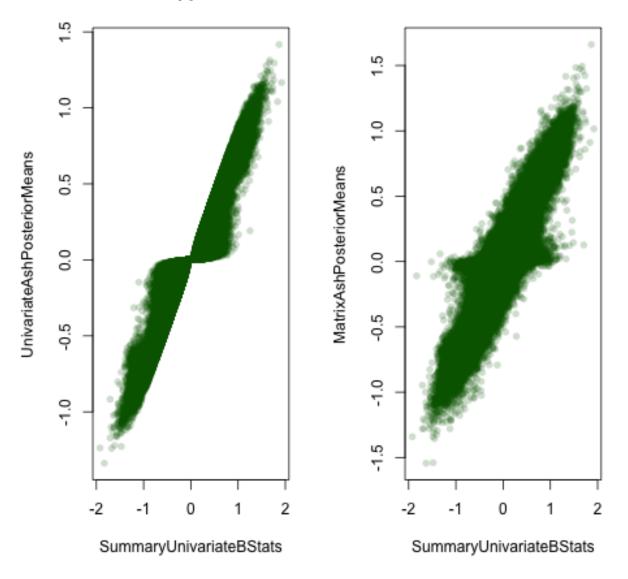
Comparing with Univariate Ash_maxb

In this document I compare with univariate ash

pdf ## 2

UnivariateAshAppliedtoAllTissueixAshPosteriorMeansvsUnivariateS



[1] "/Users/sarahurbut/gtexresults_matrixash/Analysis"

We can consider the number of statistics at each threshold using both methods:

```
thresh=0.05
sum(lfsr.mash<thresh)

## [1] 393414

sum(lfsr.ash<thresh)</pre>
```

[1] 91755

We can see that Matrix ASH identifies 4.2876573 more associations than using univariate ash. Furthemore, if we consider the number of genes with at least one LFSR less than threshold. Using

```
gene.func=function(lfsr,thresh){
sigmat=lfsr<thresh
sum(rowSums(sigmat)!=0)}
gene.func(lfsr.mash,thresh=0.05)</pre>
```

[1] 14146

```
gene.func(lfsr.ash,thresh=0.05)
```

[1] 14645

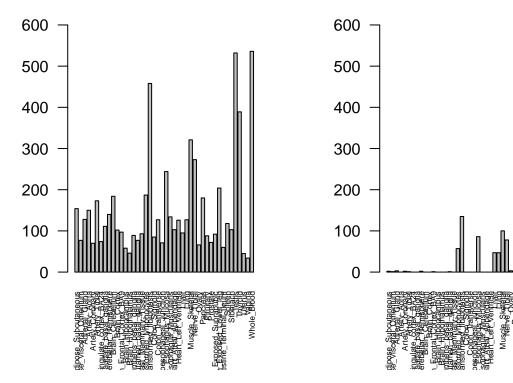
```
sum(lfsr.mash<thresh&lfsr.ash>thresh)
```

[1] 303106

Let's consider how are bias towards or against tissue speicificy changes:

```
par(mfrow=c(1,2))
plot_singletontissues(lfsr.ash,thresh,method="ASH")
plot_singletontissues(lfsr.mash,thresh,method="MATRIXASH")
```

nber of eQTL withASH<0.05 in Singl of eQTL withMATRIXASH<0.05 in S



Let's also count how many associations we might count as inconsistent by simply counting the number of times the signs differed in a vector of posterior means for a given gene snp pair:

```
inconsistent.func=function(posterior.means,lfsr,thresh=0.05){
h=apply(posterior.means,1,function(p){
  pos=sum(p>0);neg=sum(p<0);pos*neg!=0})
sum(h=="TRUE")}
inconsistent.func(pm.ash.beta,lfsr.ash)</pre>
```

[1] 14557

```
inconsistent.func(pm.mash.beta,lfsr.mash)
```

[1] 9597

If we restrict our analysis to only those considered significant at an lfsr threshold using ash computed LFSR, we have a smaller number, but this is simply because so many fewer associations are called significant using ash. Using the MASH lfsrs, the results are identical to using the univariate z statistics. WE can compare to the matrix ash posterior means as well.

