# Machine Learning in Bioinf<u>ormatics</u>

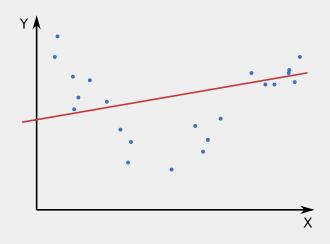
MODEL SELECTION AND REGULARIZATION

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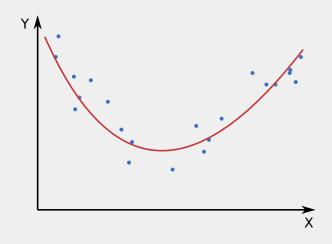
S.3 - eScience Federal Institute for Materials Research and Testing (BAM)

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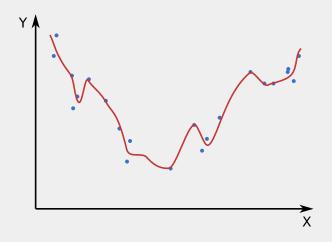
## MODEL SELECTION PROBLEM



Linear model class



Quadratic model class



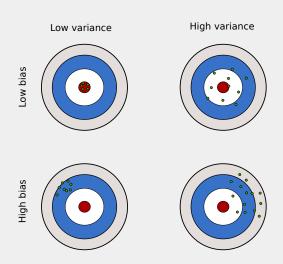
Polynomial model class

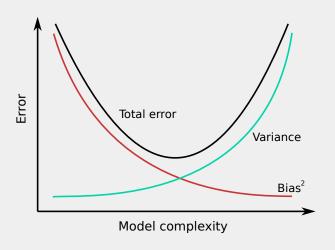
# BIAS-VARIANCE DECOMPOSITION AND TRADEOFF

- Let  $\mathbf{Y}$ ,  $\mathbf{X}$  and  $\epsilon$  be random variables such that  $\mathbf{Y} = f(\mathbf{X}) + \epsilon$ , with  $\mathbb{E}[\epsilon] = \mathbf{0}$  and  $\text{var}[\epsilon] = \sigma^2$
- Assume that  $\hat{f}_D$  has been estimated on some training data D = (X, y), where X is a matrix of n observations from  $\mathbf{X}$  and y a vector of n observations from  $\mathbf{Y}$
- At a query point *x* we have

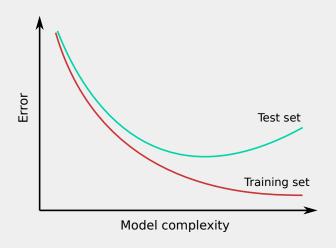
$$\mathbb{E}_{\mathbf{Y},D}[(\mathbf{Y} - \hat{f}_D(\mathbf{X}))^2] = \underbrace{[\mathbb{E}_D \hat{f}_D(\mathbf{X}) - f(\mathbf{X})]^2}_{\text{Bias}^2} + \underbrace{\mathbb{E}_D[\hat{f}_D(\mathbf{X}) - \mathbb{E}_D \hat{f}_D(\mathbf{X})]^2}_{\text{Variance}} + \sigma^2$$

- bias: Is there a bias towards a particular kind of solution (e.g. linear model)? (inductive bias)
- variance: How much does the estimated model change if you train on a different data set? (overfitting)





<sup>&</sup>lt;sup>o</sup>Note that here we average over multiple data sets. On a single data set we might observe bumps when increasing model complexity



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### **BIAS-VARIANCE DECOMPOSITION - LESSIONS LEARNED**

- Every model comes with a bias
- More complex models have a smaller bias but larger variance
- A bias is required to reduce the variance, but introducing a good bias requires domain knowledge
- Classical statistics often uses unbiased estimators, which is nowadays often questioned
- Keep in mind: There is no free lunch!¹

<sup>&</sup>lt;sup>1</sup>The *no free lunch theorem* [Wolpert and Macready, 1997] tells us that there exists no generic model that works well on all domains, but we need to tailor our models to the data at hand in order to introduce a model bias, which reduces variance.

