Change-point Detection on a Tree to Study Evolutionary Adaptation from Present-day Species

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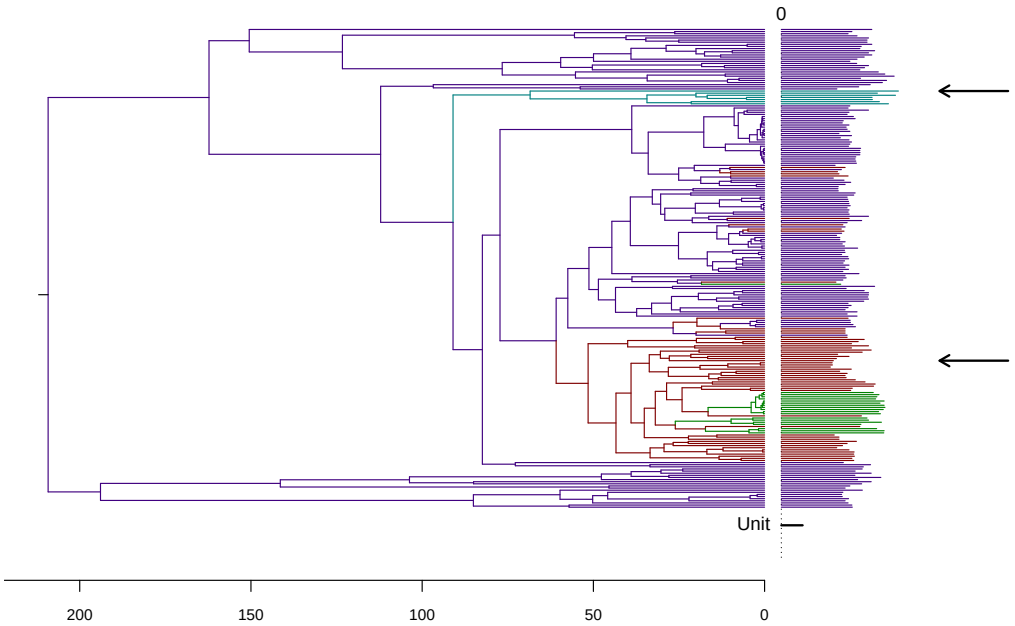
23 June 2016

Introduction



Dermochelys Coriacea

Homopus Areolatus

Turtles phylogenetic tree with habitats.
(Jaffe et al., 2011).

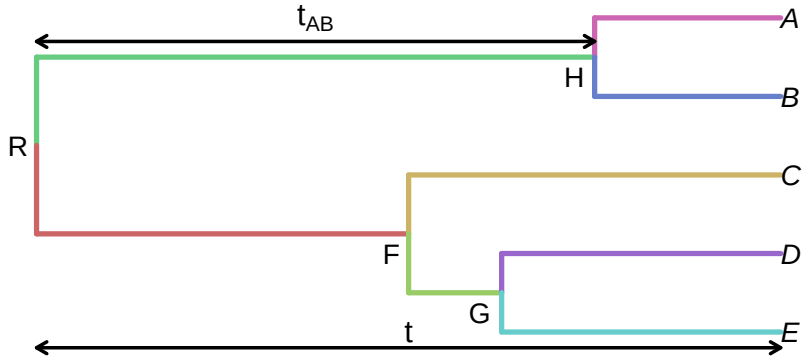
- How can we explain the diversity, while accounting for the phylogenetic correlations ?
- Modelling: a shifted stochastic process on the phylogeny.

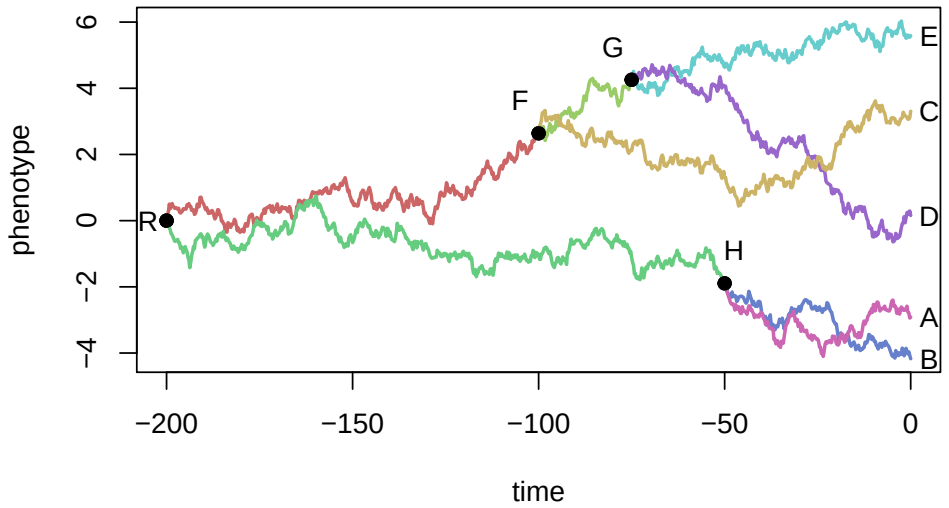
Outline

1. Stochastic Processes on Trees
2. Identifiability Problems and Counting Issues
3. Statistical Inference
4. Turtles Data Set

Stochastic Process on a Tree

(Felsenstein, 1985)





The tree is known.

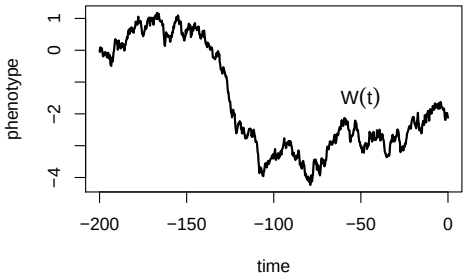
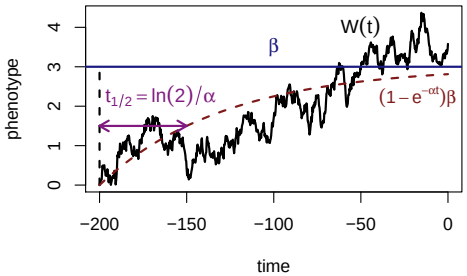
Only *tip* trait values are observed.

Brownian Motion:

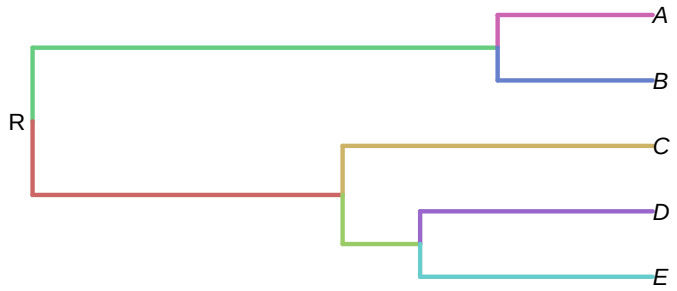
$$\mathbb{V}\mathrm{ar}\left[A \mid R\right] = \sigma^2 t$$

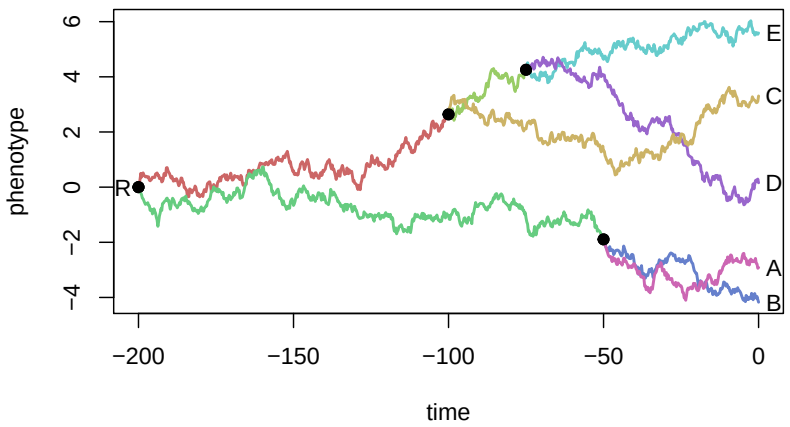
$$\mathbb{C}\mathrm{ov}\left[A; B \mid R\right] = \sigma^2 t_{AB}$$

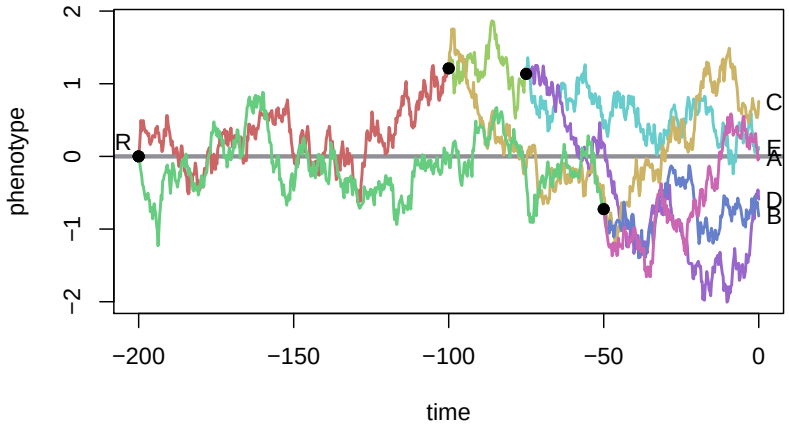
Brownian Motion vs Ornstein-Uhlenbeck

	Equation	Stationary State	Variance
	$dW(t) = \sigma dB(t)$	None.	$\sigma_{ij} = \sigma^2 t_{ij}$
	$dW(t) = \sigma dB(t)$ $+\alpha[\beta(t) - W(t)]dt$	$\begin{cases} \mu = \beta_0 \\ \gamma^2 = \dfrac{\sigma^2}{2\alpha} \end{cases}$	$\sigma_{ij} = \gamma^2 e^{-\alpha(t_i + t_j)} \times (e^{2\alpha t_{ij}} - 1)$

Shifts







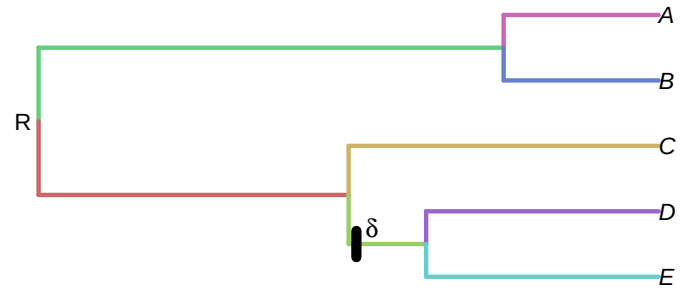
BM Shifts in the **mean**:

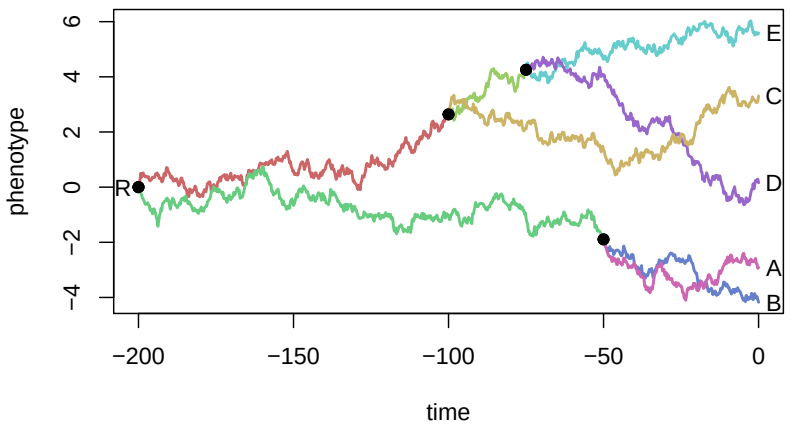
$$m_{\text{child}} = m_{\text{parent}} + \delta$$

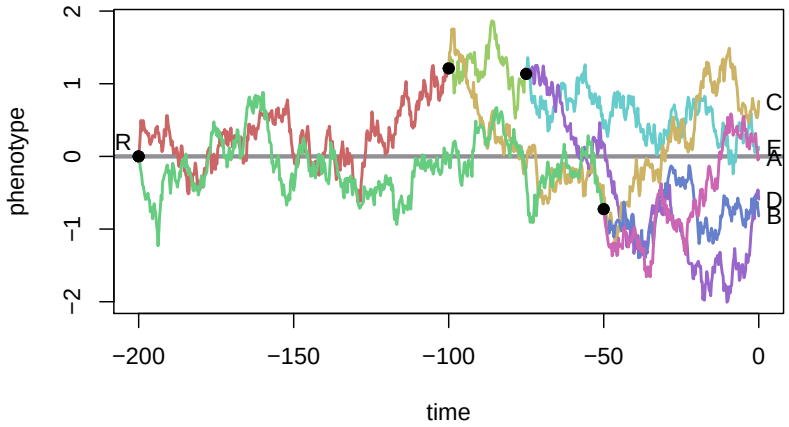
OU Shifts in the **optimal value**:

$$\beta_{\text{child}} = \beta_{\text{parent}} + \delta$$

Shifts







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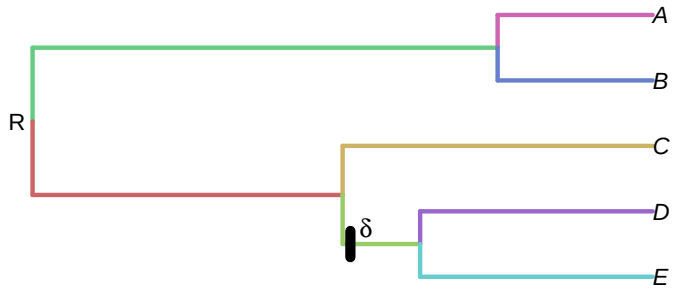

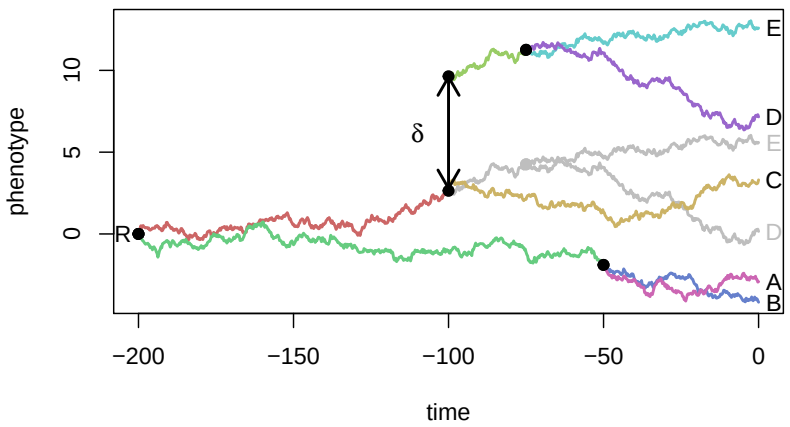

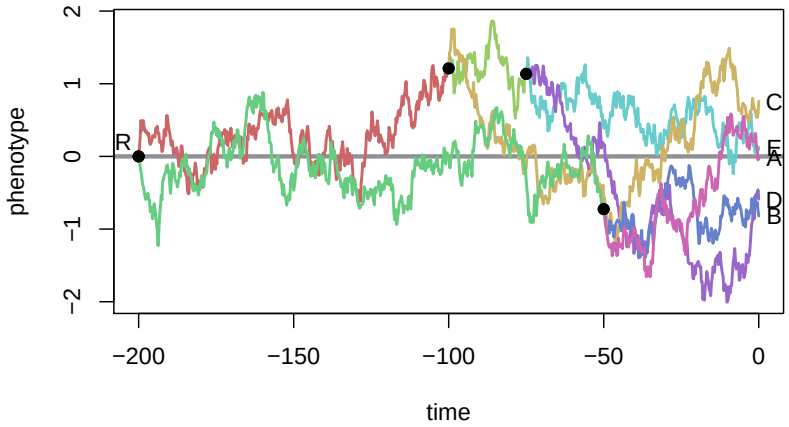

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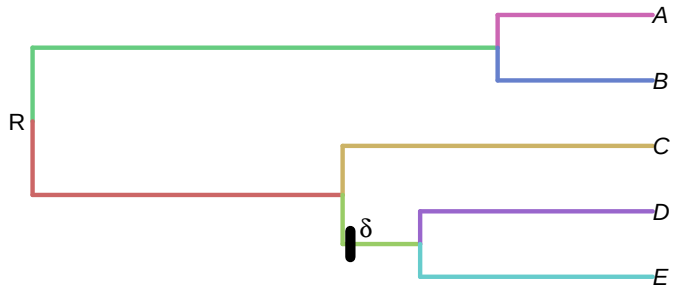