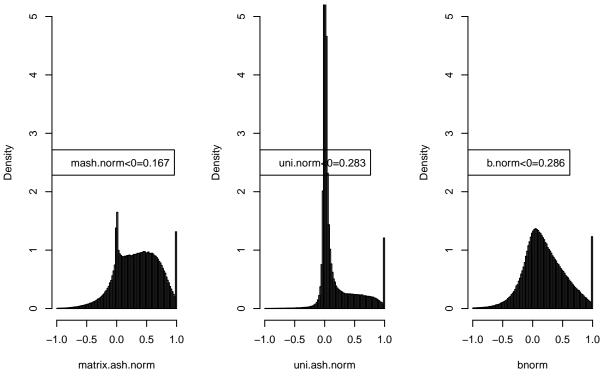
```
thresh=0.05

z=sapply(seq(1:nrow(pm.ash.beta)),function(x){
```

```
l=lfsr.mash[x,];p=pm.ash.beta[x,];plow=p[which(l<thresh)];##grab only those posterior means that are 's
if(length(plow) == 0) {return("FALSE")} ##for ones who show no significants, they can't be heterogenous
else{pos=sum(plow>0);neg=sum(plow<0);pos*neg!=0}})</pre>
sum(z==TRUE)
## [1] 9655
z=sapply(seq(1:nrow(pm.mash.beta)),function(x){
l=lfsr.mash[x,];p=pm.mash.beta[x,];plow=p[which(1<thresh)];##grab only those posterior means that are '
if(length(plow) == 0) {return("FALSE")} ##for ones who show no significants, they can't be heterogenous
else{pos=sum(plow>0);neg=sum(plow<0);pos*neg!=0}})</pre>
sum(z==TRUE)
## [1] 3180
We can also consider a histogram of normalized posterior means:
maxb=read.table("../Data/maxbetahats.txt")
bnorm=het.norm(effectsize = maxb)
sum(bnorm>0)/length(bnorm)
## [1] 0.7144276
uni.ash.norm=het.norm(effectsize = pm.ash.beta)
sum(uni.ash.norm>0)/length(uni.ash.norm)
## [1] 0.7170696
matrix.ash.norm=het.norm(effectsize = pm.mash.beta)
sum(matrix.ash.norm>0)/length(matrix.ash.norm)
## [1] 0.8330297
We might also be interested in the proportion that are greater than a particular threshold of max effect. This
sum(bnorm>0.5)/length(bnorm)
## [1] 0.1899479
sum(uni.ash.norm>0.5)/length(uni.ash.norm)
## [1] 0.1163166
sum(matrix.ash.norm>0.5)/length(matrix.ash.norm)
## [1] 0.3540626
Now, let's plot:
```

```
par(mfrow=c(1,3))
hist(matrix.ash.norm,freq=FALSE,ylim=c(0,5),nclass=100,main="B_jr|Data/B_jr[which.max(abs(B_jr|Data))],legend("left",legend=paste0("mash.norm<0=",round((sum(matrix.ash.norm<0)/length(matrix.ash.norm)),3)))
hist(uni.ash.norm,freq=FALSE,ylim=c(0,5),nclass=100,main="B_jr|Data/B_jr[which.max(abs(B_jr|Data))],ASH legend("left",legend=paste0("uni.norm<0=",round((sum(uni.ash.norm<0)/length(matrix.ash.norm)),3)))
hist(bnorm,freq=FALSE,ylim=c(0,5),nclass=100,main="B_jr|Data/B_jr[which.max(abs(B_jr|Data))],Z_raw") legend("left",legend=paste0("b.norm<0=",round((sum(bnorm<0)/length(matrix.ash.norm)),3)))</pre>
```

