

# Multivariate normal: the precision matrix

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January 12, 2026

See [here](#) for a PDF version of this vignette.

## Prerequisites

You should be familiar with the [multivariate normal distribution](#) and the idea of conditional independence, particularly as illustrated by a [Markov chain](#).

## Overview

This vignette introduces the precision matrix of a multivariate normal. It also illustrates its key property: the zeros of the precision matrix correspond to conditional independencies of the variables.

## Definition, and statement of key property

Let  $X = (X_1, \dots, X_n)$  be a multivariate normal random variable with covariance matrix  $\Sigma$ .

The precision matrix,  $\Omega$ , is simply defined to be the inverse of the covariance matrix:

$$\Omega := \Sigma^{-1}.$$

The key property of the precision matrix is that its zeros tell you about conditional independence. Specifically,

$\Omega_{ij} = 0$  if and only if  $X_i$  and  $X_j$  are conditionally independent given all other coordinates of  $X$ .

It may help to compare this with the analogous property of the covariance matrix:

$\Sigma_{ij} = 0$  if and only if  $X_i$  and  $X_j$  are independent.

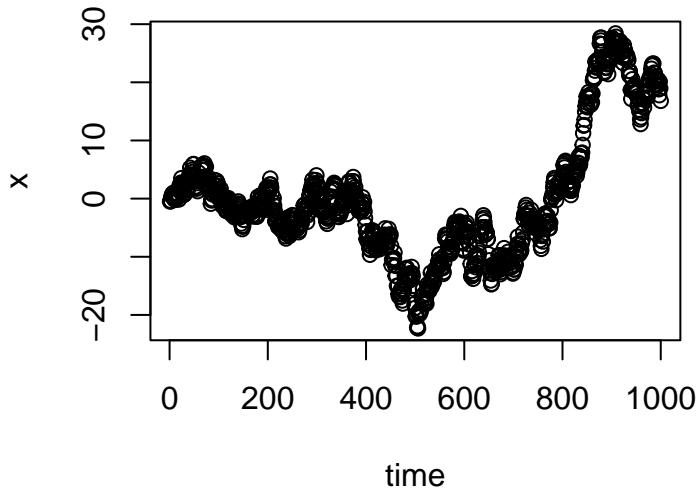
That is, whereas zeros of the covariance matrix tell you about *independence*, zeros of the precision matrix tell you about *conditional independence*.

## Example: a normal Markov chain

Consider a Markov chain  $X_1, X_2, \dots$ , where the transitions are given by  $X_{t+1} | X_t \sim N(X_t, 1)$ . You might think of this Markov chain as a type of “random walk”: given the current state, the next state is obtained by adding a random normal value (with mean 0 and variance 1).

The following code simulates a realization of this Markov chain, starting from an initial state  $X_1 \sim N(0, 1)$ , and plots it.

```
set.seed(100)
sim_normal_mc <- function (T = 1000) {
  x <- rep(0, T)
  x[1] <- rnorm(1)
  for (t in 2:T)
    x[t] <- x[t-1] + rnorm(1)
  return(x)
}
plot(sim_normal_mc(1000), xlab = "time", ylab = "x")
```



### The normal Markov chain as a multivariate normal

If you think a little, you should be able to see that the above random walk simulation is actually simulating from a 1000-dimensional multivariate normal distribution!

*Why?* Well, let's write each of the  $N(0, 1)$  variables generated using `rnorm()` in our code as  $Z_1, Z_2, \dots$ . Then we have

$$\begin{aligned} X_1 &= Z_1 \\ X_2 &= X_1 + Z_2 = Z_1 + Z_2 \\ X_3 &= X_2 + Z_3 = Z_1 + Z_2 + Z_3, \end{aligned}$$

and so on.

So we can write  $X = AZ$ , where  $A$  is the following  $1000 \times 1000$  matrix:

$$A = \begin{pmatrix} 1 & 0 & 0 & 0 & \cdots \\ 1 & 1 & 0 & 0 & \cdots \\ 1 & 1 & 1 & 0 & \cdots \\ \vdots & & & & \end{pmatrix}.$$