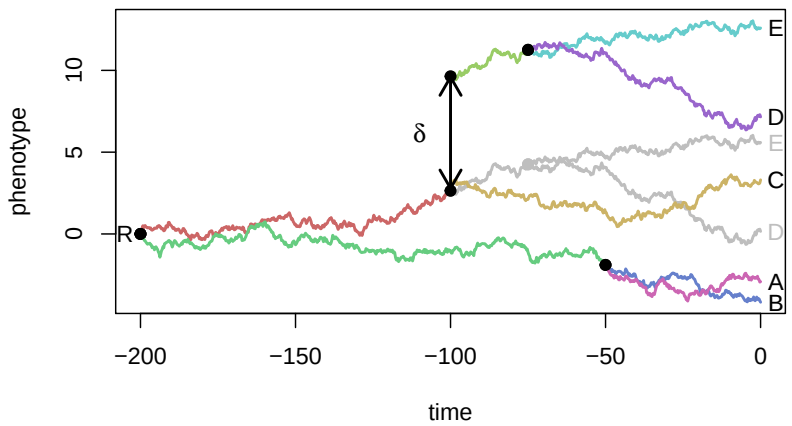

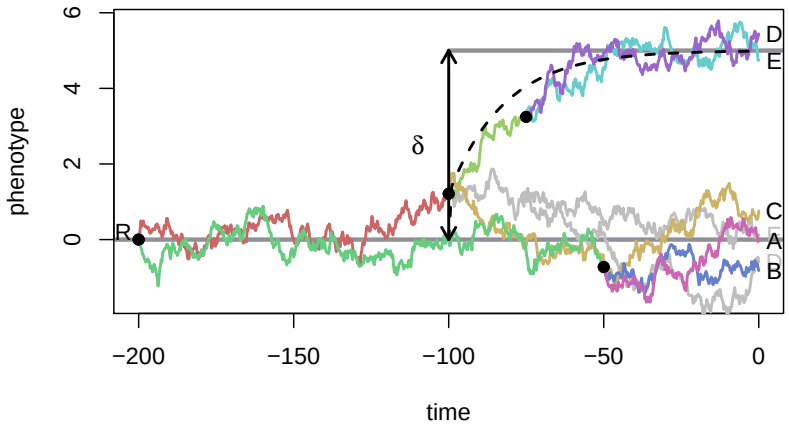

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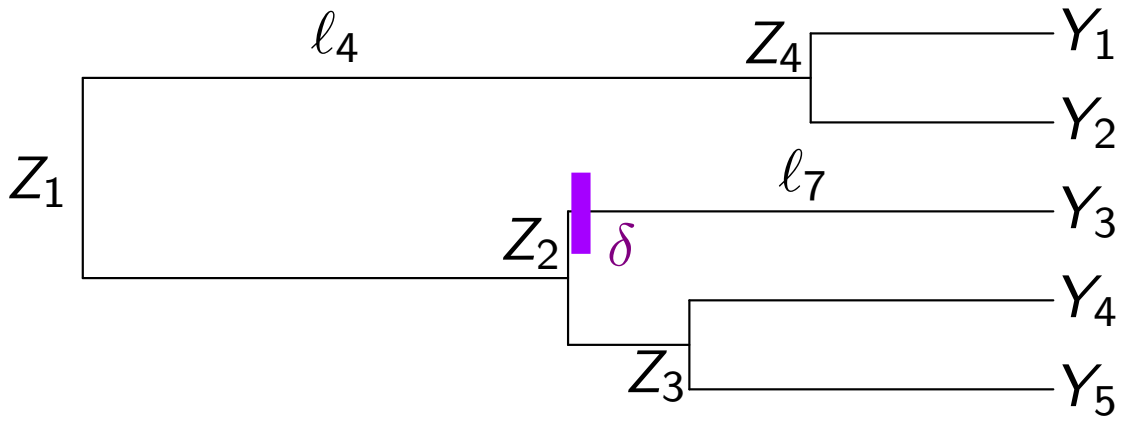
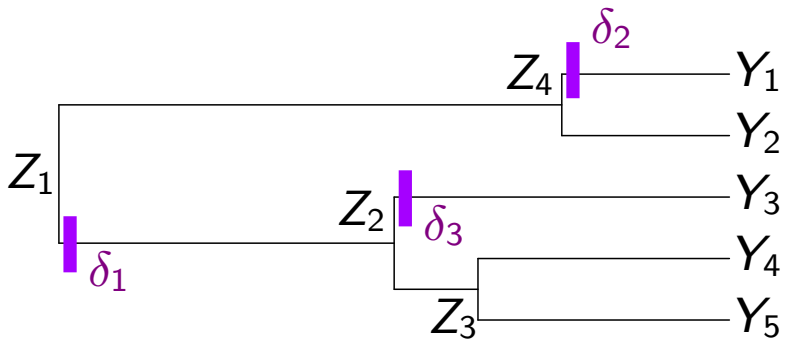
Incomplete Data Model



$$\textit{BM}\quad Z_4|Z_1 \sim \mathcal{N}\Big(Z_1 \quad\quad, \sigma^2\ell_4\Big)$$
$$Y_3|Z_2 \sim \mathcal{N}\Big(Z_2+\delta, \sigma^2\ell_7\Big)$$

$$\textit{OU}\quad Y_3|Z_2 \sim \mathcal{N}\Big(Z_2 e^{-\alpha\ell_7} + (1-e^{-\alpha\ell_7})(\beta_{Z_2}+\delta), \frac{\sigma^2}{2\alpha}(1-e^{-2\alpha\ell_7})\Big)$$

Linear Regression Model

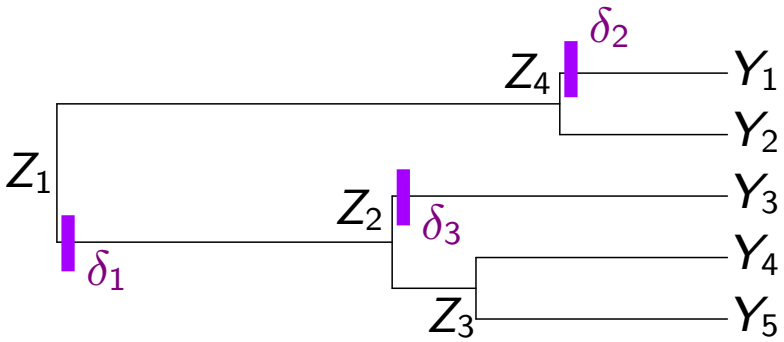


$$\Delta = \begin{pmatrix} \mu \\ \delta_1 \\ 0 \\ 0 \\ \delta_2 \\ 0 \\ \delta_3 \\ 0 \\ 0 \end{pmatrix} \qquad T\Delta = \begin{pmatrix} \mu + \delta_2 \\ \mu \\ \mu + \delta_1 + \delta_3 \\ \mu + \delta_1 \\ \mu + \delta_1 \end{pmatrix}$$

$$T = \begin{array}{c} \\ Y_1 \\ Y_2 \\ Y_3 \\ Y_4 \\ Y_5 \end{array} \begin{array}{c} \begin{array}{ccccccccc} Z_1 & Z_2 & Z_3 & Z_4 & Y_1 & Y_2 & Y_3 & Y_4 & Y_5 \end{array} \\ \begin{pmatrix} 1 & 0 & 0 & 1 & 1 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 1 & 0 & 1 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\ 1 & 1 & 1 & 0 & 0 & 0 & 0 & 1 & 0 \\ 1 & 1 & 1 & 0 & 0 & 0 & 0 & 0 & 1 \end{pmatrix} \end{array}$$

$$BM: \quad Y = T\Delta^{BM} + E^{BM}$$

Linear Regression Model



$$\Delta = \begin{pmatrix} \lambda \\ \delta_1 \\ 0 \\ 0 \\ \delta_2 \\ 0 \\ \delta_3 \\ 0 \\ 0 \end{pmatrix} \qquad TW(\alpha)\Delta = \begin{pmatrix} \lambda + w_5\delta_2 \\ \lambda \\ \lambda + w_2\delta_1 + w_7\delta_3 \\ \lambda + w_2\delta_1 \\ \lambda + w_2\delta_1 \end{pmatrix}$$

$$W(\alpha) = \text{Diag}(1 - e^{-\alpha(h - t_{\text{pa}(i)})}, 1 \leq i \leq m+n)$$

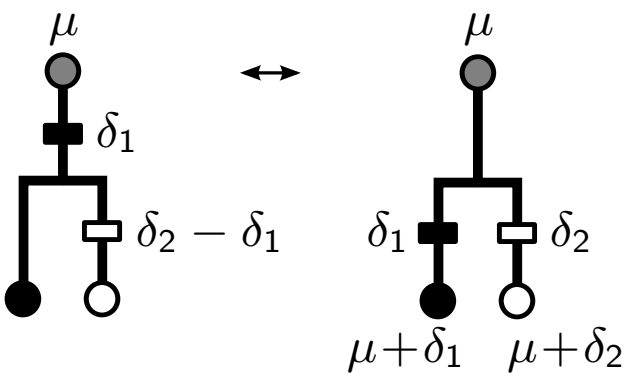
$$\lambda = \mu e^{-\alpha h} + \beta_0(1 - e^{-\alpha h})$$

$$BM: \quad Y = T\Delta^{BM} + E^{BM}$$

$$OU: \quad Y = TW(\alpha)\Delta^{OU} + E^{OU}$$

Equivalencies

- Number of shifts K fixed, several equivalent solutions.



- Problem of over-parametrization: parsimonious configurations.

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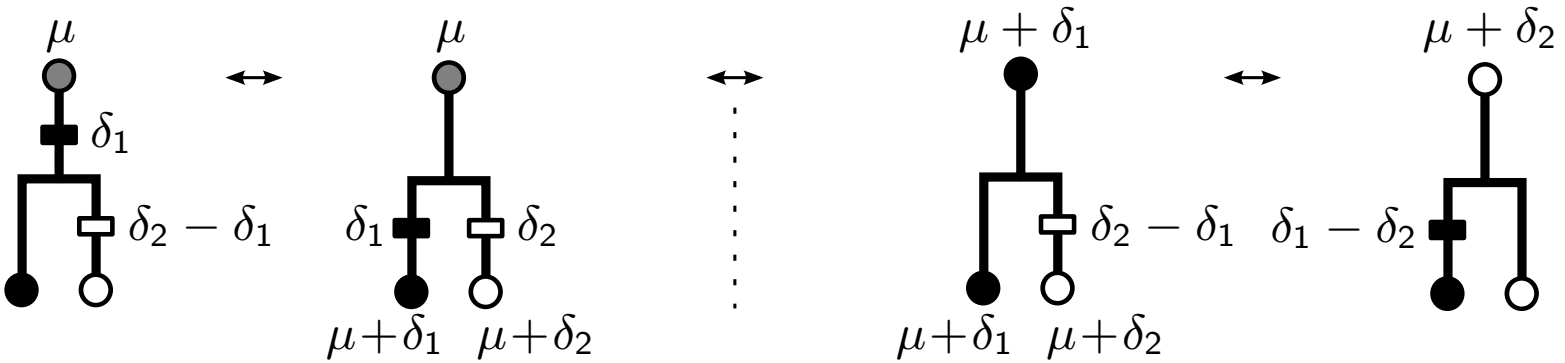

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Parsimonious Solution : Definition

Definition (Parsimonious Allocation)

A coloring of the tips being given, a *parsimonious* allocation of the shifts is such that it has a minimum number of shifts.

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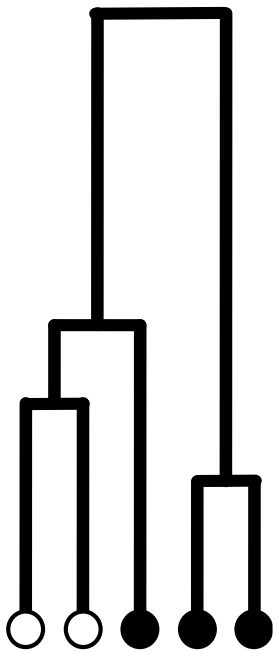

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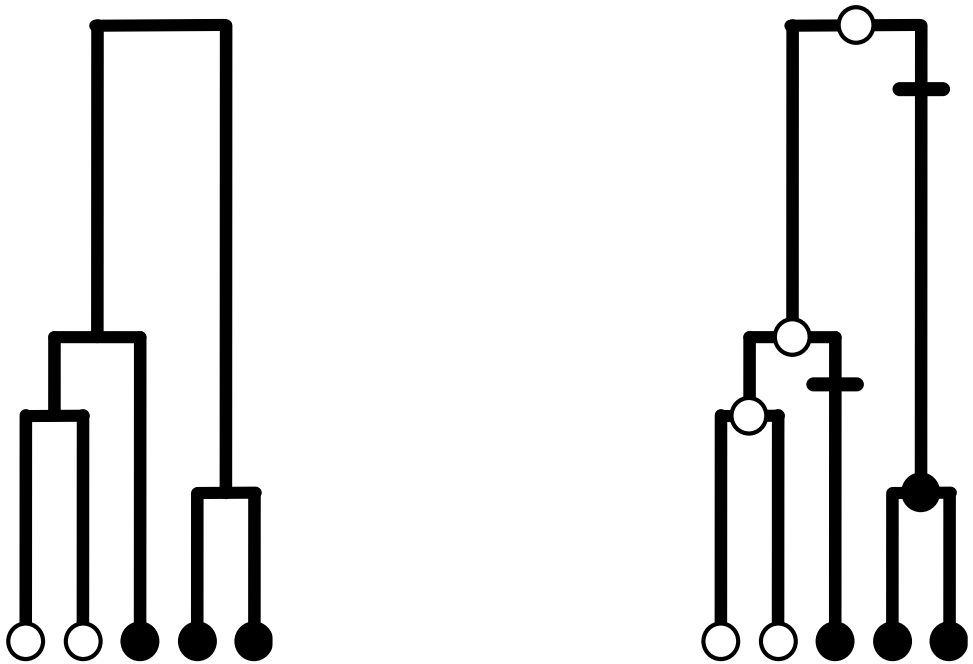

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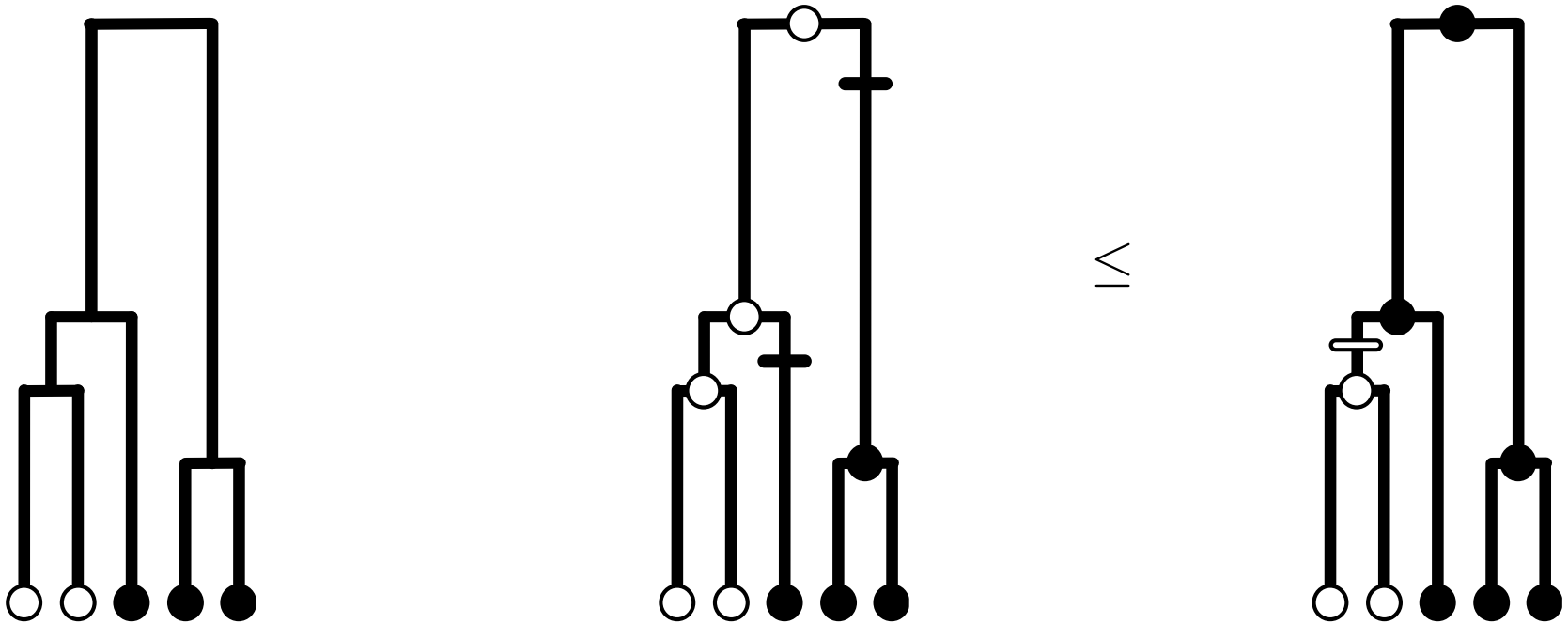

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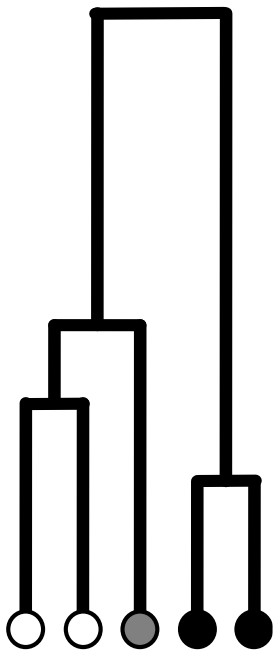

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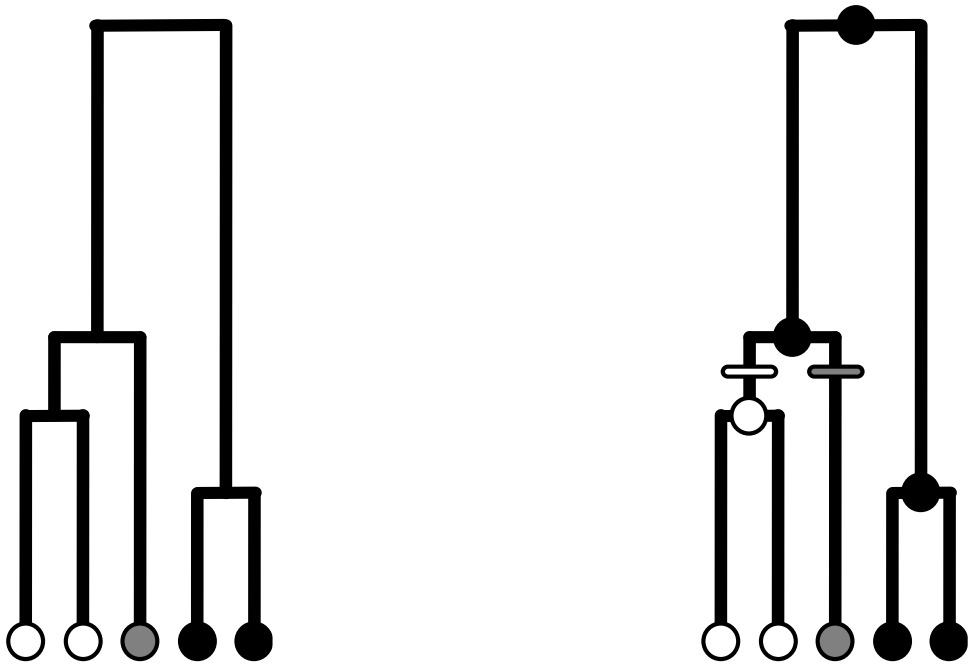

