
Redes de Computadores

Delay Models in Computer Networks

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-
- » *What are the common multiplexing strategies?*
 - » *What is a Poisson process?*
 - » *What is the Little theorem?*
 - » *What is a queue?*
 - » *What is the meaning of service time $1/\mu$ in a queue of packets?*
 - » *What is the meaning of traffic intensity ρ in a queue model?*
 - » *What is the probability of a M/M/1 queue being in a given state n ?*
 - » *What is the mean number of clients in a M/M/1 queue? What is the mean waiting time in a M/M/1 queue? What is the relationship between N and ρ in a M/M/1 queue?*
 - » *What are the differences between M/M/1 and M/G/1 queues? How to estimate mean number of packets and mean delay in a M/G/1 queue?*
 - » *How to model a network of transmission lines? How to calculate the mean number of packets and mean delay in this case?*
 - » *What is a Jackson Network? Why is it important?*

Multiplexing Traffic on a Link



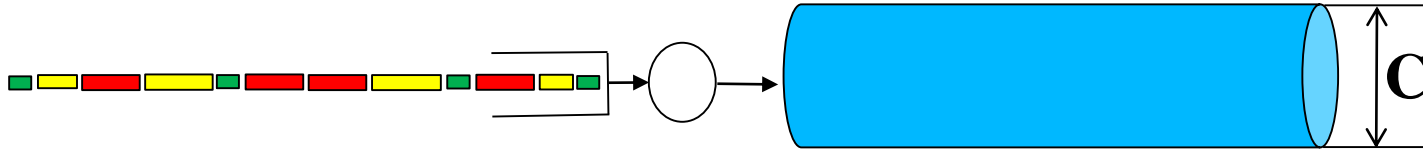
- ♦ Communication link
 - » Bit pipe with a given capacity C (bit/s)
 - » Link capacity \rightarrow rate at which bits are transmitted to the link
 - » Link may transport multiplexed traffic streams

- ♦ Multiplexing strategies
 - » Statistical Multiplexing
 - » Frequency Division Multiplexing
 - » Time Division Multiplexing

- ♦ Multiplexing strategy affects traffic delay

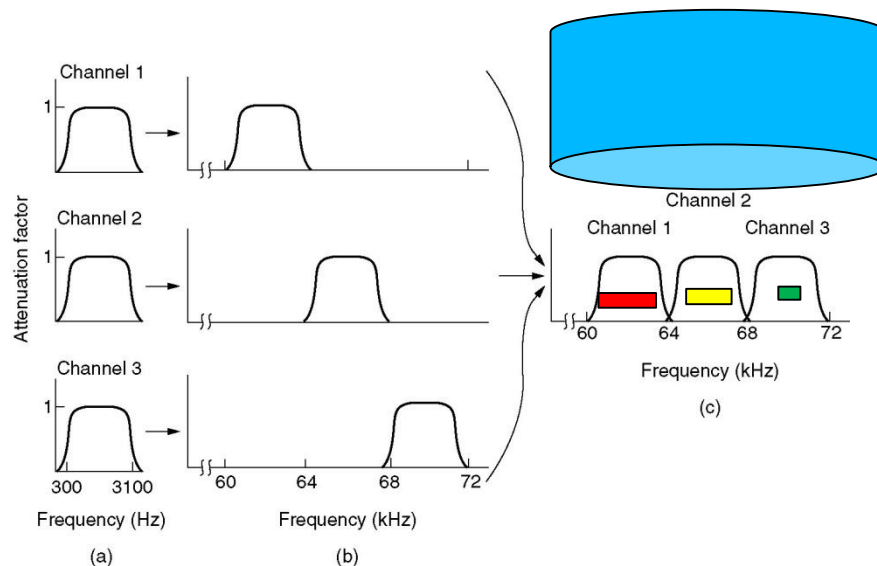
Statistical Multiplexing

- ◆ Packets of all traffic streams merged in a single queue
- ◆ Packets transmitted on a first-come first-served basis
- ◆ Time required to transmit a packet of length $L \rightarrow T_{\text{frame}} = L/C$



