementation
- Efficient algorithms
- Applicable to virology

Perspectives:

- Develop HMC
- Identifiability
- Other questions: geographical spread, comparative studies, ...

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Thank you for listening









Kasteel van Arenberg





Appendices