# Credibility theory features of actuar

Christophe Dutang Université du Mans

Vincent Goulet Université Laval

Xavier Milhaud Université Claude Bernard Lyon 1

Mathieu Pigeon Université du Québec à Montréal

#### 1 Introduction

Credibility models are actuarial tools to distribute premiums fairly among a heterogeneous group of policyholders (henceforth called *entities*). More generally, they can be seen as prediction methods applicable in any setting where repeated measures are made for subjects with different risk levels.

The credibility theory features of **actuar** consist of matrix hachemeister containing the famous data set of Hachemeister (1975) and function cm to fit hierarchical (including Bühlmann, Bühlmann-Straub) and regression credibility models. Furthermore, function simul can simulate portfolios of data satisfying the assumptions of the aforementioned credibility models; see the "simulation" vignette for details.

#### 2 Hachemeister data set

The data set of Hachemeister (1975) consists of private passenger bodily injury insurance average claim amounts, and the corresponding number of claims, for five U.S. states over 12 quarters between July 1970 and June 1973. The data set is included in the package in the form of a matrix with 5 rows and 25 columns. The first column contains a state index, columns 2–13 contain the claim averages and columns 14–25 contain the claim numbers:

```
\begin{Sinput}
> data(hachemeister)
> hachemeister
\end{Sinput}
\begin{Soutput}
     state ratio.1 ratio.2 ratio.3 ratio.4 ratio.5 ratio.6
                                          2051
                                                   2079
[1,]
               1738
                        1642
                                 1794
                                                            2234
[2,]
          2
               1364
                        1408
                                 1597
                                          1444
                                                   1342
                                                            1675
[3,]
          3
               1759
                        1685
                                 1479
                                          1763
                                                   1674
                                                            2103
[4,]
          4
               1223
                        1146
                                 1010
                                          1257
                                                   1426
                                                            1532
          5
[5,]
               1456
                        1499
                                 1609
                                          1741
                                                   1482
                                                            1572
     ratio.7 ratio.8 ratio.9 ratio.10 ratio.11 ratio.12
[1,]
        2032
                 2035
                          2115
                                    2262
                                              2267
                                                        2517
[2,]
        1470
                 1448
                          1464
                                    1831
                                              1612
                                                        1471
[3,]
        1502
                 1622
                          1828
                                    2155
                                              2233
                                                        2059
[4,]
        1953
                          1343
                                    1243
                                              1762
                                                        1306
                 1123
[5,]
        1606
                 1735
                          1607
                                    1573
                                              1613
                                                        1690
     weight.1 weight.2 weight.3 weight.4 weight.5 weight.6
[1,]
          7861
                    9251
                              8706
                                        8575
                                                  7917
[2,]
          1622
                    1742
                              1523
                                        1515
                                                  1622
                                                            1602
[3,]
          1147
                    1357
                              1329
                                        1204
                                                   998
                                                            1077
[4,]
           407
                     396
                               348
                                         341
                                                             328
                                                   315
[5,]
          2902
                    3172
                              3046
                                        3068
                                                            2910
                                                  2693
     weight.7 weight.8 weight.9 weight.10 weight.11
[1,]
          9456
                    8003
                              7365
                                         7832
                                                    7849
[2,]
          1964
                    1515
                              1527
                                         1748
                                                    1654
[3,]
          1277
                    1218
                               896
                                         1003
                                                    1108
[4,]
           352
                    331
                               287
                                          384
                                                     321
[5,]
          3275
                    2697
                              2663
                                         3017
                                                    3242
     weight.12
[1,]
           9077
[2,]
           1861
[3,]
           1121
[4,]
            342
[5,]
           3425
\end{Soutput}
```

## 3 Hierarchical credibility model

The linear model fitting function of R is named 1m. Since credibility models are very close in many respects to linear models, and since the credibility model fitting function of **actuar** borrows much of its interface from 1m, we named the credibility function cm.

Function cm acts as a unified interface for all credibility models supported by the package. Currently, these are the unidimensional models of Bühlmann (1969) and Bühlmann and Straub (1970), the hierarchical model of Jewell (1975) (of which the first two are special cases) and the regression model of Hachemeister (1975), optionally with the intercept at the barycenter of time (Bühlmann and Gisler, 2005, Section 8.4). The modular design of cm makes it easy to add new models if desired.

This subsection concentrates on usage of cm for hierarchical models.