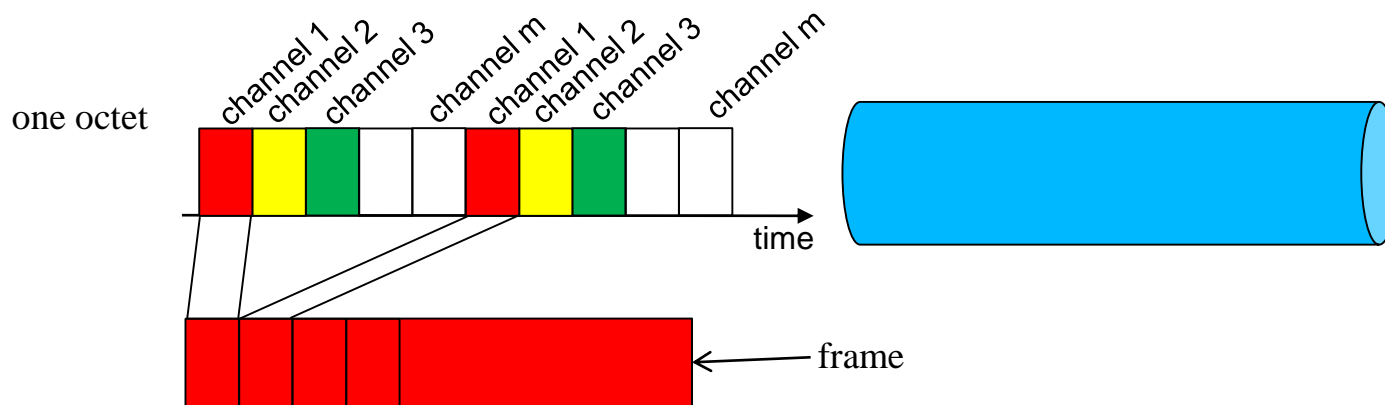
FDM – Frequency Division Multiplexing

- ♦ Link capacity C subdivided into m portions
- ♦ Channel bandwidth W subdivided into m channels of W/m Hz
- ♦ Capacity of each channel $\rightarrow C/m$
- ♦ Time required to transmit a packet of length $L \rightarrow T_{\text{frame}} = Lm/C$

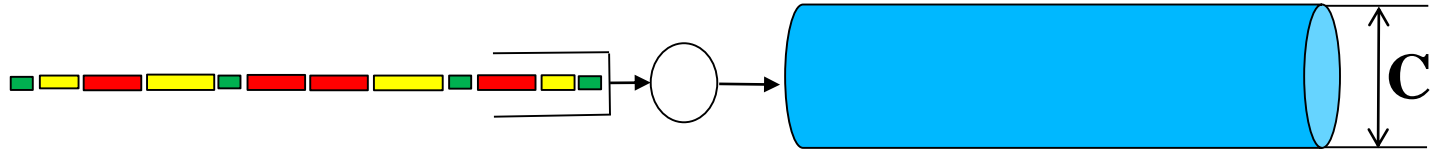


TDM – Time Division Multiplexing

- ♦ Time axis divided into m slots of fixed length
(usually one octet long)
- ♦ Communication \rightarrow m channels with capacity C/m
- ♦ Time required to transmit a packet of length $L \rightarrow T_{\text{frame}} = Lm/C$



Delay on Computer Networks



♦ Delay

- » Important performance parameter in computer networks
- » Characterized using queue models

♦ Queue model

- » Customers arrive at random times to obtain service
- » Customer → packet to be transmitted through a link
- » Serve a packet = transmit a packet
- » Service time → **packet transmission time** = $T_{\text{pac}(\text{frame})} = L/C$

♦ Queue models enable the quantification of

- » Average number of customers/packets in the network
- » Average delay per packet → waiting plus service times

Computer Networks Modeled as Queue Networks

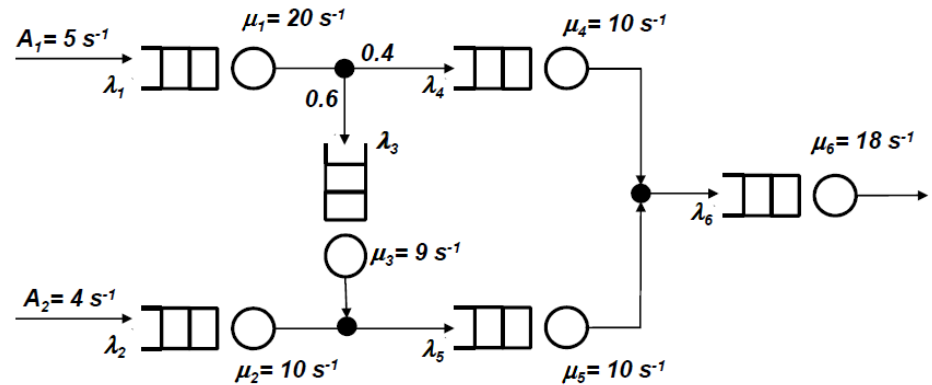
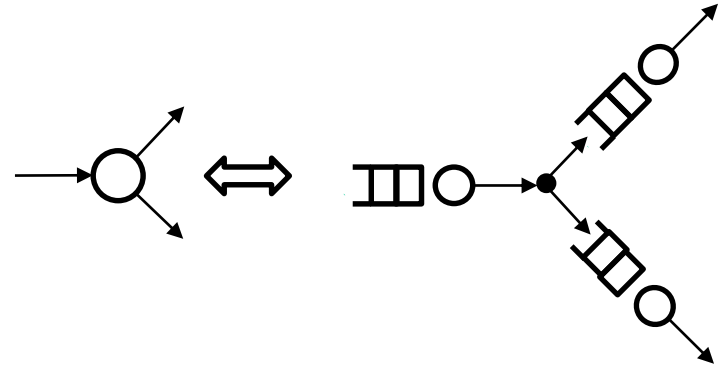
Mobile network

