

▼ Practice 2 PAN

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▼ Q1

From the statement we know that $X_i \in \{0, 1\}$. Such as it is shown in the randomize mechanism, Y_i (RR) can be either the real answer (X_i) or an altered answer ($1 - X_i$) with a certain probability ($1 - \gamma$, in the case of the strategy). If $X_i = 1$, then $1 - X_i = 0$ and viceversa, it seems that the strategy success defining the mechanism.

Now, let's focus at how it works the mechanism:

- After the first coin we have a 50% of chances of obtaining truth.
- After the second coin, the result are going to be true in half of the cases.

Therefore, we know that the final result are going to be tampered in 25% of the cases, i.e. $\gamma = \frac{3}{4}$.

```
p = 0.5 #Probabilidad de haber copiado
gamma = 0.5 + p*0.25 + (1-p)*0.25
print(gamma)
```

0.75

▼ Q2

```
import numpy as np

def avg(v):
    return sum(v)/len(v)

def rr(X):
    Y = []
    for i in X:
        coin = np.random.binomial(n=1, p=0.5)
        # coin = 1 => Face
        if not coin:
            Y.append(i)
        else:
            coin = np.random.binomial(n=1, p=0.5)
            if coin:
                Y.append(0) # Not cheated
            else:
                Y.append(1) # Cheated
```

```

return np.array(Y)

n = 100000
p = 0.2
X = np.random.binomial(n=1, p=p, size=n)

Y = rr(X)

gamma_hat = 1 - avg(Y ^ X)
gamma_hat

0.74979

```

▼ Q3

The value q shows the probability that a random answer is *CHEATED* or not. We already know that:

- In the 50% of the cases the answer does not change and it has a probability p that be *CHEATED*.
- In the other 50%, the probability to be *CHEATED* is 50%.

Thus, we have to take into account two probabilities:

- Do not change the answer and cheat: γp
- Change the answer and do not cheat: $(1 - \gamma) \cdot (1 - p)$

Which gives rise to:

$$\begin{aligned}
 q &= \gamma p + (1 - \gamma) \cdot (1 - p) = 2\gamma p + 1 - p \\
 &\quad - \gamma \iff \\
 \iff \hat{p} &= \frac{q + \gamma - 1}{2\gamma - 1}
 \end{aligned}$$

If we apply the result of the last question ($\gamma = \frac{3}{4}$):

$$\begin{aligned}
 q &= \frac{1}{2}p + \frac{1}{4} \\
 \hat{p} &= 2q - \frac{1}{2}
 \end{aligned}$$

▼ Q4

```

import matplotlib.pyplot as plt

differences = []
n = 10000

```

```

p = 0.2
for i in range(50):
    X = np.random.binomial(n=1, p=p, size=n)
    Y = rr(X)
    q = sum(Y)/len(Y)
    p_hat = 2*q - 1/2
    differences.append(p-p_hat)

print("Average of results: "+str(avg(differences)))

```

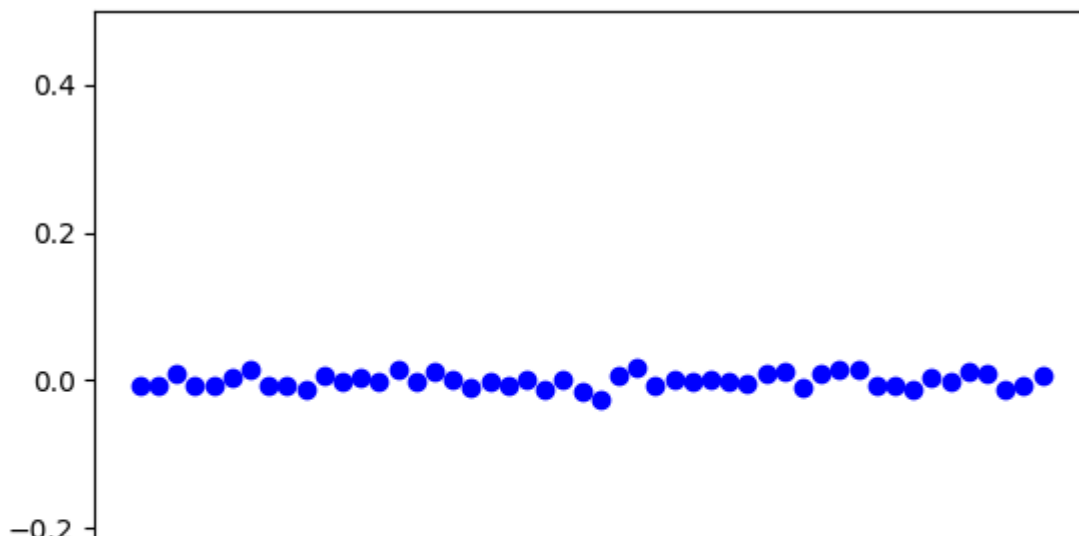
Average of results: -0.00011999999999999567

```

# Plot
plt.ylim(-0.5,0.5)
plt.scatter(range(len(differences)), differences, marker='o', color='b')