There are some variations in the formulas of the hierarchical model in the literature. We compute the credibility premiums as given in Bühlmann and Jewell (1987) or Bühlmann and Gisler (2005). We support three types of estimators of the between variance structure parameters: the unbiased estimators of Bühlmann and Gisler (2005) (the default), the slightly different version of Ohlsson (2005) and the iterative pseudo-estimators as found in Goovaerts and Hoogstad (1987) or Goulet (1998).

Consider an insurance portfolio where contracts are classified into cohorts. In our terminology, this is a two-level hierarchical classification structure. The observations are claim amounts  $S_{ijt}$ , where index  $i=1,\ldots,I$  identifies the cohort, index  $j=1,\ldots,J_i$  identifies the contract within the cohort and index  $t=1,\ldots,n_{ij}$  identifies the period (usually a year). To each data point corresponds a weight — or volume –  $w_{ijt}$ . Then, the best linear prediction for the next period outcome of a contract based on ratios  $X_{ijt} = S_{ijt}/w_{ijt}$  is

$$\hat{\pi}_{ij} = z_{ij} X_{ijw} + (1 - z_{ij}) \hat{\pi}_i 
\hat{\pi}_i = z_i X_{izw} + (1 - z_i) m$$
(1)

with the credibility factors

$$z_{ij} = rac{w_{ij\Sigma}}{w_{ijk\Sigma} + s^2/a},$$
  $w_{ij\Sigma} = \sum_{t=1}^{n_{ij}} w_{ijt}$   $z_i = rac{z_{i\Sigma}}{z_{i\Sigma} + a/b},$   $z_{i\Sigma} = \sum_{j=1}^{J_i} z_{ij}$ 

and the weighted averages

$$X_{ijw} = \sum_{t=1}^{n_{ij}} \frac{w_{ijt}}{w_{ij\Sigma}} X_{ijt}$$
$$X_{izw} = \sum_{j=1}^{J_i} \frac{z_{ij}}{z_{i\Sigma}} X_{ijw}.$$

The estimator of  $s^2$  is

$$\hat{s}^2 = \frac{1}{\sum_{i=1}^{I} \sum_{j=1}^{J_i} (n_{ij} - 1)} \sum_{i=1}^{I} \sum_{j=1}^{J_i} \sum_{t=1}^{n_{ij}} w_{ijt} (X_{ijt} - X_{ijw})^2.$$
 (2)

The three types of estimators for parameters a and b are the following. First, let

$$A_{i} = \sum_{j=1}^{J_{i}} w_{ij\Sigma} (X_{ijw} - X_{iww})^{2} - (J_{i} - 1)s^{2} \qquad c_{i} = w_{i\Sigma\Sigma} - \sum_{j=1}^{J_{i}} \frac{w_{ij\Sigma}^{2}}{w_{i\Sigma\Sigma}}$$

$$B = \sum_{i=1}^{I} z_{i\Sigma} (X_{izw} - \bar{X}_{zzw})^{2} - (I - 1)a \qquad d = z_{\Sigma\Sigma} - \sum_{i=1}^{I} \frac{z_{i\Sigma}^{2}}{z_{\Sigma\Sigma}},$$

with

$$\bar{X}_{zzw} = \sum_{i=1}^{I} \frac{z_{i\Sigma}}{z_{\Sigma\Sigma}} X_{izw}.$$
 (3)

(Hence,  $\mathrm{E}[A_i]=c_ia$  and  $\mathrm{E}[B]=db$ .) Then, the Bühlmann–Gisler estimators are

$$\hat{a} = \frac{1}{I} \sum_{i=1}^{I} \max \left( \frac{A_i}{c_i}, 0 \right) \tag{4}$$

$$\hat{b} = \max\left(\frac{B}{d}, 0\right),\tag{5}$$

the Ohlsson estimators are

$$\hat{a}' = \frac{\sum_{i=1}^{I} A_i}{\sum_{i=1}^{I} c_i} \tag{6}$$

$$\hat{b}' = \frac{B}{d} \tag{7}$$

and the iterative (pseudo-)estimators are

$$\tilde{a} = \frac{1}{\sum_{i=1}^{I} (J_i - 1)} \sum_{i=1}^{I} \sum_{j=1}^{J_i} z_{ij} (X_{ijw} - X_{izw})^2$$
(8)

$$\tilde{b} = \frac{1}{I - 1} \sum_{i=1}^{I} z_i (X_{izw} - X_{zzw})^2, \tag{9}$$

where

$$X_{zzw} = \sum_{i=1}^{I} \frac{z_i}{z_{\Sigma}} X_{izw}.$$
 (10)

Note the difference between the two weighted averages (3) and (10). See Belhadj et al. (2009) for further discussion on this topic.