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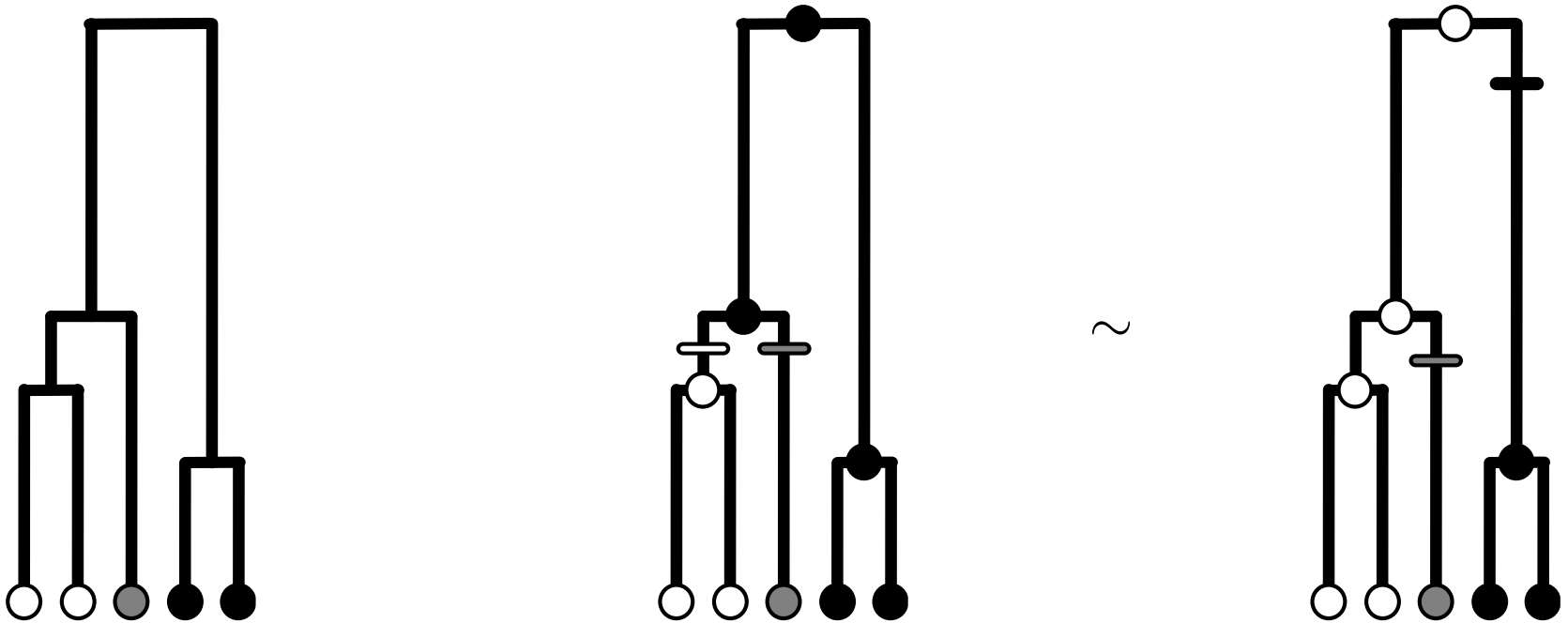


Equivalent Parsimonious Allocations

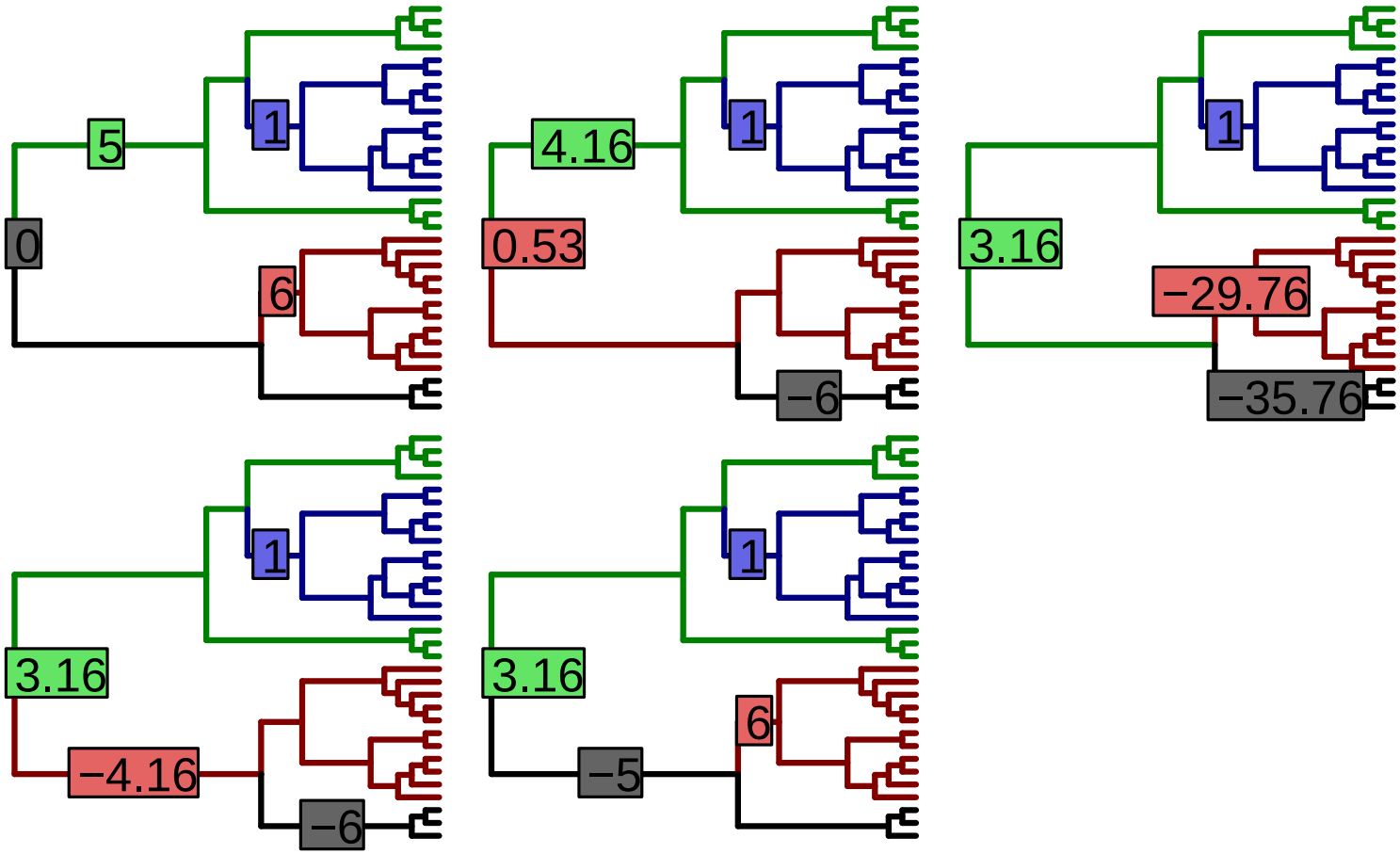
Definition (Equivalency)

Two allocations are said to be *equivalent* (noted $\sim$) if they are both parsimonious and give the same colors at the tips.

Find one solution Several existing Dynamic Programming algorithms (Fitch, Sankoff, see Felsenstein, 2004).

Enumerate all solutions New recursive algorithm, adapted from previous ones (and implemented in R).

Equivalent Parsimonious Solutions for an OU Model.



Equivalent allocations and values of the shifts - OU.

Collection of Models

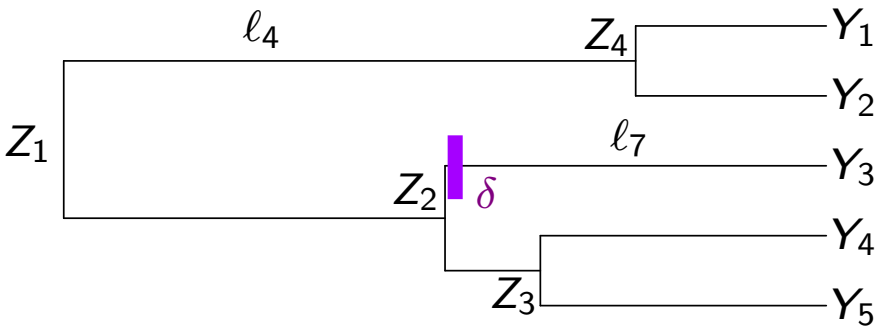
New Problem Number of Equivalence Classes: $\left|\mathcal{S}_K^{PI}\right|$?

- $\left|\mathcal{S}_K^{PI}\right| \leq \binom{m+n-1}{K} = \binom{\#\text{ of edges}}{\#\text{ of shifts}}$
- A recursive algorithm to compute $\left|\mathcal{S}_K^{PI}\right|$ (implemented in R).

$\mapsto$ Generally dependent on the topology of the tree.

- Binary tree: $\left|\mathcal{S}_K^{PI}\right| = \binom{2n-2-K}{K} = \binom{\#\text{ of edges}-\#\text{ of shifts}}{\#\text{ of shifts}}$

EM Algorithm: number of shifts K fixed



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$$Y_3|Z_2 \sim \mathcal{N}\left(Z_2 + \delta, \sigma^2 \ell_7\right)$$

$$p_\theta(Z, Y) = p_\theta(Z_1) \prod_{1<j\leq m} p_\theta(Z_j|Z_{\text{parent}(j)}) \prod_{1\leq i\leq n} p_\theta(Y_i|Z_{\text{parent}(i)})$$

EM Algorithm $\log p_\theta(Y) = \mathbb{E}_\theta[\log p_\theta(Z, Y) \mid Y] - \mathbb{E}_\theta[\log p_\theta(Z) \mid Y]$

E step Given θ^h, compute $p_{\theta^h}(Z \mid Y)$

M step $\theta^{h+1} = \text{argmax}_\theta \, \mathbb{E}_{\theta^h}[\log p_\theta(Z, Y) \mid Y]$

Initialization Lasso

Model Selection on K

Assumption α fixed

$$Y = TW(\alpha)\Delta + \gamma E \qquad , \qquad E \sim \mathcal{N}(0, V(\alpha))$$

Models

$\eta \in \bigcup_{K=0}^{p-1} \mathcal{S}_K^{PI}$: Identifiable parsimonious allocations of shifts

EM Estimators

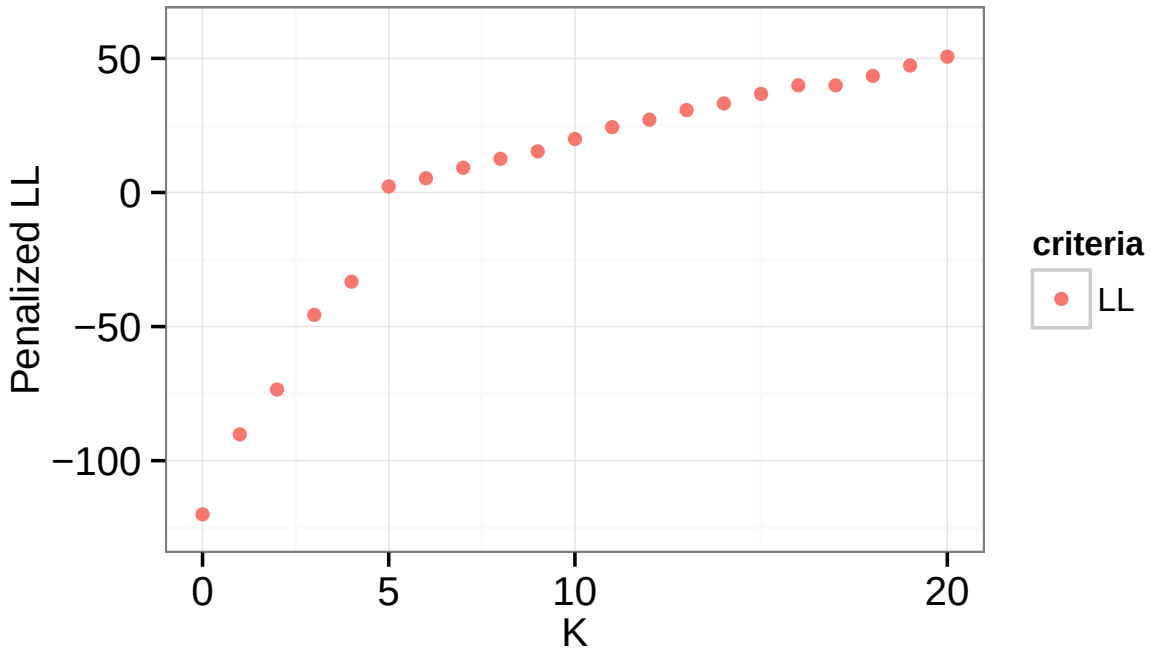
$$\hat{Y}_K = \underset{\eta \in \mathcal{S}_K^{PI}}{\operatorname{argmin}} \left\| Y - \hat{Y}_\eta \right\|_V^2$$

Oracle

$$\inf_{\eta \in \bigcup_{K=0}^{p-1} \mathcal{S}_K^{PI}} \left\| \mathbb{E}[Y] - Y_\eta^* \right\|_V^2$$

Model Selection: Penalized Likelihood

$$\text{Idea} \quad \hat{K} = \underset{0 \leq K \leq p-1}{\operatorname{argmax}} \left\{ \frac{n}{2} \log \left(\frac{1}{n} \left\| Y - \hat{Y}_K \right\|_V^2 \right) - \frac{1}{2} \operatorname{pen}'(K) \right\}$$



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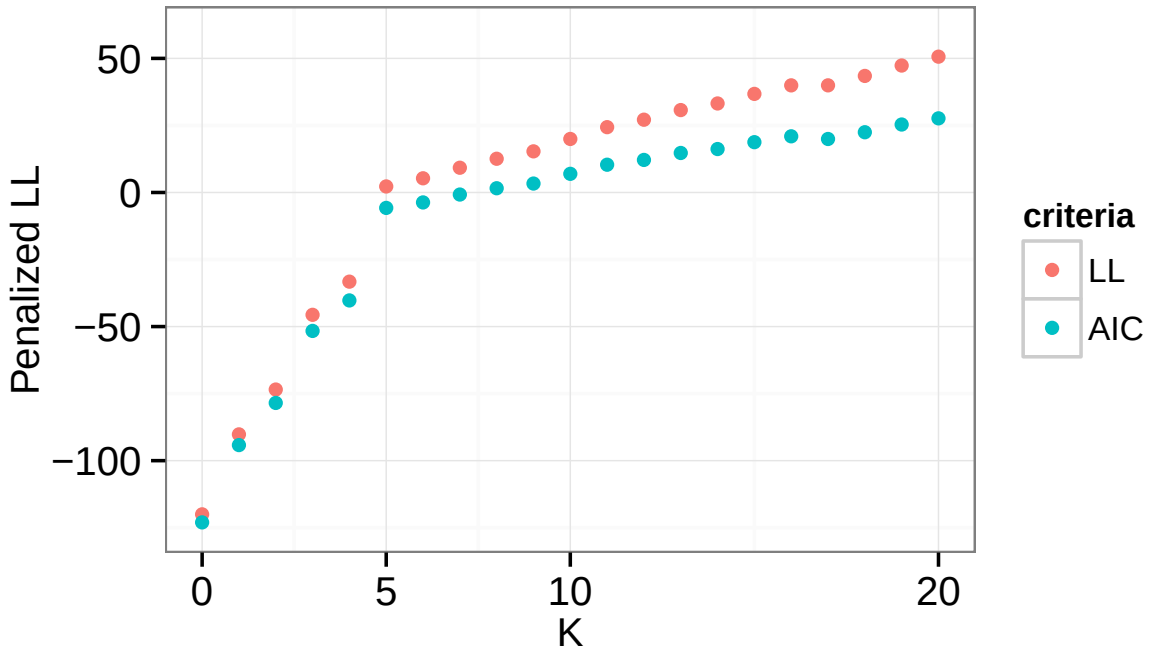

Penalties:

AIC $K + 3$

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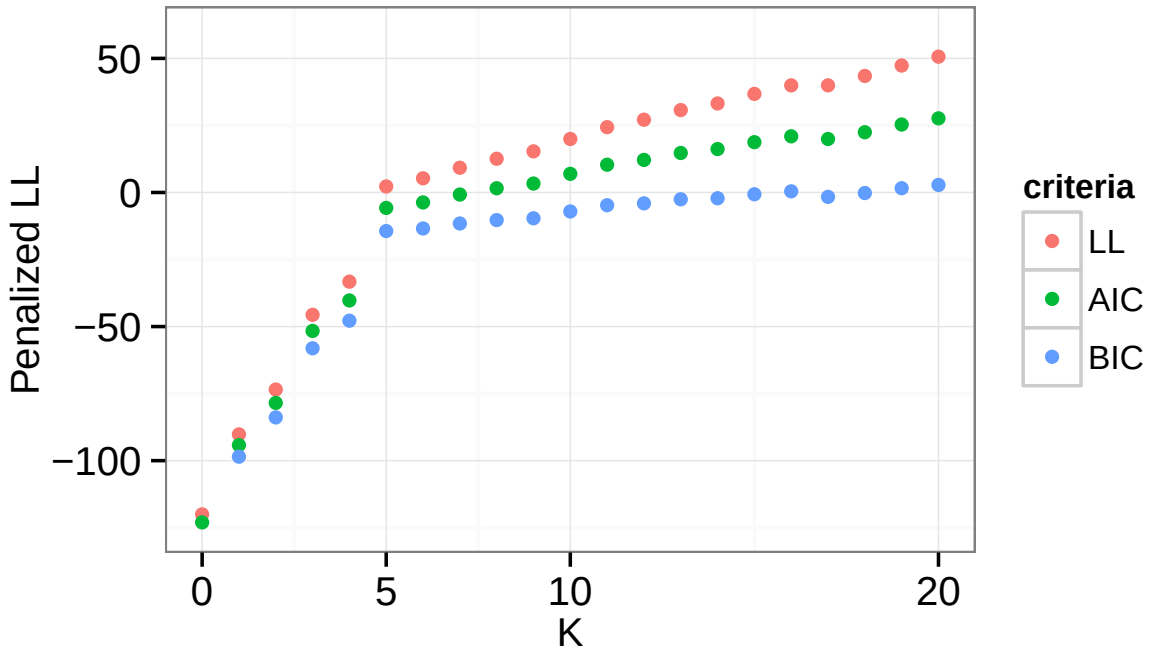

Penalties:

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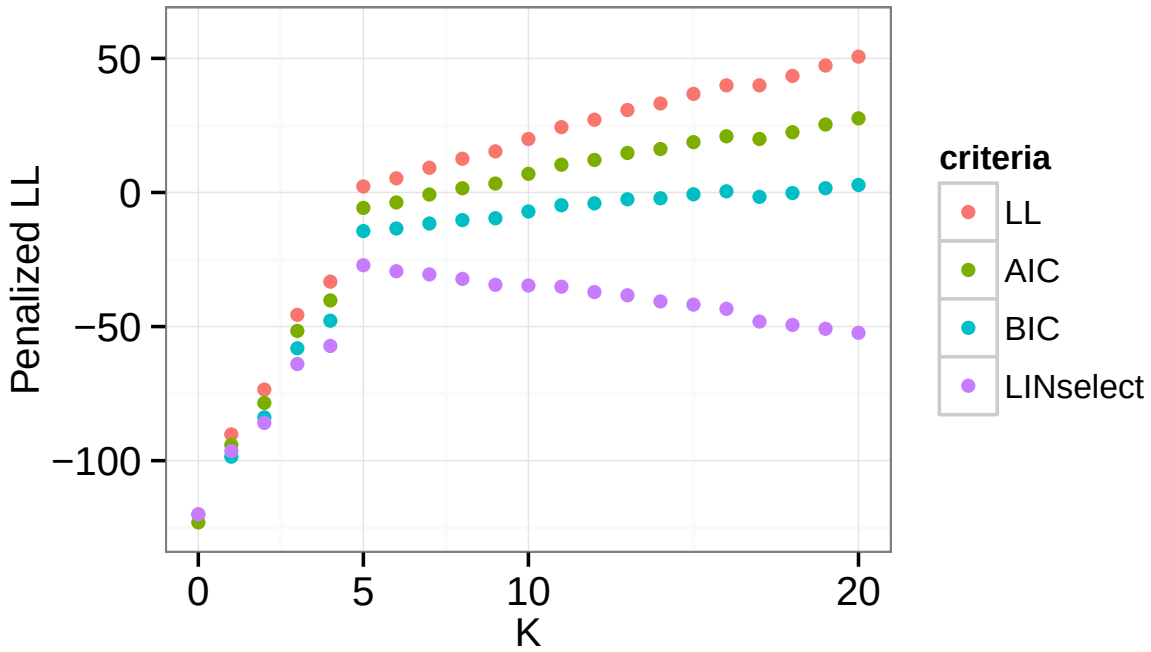
BIC $\frac{1}{2}(K + 3)\log(n)$