### **COMPLEXITY MEASURES**

### COMPLEXITY OF CLASSIFIERS - VC DIMENSION

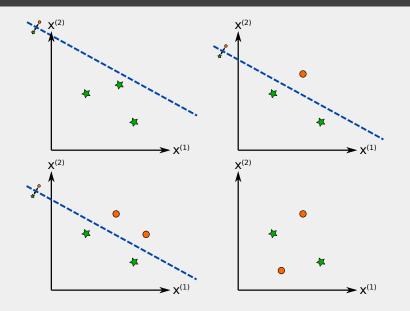
### VC-Dimension (Vapnik Chervonenkis)

Let  $\mathbb{F}_p$  be a set of classifiers on an n-dimensional input space. The VC-dimension  $VC(\mathbb{F}_p)$  is defined as the maximum number of points that can be correctly classified by at least one member of  $\mathbb{F}_p$ .

#### ■ Examples:

- ► Linear classifier on  $\mathbb{R}^p$ : VC = p + 1
- ightharpoonup SVM with RBF kernel:  $VC = \infty$
- Neural network with  $n_e$  edges,  $n_v$  nodes and sigmoid activation function:  $\Omega(n_e^2) < \text{VC} < \mathcal{O}(n_e^2 n_v^2)$  [Shalev-Shwartz and Ben-David, 2014, Section 20.4]

### COMPLEXITY OF CLASSIFIERS - VC DIMENSION



### Degrees of Freedom (DF) [Efron, 1986]

The degrees of freedom of an estimate  $\hat{y} = \hat{f}(X)$  is defined as

$$df(\hat{y}) = \frac{1}{\sigma^2} \sum_{i=1}^n cov(\hat{y}_i, y_i) = \frac{1}{\sigma^2} tr cov(\hat{y}, y),$$

#### where

- X denotes a fixed set of n covariates of dimension p
- $\blacksquare$   $y = (y_1, \dots, y_n)$  is a vector of n observations from

$$\mathbf{Y} = f(X) + \epsilon$$

for some function f, assuming  $\mathbb{E}[\epsilon] = 0$  and  $var[\epsilon] = \sigma^2$ 

<sup>&</sup>lt;sup>1</sup>df is normalized by the magnitude of the aleatory uncertainty ( $\sigma^2$ )

Degrees of freedom for the OLS estimate:

$$\begin{split} \mathrm{df}(\hat{y}) &= \frac{1}{\sigma^2} \operatorname{tr} \operatorname{cov}(\hat{y}, y) \\ &= \frac{1}{\sigma^2} \operatorname{tr} \operatorname{cov} \left( X (X^\top X)^{-1} X^\top y, y \right) \\ &= \frac{1}{\sigma^2} \operatorname{tr} \left( X (X^\top X)^{-1} X^\top \right) \operatorname{cov}(y, y) \\ &= \operatorname{tr} \left( X (X^\top X)^{-1} X^\top \right) \\ &= p \end{split}$$

- $df(\hat{y}) = p$ , i.e. the number of parameters, assuming independent feature vectors (i.e. columns of X)
- This result holds for p < n

 $<sup>{}^{1}</sup>X(X^{\top}X)^{-1}X^{\top}$  is the hat matrix  $H \in \mathbb{R}^{n \times n}$ , hence  $\mathrm{df}(\hat{y}) = \mathrm{rank}(H)$ 

