

Multivariate normal: the precision matrix

Matthew Stephens

University of Chicago

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See [here](#) for a PDF version of this vignette.

Prerequisites

You should be familiar with the [multivariate normal distribution](#) and the idea of conditional independence, particularly as illustrated by a [Markov chain](#).

Overview

This vignette introduces the precision matrix of a multivariate normal. It also illustrates its key property: the zeros of the precision matrix correspond to conditional independencies of the variables.

Definition, and statement of key property

Let $X = (X_1, \dots, X_n)$ be a multivariate normal random variable with covariance matrix Σ .

The precision matrix, Ω , is simply defined to be the inverse of the covariance matrix:

$$\Omega := \Sigma^{-1}.$$

The key property of the precision matrix is that its zeros tell you about conditional independence. Specifically,

$\Omega_{ij} = 0$ if and only if X_i and X_j are conditionally independent given all other coordinates of X .

It may help to compare this with the analogous property of the covariance matrix:

$$\Sigma_{ij} = 0 \text{ if and only if } X_i \text{ and } X_j \text{ are independent.}$$

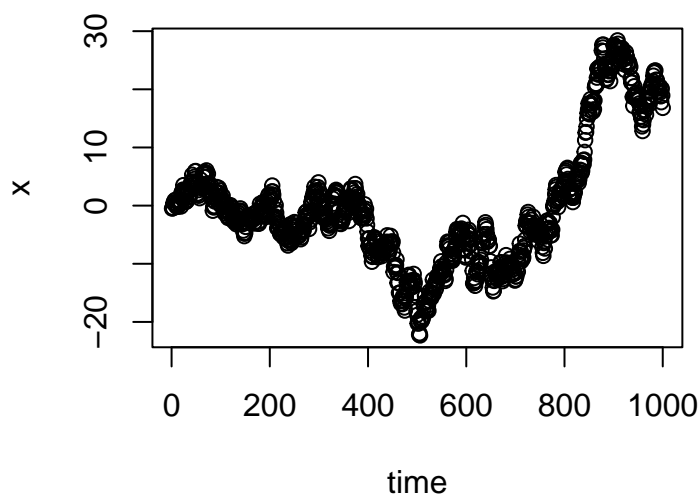
That is, whereas zeros of the covariance matrix tell you about *independence*, zeros of the precision matrix tell you about *conditional independence*.

Example: a normal Markov chain

Consider a Markov chain X_1, X_2, \dots , where the transitions are given by $X_{t+1} \mid X_t \sim N(X_t, 1)$. You might think of this Markov chain as a type of “random walk”: given the current state, the next state is obtained by adding a random normal value (with mean 0 and variance 1).

The following code simulates a realization of this Markov chain, starting from an initial state $X_1 \sim N(0, 1)$, and plots it.

```
set.seed(100)
sim_normal_mc <- function (T = 1000) {
  x <- rep(0, T)
  x[1] <- rnorm(1)
  for (t in 2:T)
    x[t] <- x[t-1] + rnorm(1)
  return(x)
}
plot(sim_normal_mc(1000), xlab = "time", ylab = "x")
```



The normal Markov chain as a multivariate normal

If you think a little, you should be able to see that the above random walk simulation is actually simulating from a 1000-dimensional multivariate normal distribution!

Why? Well, let's write each of the $N(0, 1)$ variables generated using `rnorm()` in our code as Z_1, Z_2, \dots . Then we have

$$X_1 = Z_1$$

$$X_2 = X_1 + Z_2 = Z_1 + Z_2$$

$$X_3 = X_2 + Z_3 = Z_1 + Z_2 + Z_3,$$

and so on.