Global ISP

Home network

Regional ISP

Institutional network



Poisson Distribution and Poisson Process

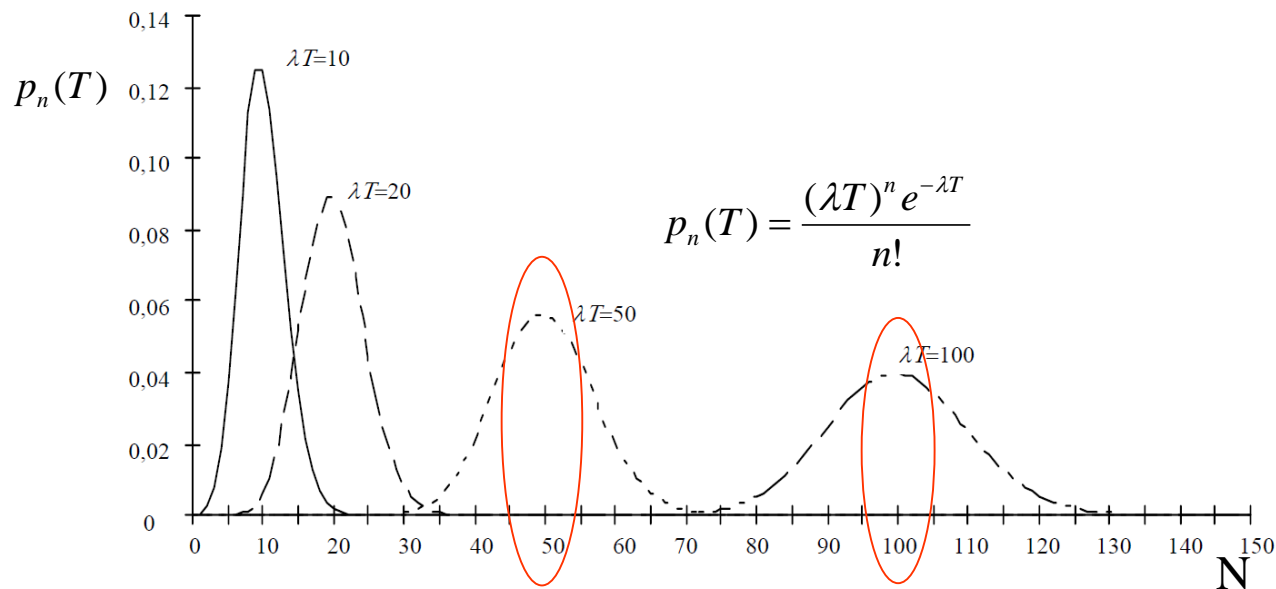
- ♦ **Poisson distribution** with parameter m

$$P[N = n] = p_n = \frac{m^n e^{-m}}{n!}, \quad n = 0, 1, \dots \quad E[N] = \text{Var}[N] = m$$

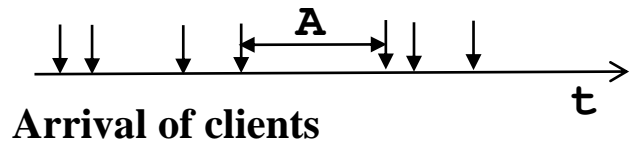
- ♦ **Poisson process**

» $\lambda T = m$, (e.g. $\lambda \rightarrow$ arrivals/s)

» $P[\text{n arrivals in interval } T] = p_n(T) = p_n = \frac{(\lambda T)^n e^{-\lambda T}}{n!} \quad E[N] = \text{Var}[N] = \lambda T$



Inter-Arrival Interval A – Statistical Characterization



A – time interval between the
arrival of consecutive clients

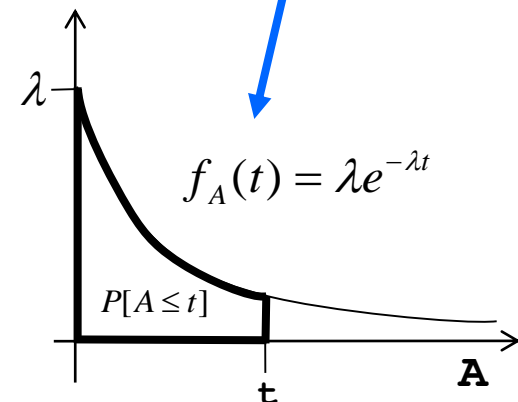
$$F_A(t) = P[A \leq t] = 1 - P[A > t] = 1 - p_0(t) = 1 - e^{-\lambda t}$$

$$f_A(t) = pdf = \frac{\partial F_A(t)}{\partial t} = \lambda e^{-\lambda t}$$

Exponential distribution

$$E[A] = 1/\lambda$$

$$Var[A] = 1/\lambda^2$$



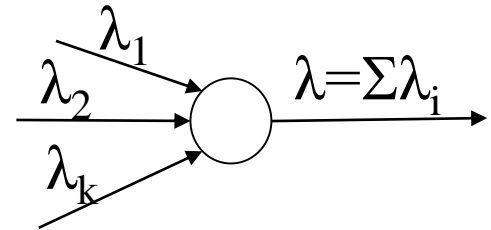
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- ♦ What is the difference between
Deterministic arrivals and
Poisson arrivals?