■ Ridge regression is defined as

$$\hat{\theta} = \operatorname*{arg\,min}_{\theta} \|\mathbf{y} - \mathbf{X}\boldsymbol{\theta}\|_{\mathbf{2}}^{2} + \lambda \, \|\boldsymbol{\theta}\|_{\mathbf{2}}^{2}$$

for some regularization strength  $\lambda \geq 0$ 

■ The ridge estimator has

$$\mathsf{df}(\hat{y}) = \sum_{j=1}^p \frac{d_j^2}{d_j^2 + \lambda}$$

degrees of freedom, where  $(d_j)_j$  are the singular values of X

■ Increasing  $\lambda$  decreases model complexity

- There is some criticism about used DF as measure of model complexity [Janson et al., 2015]
- In some cases, we also need *X* to be random [Luan et al., 2021]
- We will see other measures when turning to model selection

### MODEL SELECTION APPROACHES

- A measure of accuracy or fit, such as the mean squared error (MSE), is not enough: Increasing model complexity will always lead to a better fit
- Estimating a model requires to minimize both
  - ▶ in-sample-error (loss on training data), and
  - out-of-sample-error (generalization error)
- Cross-validation (CV) estimates generalization error on left-out samples<sup>2</sup>
- Traditional statistics: Combine measure of accuracy (in-sample-error) with a penalty for complexity

<sup>&</sup>lt;sup>2</sup>Heavy hyperparameter tuning using CV can lead to overfitting and requires to select a final holdout set

### MODEL SELECTION APPROACHES - LOO-CV

- Leave-one-out Cross-Validation (LOO-CV) at iteration i = 1, 2, ..., n:
  - ► Compute estimate on data set without the *i*-th sample
  - ► Compute prediction error on the *i*-th sample
- Report the average prediction error over all *n* samples
- PRESS statistic (predicted residual error sum of squares):

PRESS = 
$$\sum_{i=1}^{n} (y_i - \hat{y}_{-i})^2$$

where  $\hat{y}_{-i}$  is the prediction for the *i*-th sample where the model has been estimated on all but the *i*-th sample

### MODEL SELECTION APPROACHES - PRESS

- LOO-CV is very costly for large data sets and complex models
- k-fold CV with k = 5 or k = 10 is often used in practice
- For (ridge) linear regression with mean squared error we can efficiently compute LOO-CV [Cook, 1977]

PRESS = 
$$\sum_{i=1}^{n} (y_i - \hat{y}_{-i})^2$$
$$= \sum_{i=1}^{n} \frac{(y_i - \hat{y}_i)^2}{(1 - H_{ii})^2}$$

■ The matrix

$$H = X(X^{\top}X + \lambda I)^{-1}X^{\top}$$

is called the hat matrix, because it puts a hat on y, i.e.  $\hat{y} = Hy$ 

### MODEL SELECTION APPROACHES

- LOO-CV is computationally very expensive
- *k*-fold CV is cheaper, but uses a large fraction of the data for testing
- Model performance could be better if this data was used for training
- Overfitting if we use CV for testing too many models (requires final hold out data)
- Can we do model selection by using all data for training?

### MODEL SELECTION APPROACHES - DF

Assume again the following model

$$\mathbf{Y} = f(X) + \epsilon$$

where  $X \in \mathbb{R}^{n \times p}$  is a fixed set of n predictors and  $\mathbf{Y} \in \mathbb{R}^n$ 

- Setup is very similar to the bias-variance decomposition, but X is now fixed
- Let  $\mathbf{Y}_t \in \mathbb{R}^n$  a vector of n independent observations and  $\hat{f}_{\mathbf{Y}_t}$  an estimate on the training set  $(X, \mathbf{Y}_t)$ , then [Efron, 1986]

$$\underbrace{\mathbb{E}_{\mathbf{Y},\mathbf{Y}_{t}} \left\| \mathbf{Y} - \hat{f}_{\mathbf{Y}_{t}}(X) \right\|_{2}^{2}}_{\text{expected prediction error}} = \underbrace{\mathbb{E}_{\mathbf{Y}_{t}} \left\| \mathbf{Y}_{t} - \hat{f}_{\mathbf{Y}_{t}}(X) \right\|_{2}^{2}}_{\text{expected training error}} + 2\sigma^{2} \operatorname{df}(\hat{f})$$