Model Selection: Penalized Likelihood

Idea $$\hat{K} = \underset{0 \leq K \leq p-1}{\operatorname{argmax}} \left\{ \frac{n}{2} \log \left(\frac{1}{n} \left\| Y - \hat{Y}_K \right\|_V^2 \right) - \frac{1}{2} \operatorname{pen}'(K) \right\}$$



Penalties:

AIC $K + 3$

BIC $\frac{1}{2}(K+3)\log(n)$

LINselect $\operatorname{pen}(n, K, |\mathcal{S}_K^{PI}|)$

Proposition: LINselect Penalty

Proposition (Form of the Penalty and guarantees (α known))

Under our setting: $Y = TW(\alpha)\Delta + \gamma E$ *with* $E \sim \mathcal{N}(0, V)$, *define the penalty:*

$$\text{pen}(K) = A\frac{n-K-1}{n-K-2}\,\text{EDkhi}\left[K+2, n-K-2, exp\left(-\log\left|\mathcal{S}_K^{Pl}\right| - 2\log(K+2)\right)\right]$$

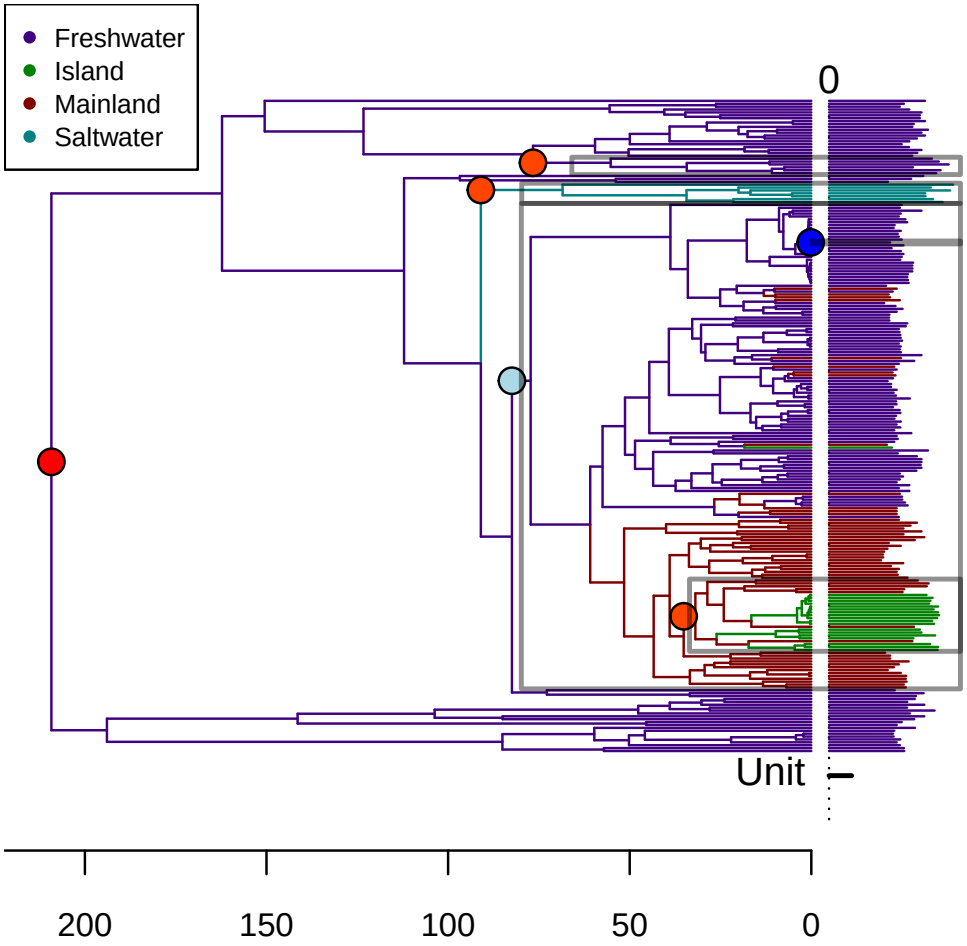
If $\kappa < 1$, *and* $p \leq \min\left(\frac{\kappa n}{2+\log(2)+\log(n)}, n-7\right)$, *we get:*

$$\mathbb{E}\left[\frac{\left\|\mathbb{E}[Y] - \hat{Y}_{\hat{K}}\right\|_V^2}{\gamma^2}\right] \leq C(A,\kappa)\inf_{\eta\in\mathcal{M}}\left\{\frac{\left\|\mathbb{E}[Y] - Y_\eta^*\right\|_V^2}{\gamma^2} + (K_\eta + 2)(3+\log(n))\right\}$$

with $C(A,\kappa)$ *a constant depending on* A *and* κ *only.*

Based on Baraud et al. (2009) +

Turtles Dataset

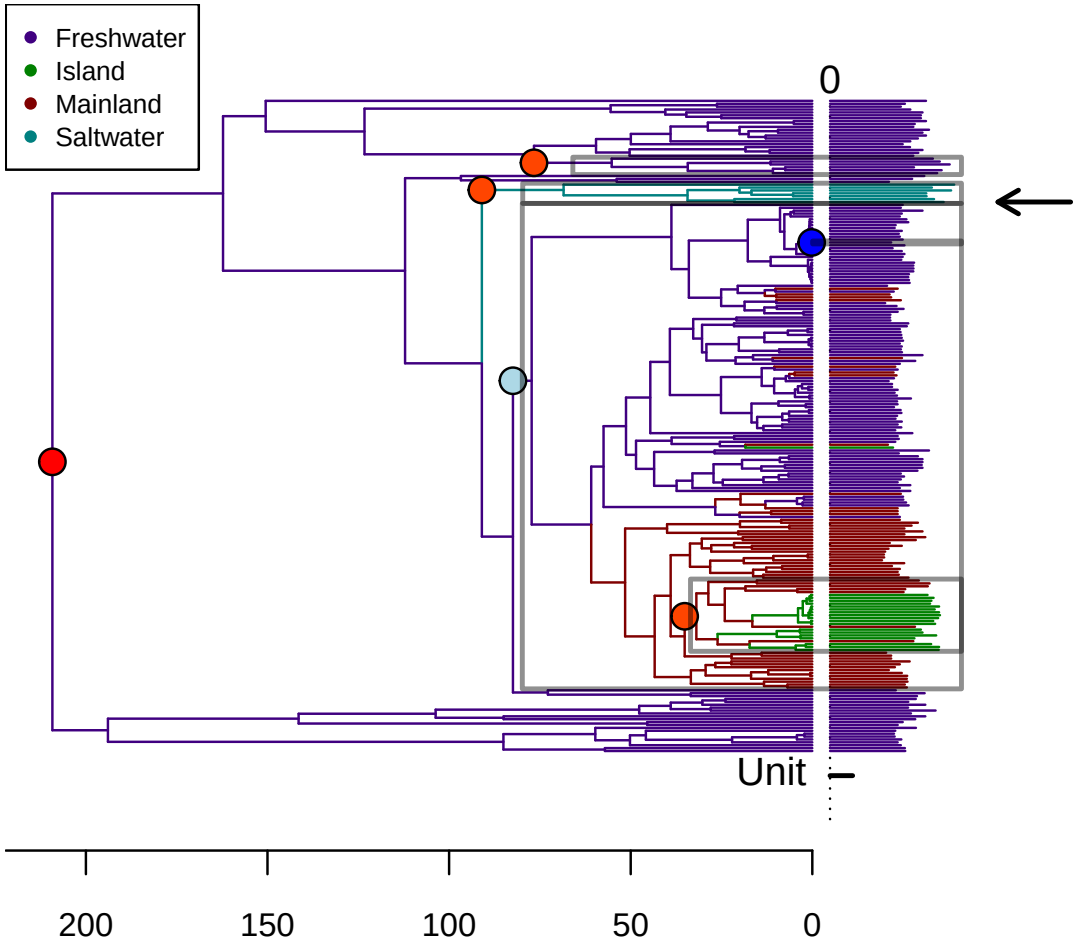


Colors: habitats.
Boxes: selected EM regimes.

	Habitat	EM
No. of shifts	16	5
No. of regimes	4	6
lnL	-133.86	-97.59
$\ln 2/\alpha$ (%)	7.44	5.43
$\sigma^2/2\alpha$	0.33	0.22
CPU t (min)	65.25	134.49

(Jaffe et al., 2011)

Turtles Dataset



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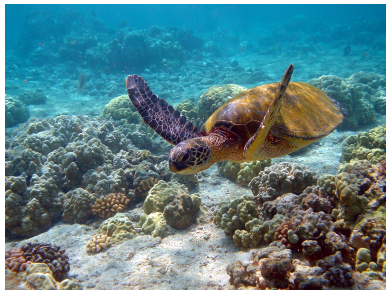
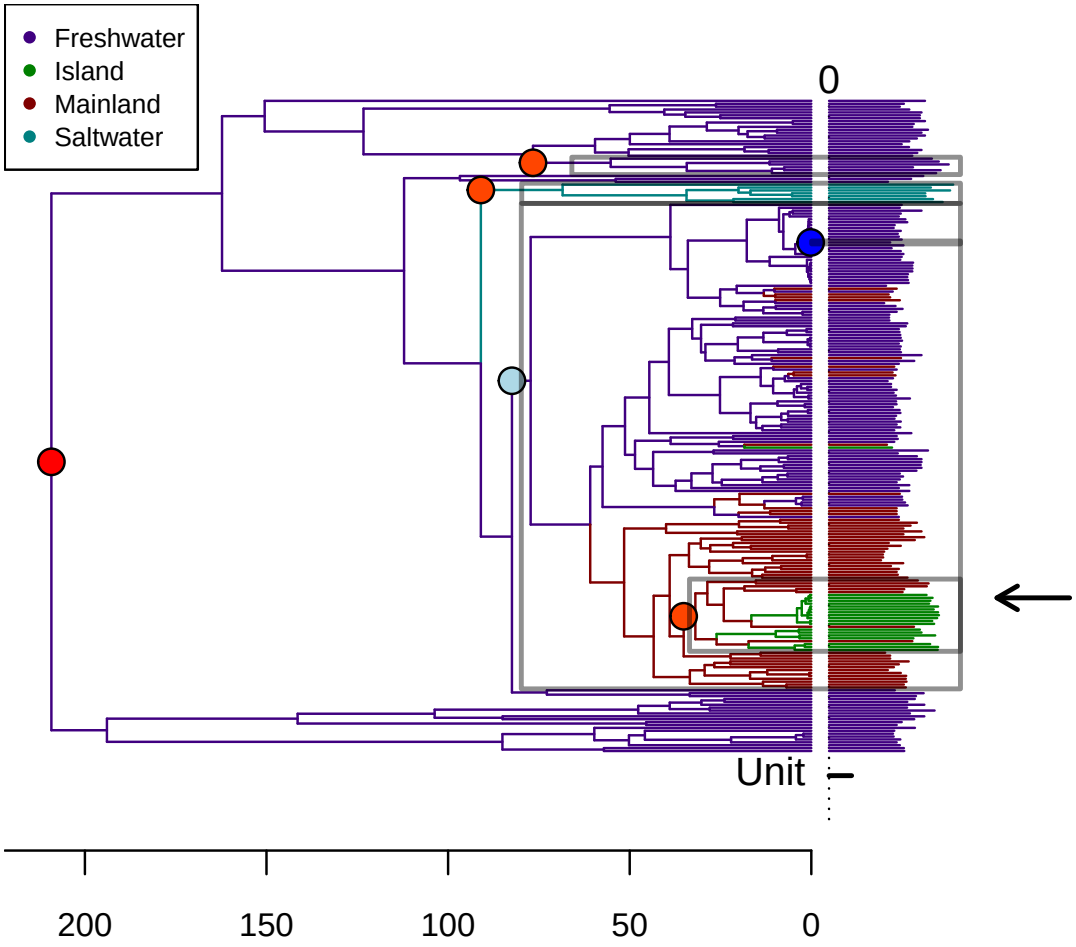


Chelonia mydas

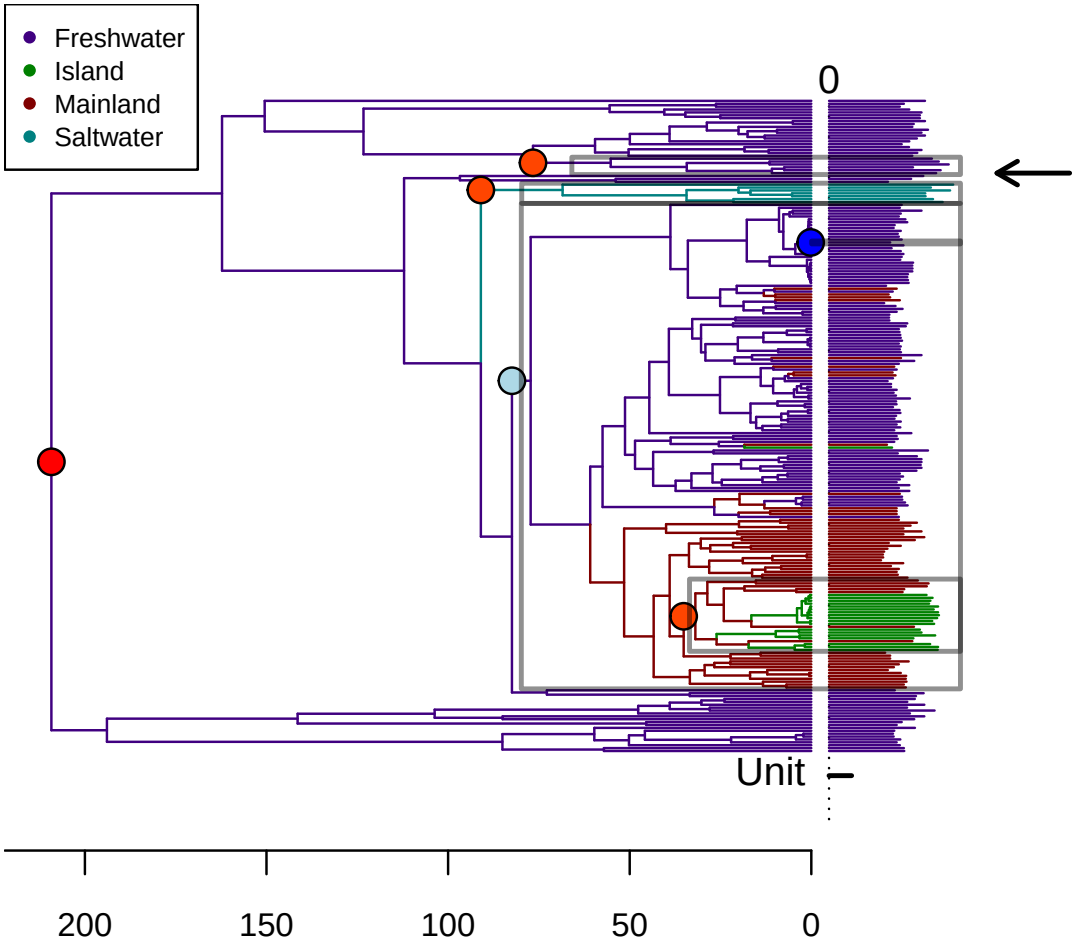
Turtles Dataset



Colors: habitats.
Boxes: selected EM regimes.

Geochelone nigra abingdoni

Turtles Dataset



Colors: habitats.
Boxes: selected EM regimes.

Chitra indica

Conclusion and Perspectives

A general inference framework for trait evolution models.

Conclusions

- Identifiability can be assessed.
- An EM can be written to maximize likelihood.
- Model selection for a non-iid framework.

R codes Available on GitHub:

`https://github.com/pbastide/PhylogeneticEM`

Perspectives

- Multivariate traits.
- Deal with uncertainty (tree, data).
- Phylogenetic networks.

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- "Florida Box Turtle Digon3a", "Jonathan Zander (Digon3)" derivative work: Materialscientist

Thank you for listening

pbastide.github.io

Appendices

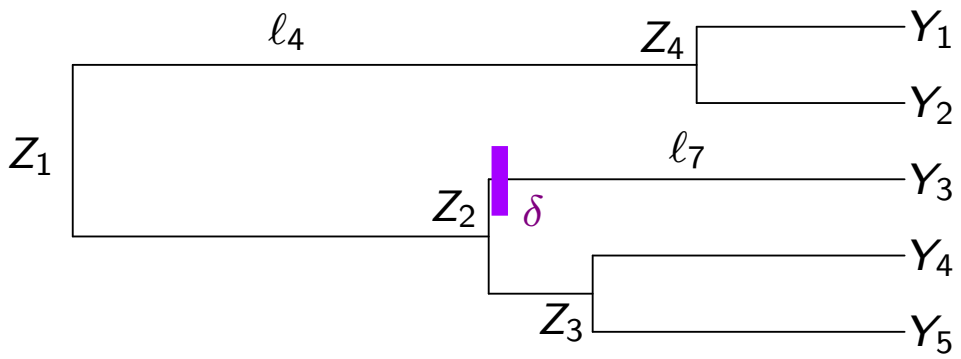
5 Inference
- EM
- Model Selection

6 Identifiability Issues
- OU $\Longleftrightarrow$ BM
- Cardinal of Equivalence Classes
- Number of Tree Compatible Clustering

7 Simulations Results

8 Multivariate
- Models
- Inference

E step



Compute the following quantities:

$$\mathbb{E}^{(h)}[Z_j \mid Y],\ \mathbb{V}\mathrm{ar}^{(h)}[Z_j \mid Y],\ \mathbb{C}\mathrm{ov}^{(h)}\left[Z_j, Z_{\mathrm{parent}(j)} \mid Y\right]$$

- Using Gaussian properties. Need to invert matrices: complexity in $O(n^3)$.
- Using Gaussian properties **and** the tree structure: "Upward-Downward" algorithm. Complexity in $O(n)$.

M Step

Maximize:

$$\mathbb{E}\left[\log p_\theta(X) \mid Y\right] = -\sum_{j=2}^{m+n} C_j(\alpha, \text{shifts}) + \mathcal{F}^{(h)}\left(\mu, \gamma^2, \sigma^2, \alpha\right)$$

- μ, γ^2, σ^2: simple maximization
- Discrete location of K shifts
  - $\mapsto$ Exact and fast for the BM
- α: numerical maximization and/or on a grid
  - $\mapsto$ Generalized EM

Initialization

Shifts : Lasso regression.

$$\hat{\Delta} = \underset{\Delta}{\operatorname{argmin}} \left\{ \| Y - TW(\alpha)\Delta \|^2_{\Sigma^{-1}_{YY}} + \lambda \| \Delta_{-1} \|_1 \right\}$$

- Initialize $\Sigma_{YY}(\alpha)$, then estimate Δ with a Gauss Lasso procedure, using a Cholesky decomposition.