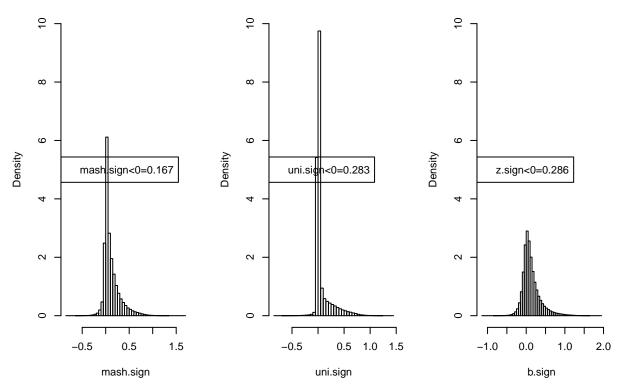
eta/B_jr[which.max(abs(B_jr|DatData/B_jr[which.max(abs(B_jr|Daeata/B_jr[which.max(abs(B_jr|Dat



```
mash.sign=sign.norm(effectsize = pm.mash.beta)
uni.sign=sign.norm(effectsize = pm.ash.beta)
b.sign=sign.norm(effectsize = maxb)

par(mfrow=c(1,3))
hist(mash.sign,freq=FALSE,ylim=c(0,10),main="B_jr|Data/sign(B_jr[which.max(abs(Z_jr|Data))],MASH",nclasslegend("left",legend=paste0("mash.sign<0=",round((sum(matrix.ash.norm<0))/length(matrix.ash.norm)),3)))
hist(uni.sign,freq=FALSE,ylim=c(0,10),main="B_jr|Data/sign(B_jr[which.max(abs(Z_jr|Data))],ASH",nclasslegend("left",legend=paste0("uni.sign<0=",round((sum(uni.ash.norm<0))/length(matrix.ash.norm)),3)))
hist(b.sign,freq=FALSE,ylim=c(0,10),main="B_jr|Data/sign(B_jr[which.max(abs(Z_jr|Data))],Z.raw)",nclasslegend("left",legend=paste0("z.sign<0=",round((sum(bnorm<0))/length(matrix.ash.norm)),3)))</pre>
```

$a/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(Z_jr|Dta/sign(B_jr[which.max(abs(A_jr[which.m$



Understandably, there are no examples in which the sign of the posterior mean using univariate ash is flipped because there is no sharing of information across tissues, while with matrix ash, the sign changes about 18% of the time:

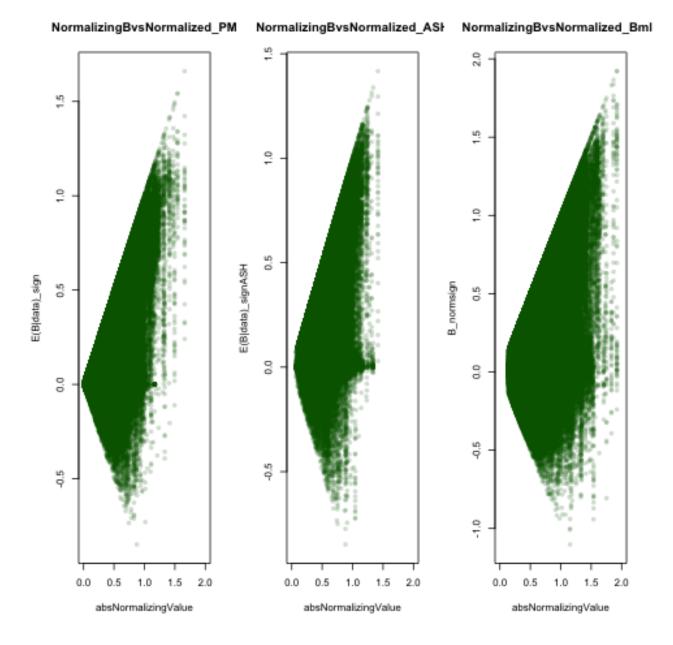
```
b.stat=maxb
sum2=b.stat*pm.mash.beta
sum(sum2<0)/length(unlist(sum2))</pre>
```

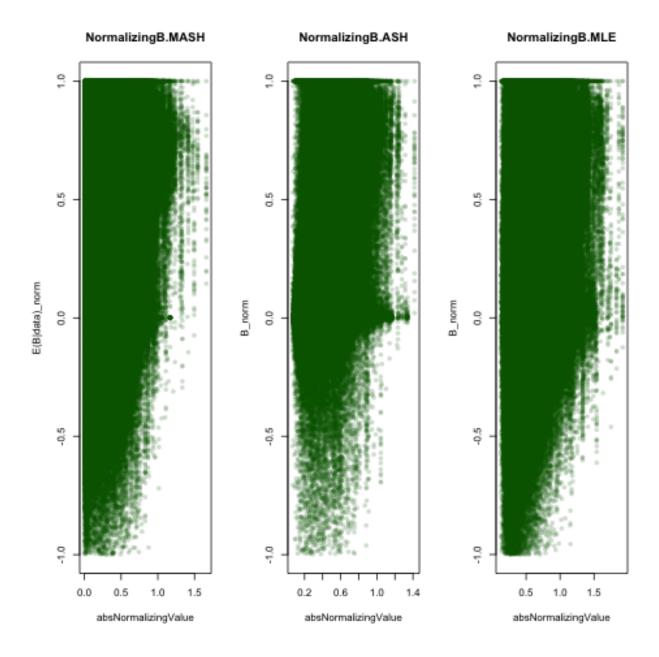
[1] 0.1842254

```
sumash=b.stat*pm.ash.beta
sum(sumash<0)/length(unlist(sum2))</pre>
```