### MODEL SELECTION APPROACHES - DF

■ This motivates the following model selection criterium [Mallows, 2000]

$$\underbrace{\left\|y_t - \hat{f}_{y_t}(X)\right\|_2^2}_{\text{training error}} + \underbrace{2\sigma^2 \operatorname{df}(\hat{f})}_{\text{complexity penalty}}$$

- The more complex a model, the larger the penalty
- If two models fit the data equally well, we select the simpler one (Occam's razor)

### MODEL SELECTION APPROACHES - BAYES APPROACH

- Assume we have a set of models  $(m_i)_i$
- In a probabilistic setting we evaluate the probability of a model  $m_i$  given data x, i.e. using Bayes theorem

$$\operatorname{pr}(m_i \mid x) = \frac{\operatorname{pr}(x \mid m_i) \operatorname{pr}(m_i)}{\sum_j \operatorname{pr}(x \mid m_j) \operatorname{pr}(m_j)} = \frac{\operatorname{pr}(x \mid m_i) \operatorname{pr}(m_i)}{\operatorname{pr}(x)}$$

■ We compare two models  $m_i$  and  $m_j$  using

$$\frac{\operatorname{pr}(m_i \mid x)}{\operatorname{pr}(m_j \mid x)} = \frac{\frac{\operatorname{pr}(x \mid m_i)\operatorname{pr}(m_i)}{\operatorname{pr}(x)}}{\frac{\operatorname{pr}(x \mid m_j)\operatorname{pr}(m_j)}{\operatorname{pr}(x)}} = \frac{\operatorname{pr}(x \mid m_i)\operatorname{pr}(m_i)}{\operatorname{pr}(x \mid m_j)\operatorname{pr}(m_j)}$$

because pr(x) drops

### MODEL SELECTION APPROACHES - BAYES FACTOR

■ With a uniform prior over models we arrive at the Bayes factor [Kass and Raftery, 1995]

$$\frac{\operatorname{pr}(x\mid m_i)}{\operatorname{pr}(x\mid m_j)}$$

■ Hence, in Bayesian model selection, we evaluate a model *m* based on its *marginal likelihood* 

$$\operatorname{pr}(x \mid m) = \int_{\theta} \operatorname{pr}(x \mid \theta, m) \operatorname{pr}(\theta \mid m) d\theta$$

where  $\theta$  are the model parameters

■ The marginal likelihood is often difficult to evaluate, even numerically!

### MODEL SELECTION APPROACHES - BIC

- The marginal likelihood is tractable only for very simple models
- As an alternative, we use approximations of the marginal likelihood
- The Bayes information criterion (BIC) is such an approximation. Let x contain n samples and assume that  $n \gg p$ , then

$$\operatorname{pr}(x \mid m) \approx \exp\left\{-\frac{1}{2}\operatorname{BIC}(x; m)\right\}$$
$$\operatorname{BIC}(x; m) = -2\log\operatorname{pr}(x \mid \hat{\theta}, m) + p\log(n)$$

where  $\hat{\theta}$  refers to the maximum likeklihood estimate and p to the number of parameters

### MODEL SELECTION APPROACHES - BIC

- lacksquare Let **Y** and  $\epsilon$  be two random variables such that lacksquare = f(lacksquare  $)+\epsilon$
- Let  $f_{\hat{\theta}}$  denote a maximum likelihood estimate on some training data
- For  $\epsilon \sim \text{Normal}(O, \sigma^2)$  the BIC is related to the mean squared error with complexity penalty

$$\mathrm{BIC}(x; m) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (y_i - f_{\hat{\theta}}(x_i))^2 + p \log(n) + C_n$$
$$\propto \frac{1}{\sigma^2} \left\| y - f_{\hat{\theta}}(x) \right\|_2^2 + p \log(n)$$

where  $C_n$  is a constant depending on n, which can be dropped for model comparison