+

- λ chosen to get K shifts.

The selection strength α : Initialization using couples of tips.

back

Cholesky Decomposition

The problem is:

$$\hat{\Delta} = \underset{\Delta}{\operatorname{argmin}} \left\{ \| Y - R\Delta \|^2_{\Sigma_{YY}} + \lambda \left| \Delta_{-1} \right|_1 \right\}$$

Cholesky decomposition of Σ_{YY}:

$$\Sigma_{YY} = LL^T \ , \ L \text{ a lower triangular matrix}$$

Then:

$$\| Y - R\Delta \|^2_{\Sigma_{YY}} = \left\| L^{-1}Y - L^{-1}R\Delta \right\|^2$$

And if $Y' = L^{-1}Y$ and $R' = L^{-1}R$, the problem becomes:

$$\hat{\Delta} = \underset{\Delta}{\operatorname{argmin}} \left\{ \left\| Y' - R'\Delta \right\|^2 + \lambda \left| \Delta_{-1} \right|_1 \right\}$$

Gauss Lasso

Let $\hat{m}_\lambda$ be the set of selected variables (including the root). Then:

$$\hat{\Delta}^{\text{Gauss}} = \Pi_{\hat{F}_\lambda}(Y') \text{ with } \hat{F}_\lambda = \text{Span}\{R'_j : j \in \hat{m}_\lambda\}$$

back

Goal and Notations

Data A process on a tree with the following structure:

$$\forall j > 1, \quad X_j | X_{\mathsf{pa}(j)} \sim \mathcal{N}\left(m_j(X_{\mathsf{pa}(j)}) = q_j X_{\mathsf{pa}(j)} + r_j, \sigma_j^2\right)$$

BM:
$$\begin{cases} q_j = 1 \\ r_j = \sum_k \mathbb{I}\{\tau_k = b_j\}\delta_k \\ \sigma_j^2 = \ell_j \sigma^2 \end{cases}$$

OU:
$$\begin{cases} q_j = e^{-\alpha \ell_j} \\ r_j = \beta^{\mathsf{pa}(j)}(1 - e^{-\alpha \ell_j}) + \sum_k \mathbb{I}\{\tau_k = b_j\}\delta_k \left(1 - e^{-\alpha(1-\nu_k)\ell_j}\right) \\ \sigma_j^2 = \dfrac{\sigma^2}{2\alpha}(1 - e^{-2\alpha \ell_j}) \end{cases}$$

Goal Compute the following quantities, at every node j:

$$\mathbb{V}\mathrm{ar}^{(h)}\left[Z_j \mid Y\right], \mathbb{C}\mathrm{ov}^{(h)}\left[Z_j, Z_{\mathsf{pa}(j)} \mid Y\right], \mathbb{E}^{(h)}[Z_j \mid Y]$$

Upward

Goal Compute for a vector of tips, given their common ancestor:

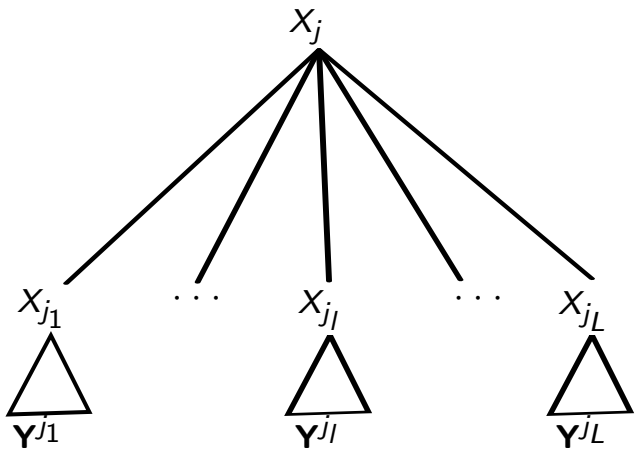
$$f_{\mathbf{Y}^j|X_j}(\mathbf{Y}^j; a) = A_j(\mathbf{Y}^j)\Phi_{M_j(\mathbf{Y}^j),S_j^2(\mathbf{Y}^j)}(a)$$

Initialization For tips: $f_{Y_i|Y_i}(Y_i; a) = \Phi_{Y_i,0}(a)$

Propagation

$$f_{\mathbf{Y}^j|X_j}(\mathbf{Y}^j; a) = \prod_{l=1}^{L} f_{\mathbf{Y}^{j_l}|X_j}(\mathbf{Y}^{j_l}; a)$$

$$f_{\mathbf{Y}^{j_l}|X_j}(\mathbf{Y}^{j_l}; a) = \int_{\mathbb{R}} f_{\mathbf{Y}^{j_l}|X_{j_l}}(\mathbf{Y}^{j_l}; b) f_{X_{j_l}|X_j}(b; a)db$$



Root Node and Likelihood At the root:

$$f_{X_1|\mathbf{Y}}(a; \mathbf{Y}) \propto f_{\mathbf{Y}|X_1}(\mathbf{Y}; a) f_{X_1}(a)$$

$$\begin{cases} \mathbb{V}\mathrm{ar}[X_1 \mid \mathbf{Y}] = \left(\frac{1}{\gamma^2} + \frac{1}{S_1^2(\mathbf{Y})}\right)^{-1} \\ \mathbb{E}[X_1 \mid \mathbf{Y}] = \mathbb{V}\mathrm{ar}[X_1 \mid \mathbf{Y}]\left(\frac{\mu}{\gamma^2} + \frac{M_1(\mathbf{Y})}{S_1^2(\mathbf{Y})}\right) \end{cases}$$

Downward

Compute $E_j = \mathbb{E}\left[X_j \mid \mathbf{Y}\right]$, $V_j^2 = \mathbb{V}\mathrm{ar}\left[X_j \mid \mathbf{Y}\right]$, $C^2_{j,\mathrm{pa}(j)} = \mathbb{C}\mathrm{ov}\left[X_j; X_{\mathrm{pa}(j)} \mid \mathbf{Y}\right]$

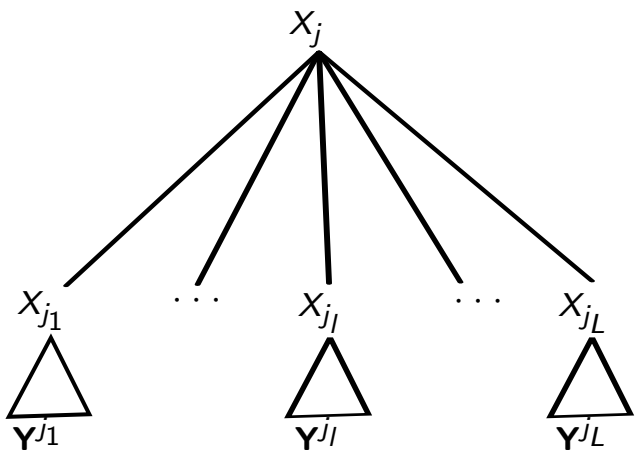
Initialization Last step of Upward.

Propagation

$$f_{X_{\mathrm{pa}(j)},X_j|\mathbf{Y}}(a,b;\mathbf{Y}) = f_{X_{\mathrm{pa}(j)}|\mathbf{Y}}(a;\mathbf{Y}) f_{X_j|X_{\mathrm{pa}(j)},\mathbf{Y}}(b;a,\mathbf{Y})$$

$$f_{X_j|X_{\mathrm{pa}(j)},\mathbf{Y}}(b;a,\mathbf{Y}) = f_{X_j|X_{\mathrm{pa}(j)},\mathbf{Y}^j}(b;a,\mathbf{Y}^j)$$

$$\propto f_{X_j|X_{\mathrm{pa}(j)}}(b;a) f_{\mathbf{Y}^j|X_j}(\mathbf{Y}^j;b)$$



Formulas

Upward

$$\begin{cases} S_j^2(\mathbf{Y}^j) = \left(\sum_{l=1}^{L} \dfrac{q_{j_l}^2}{S_{j_l}^2(\mathbf{Y}^{j_l}) + \sigma_{j_l}^2} \right)^{-1} \\ M_j(\mathbf{Y}^j) = S_j^2(\mathbf{Y}^j) \sum_{l=1}^{L} q_{j_l} \dfrac{M_{j_l}(\mathbf{Y}^{j_l}) - r_{j_l}}{S_{j_l}^2(\mathbf{Y}^{j_l}) + \sigma_{j_l}^2} \end{cases}$$

Downward

$$\begin{cases} C_{j,\mathrm{pa}(j)}^2 = q_j \dfrac{S_j^2(\mathbf{Y}^j)}{S_j^2(\mathbf{Y}^j) + \sigma_j^2} V_{\mathrm{pa}(j)}^2 \\ E_j = \dfrac{S_j^2(\mathbf{Y}^j)(q_j E_{\mathrm{pa}(j)} + r_j) + \sigma_j^2 M_j(\mathbf{Y}^j)}{S_j^2(\mathbf{Y}^j) + \sigma_j^2} \\ V_j^2 = \dfrac{S_j^2(\mathbf{Y}^j)}{S_j^2(\mathbf{Y}^j) + \sigma_j^2} \left(\sigma_j^2 + p_j^2 \dfrac{S_j^2(\mathbf{Y}^j)}{S_j^2(\mathbf{Y}^j) + \sigma_j^2} V_{\mathrm{pa}(j)}^2 \right) \end{cases}$$

back

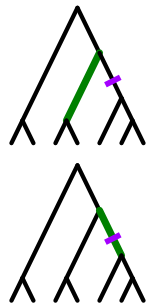
M Step: Segmentation

$$C_j(\alpha, \tau, \delta) = \sigma_j^{-2} \left(\mathbb{E}\left[X_j \mid Y\right] - q_j \mathbb{E}\left[X_{\mathrm{pa}(j)} \mid Y\right] - r_j - s_j \sum_k \mathbb{I}\{\tau_k = b_j\}\delta_k \right)^2$$

BM : $r_j = 0$, each cost is independent.

$$C_j^0(\alpha) = \sigma_j^{-2} \left(\mathbb{E}\left[X_j \mid Y\right] - q_j \mathbb{E}\left[X_{\mathrm{pa}(j)} \mid Y\right] \right)^2$$

$$C_j^1(\alpha, \tau, \delta) = \sigma_j^{-2} \left(\mathbb{E}\left[X_j \mid Y\right] - q_j \mathbb{E}\left[X_{\mathrm{pa}(j)} \mid Y\right] - s_j \sum_k \mathbb{I}\{\tau_k = b_j\}\delta_k \right)^2$$



Algorithm:

1. Find the K branches $j_1, \ldots, j_K$ with largest C_j^0;
2. Allocate one change point in the first K branches;
3. For each of these branches, set $\delta_{j_k}^{(h+1)}$ so that $C_j^1(\tau, \delta) = 0$

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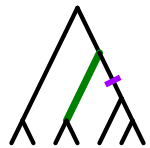
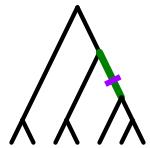


$$C_j^1(\alpha, \tau, \delta) = \sigma_j^{-2} \left(\mathbb{E}\left[X_j \mid Y\right] - q_j \mathbb{E}\left[X_{\mathsf{pa}(j)} \mid Y\right] - s_j \sum_k \mathbb{I}\{\tau_k = b_j\}\delta_k \right)^2$$



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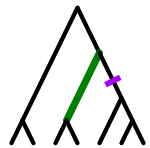

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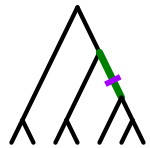

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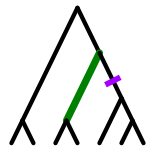

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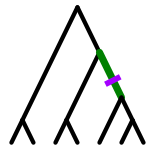

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OU : $r_j = \beta^{\mathsf{pa}(j)}$, a cost depends on all its parents.

- Exact minimization: too costly.
- Need of an heuristic.
- Idea: rewrite as a least square:

$$\|D - AU\Delta\|^2$$

with D a vector of size $n+m$, A a diagonal matrix of size $n+m$, Δ the vector of shifts and U the incidence matrix of the tree.

- Then use Stepwise selection or LASSO.

back

Model Selection on K

$$\hat{K} = \underset{0 \leq K \leq p-1}{\operatorname{argmin}} \frac{n}{2} \log\left(\frac{\|Y - \hat{s}_K\|_V^2}{n}\right) + \frac{1}{2}\operatorname{pen}'(K)$$

$$= \underset{0 \leq K \leq p-1}{\operatorname{argmin}} \|Y - \hat{s}_K\|_V^2 \left(1 + \frac{\operatorname{pen}(K)}{n - K - 1}\right)$$

Definition (Baraud et al. (2009))

Let D, $N > 0$, and $X_D \sim \chi^2(D)$, $X_N \sim \chi^2(N)$, $X_D \perp X_N$.

$$\text{Dkhi}[D, N, x] = \frac{1}{\mathbb{E}[X_D]} \mathbb{E}\left[\left(X_D - x\frac{X_N}{N}\right)_+\right], \quad \forall x > 0$$

$$\text{Dkhi}[D, N, \text{EDkhi}[D, N, q]] = q, \quad \forall 0 < q \leq 1$$

Model Selection with Unknown Variance

Theorem (Baraud et al. (2009))

Under the following setting:

$$Y' = \mathbb{E}\left[Y'\right] + \gamma E' \quad \textit{with} \quad E' \sim \mathcal{N}(0, I_n) \quad \textit{and} \quad \mathcal{S}' = \{S'_\eta, \eta \in \mathcal{M}\}$$

If $D_\eta = \text{Dim}(S'_\eta)$, $N_\eta = n - D_\eta \geq 7$, $\max(L_\eta, D_\eta) \leq \kappa n$, *with* $\kappa < 1$, *and:*

$$\Omega' = \sum_{\eta \in \mathcal{M}} (D_\eta + 1) e^{-L_\eta} < +\infty$$

$$\textit{If:} \quad \hat{\eta} = \underset{\eta \in \mathcal{M}}{\text{argmin}} \left\| Y' - \hat{Y}'_\eta \right\|^2 \left(1 + \frac{\text{pen}(\eta)}{N_\eta}\right)$$

$$\textit{with:} \quad \text{pen}(\eta) = \text{pen}_{A,\mathcal{L}}(\eta) = A \frac{N_\eta}{N_\eta - 1} \text{EDkhi}[D_\eta + 1, N_\eta - 1, e^{-L_\eta}] \quad , \quad A > 1$$

$$\textit{Then:} \ \mathbb{E}\left[\frac{\left\| \mathbb{E}\left[Y'\right] - \hat{Y}'_{\hat{\eta}} \right\|^2}{\gamma^2}\right] \leq C(A, \kappa) \left[\inf_{\eta \in \mathcal{M}} \left\{ \frac{\left\| \mathbb{E}\left[Y'\right] - Y'_\eta \right\|^2}{\gamma^2} + \max(L_\eta, D_\eta) \right\} + \Omega' \right]$$

IID Framework ($\alpha = 0$)

$$\text{Assume} \quad K_\eta = D_\eta - 1 \leq p - 1 \leq n - 8, \quad \forall \eta \in \mathcal{M}$$

Then:

$$\begin{aligned}
\Omega' &= \sum_{\eta \in \mathcal{M}} (D_\eta + 1) e^{-L_\eta} = \sum_{\eta \in \mathcal{M}} (K_\eta + 2) e^{-L_\eta} \\
&= \sum_{K=0}^{p-1} \left| \mathcal{S}_K^{PI} \right| (K+2) e^{-L_K} = \sum_{K=0}^{p-1} \left| \mathcal{S}_K^{PI} \right| (K+2) e^{-(\log \left| \mathcal{S}_K^{PI} \right| + 2\log(K+2))} \\
&= \sum_{K=0}^{p-1} \frac{1}{K+2} \leq \log(p) \leq \log(n)
\end{aligned}$$

And:

$$L_K \leq \log \binom{n+m-1}{K} + 2\log(K+2) \leq K \log(n+m-1) + 2(K+1) \leq p(2 + \log(2n-2))$$

Hence, if $p \leq \min\left(\frac{\kappa n}{2 + \log(2) + \log(n)}, n - 7 \right)$, then $\max(L_\eta, D_\eta) \leq \kappa n$ for any $\eta \in \mathcal{M}$.

Non-IID Framework ($\alpha \neq 0$)

Cholesky decomposition: $V = LL^T \quad Y' = L^{-1}Y \quad s' = L^{-1}s \quad E' = L^{-1}E$

$$Y' = \mathbb{E}\left[Y'\right] + \gamma E', \text{ with: } E' \sim \mathcal{N}(0, I_n)$$

$$S'_\eta = L^{-1}S_\eta, \qquad \hat{Y}'_\eta = \text{Proj}_{S'_\eta} Y' = \underset{a' \in S'_\eta}{\text{argmin}} \left\| Y - La' \right\|_V^2 = L^{-1}\hat{Y}_\eta$$

$$\left\| \mathbb{E}\left[Y\right] - \hat{Y}_{\hat{\eta}} \right\|_V^2 = \left\| \mathbb{E}\left[Y'\right] - \hat{Y}'_{\hat{\eta}} \right\|^2, \qquad \left\| Y - \hat{Y}_\eta \right\|_V^2 = \left\| Y' - \hat{Y}'_\eta \right\|^2$$

$$\text{Crit}_{MC}(\eta) = \left\| Y' - \hat{Y}'_\eta \right\|^2 \left(1 + \frac{\text{pen}_{A,\mathcal{L}}(\eta)}{N_\eta}\right) = \left\| Y - \hat{Y}_\eta \right\|_V^2 \left(1 + \frac{\text{pen}_{A,\mathcal{L}}(\eta)}{N_\eta}\right)$$

back

OU $\iff$ BM

Expectations

$$\mathbb{E}[Y \mid X_1 = \mu] = T \underbrace{W(\alpha)\Delta^{OU}}_{\Delta^{BM}}$$

Remark: $\mu^{BM} = \lambda^{OU} = \mu e^{-\alpha h} + \beta_0(1 - e^{-\alpha h})$

Variance

$$\mathbb{C}\text{ov}[Y_i;\, Y_j \mid X_1 = \mu] = \sigma^2 \times \underbrace{\frac{1}{2\alpha} e^{-2\alpha h}(e^{2\alpha t_{ij}} - 1)}_{t'_{ij}}$$

OU $\iff$ BM on a re-scaled tree with $t' = e^{-2\alpha h}(e^{2\alpha t} - 1)$

OU $\iff$ BM

OU $\iff$ BM on a re-scaled tree with $t' = e^{-2\alpha h}(e^{2\alpha t} - 1)$

Remarks:

- This only works for an *ultrametric* tree.
- The laws of the internal nodes is changed.