Let's take a look at what the covariance matrix  $\Sigma$  looks like. (We can get a good idea from looking

at the top left corner of the matrix.)

```
A <- matrix(0,1000,1000)
for(i in 1:1000)
  A[i,] <- c(rep(1,i),rep(0,1000 - i))
Sigma <- A %*% t(A)
Sigma[1:10,1:10]
#      [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10]
# [1,]    1    1    1    1    1    1    1    1    1    1
# [2,]    1    2    2    2    2    2    2    2    2    2
# [3,]    1    2    3    3    3    3    3    3    3    3
# [4,]    1    2    3    4    4    4    4    4    4    4
# [5,]    1    2    3    4    5    5    5    5    5    5
# [6,]    1    2    3    4    5    6    6    6    6    6
# [7,]    1    2    3    4    5    6    7    7    7    7
# [8,]    1    2    3    4    5    6    7    8    8    8
# [9,]    1    2    3    4    5    6    7    8    9    9
# [10,]   1    2    3    4    5    6    7    8    9    10
```

Now let us examine the precision matrix,  $\Omega$ , which recall is the inverse of  $\Sigma$ . Again we just show the top left corner of the precision matrix here.

```
Omega <- chol2inv(chol(Sigma))
Omega[1:10,1:10]
#      [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10]
# [1,]    2   -1    0    0    0    0    0    0    0    0
# [2,]   -1    2   -1    0    0    0    0    0    0    0
# [3,]    0   -1    2   -1    0    0    0    0    0    0
# [4,]    0    0   -1    2   -1    0    0    0    0    0
# [5,]    0    0    0   -1    2   -1    0    0    0    0
# [6,]    0    0    0    0   -1    2   -1    0    0    0
# [7,]    0    0    0    0    0   -1    2   -1    0    0
# [8,]    0    0    0    0    0    0   -1    2   -1    0
# [9,]    0    0    0    0    0    0    0   -1    2   -1
# [10,]   0    0    0    0    0    0    0    0   -1    2
```

Notice all the zeros in the precision matrix. This is because of the conditional independencies that occur in a Markov chain. In a Markov chain (*any* Markov chain), the conditional distribution of  $X_t$  given the other  $X_s$  ( $s \neq t$ ) depends only on its neighbors  $X_{t-1}$  and  $X_{t+1}$ . That is,  $X_t$  is conditionally independent of all other  $X_s$  given  $X_{t-1}$  and  $X_{t+1}$ . This is exactly what we are seeing in the precision matrix above: the non-zero elements of the  $t$ th row are at coordinates  $t-1, t$  and  $t+1$ .

### Addendum: interpretation of $\Omega$ in terms of conditional mean of $X_i$

The following fact is also useful, both in practice and for intuition.

Suppose  $X \sim N_r(0, \Omega^{-1})$ , where the subscript  $r$  indicates that  $X$  is  $r$ -variate.

Let  $Y_1$  denote the first coordinate of  $X$ , and let  $Y_2$  denote the remaining coordinates, that is,  $Y_2 := (X_2, \dots, X_r)$ . Further let  $\Omega_{12}$  denote the  $1 \times (r - 1)$  submatrix of  $\Omega$  that consists of row 1 and columns 2 through  $r$ .

The conditional distribution of  $Y_1 | Y_2$  is (univariate) normal with mean

$$E[Y_1 | Y_2] = -\Omega_{12}Y_2/\Omega_{11}$$

and variance  $1/\Omega_{11}$ .

Of course, there is nothing special about  $X_1$ : a similar result applies for any  $X_i$ . You just have to replace  $\Omega_{11}$  with  $\Omega_{ii}$  and define  $\Omega_{12}$  to be the  $i$ th row of  $\Omega$  with all columns except column  $i$ .

## Application

An application of this is imputation of missing values: suppose one of the  $X$  values is missing, say  $X_i$  is missing, but you know the covariance matrix and all the other  $X$  values. Then you could impute  $X_i$  by its conditional mean, which is a simple linear combination of the other values that can be read directly off the  $i$ th row of the precision matrix. This idea is the essence of [Kriging](#).

## Example

Consider the Markov chain above. The conditional distribution of  $X_1$  given all other  $X$  values is given by

$$X_1 | X_2, X_3, \dots, X_{1000} \sim N(X_2/2, 1/2).$$

And the conditional distribution of  $X_2$  given all other  $X$  values is

$$X_2 | X_1, X_3, X_4, \dots, X_{1000} \sim N((X_1 + X_3)/2, 1/2).$$

And similarly for  $X_i, i = 3, \dots, 1000$ . The intuition is that, if we wanted to guess what the value of  $X_i$  were given all other  $X$ 's, the best guess would be the average of its neighbours.