Markov Process - Properties

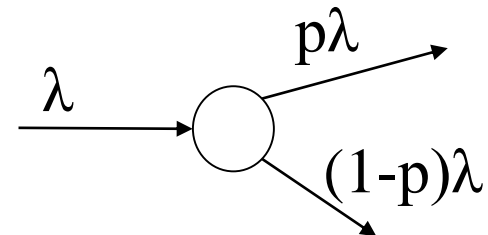
♦ Merging Property

- » A_1, A_2, \dots, A_k are independent Poisson Processes with rates $\lambda_1, \lambda_2, \dots, \lambda_k$
- » $A = \sum A_i$ still is a Poisson process, with rate $\lambda = \sum \lambda_i$



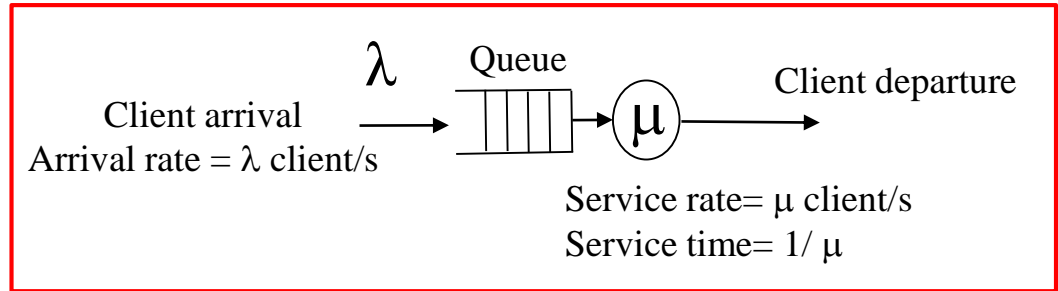
♦ Splitting property

- » Packets arrive to a router according to a Poisson Process (A, λ)
- » They are routed randomly to two output lines with probabilities **p** and **1-p**
- » Packets leaving the router still are Poisson Processes, characterized by $(A, p\lambda)$ and $(A, (1-p)\lambda)$



Queue Model

- ♦ Queue – model used for
 - » Customers waiting in line
 - » Packets in a network
- ♦ Used to determine
 - » Average number of clients in the system $\rightarrow N$
 - » Average delay experienced by a client $\rightarrow T$
- ♦ Queue characterized in terms of
 - » λ - arrival rate of client (average number of clients per time unit)
 - » μ - service rate (average number of clients the server processes per time unit)
 - » $\rho = \lambda / \mu$ – traffic intensity (occupation of the server)
- ♦ Kendall notation $\rightarrow \mathbf{A/S/s/K}$
 - » A – arrival statistical process
 - » S – service statistical process
 - » s – number of servers
 - » K – capacity of the system in buffers



Little's Theorem

◆ $N = \lambda T$

- » N - average number of clients in a system
- » T - average amount of time a client spends in the system
- » λ - arrival rate of clients to the system

◆ $T = T_w + T_s$

- » T_w - time a client waits in the queue for being served
- » T_s - service time

◆ $N = N_w + N_s$

- » N_w - number of clients waiting in the queue for being served
- » N_s - number of clients being served

◆ $N_w = \lambda T_w$

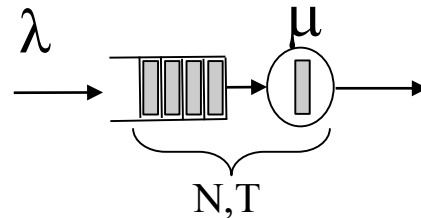
$$N_w = \lambda T_w \rightarrow T_w = N_w / \lambda$$

- ♦ The (mean) time a client has to wait before being served (T_w) depends on the number of clients waiting (N_w) and on the arrival rate of clients (λ)
- ♦ No dependence on the service rate?!
- ♦ Can you explain it?