Finally, the estimator of the collective mean m is  $\hat{m} = X_{zzw}$ .

The credibility modeling function cm assumes that data is available in the format most practical applications would use, namely a rectangular array (matrix or data frame) with entity observations in the rows and with one or more

classification index columns (numeric or character). One will recognize the output format of simul and its summary methods.

Then, function cm works much the same as 1m. It takes in argument: a formula of the form ~ terms describing the hierarchical interactions in a data set; the data set containing the variables referenced in the formula; the names of the columns where the ratios and the weights are to be found in the data set. The latter should contain at least two nodes in each level and more than one period of experience for at least one entity. Missing values are represented by NAs. There can be entities with no experience (complete lines of NAs).

In order to give an easily reproducible example, we group states 1 and 3 of the Hachemeister data set into one cohort and states 2, 4 and 5 into another. This shows that data does not have to be sorted by level. The fitted model using the iterative estimators is:

```
\begin{Sinput}
> X < - cbind(cohort = c(1, 2, 1, 2, 2), hachemeister)
> fit <- cm(~cohort + cohort:state, data = X,
            ratios = ratio.1:ratio.12,
+
            weights = weight.1:weight.12,
            method = "iterative")
> fit
\end{Sinput}
\begin{Soutput}
Call:
cm(formula = ~cohort + cohort:state, data = X, ratios = ratio.1:ratio.12,
    weights = weight.1:weight.12, method = "iterative")
Structure Parameters Estimators
  Collective premium: 1746
  Between cohort variance: 88981
  Within cohort/Between state variance: 10952
  Within state variance: 139120026
\end{Soutput}
```

The function returns a fitted model object of class "cm" containing the estimators of the structure parameters. To compute the credibility premiums, one calls a method of predict for this class:

```
\begin{Sinput}
> predict(fit)
\end{Sinput}
\begin{Soutput}
$cohort
[1] 1949 1543

$state
[1] 2048 1524 1875 1497 1585
\end{Soutput}
```

One can also obtain a nicely formatted view of the most important results with a call to summary:

```
\begin{Sinput}
> summary(fit)
\end{Sinput}
\begin{Soutput}
Call:
cm(formula = ~cohort + cohort:state, data = X, ratios = ratio.1:ratio.12,
    weights = weight.1:weight.12, method = "iterative")
Structure Parameters Estimators
 Collective premium: 1746
 Between cohort variance: 88981
 Within cohort/Between state variance: 10952
 Within state variance: 139120026
Detailed premiums
 Level: cohort
   cohort Indiv. mean Weight Cred. factor Cred. premium
          1967
                      1.407 0.9196
                                           1949
    2
           1528
                      1.596 0.9284
                                           1543
 Level: state
    cohort state Indiv. mean Weight Cred. factor
          1
                2061
                           100155 0.8874
    2
          2
                1511
                             19895 0.6103
          3
                1806
    1
                            13735 0.5195
   2
                1353
                              4152 0.2463
           5
                1600
                              36110 0.7398
   Cred. premium
    2048
    1524
    1875
    1497
    1585
\end{Soutput}
```

The methods of predict and summary can both report for a subset of the levels by means of an argument levels. For example:

```
\begin{Sinput}
> summary(fit, levels = "cohort")
\end{Sinput}
\begin{Soutput}
Call:
cm(formula = ~cohort + cohort:state, data = X, ratios = ratio.1:ratio.12,
   weights = weight.1:weight.12, method = "iterative")
Structure Parameters Estimators
 Collective premium: 1746
 Between cohort variance: 88981
 Within cohort variance: 10952
Detailed premiums
 Level: cohort
   cohort Indiv. mean Weight Cred. factor Cred. premium
                     1.407 0.9196
          1967
                                          1949
    2
                      1.596 0.9284
                                           1543
\end{Soutput}
\begin{Sinput}
> predict(fit, levels = "cohort")
\end{Sinput}
\begin{Soutput}
$cohort
[1] 1949 1543
\end{Soutput}
```

The results above differ from those of Goovaerts and Hoogstad (1987) for the same example because the formulas for the credibility premiums are different.