```

<matplotlib.collections.PathCollection at 0x792fd52f1990>



Q5

Our goal in this exercise is compute $Var[p - \hat{p}]$. Knowing that p is the average of X and the expectation for \hat{p} ($E[\hat{p}] = p$), its variance in 0 $\Rightarrow Var[p - \hat{p}] = Var[\hat{p}]$.

- From Q5: $\hat{p} = \frac{q+\gamma-1}{2\gamma-1}$

$$\begin{aligned} \text{Var}[\hat{p}] &= \text{Var} \left[\frac{q + \gamma - 1}{2\gamma - 1} \right] \\ &= \left(\frac{1}{2\gamma - 1} \right)^2 \cdot (\text{Var}[q] + \text{Var}[\gamma - 1]) \\ &=^{*1,2} \end{aligned}$$

$$= \left(\frac{1}{2\gamma - 1} \right)^2 \frac{q \cdot (1 - q)}{n} =^{*3}$$

$$\begin{aligned} &= \left(\frac{1}{2\gamma - 1} \right)^2 \\ &\cdot \frac{(2\gamma p + 1 - p - \gamma)(-2\gamma p + p + \gamma)}{n} \end{aligned}$$

$$\begin{aligned} \bullet \text{ For } \gamma = \frac{3}{4}: \text{Var}[p - \hat{p}] &= 4 \cdot \frac{-4p^2 + 4p + 3}{4n} \\ &= \frac{-4p^2 + 4p + 3}{n} \end{aligned}$$

(*₁) As long as $\gamma - 1$ is a constant $\Rightarrow \text{Var}[\gamma - 1] = 0$.

(*₂)

$$\text{Var}[q] = \text{Var} \left[\frac{1}{n} \sum_{i=1}^n Y_i \right] = \frac{1}{n^2}$$

$$\sum_{i=1}^n \text{Var}[Y_i] =^{*4}$$

$$= \frac{1}{n} \text{Var}[Y_i] =^{*5} \frac{q \cdot (1 - q)}{n}$$

(*₃) In Q3 we obtained that $q = \gamma p + (1 - \gamma) \cdot (1 - p) = 2\gamma p + 1 - p - \gamma$

(*₄) The variables Y_i are independent.

(*₅)

$$\begin{aligned} \text{Var}[Y_i] &= E[Y_i^2] - E[Y_i]^2 = q - q^2 \\ &= q(1 - q) \end{aligned}$$

$$\begin{aligned} \bullet \text{ } E[Y_i] &= P[Y_i = 1] \cdot 1 + P[Y_i = 0] \\ &\cdot 0 = P[Y_i = 1] = q \end{aligned}$$

- $E[Y_i^2] = P[Y_i^2 = 1] \cdot 1$
 $+ P[Y_i^2 = 0] \cdot 0 =$
 $= P[Y_i^2 = 1] = P[Y_i = 1] = q$

-

When n grows the second part of $Var[p - \hat{p}]$ tends to decrease in absolute value (the result is closest to 0). Therefore:

$$\lim_{n \rightarrow \infty} (Var[p - \hat{p}]) = 0$$

I.e. the variance of is null and as long as $E[\hat{p}] = p$, then it means that for $n \rightarrow \infty \Rightarrow \hat{p} = p$.