- This is *not* the following standart time transformation:

Lemma (Brownian Solution for the OU)

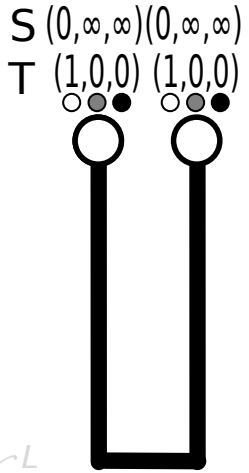
The stochastic process defined by:

$$X_t = X_0 e^{-\alpha t} + \beta(1 - e^{-\alpha t}) + \frac{\sigma}{\sqrt{2\alpha}} e^{-\alpha t} B_{e^{2\alpha t}-1}$$

is an OU, solution of the EDS $dX_t = \alpha(\beta - X_t) + \sigma dB_t$.

back

Cardinal of Equivalence Classes



Initialization For tips

Propagation

$$\mathcal{K}_k^l = \underset{1\leq p\leq K}{\operatorname{argmin}} \left\{ S_{i_l}(p) + \mathbb{I}\{p \neq k\} \right\}$$

$$S_i(k) = \sum_{l=1}^{L} S_{i_l}(p_l) + \mathbb{I}\{p_l \neq k\} \ , \ \forall (p_1, \dots p_L) \in \mathcal{K}_k^1 \times \dots \times \mathcal{K}_k^L$$

$$T_i(k) = \sum_{(p_1,\dots p_L)\in \mathcal{K}_k^1 \times \dots \times \mathcal{K}_k^L} \prod_{l=1}^{L} T_{i_l}(p_l) = \prod_{l=1}^{L} \sum_{p_l \in \mathcal{K}_k^l} T_{i_l}(p_l)$$

Termination Sum on the root vector

back

Cardinal of Equivalence Classes

Initialization For tips

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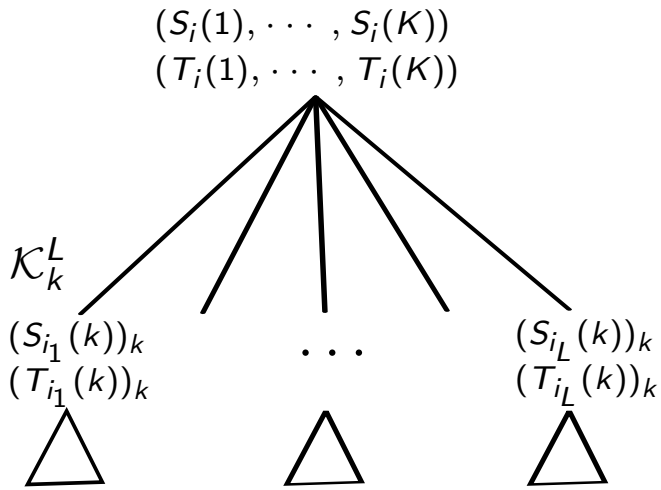

Termination Sum on the root vector

back

Cardinal of Equivalence Classes

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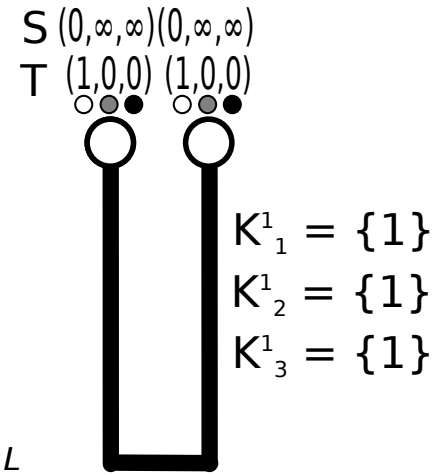


○ S(1) = 0 + 0 ; T(1) = 1 x 1

● S(2) = 0 + 1 ; T(2) = 1 x 1

● S(3) = 0 + 1 ; T(3) = 1 x 1

Termination Sum on the root vector

back

Cardinal of Equivalence Classes

Initialization For tips

Propagation

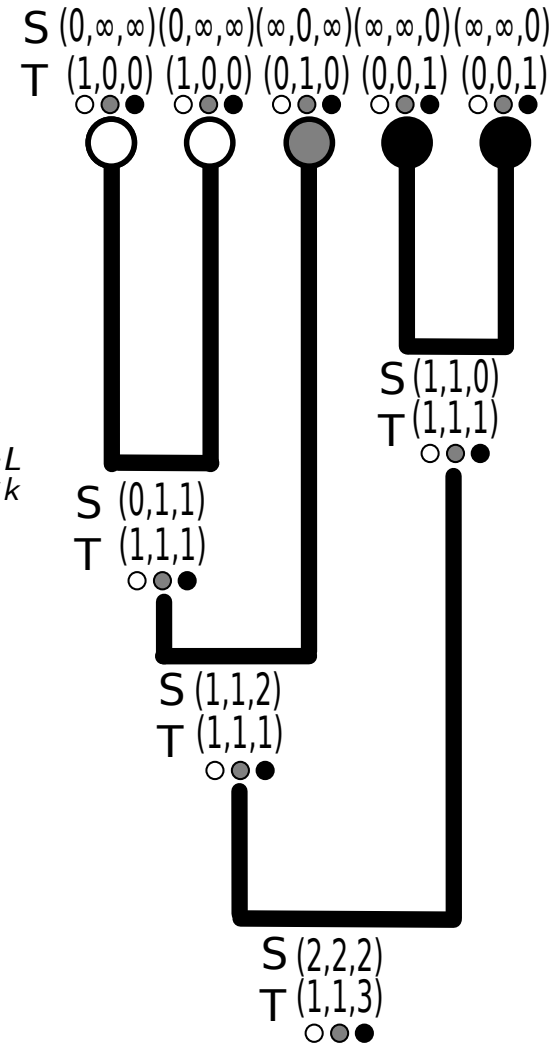
$$\mathcal{K}_k^l = \underset{1 \leq p \leq K}{\text{argmin}} \left\{ S_{i_l}(p) + \mathbb{I}\{p \neq k\} \right\}$$

$$S_i(k) = \sum_{l=1}^{L} S_{i_l}(p_l) + \mathbb{I}\{p_l \neq k\} \ , \ \forall (p_1, \dots p_L) \in \mathcal{K}_k^1 \times \dots \times \mathcal{K}_k^L$$

$$T_i(k) = \sum_{(p_1, \dots p_L) \in \mathcal{K}_k^1 \times \dots \times \mathcal{K}_k^L} \prod_{l=1}^{L} T_{i_l}(p_l) = \prod_{l=1}^{L} \sum_{p_l \in \mathcal{K}_k^l} T_{i_l}(p_l)$$

Termination Sum on the root vector

back



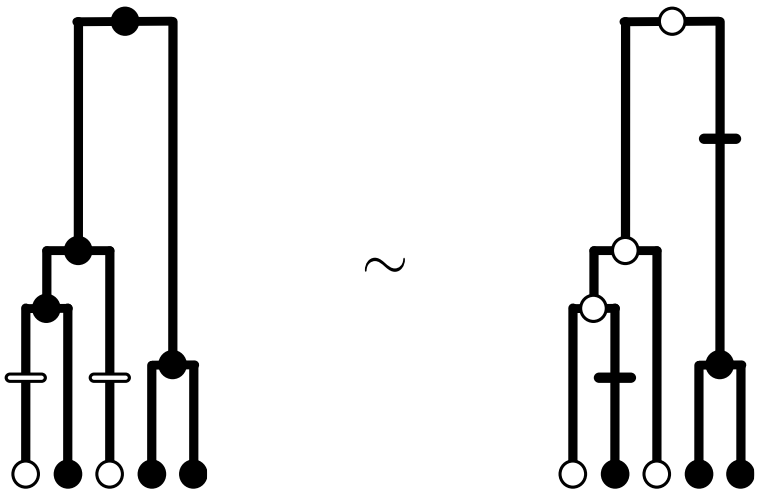
Linking Shifts and Clustering

Assumption "No Homoplasy": 1 shift = 1 new color

Proposition "K shifts $\iff$ $K+1$ clusters"

Linking Shifts and Clustering

Assumption "No Homoplasy": 1 shift = 1 new color

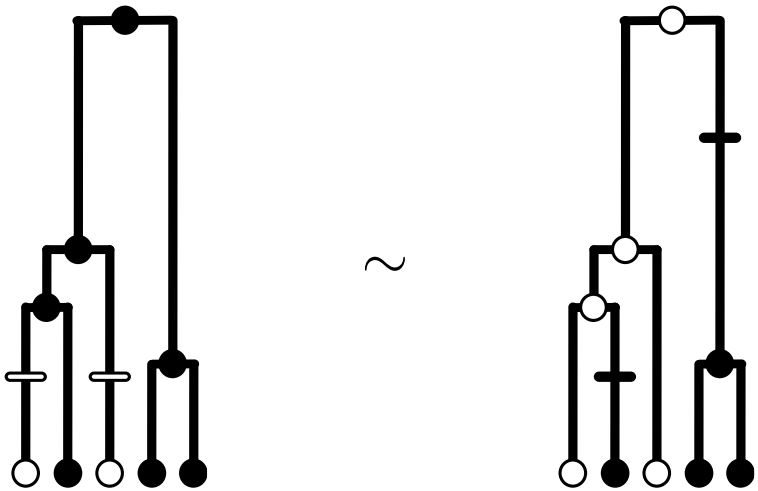


The No Homoplasy hypothesis is not respected.

Proposition "K shifts $\iff$ $K+1$ clusters"

Linking Shifts and Clustering

Assumption "No Homoplasy": 1 shift = 1 new color

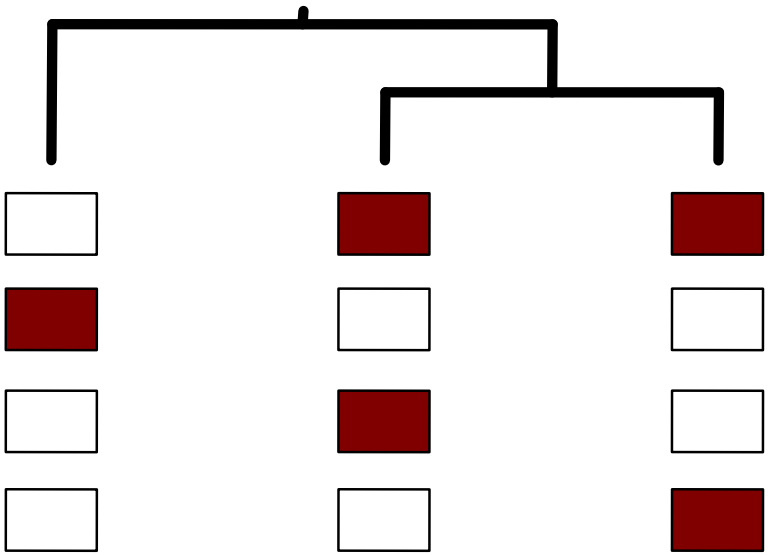


The No Homoplasy hypothesis is not respected.

Proposition "K shifts $\iff$ $K+1$ clusters"

Definitions

- $\mathcal{T}$ a rooted tree with n tips
- $N_K^{(\mathcal{T})} = |\mathcal{C}_K|$ the number of possible partitions of the tips in K clusters
- $A_K^{(\mathcal{T})}$ the number of possible *marked* partitions



Partitions in two groups for a binary tree with 3 tips

Difference between $N_2^{(\mathcal{T}_3)}$ and $A_2^{(\mathcal{T}_3)}$:

- $N_2^{(\mathcal{T}_3)} = 3$: partitions 1 and 2 are equivalent
- $A_2^{(\mathcal{T}_3)} = 4$: one marked color ("white = ancestral state")

General Formula (Binary Case)

If $\mathcal{T}$ is a binary tree, consider T_ℓ and $\mathcal{T}_r$ the left and right sub-trees of $\mathcal{T}$. Then:

$$\begin{cases} N_K^{(\mathcal{T})} = \displaystyle\sum_{k_1+k_2=K} N_{k_1}^{(\mathcal{T}_\ell)} N_{k_2}^{(\mathcal{T}_r)} + \sum_{k_1+k_2=K+1} A_{k_1}^{(\mathcal{T}_\ell)} A_{k_2}^{(\mathcal{T}_r)} \\ A_K^{(\mathcal{T})} = \displaystyle\sum_{k_1+k_2=K} A_{k_1}^{(\mathcal{T}_\ell)} N_{k_2}^{(\mathcal{T}_r)} + N_{k_1}^{(\mathcal{T}_\ell)} A_{k_2}^{(\mathcal{T}_r)} + \sum_{k_1+k_2=K+1} A_{k_1}^{(\mathcal{T}_\ell)} A_{k_2}^{(\mathcal{T}_r)} \end{cases}$$

We get:

$$N_{K+1}^{(\mathcal{T})} = N_{K+1}^{(n)} = \binom{2n-2-K}{K} \quad \text{and} \quad A_{K+1}^{(\mathcal{T})} = A_{K+1}^{(n)} = \binom{2n-1-K}{K}$$

Recursion Formula (General Case)

If we are at a node defining a tree $\mathcal{T}$ that has p daughters, with sub-trees $\mathcal{T}_1, \ldots, \mathcal{T}_p$, then we get the following recursion formulas:

$$\begin{cases} N_K^{(\mathcal{T})} = \displaystyle\sum_{\substack{k_1+\cdots+k_p=K \\ k_1,\ldots,k_p\geq 1}} \prod_{i=1}^{p} N_{k_i}^{(\mathcal{T}_i)} + \sum_{\substack{I\subset[\![1,p]\!] \\ |I|\geq 2}} \sum_{\substack{k_1+\cdots+k_p=K+|I|-1 \\ k_1,\ldots,k_p\geq 1}} \prod_{i\in I} A_{k_i}^{(\mathcal{T}_i)} \prod_{i\notin I} N_{k_i}^{(\mathcal{T}_i)} \\ A_K^{(\mathcal{T})} = \displaystyle\sum_{\substack{I\subset[\![1,p]\!] \\ |I|\geq 1}} \sum_{\substack{k_1+\cdots+k_p=K+|I|-1 \\ k_1,\ldots,k_p\geq 1}} \prod_{i\in I} A_{k_i}^{(\mathcal{T}_i)} \prod_{i\notin I} N_{k_i}^{(\mathcal{T}_i)} \end{cases}$$

No general formula. The result depends on the topology of the tree.

back

Simulations Design

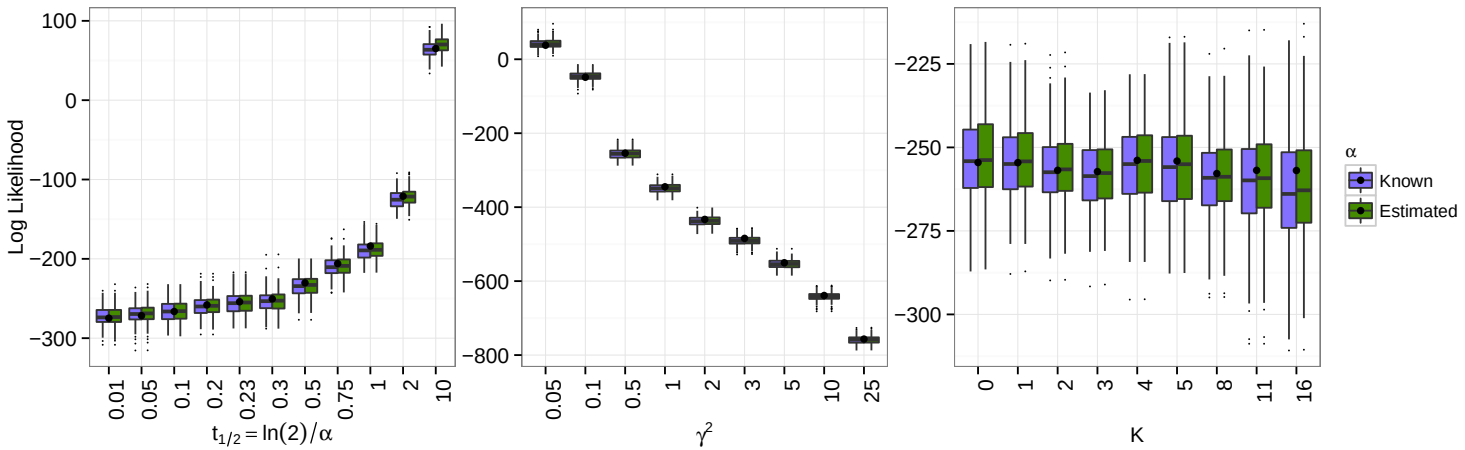
(Uyeda and Harmon, 2014)

- Topology of the tree fixed (unit height, $\lambda = 0.1$, with 64, 128, 256 taxa).
- Initial optimal value fixed: $\beta_0 = 0$
- One "base" scenario $\alpha_b = 3$, $\gamma_b^2 = 0.5$, $K_b = 5$.
- $\alpha \in \log(2)/\{0.01, 0.05, 0.1, 0.2, 0.23, 0.3, 0.5, 0.75, 1, 2, 10\}$.
- $\gamma^2 \in \{0.3, 0.6, 3, 6, 12, 18, 30, 60, 150\}/(2\alpha_b)$.
- $K \in \{0, 1, 2, 3, 4, 5, 8, 11, 16\}$.
- Shifts values $\sim \frac{1}{2}\mathcal{N}(4, 1) + \frac{1}{2}\mathcal{N}(-4, 1)$
- Shifts randomly placed at regular intervals separated by 0.1 unit length.
- $n = 200$ repetitions : 16200 configurations.