Little's Theorem

- ♦ Can be applied to a single Queue



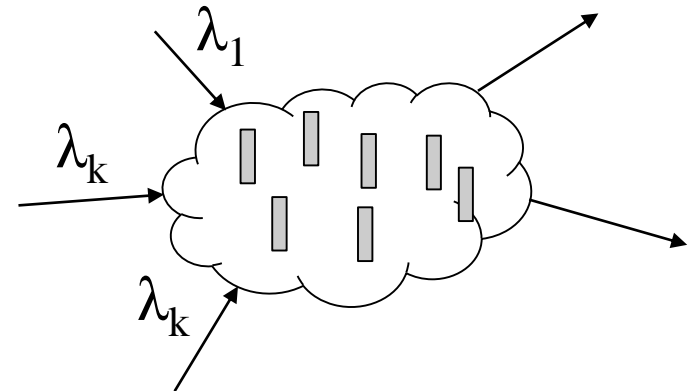
- ♦ Can be applied to a complex system

- » For each stream $i \rightarrow N_i = \lambda_i T_i$

- » For the system:

$$\lambda = \sum \lambda_i \quad N = \sum N_i$$

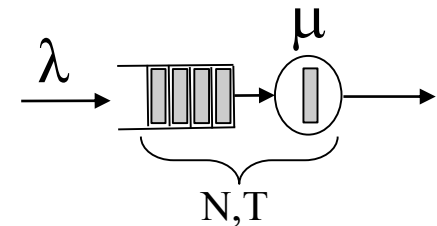
$$T = (\sum N_i) / (\sum \lambda_i) \rightarrow T = N / \lambda$$



M/M/1 Queue

♦ M/M/1

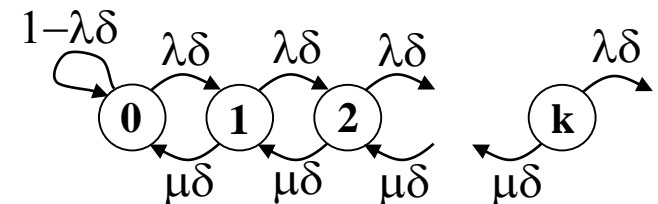
- » Poisson arrival, exponential service time



♦ Modeled by a Markov Chain

- » State \mathbf{k} - k clients in the queue
- » $p(i, j)$ – probability of transition from state i to state j
- » When $\delta \rightarrow 0$

$$\begin{aligned} p(i, i+1) &= \lambda\delta & p(i, i-1) &= \mu\delta \\ p(i, i) &= 1 - \lambda\delta - \mu\delta & p(0, 0) &= 1 - \lambda\delta \\ p(i, j) &= 0 \text{ for other values } i, j \end{aligned}$$



» Birth-death chain

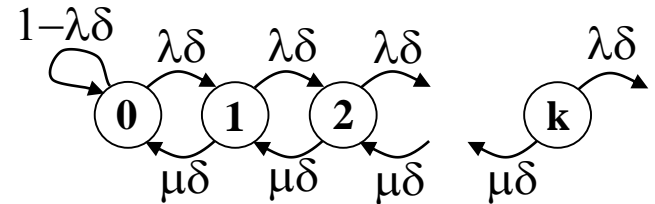
- Transitions between adjacent states
- $\lambda\delta$ and $\mu\delta$ become flow rates between states

$$\begin{aligned} p(i, i+1) &= p_1(\delta) = (\lambda\delta)e^{-\lambda\delta} \approx \lambda\delta \\ p(0, 0) &= p_0(\delta) = e^{-\lambda\delta} \approx 1 - \lambda\delta \end{aligned}$$

M/M/1 Queue – Equilibrium Analysis

- ♦ $P(j)$ – probability of the Markov chain be in state j
- ♦ Markov Chain - global balance equations

$$P(j) \sum_{i \neq j}^{\infty} p(j, i) = \sum_{i \neq j}^{\infty} P(i) p(i, j)$$



- ♦ In the case of M/M/1

$$P(0)\lambda\delta = P(1)\mu\delta \Rightarrow P(1) = \rho P(0)$$

$$P(2) = \rho P(1) = \rho^2 P(0)$$

$$P(n) = \rho^n P(0)$$

$$\sum_{i=0}^{\infty} P(i) = 1$$

$$\sum_{i=0}^{\infty} \rho^i P(0) = \frac{P(0)}{1 - \rho} = 1$$

$$P(0) = 1 - \rho$$

$$P(n) = \rho^n (1 - \rho)$$

M/M/1 Queue

- ♦ Average Queue size N

$$N = \sum_{n=0}^{\infty} nP(n) = \sum_{n=0}^{\infty} n\rho^n(1-\rho) = \frac{\rho}{1-\rho} \quad N = \sum_{n=0}^{\infty} nP(n) = \frac{\rho}{1-\rho} = \frac{\lambda/\mu}{1-\lambda/\mu} = \frac{\lambda}{\mu-\lambda}$$

- ♦ Average amount of time the client spends in the system, T

» Little's formula, $T=N/\lambda \quad \Rightarrow \quad T = \frac{1}{\mu-\lambda}$