▼ Q6

```
def compute_emp_and_th_vars(p: float, n: int, times: int):
    p_array = []

    for i in range(times):
        X = np.random.binomial(n=1, p=p, size=n)
        p_array.append(avg(X))

    avg_p = avg(p_array)
    emp_var = sum((p_array - avg_p) ** 2)/n
    th_var = (-4*p**2 + p + 3)/n

    return emp_var, th_var

p=0.2; times=5000; n = 1000

emp_var, th_var = compute_emp_and_th_vars(p, n, times)
print('Result for n=%i: \n'
      '\t - Empirical variance: %f\n'
      '\t - Theoretical variance: %f\n' % (n, emp_var, th_var))

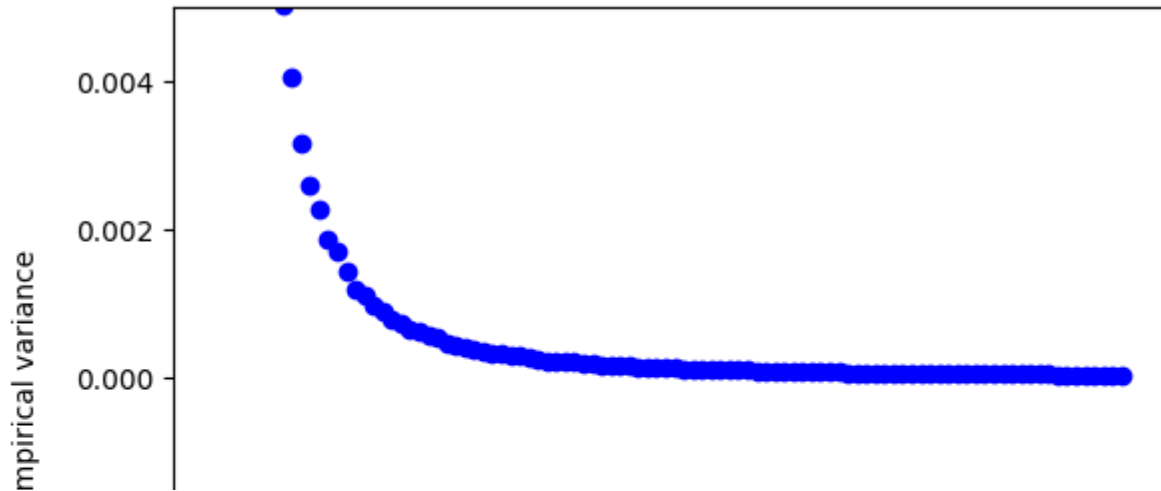
n_array = np.linspace(50, 5000, 100)
p_plot = []
for i in n_array:
    emp_var, _ = compute_emp_and_th_vars(p=p, n=int(i), times=times)
    p_plot.append(emp_var)

plt.xlabel('n values')
plt.ylabel('Empirical variance')
plt.scatter(n_array, p_plot, marker='o', color='b', label='p-p_hat')
plt.ylim(-0.005,0.005)
```

Result for n=1000:

- Empirical variance: 0.000809
- Theoretical variance: 0.003040

(-0.005, 0.005)



▼ Q7

On the one hand, when $\gamma = 1$ the result is deterministic and is the truth, and viceversa for $\gamma = 0$ (the result is always tamped):

$$\begin{aligned}
 q &= \gamma p + (1 - \gamma) \cdot (1 - p) \\
 \text{Var}[p - \hat{p}] &= \left(\frac{1}{2\gamma - 1} \right)^2 \\
 &\cdot \frac{(2\gamma p + 1 - p - \gamma)(-2\gamma p + p + \gamma)}{n} \\
 \gamma = 1 &\Rightarrow q = p, \quad \text{Var}[p - \hat{p}] \\
 &= \frac{p(1 - p)}{n} \\
 \gamma = 0 &\Rightarrow q = 1 - p \quad \text{Var}[p - \hat{p}] \\
 &= \frac{p(1 - p)}{n}
 \end{aligned}$$

We can recovery the number of students who *CHEATED* (the best accuracy but no privacy).

On the other mand, when $\gamma = \frac{1}{2}$ we don't have any clue about the real answers:

$$\gamma = \frac{1}{2} \Rightarrow q = \frac{1}{2}, \quad Var[p - \hat{p}] = \infty$$

We have that q is independent of p and the variance is ∞ achiving the maximum privacy but the lower accuracy.

▼ Q8

I know, as a professor, the values of Y_i when students trust me. When they don't trust me, Prof. Trustworthy give me q . The values of Y_i are anonymous so the order of the answers are not important an we can already compute the number of $Y_i = \textit{CHEATED}$ as long as we know the number of students and q :

$$\sum_{i=1}^n Y_i = q/n$$

Therefore, students does not increase their security because we can obtain the same information in both scenarios.