### MODEL SELECTION APPROACHES - FIC

- BIC assumes  $n \gg p$  and therefore depends only on the number of parameters
- Fisher Information Approximation (FIA) [Ly et al., 2017]:

$$\operatorname{FIA}(x; m) \approx \exp\left\{-\operatorname{FIA}(x; m)\right\}$$

$$\operatorname{FIA}(x; m) = \underbrace{-\log \operatorname{pr}(x \mid \hat{\theta}, m) + \frac{p}{2} \log\left(\frac{n}{2\pi}\right)}_{\text{BIC like term}} + \log C_m$$

$$C_m = \underbrace{\int_{\theta} \sqrt{\det \mathcal{I}_m(\theta)} d\theta}_{\text{Geometric complexity}}$$

where  $\mathcal{I}_m$  denotes the Fisher information matrix

■  $C_m$  is essential if  $n \gg p$  is not given [Cheema and Sugiyama, 2020]

### How do we control model complexity?

- Regularization (e.g. ridge regression):
  - Constrain the feasible set of parameter values
  - ► Keep the number of parameters in the model constant, but allow them to become zero
- Number of parameters:
  - ightharpoonup A good approximation of model complexity if n < p
  - For n > p we saw that the optimization problem has many solutions
    - In deep neural networks, the gradient descent method can act similar to a regularizer
    - Model complexity can decrease when adding more parameters (double descent)

### REGULARIZATION

### l<sub>k</sub>-PENALIZED REGRESSION

### Objective function

$$\omega(\theta) = -\log \operatorname{pr}_{\theta}(y)$$
 (maximum likelihood), or  $\omega(\theta) = \|y - X\theta\|_2^2$  (linear regression)

Regularized estimate with  $\ell_{k}$ -norm penalty

$$\hat{\theta} = \begin{cases} \underset{\theta}{\text{arg min}} & \omega(\theta) \\ \text{subject to} & \|\theta\|_k^k = \Lambda \end{cases}$$

where

$$\|\theta\|_k = \left(\sum_{j=2}^p |\theta_j|^k\right)^{1/k}$$

 $<sup>^{2}</sup>$ Remember that we do not regularize the bias or y-intercept  $heta_{0}$ 

### l<sub>k</sub>-PENALIZED REGRESSION

Identify saddle points of Lagrangian

$$\mathcal{L}(\theta, \lambda) = \omega(\theta) + \lambda(\|\theta\|_{k}^{k} - \Lambda)$$

In practice, we do not work with  $\Lambda$ , but set  $\lambda$  such that the classification performance is optimal, i.e. we work with the Lagriangian

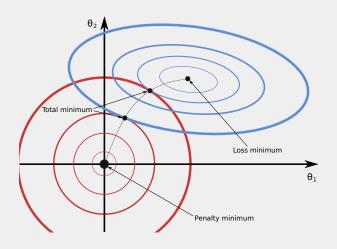
$$\hat{\theta}(\lambda) = \operatorname*{arg\,min}_{\theta} \omega(\theta) + \lambda \, \|\theta\|_{k}^{k}$$

At the optimum we must have

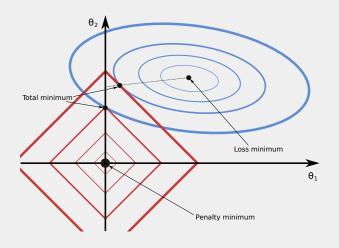
$$\nabla_{\theta} \ \omega(\theta) + \lambda \nabla_{\theta} \|\theta\|_{k}^{k} = 0$$

i.e. the gradients of  $\omega(\theta)$  and  $\lambda \|\theta\|_k^k$  must point to opposite directions