- ♦ Average waiting time $T_w \Rightarrow T_w = T - T_s = \frac{1}{\mu-\lambda} - \frac{1}{\mu} = \frac{\rho}{\mu(1-\rho)}$

- ♦ Average number of clients waiting in the queue, N_w

$$N_w = T_w \lambda = \frac{\lambda}{\mu-\lambda} - \frac{\lambda}{\mu} = N - \rho$$

M/M/1 Queue – $N=f(\rho)$

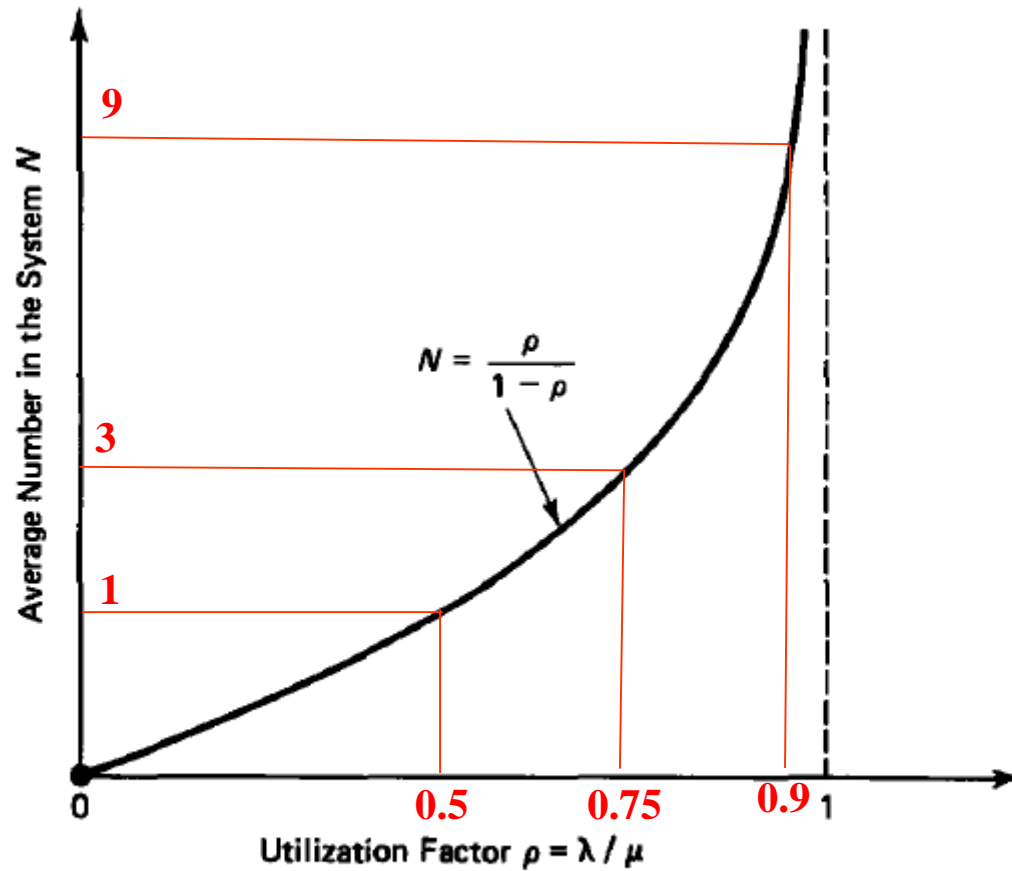



Figure 3.6 The average number in the system versus the utilization factor in the *M/M/1* system. As $\rho \rightarrow 1$, $N \rightarrow \infty$.

♦ M/M/1: $\rho=0.9 \rightarrow N=9$

♦ Why have clients to wait if the server is busy only 90% of his time? 

♦ What would happen for D/D/1, $\rho=0.9$? 

Packet Length, Service Time, Speed

- » 100 packet/s are required to be transmitted through a link
- » Packets arrive according to a Poisson process
- » Packet lengths are exponentially distributed $\rightarrow E[L]=10^4$ bit/packet
- » Link has capacity $C=10$ Mbit/s

◆ Then

- » Arrival rate: $\lambda=100$ packet/s
- » Service rate: $\mu=C/E[L]=10^7/10^4=10^3$ packet/s
- » $\rho=\lambda/\mu=0.1$, $N=\rho/(1-\rho)=1/9$, $T=N/\lambda=1/900$ s

◆ Assume now: $\lambda'=10\lambda$ and $C'=10C \rightarrow \mu'=10C/E[L]=10\mu$

- » Then $\rho'=\rho$ and $N'=N$ but $T'=N'/\lambda'=T/10$

The speed of the system increases!