### 4 Bühlmann and Bühlmann-Straub models

As mentioned above, the Bühlmann and Bühlmann–Straub models are simply one-level hierarchical models. In this case, the Bühlmann–Gisler and Ohlsson estimators of the between variance parameters are both identical to the usual Bühlmann and Straub (1970) estimator

$$\hat{a} = \frac{w_{\Sigma\Sigma}}{w_{\Sigma\Sigma}^2 - \sum_{i=1}^{I} w_{i\Sigma}^2} \left( \sum_{i=1}^{I} w_{i\Sigma} (X_{iw} - X_{ww})^2 - (I-1)\hat{s}^2 \right), \tag{11}$$

and the iterative estimator

$$\tilde{a} = \frac{1}{I - 1} \sum_{i=1}^{I} z_i (X_{iw} - X_{zw})^2$$
 (12)

is better known as the Bichsel-Straub estimator.

To fit the Bühlmann model using cm, one simply does not specify any weights:

```
\begin{Sinput}
> cm(~state, hachemeister, ratios = ratio.1:ratio.12)
\end{Sinput}
\begin{Soutput}
Call:
cm(formula = ~state, data = hachemeister, ratios = ratio.1:ratio.12)

Structure Parameters Estimators

Collective premium: 1671

Between state variance: 72310
Within state variance: 46040
\end{Soutput}
```

In comparison, the results for the Bühlmann–Straub model using the Bichsel–Straub estimator are:

## 5 Regression model of Hachemeister

The regression model of Hachemeister (1975) is a generalization of the Bühlmann–Straub model. If data shows a systematic trend, the latter model will typically under- or over-estimate the true premium of an entity. The idea of Hachemeister was to fit to the data a regression model where the parameters are a credibility weighted average of an entity's regression parameters and the group's parameters.

In order to use cm to fit a credibility regression model to a data set, one simply has to supply as additional arguments regformula and regdata. The first one is a formula of the form ~ terms describing the regression model and the second is a data frame of regressors. That is, arguments regformula and regdata are in every respect equivalent to arguments formula and data of lm, with the minor difference that regformula does not need to have a left hand side (and is ignored if present). For example, fitting the model

$$X_{it} = \beta_0 + \beta_1 t + \varepsilon_t, \quad t = 1, ..., 12$$

to the original data set of Hachemeister (1975) is done with

```
\begin{Sinput}
> fit <- cm(~state, hachemeister, regformula = ~ time,</pre>
            regdata = data.frame(time = 1:12),
            ratios = ratio.1:ratio.12,
            weights = weight.1:weight.12)
> fit
\end{Sinput}
\begin{Soutput}
Call:
cm(formula = ~state, data = hachemeister, ratios = ratio.1:ratio.12,
    weights = weight.1:weight.12, regformula = ~time, regdata = data.frame(time = 1:12))
Structure Parameters Estimators
  Collective premium: 1469 32.05
 Between state variance: 24154 2700.0
                            2700 301.8
 Within state variance: 49870187
\end{Soutput}
```

Computing the credibility premiums requires to give the "future" values of the regressors as in predict.lm:

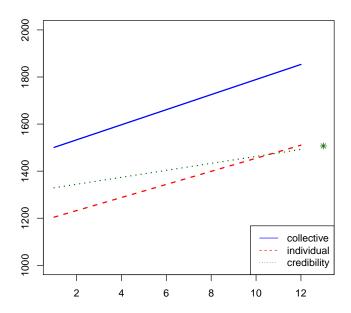


Figure 1: Collective, individual and credibility regression lines for State 4 of the Hachemeister data set. The point indicates the credibility premium.

```
\begin{Sinput}
> predict(fit, newdata = data.frame(time = 13))
\end{Sinput}
\begin{Soutput}
[1] 2437 1651 2073 1507 1759
\end{Soutput}
```

It is well known that the basic regression model has a major drawback: there is no guarantee that the credibility regression line will lie between the collective and individual ones. This may lead to grossly inadequate premiums, as Figure 1 shows.