CPU time on cluster MIGALE (Jouy-en-Josas):

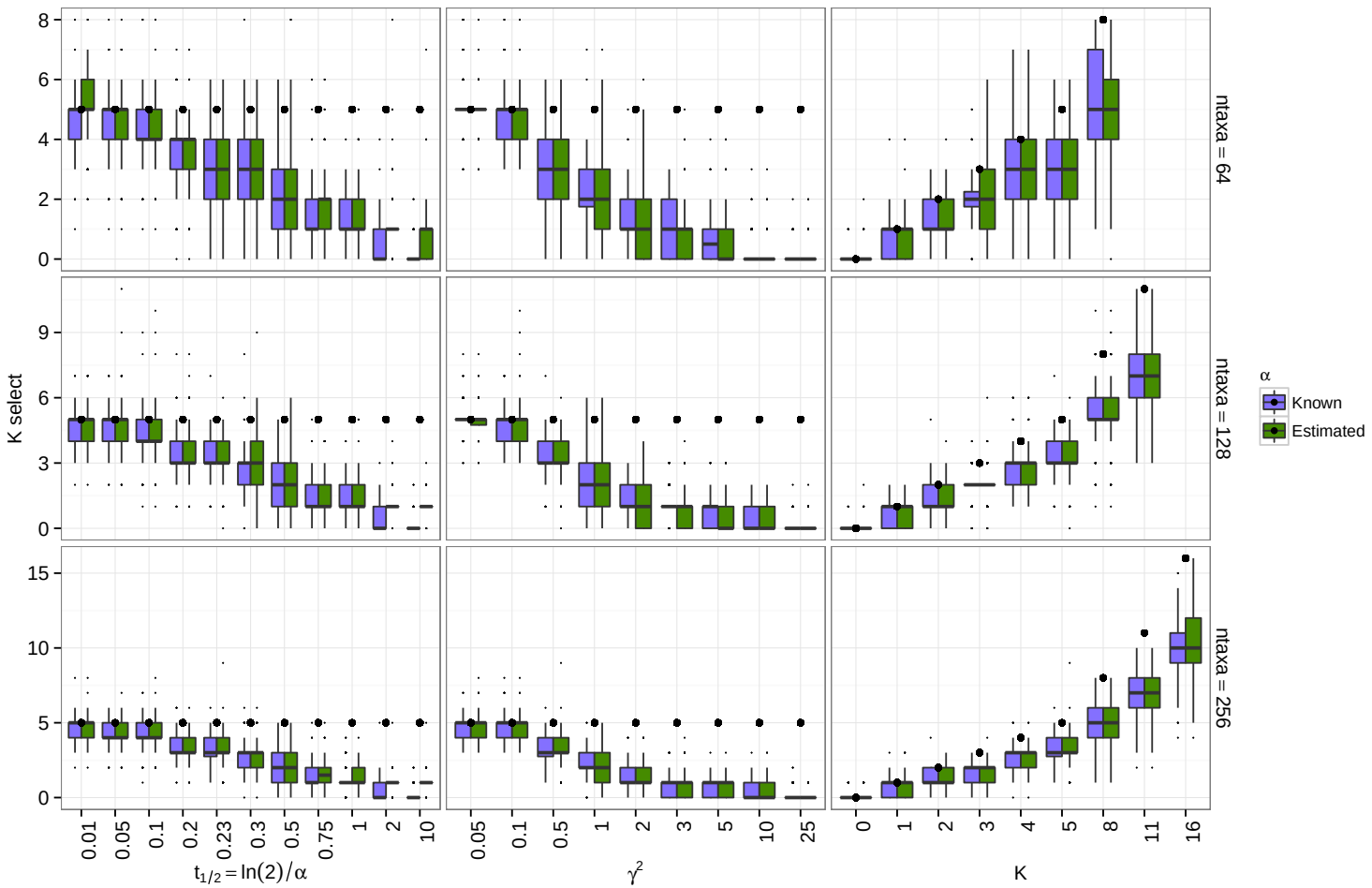
- α known: 6 minutes per estimation (66 days in total).
- α unknown: 52 minutes per estimation (570 days in total).

Log-Likelihood

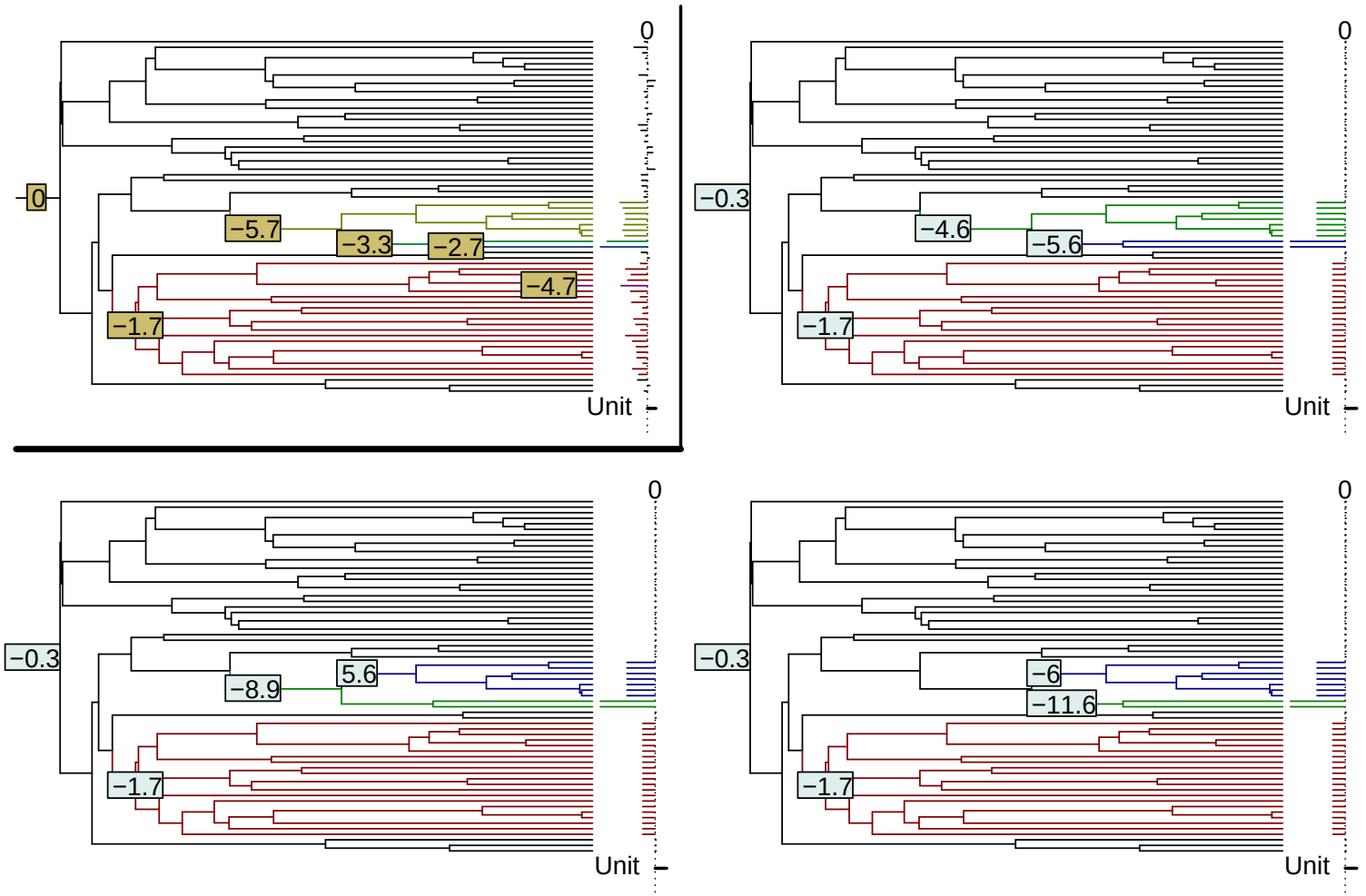


Log likelihood for a tree with 256 tips. Solid black dots are the median of the log likelihood for the true parameters.

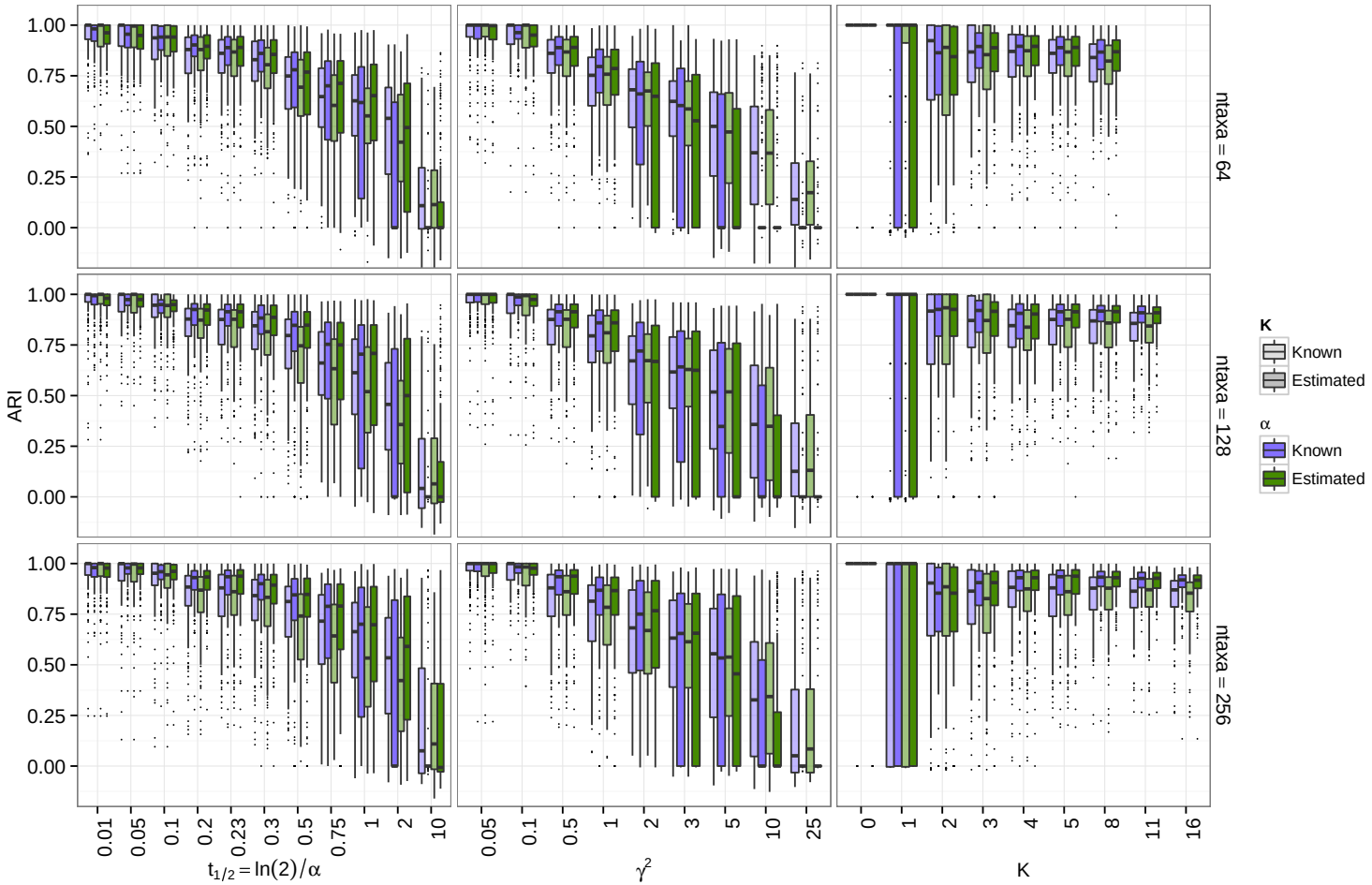
Number of Shifts



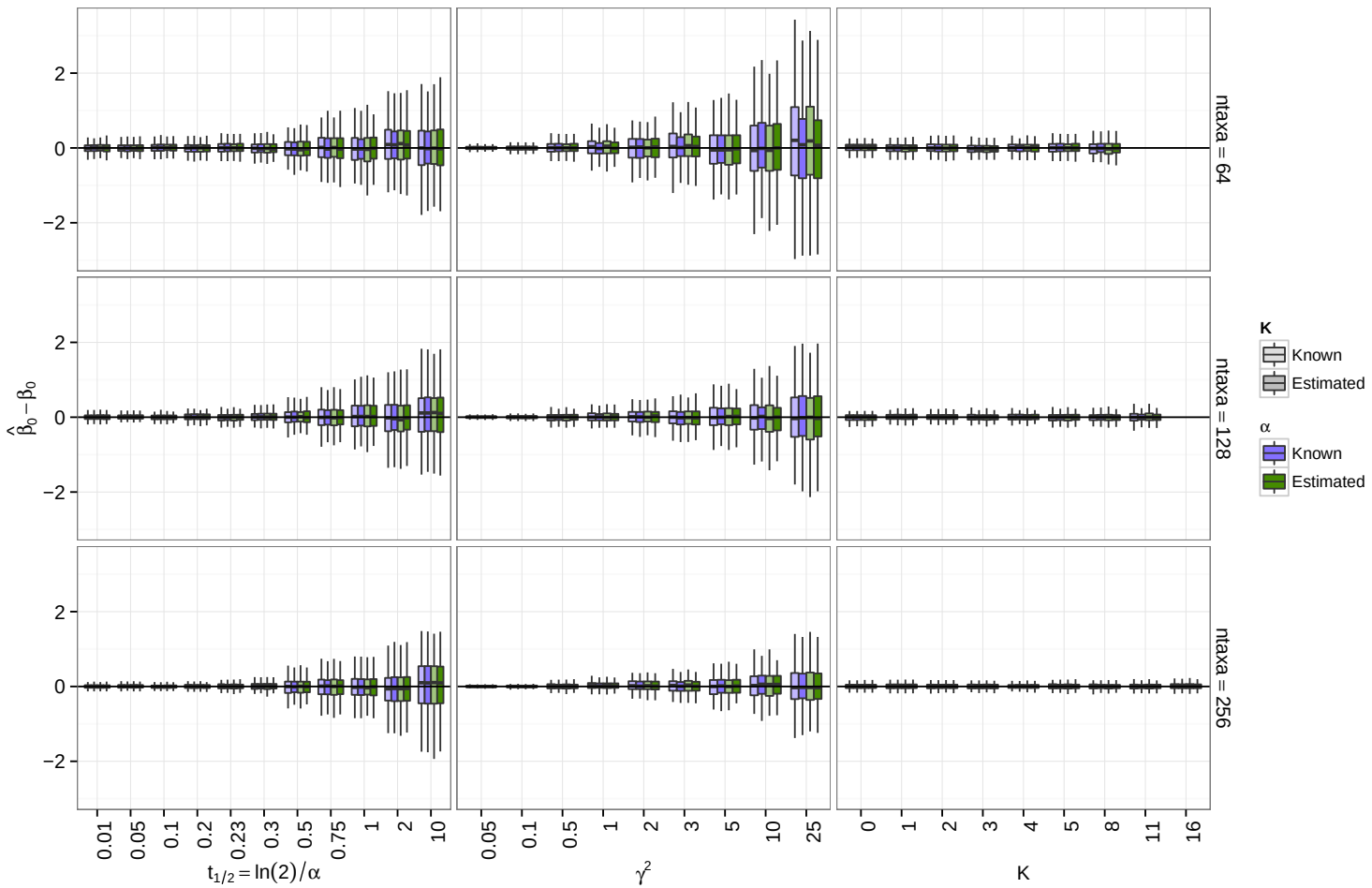
One Example


0
−5.7
−3.3
−2.7
−4.7
−1.7
Unit
0
−0.3
−4.6
−5.6
−1.7
Unit
0
−0.3
−8.9
5.6
−1.7
Unit
0
−0.3
−6
−11.6
−1.7
Unit