M/M/1/B Queue

- ♦ M/M/1 queue has limited capacity (B buffers)
 - » Packets can be lost
 - » Probability of packet being lost = $P(B)$ → Queue is full
- ♦ Analysis similar to M/M/1

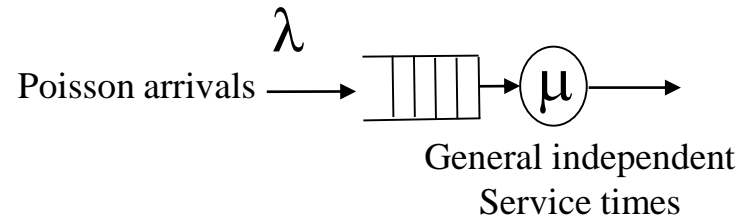
$$\sum_{i=0}^B P(i) = 1 \qquad P(n) = \rho^n P(0)$$

$$P(0) = \frac{1 - \rho}{1 - \rho^{B+1}} \qquad P(B) = \frac{(1 - \rho)\rho^B}{1 - \rho^{B+1}}$$

- ♦ Particular cases

$$\rho = 1, \quad P(B) = \frac{1}{B+1} \qquad \rho \gg 1, \quad P(B) \approx \frac{\rho - 1}{\rho} = \frac{\lambda - \mu}{\lambda}$$

M/G/1 Queue



- ♦ Poisson arrivals at rate λ
- ♦ Service time X has arbitrary distribution with given $E[X]$ and $E[X^2]$
 - » Service times Independent and Identically Distributed (IID)
 - » Independent of arrival times
 - » $E[\text{service time}] = E[X] = 1/\mu$
 - » Single Server queue

M/G/1 Queue – Pollaczek-Khinchin (P-K) Formula

$$T_w = \frac{\lambda E[X^2]}{2(1-\rho)}$$

- ♦ where $\rho = \lambda/\mu = \lambda E[X]$ = line utilization
- ♦ From Little's Theorem
 - » $N_w = \lambda T_w$
 - » $T = T_w + E[X] = T_w + 1/\mu$
 - » **$N = \lambda T = \lambda(T_w + 1/\mu) = N_w + \rho$**

M/G/1 Queue – Proof of (P-K) Formula

$$T_w = \frac{\lambda E[X^2]}{2(1-\rho)}$$

◆ Let

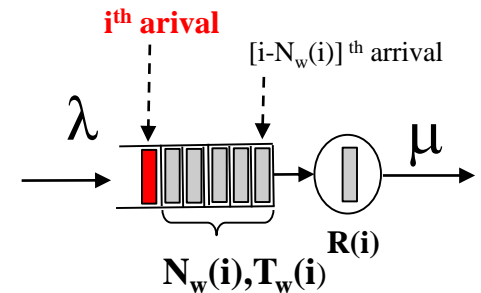
- $T_w(i)$ - waiting time in queue of i^{th} arrival
- $R(i)$ – residual service time seen by the i^{th} arrival
- $N_w(i)$ – number of clients found in queue by the i^{th} arrival
- $X(i)$ – service time of the i^{th} arrival

$$T_w(i) = \sum_{j=i-N_w(i)}^{i-1} X(j) + R(i)$$

$$E[T_w(i)] = T_w = E[N_w(i)] \times E[X(i)] + E[R(i)] = \frac{N_w}{\mu} + E[R(i)]$$

» Using Little's formula

$$T_w = \frac{\lambda T_w}{\mu} + E[R(i)] \qquad T_w = \frac{E[R(i)]}{1-\rho}$$



M/G/1 Queue – Proof of (P-K) Formula

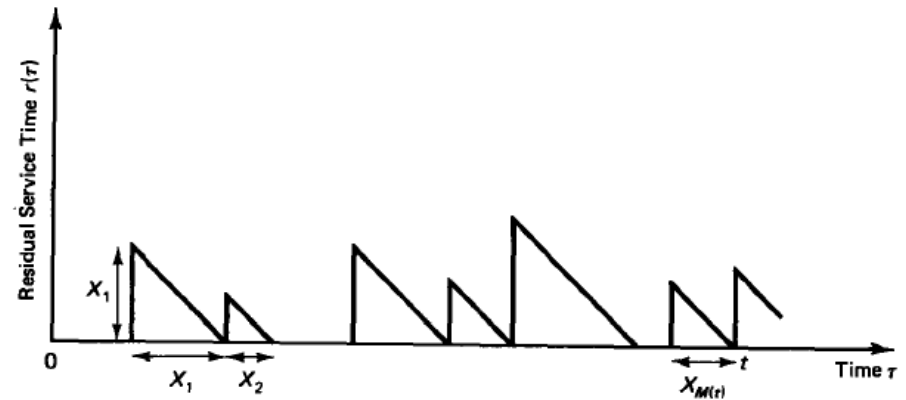


Figure 3.10 Derivation of the mean residual service time. During period $[0, t]$, the time average of the residual service time $r(\tau)$ is