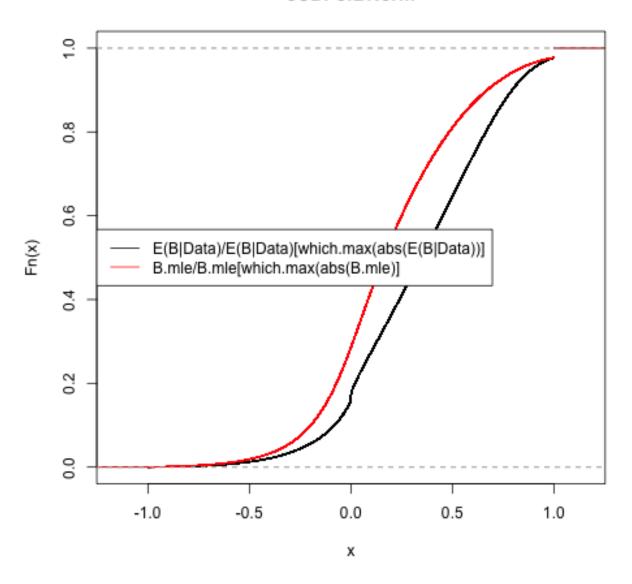
[1] 0

Let's consider the biplots of the values vs their maximums: First, I plot the biplot of the Maximum values vs their Normalized using Posterior means computed with matrix ash, univariate ash, and univariate summary statistics.





eCDFofBNorm



```
R=ncol(pm.mash.beta)
colnames(pm.mash.beta)=colnames(maxb)

max.mash=t(apply(pm.mash.beta,1,function(x){
   rep(x[which.max(abs(x))],R)
}))

max.mle=t(apply(maxb,1,function(x){
   rep(x[which.max(abs(x))],R)
}))

max.ash=t(apply(pm.ash.beta,1,function(x){
   rep(x[which.max(abs(x))],R)
}))
```

```
png("../Figures/normstuffeb.png")
par(mfrow=c(1,3))
plot(abs(max.mash), matrix.ash.norm, main="NormalizingB.MASH", xlab="absNormalizingValue", ylab="E(B|data)_
     col=rgb(0,100,0,50,maxColorValue=255), pch=16,xlim=c(0,1))
plot(abs(max.ash),uni.ash.norm,main="NormalizingB.ASH",ylab="B_norm",xlab="absNormalizingValue",col=rgb
plot(abs(max.mle),bnorm,main="NormalizingB.MLE",ylab="B_norm",xlab="absNormalizingValue",col=rgb(0,100,
dev.off()
png("../Figures/normstuffeb_sign.png")
par(mfrow=c(1,3))
plot(abs(max.mash),mash.sign,xlim=c(0,2),main="NormalizingBvsNormalized_PM",xlab="absNormalizingValue",
     col=rgb(0,100,0,50,maxColorValue=255), pch=16)
plot(abs(max.ash),uni.sign,main="NormalizingBvsNormalized_ASH",xlim=c(0,2),ylab="E(B|data)_signASH",xla
plot(abs(max.mle),b.sign,main="NormalizingBvsNormalized_Bmle",ylab="B_normsign",xlim=c(0,2),xlab="absNo
dev.off()
png("ecdf.png")
e=ecdf(ebnorm)
plot(e,main="eCDFofBNorm")
lines(ecdf(bnorm),col="red")
legend("left",lty=1,col=c("black","red"),legend=c("E(B|Data)/E(B|Data)[which.max(abs(E(B|Data))]","B.ml
dev.off()
```