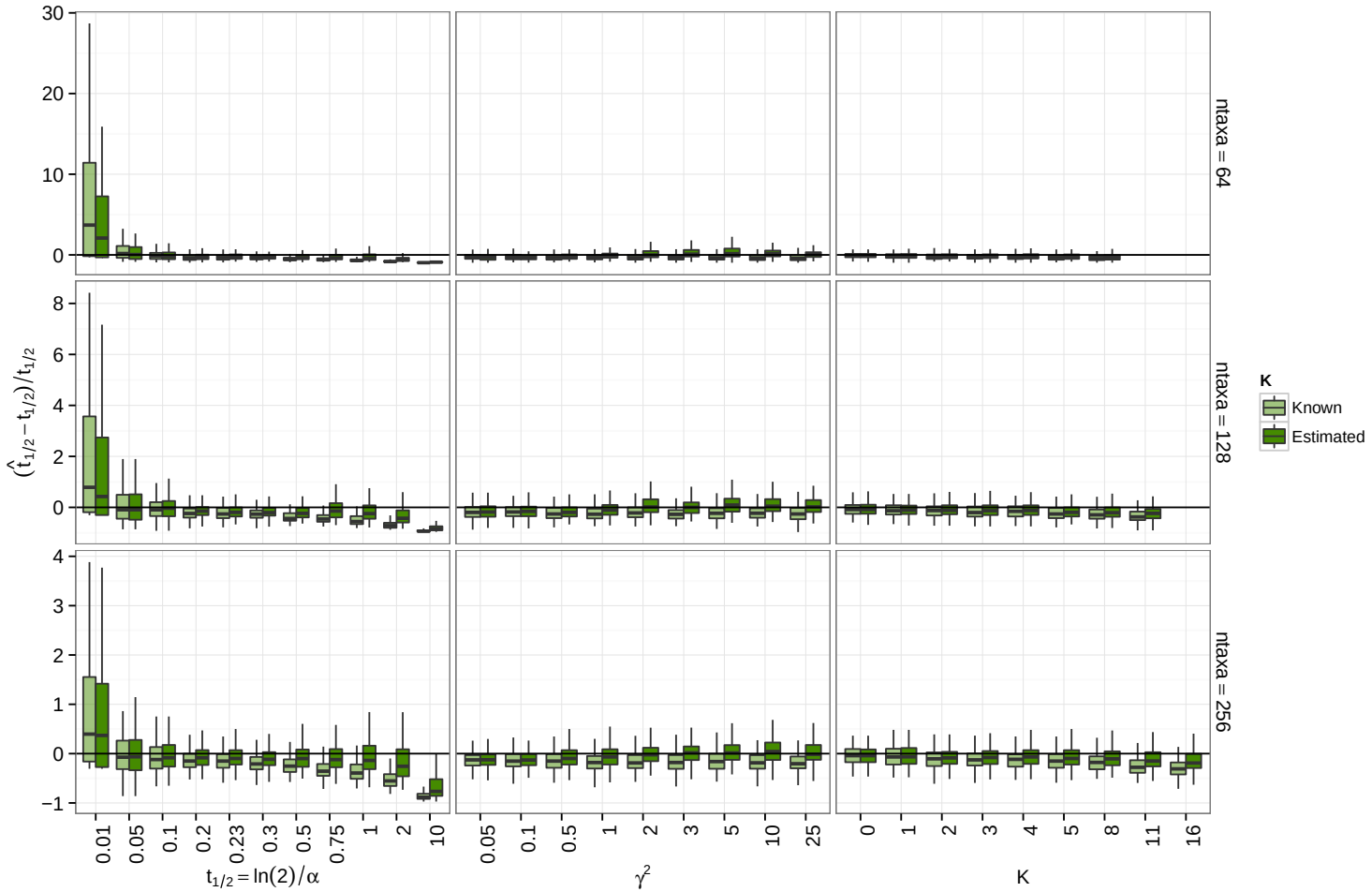
Adjusted Rand Index



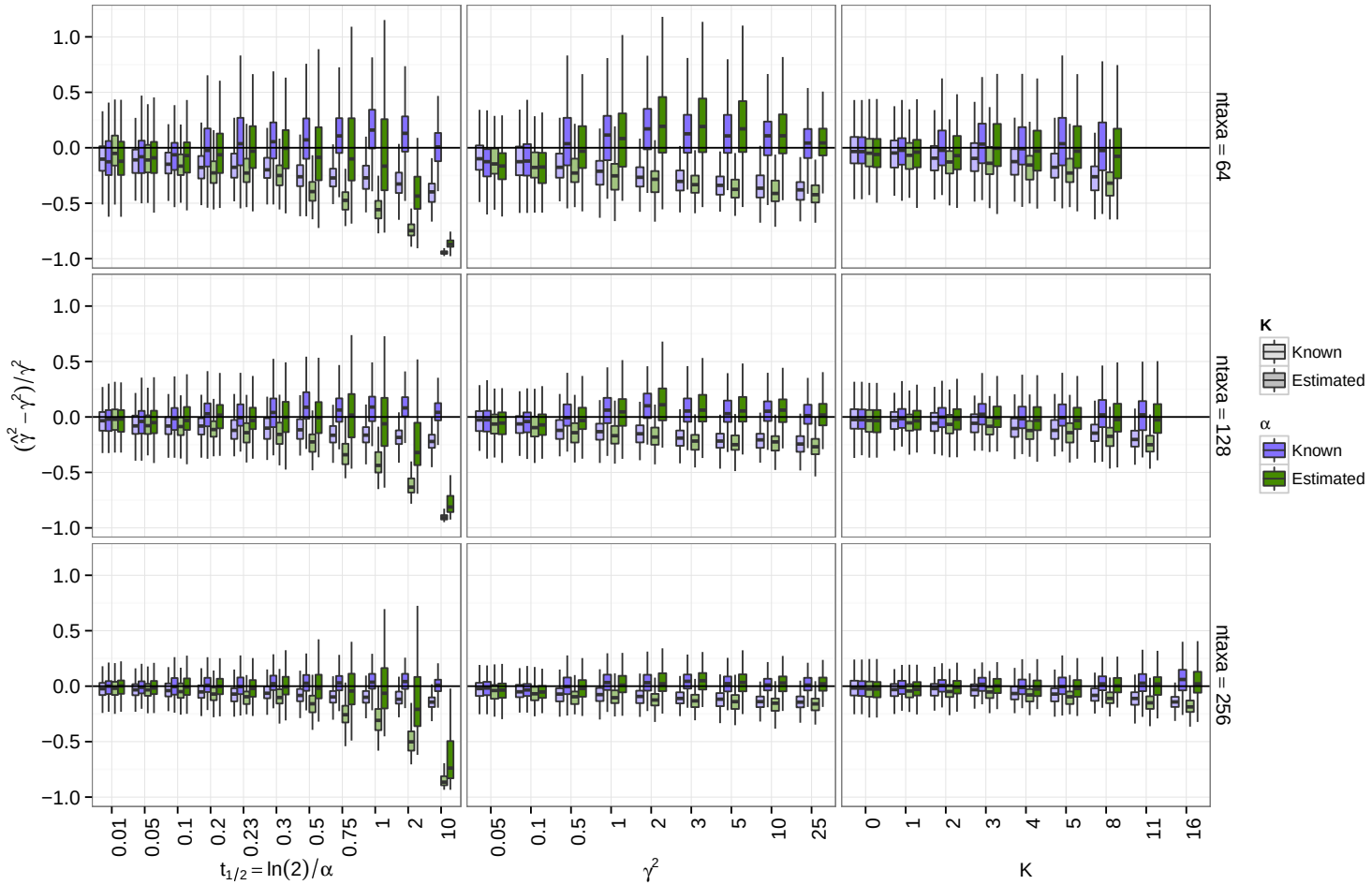
Parameters: β_0



Parameters: α



Parameters: γ^2



BM Model

Data n *vectors* of p traits at the tips: $\mathbf{Y}_i = \begin{pmatrix} Y_{i1} \\ \vdots \\ Y_{ip} \end{pmatrix}$

SDE $d\mathbf{W}(t) = \mathbf{\Sigma} d\mathbf{B}_t$, rate matrix $\mathbf{R} = \mathbf{\Sigma}\mathbf{\Sigma}^T$ $(p \times p)$

Covariances $\mathbb{C}\text{ov}\left[Y_{il}; Y_{jq}\right] = t_{ij} R_{lq}$ for i, j tips, and l, q characters

$$\mathbb{V}\text{ar}\left[\text{vec}(\mathbf{Y})\right] = \mathbf{C}_n \otimes \mathbf{R}$$

Shifts K shifts $\boldsymbol{\delta}_1, \cdots, \boldsymbol{\delta}_K$ *vectors* size p

$\mapsto$ All the characters shift at the same time

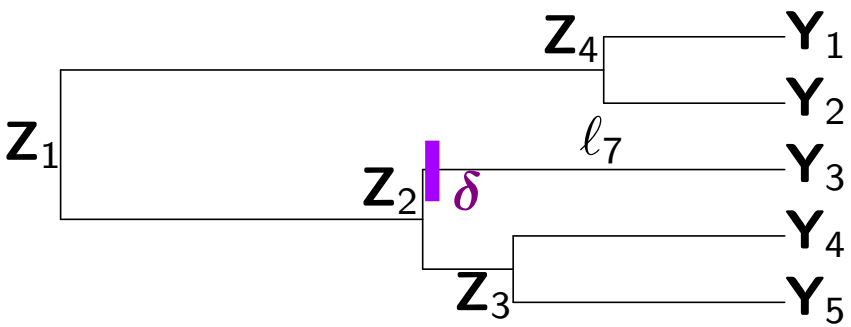
BM Model

Linear Model Representation

$$\text{vec}(\mathbf{Y}) = \text{vec}(\mathbf{\Delta T}^T) + \mathbf{E} \text{ with } \mathbf{E} \sim \mathcal{N}(0, \mathbf{V} = \mathbf{C}_n \otimes \mathbf{R})$$

Incomplete Data Representation

$$\mathbf{Y}_3 \mid \mathbf{Z}_2 \sim \mathcal{N}\Big(\mathbf{Z}_2 + \boldsymbol{\delta},\ \ell_7 \mathbf{R}\Big)$$



OU Model: General Case

Data n *vectors* of p traits at the tips: $\mathbf{Y}_i = \begin{pmatrix} Y_{i1} \\ \vdots \\ Y_{ip} \end{pmatrix}$

SDE $\mathbf{A}$ $(p \times p)$ "selection strength"

$$d\mathbf{W}(t) = -\mathbf{A}(\mathbf{W}(t) - \beta(t))dt + \mathbf{\Sigma} d\mathbf{B}_t$$

Covariances

$$\mathbb{C}\mathrm{ov}\left[\mathbf{X}_i; \mathbf{X}_j\right] = e^{-\mathbf{A}t_i}\mathbf{\Gamma} e^{-\mathbf{A}^T t_j} + e^{-\mathbf{A}(t_i - t_{ij})}\left(\int_0^{t_{ij}} e^{-\mathbf{A}v}\mathbf{\Sigma}\mathbf{\Sigma}^T e^{-\mathbf{A}^T v} dv\right) e^{-\mathbf{A}^T(t_j - t_{ij})}$$

Shifts K shifts $\boldsymbol{\delta}_1, \cdots, \boldsymbol{\delta}_K$ *vectors* size p

$\mapsto$ On the optimal values

OU Model: **A** scalar

Assumption $\mathbf{A} = \alpha \mathbf{I}_p$ "scalar"

Stationnary State $\mathbf{S} = \frac{1}{2\alpha}\mathbf{R}$

Fixed Root For i, j tips and l, q characters:

$$\mathbb{C}\text{ov}\left[Y_{il}; Y_{jq}\right] = \frac{1}{2\alpha} e^{-2\alpha h} \left(e^{2\alpha t_{ij}} - 1\right) R_{lq}$$

$\mapsto$ Can be reduced to a BM on a re-scaled tree

EM algorithm

BM Natural generalization of the univariate case.

OU M step intractable in general.

Incomplete Data Model: Can readily handle missing data.

Model Selection

- Previous criterion cannot be applied
- Solution: "Slope Heuristic"-based method
  - Massart (2007)
    - oracle inequality with known variance
    - penalty up to a multiplicative constant
  - Baudry et al. (2012)
    - Slope-heuristic method to calibrate the constant
    - Implemented in capushe (Brault et al., 2012)

Model Selection: Toy Example

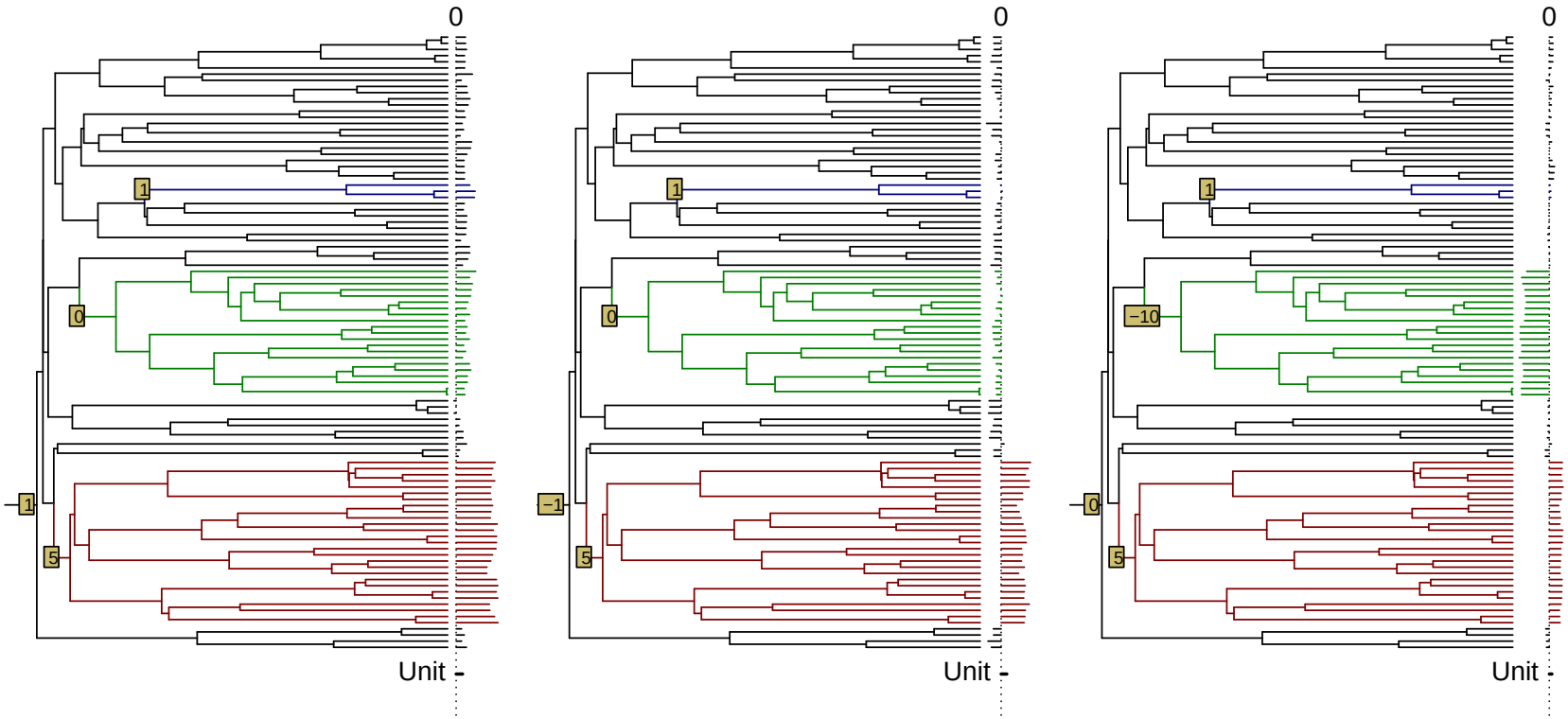


Figure: *Simulated Process.*

Model Selection: Toy Example

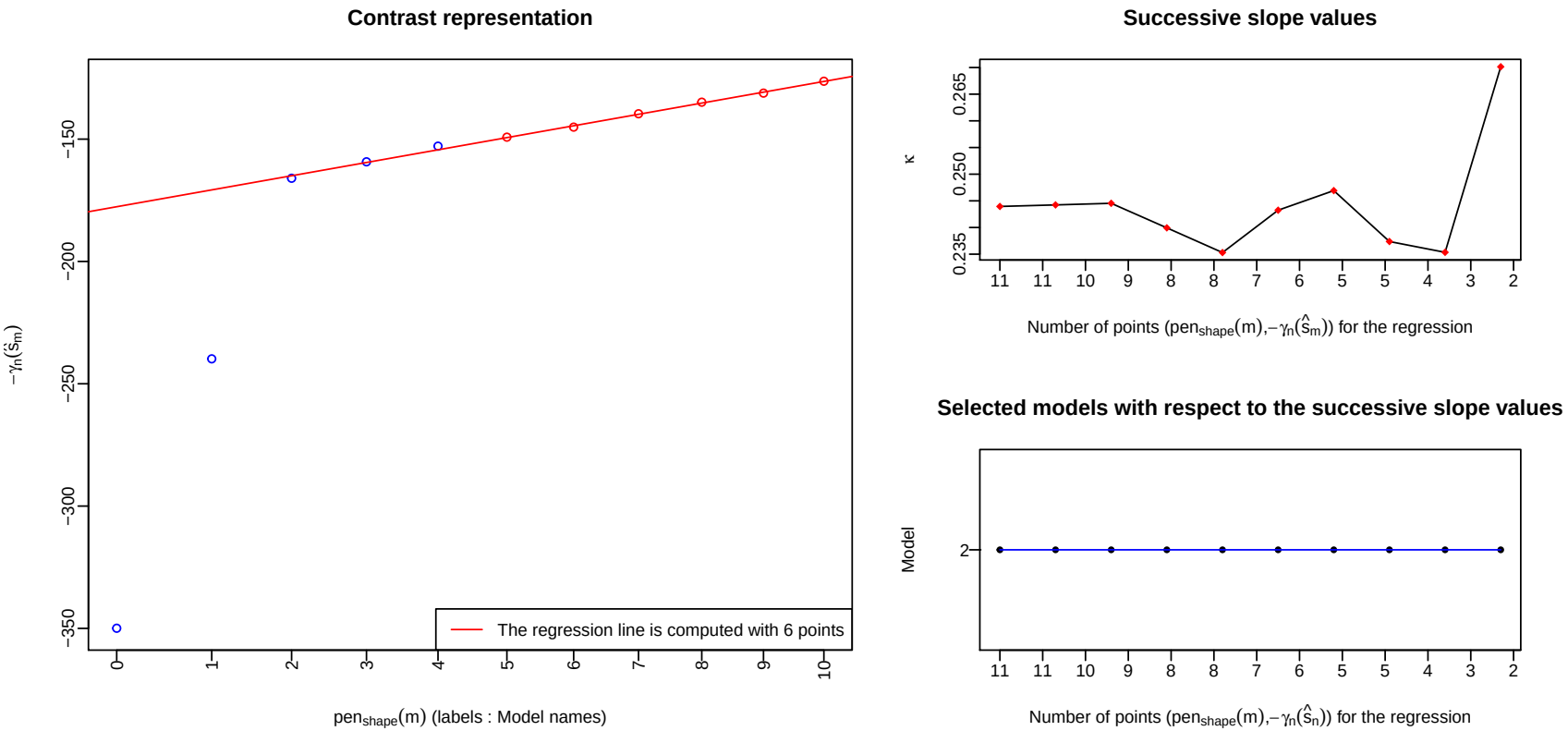


Figure: capushe *output for penalized log-likelihood.*

Model Selection: Toy Example

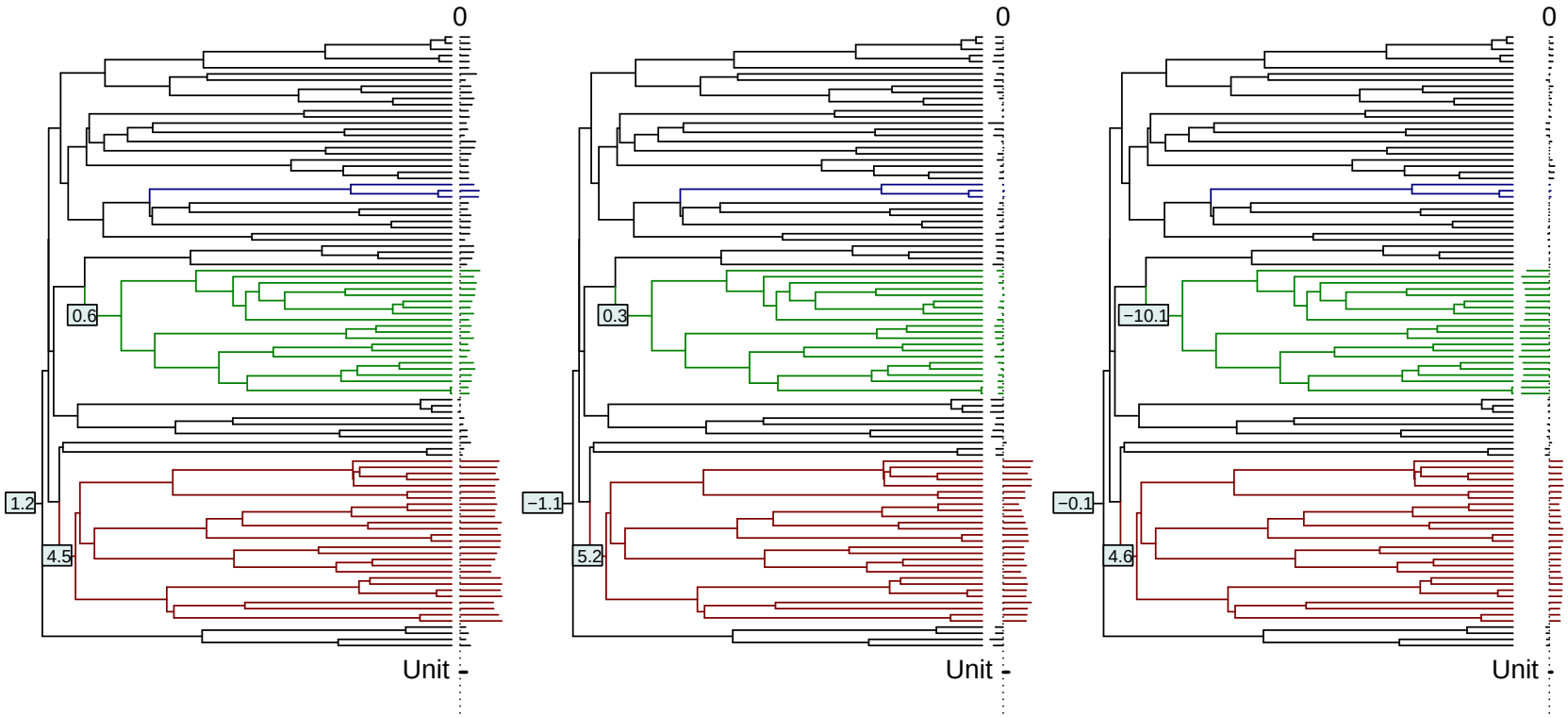


Figure: Reconstructed Process.