The solution proposed by Bühlmann and Gisler (1997) is simply to position the intercept at the barycenter of time instead of at time origin (see also Bühlmann and Gisler, 2005, Section 8.4). In mathematical terms, this essentially amounts to using an orthogonal design matrix. By setting the argument adj.intercept to TRUE in the call, cm will automatically fit the credibility regression model with the intercept at the barycenter of time. The resulting regression coefficients have little meaning, but the predictions are sensible:

```
\begin{Sinput}
> fit2 <- cm(~state, hachemeister, regformula = ~ time,</pre>
             regdata = data.frame(time = 1:12),
             adj.intercept = TRUE,
             ratios = ratio.1:ratio.12,
             weights = weight.1:weight.12)
> summary(fit2, newdata = data.frame(time = 13))
\end{Sinput}
\begin{Soutput}
Call:
cm(formula = ~state, data = hachemeister, ratios = ratio.1:ratio.12,
    weights = weight.1:weight.12, regformula = ~time, regdata = data.frame(time = 1:12),
    adj.intercept = TRUE)
Structure Parameters Estimators
  Collective premium: -1675 117.1
 Between state variance: 93783
                               0 8046
 Within state variance: 49870187
Detailed premiums
  Level: state
    state Indiv. coef. Credibility matrix Adj. coef.
                                          -2060.41
          -2062.46
                       0.9947 0.0000
            216.97
                       0.0000 0.9413
                                            211.10
    2
          -1509.28
                       0.9740 0.0000
                                          -1513.59
                       0.0000 0.7630
             59.60
                                              73.23
    3
          -1813.41
                       0.9627 0.0000
                                          -1808.25
            150.60
                       0.0000 0.6885
                                            140.16
          -1356.75
                       0.8865 0.0000
                                          -1392.88
             96.70
                       0.0000 0.4080
                                            108.77
    5
          -1598.79
                       0.9855 0.0000
                                          -1599.89
             41.29
                       0.0000 0.8559
                                             52.22
    Cred. premium
    2457
    1651
    2071
    1597
    1698
\end{Soutput}
```

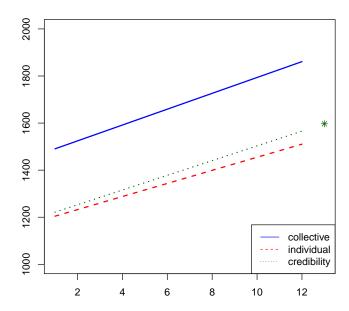


Figure 2: Collective, individual and credibility regression lines for State 4 of the Hachemeister data set when the intercept is positioned at the barycenter of time. The point indicates the credibility premium.

Figure 2 shows the beneficient effect of the intercept adjustment on the premium of State 4.

## References

- H. Belhadj, V. Goulet, and T. Ouellet. On parameter estimation in hierarchical credibility. *ASTIN Bulletin*, 39(2), 2009.
- H. Bühlmann. Experience rating and credibility. *ASTIN Bulletin*, 5:157–165, 1969.
- H. Bühlmann and A. Gisler. Credibility in the regression case revisited. *ASTIN Bulletin*, 27:83–98, 1997.
- H. Bühlmann and A. Gisler. *A course in credibility theory and its applications*. Springer, 2005. ISBN 3-5402575-3-5.

- H. Bühlmann and W. S. Jewell. Hierarchical credibility revisited. *Bulletin of the Swiss Association of Actuaries*, 87:35–54, 1987.
- H. Bühlmann and E. Straub. Glaubgwürdigkeit für Schadensätze. *Bulletin of the Swiss Association of Actuaries*, 70:111–133, 1970.
- M. J. Goovaerts and W. J. Hoogstad. *Credibility theory*. Number 4 in Surveys of actuarial studies. Nationale-Nederlanden N.V., Netherlands, 1987.
- V. Goulet. Principles and application of credibility theory. *Journal of Actuarial Practice*, 6:5–62, 1998.
- C. A. Hachemeister. Credibility for regression models with application to trend. In *Credibility, theory and applications*, Proceedings of the berkeley Actuarial Research Conference on Credibility, pages 129–163, New York, 1975. Academic Press.
- W. S. Jewell. The use of collateral data in credibility theory: a hierarchical model. *Giornale dell'Istituto Italiano degli Attuari*, 38:1–16, 1975.
- E. Ohlsson. Simplified estimation of structure parameters in hierarchical credibility. Presented at the Zurich ASTIN Colloquium, 2005. URL http://www.actuaries.org/ASTIN/Colloquia/Zurich/Ohlsson.pdf.