So we can write $X = AZ$, where A is the following 1000×1000 matrix:

$$A = \begin{pmatrix} 1 & 0 & 0 & 0 & \dots \\ 1 & 1 & 0 & 0 & \dots \\ 1 & 1 & 1 & 0 & \dots \\ \vdots & & & & \end{pmatrix}.$$

Let's take a look at what the covariance matrix Σ looks like. (We can get a good idea from looking

at the top left corner of the matrix.)

```
A <- matrix(0,1000,1000)
for(i in 1:1000)
  A[i,] <- c(rep(1,i),rep(0,1000 - i))
Sigma <- A %*% t(A)
Sigma[1:10,1:10]
```

#	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]	[,9]	[,10]
# [1,]	1	1	1	1	1	1	1	1	1	1
# [2,]	1	2	2	2	2	2	2	2	2	2
# [3,]	1	2	3	3	3	3	3	3	3	3
# [4,]	1	2	3	4	4	4	4	4	4	4
# [5,]	1	2	3	4	5	5	5	5	5	5
# [6,]	1	2	3	4	5	6	6	6	6	6
# [7,]	1	2	3	4	5	6	7	7	7	7
# [8,]	1	2	3	4	5	6	7	8	8	8
# [9,]	1	2	3	4	5	6	7	8	9	9
# [10,]	1	2	3	4	5	6	7	8	9	10

Now let us examine the precision matrix, Ω , which recall is the inverse of Σ . Again we just show the top left corner of the precision matrix here.

```
Omega <- chol2inv(chol(Sigma))
Omega[1:10,1:10]
```

#	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]	[,9]	[,10]
# [1,]	2	-1	0	0	0	0	0	0	0	0
# [2,]	-1	2	-1	0	0	0	0	0	0	0
# [3,]	0	-1	2	-1	0	0	0	0	0	0
# [4,]	0	0	-1	2	-1	0	0	0	0	0
# [5,]	0	0	0	-1	2	-1	0	0	0	0
# [6,]	0	0	0	0	-1	2	-1	0	0	0
# [7,]	0	0	0	0	0	-1	2	-1	0	0
# [8,]	0	0	0	0	0	0	-1	2	-1	0
# [9,]	0	0	0	0	0	0	0	-1	2	-1
# [10,]	0	0	0	0	0	0	0	0	-1	2

Notice all the zeros in the precision matrix. This is because of the conditional independencies that occur in a Markov chain. In a Markov chain (*any* Markov chain), the conditional distribution of X_t given the other X_s ($s \neq t$) depends only on its neighbors X_{t-1} and X_{t+1} . That is, X_t is conditionally independent of all other X_s given X_{t-1} and X_{t+1} . This is exactly what we are seeing in the precision matrix above: the non-zero elements of the t th row are at coordinates $t - 1, t$ and $t + 1$.

Addendum: interpretation of Ω in terms of conditional mean of X_i

The following fact is also useful, both in practice and for intuition.

Suppose $X \sim N_r(0, \Omega^{-1})$, where the subscript r indicates that X is r -variate.

Let Y_1 denote the first coordinate of X , and let Y_2 denote the remaining coordinates, that is, $Y_2 := (X_2, \dots, X_r)$. Further let Ω_{12} denote the $1 \times (r - 1)$ submatrix of Ω that consists of row 1 and columns 2 through r .

The conditional distribution of $Y_1 \mid Y_2$ is (univariate) normal with mean

$$E[Y_1 \mid Y_2] = -\Omega_{12}Y_2/\Omega_{11}$$

and variance $1/\Omega_{11}$.

Of course, there is nothing special about X_1 : a similar result applies for any X_i . You just have to replace Ω_{11} with Ω_{ii} and define Ω_{i2} to be the i th row of Ω with all columns except column i .

Application

An application of this is imputation of missing values: suppose one of the X values is missing, say X_i is missing, but you know the covariance matrix and all the other X values. Then you could impute X_i by its conditional mean, which is a simple linear combination of the other values that can be read directly off the i th row of the precision matrix. This idea is the essence of [Kriging](#).

Example

Consider the Markov chain above. The conditional distribution of X_1 given all other X values is given by

$$X_1 \mid X_2, X_3, \dots, X_{1000} \sim N(X_2/2, 1/2).$$

And the conditional distribution of X_2 given all other X values is

$$X_2 \mid X_1, X_3, X_4, \dots, X_{1000} \sim N((X_1 + X_3)/2, 1/2).$$

And similarly for $X_i, i = 3, \dots, 1000$. The intuition is that, if we wanted to guess what the value of X_i were given all other X 's, the best guess would be the average of its neighbours.