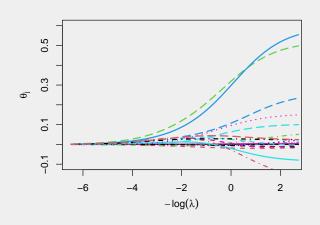
### **REGULARIZATION - K=2**



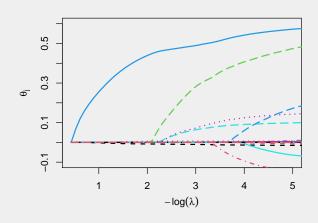
### **REGULARIZATION - K=1**



# REGULARIZATION PATHS - K=2

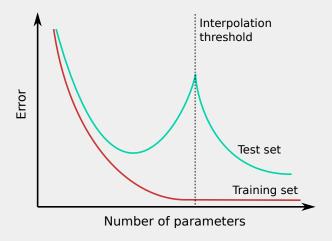


# **REGULARIZATION PATHS - K=1**

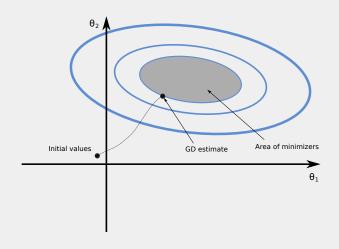


# IMPLICIT REGULARIZATION AND DOUBLE DESCENT

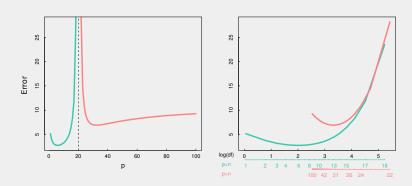
# IMPLICIT REGULARIZATION - DOUBLE DESCENT



# IMPLICIT REGULARIZATION - DOUBLE DESCENT

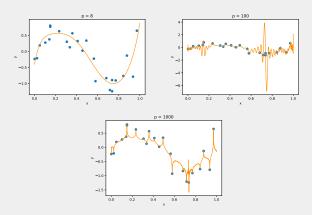


# MINIMUM $\ell_2$ -NORM ESTIMATE - DF



 $<sup>^{2}</sup>$ Requires a more advanced definition of DF that treats X as random variable [Luan et al., 2021]

# **IMPLICIT REGULARIZATION**



**Figure:** Fitting degree d=p-1 Legendre polynomials. For p>n the solution with the smallest  $\ell_2$ -norm is used.

<sup>&</sup>lt;sup>2</sup>Legendre polynomials are quite useful, since their absolute value is bounded by one.

# TAKE HOME MESSAGES

- Expected performance is the sum of training performance and model complexity
- Complex models require regularization to prevent overfitting
- The number of parameters does not correspont to the complexity of a model
- Increasing the number of features can reduce model complexity if a min- $\ell_2$ -norm estimator is used
- If we have complex data and cannot make any assumptions on the generating process, we might be better off with an overparametrized model using regularization (success behind deep learning)

### MORE REFERENCES

- Akaike information criterion (AIC) [Akaike, 1974, Cavanaugh and Neath, 2019]
- Bayesian information criterion (BIC) [Schwarz, 1978]
- Deviance information criterion (DIC) [Spiegelhalter et al., 2002]
- Fisher Information Approximation (FIA) [Rissanen, 1996, Grünwald, 2007, Cheema and Sugiyama, 2020]
- Degrees of freedom (DF) [Tibshirani, 2015, Gao and Jojic, 2016, Luan et al., 2021]
- Implicit regularization and double descent [Hastie et al., 2022, Luan et al., 2021, Derezinski et al., 2020, Kobak et al., 2020]

# READING

■ Sections 3.4, 7.3, 7.6, 7.7 and 7.9 [Hastie et al., 2009]

# THE END

"All models are wrong, but some are useful." [Moody, 1991]

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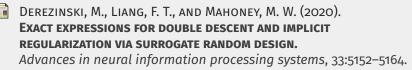
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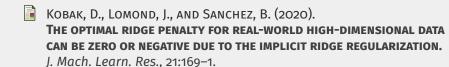
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