$M(t)$ – number of clients served by time t

$$E[R(i)] = R_t = \frac{1}{t} \int_0^t r(\tau) d\tau = \frac{1}{t} \sum_{i=1}^{M(t)} \frac{X_i^2}{2} = \frac{M(t)}{2t} \sum_{i=1}^{M(t)} \frac{X_i^2}{M(t)}$$

$$t \rightarrow \infty, \quad \frac{M(t)}{t} = \lambda = \text{arrival rate} = \text{departure rate}$$

$$E[R(i)] = \frac{\lambda}{2} \sum_{i=1}^{M(t)} \frac{X_i^2}{M(t)} = \frac{\lambda}{2} \times E[X^2]$$

$$T_w = \frac{E[R(i)]}{1 - \rho}$$

$$T_w = \frac{\lambda E[X^2]}{2(1 - \rho)}$$

M/G/1 Examples

- ♦ Case M/M/1

- » $E[X] = 1/\mu$; $E[X^2] = 2/\mu^2$

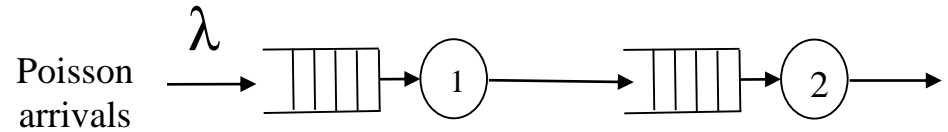
$$T_w = \frac{\lambda}{\mu^2(1-\rho)} = \frac{\rho}{\mu(1-\rho)}$$

- ♦ Case M/D/1

- » Deterministic, constant service time $1/\mu$

- » $E[X] = 1/\mu$; $E[X^2] = 1/\mu^2$

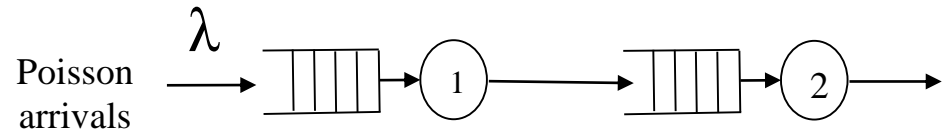
$$T_w = \frac{\lambda}{2\mu^2(1-\rho)} = \frac{\rho}{2\mu(1-\rho)}$$



- ♦ Assume Queue 1 is M/D/1.
- ♦ Can the arrival of packets to Queue 2 be described as a Poisson process?

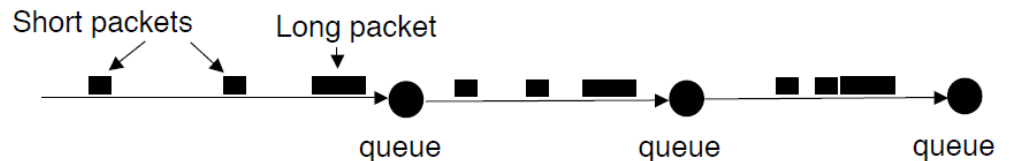
Networks of Transmission Lines - Problems

♦ Case 1



- » Arrival to $Q_1 \rightarrow$ Poisson, λ
- » Assume constant packet length $\rightarrow Q_1 = M/D/1$
- » Arrival to Q_2 is not Poisson; $\lambda_2 < \mu_2 \rightarrow 1/\lambda_2 > 1/\mu_2$
 \rightarrow no waiting at Q_2

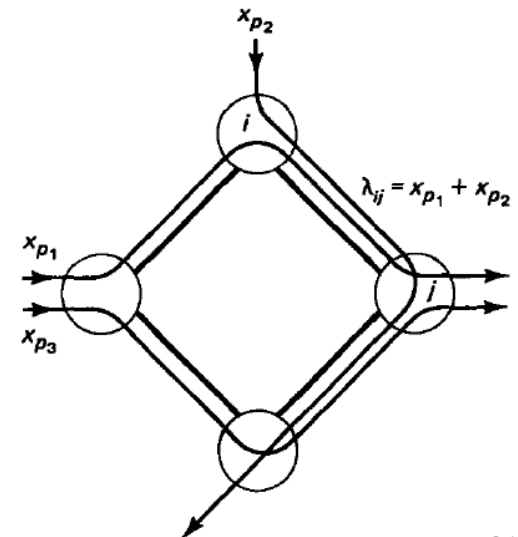
♦ Case 2



- » $Q_1 = M/M/1$
- » arrival to Q_2 strongly related to packet length
- » long packets require long service at each node
- » shorter packets will catch up long packets \rightarrow interarrival times change
 $\rightarrow Q_2$ cannot be modeled as $M/M/1$

Kleinrock Independence Approximation

- ◆ Merging several packet streams on a transmission line
restores independence of interarrival times and packet lengths
- ◆ M/M/1 can be used to model each communication link
- ◆ Approximation good for
 - » systems involving Poisson stream arrivals at the entry points
 - » packet lengths nearly exponentially distributed
 - » densely connected networks
 - » Moderate to heavy traffic loads



Kleinrock Independence Approximation

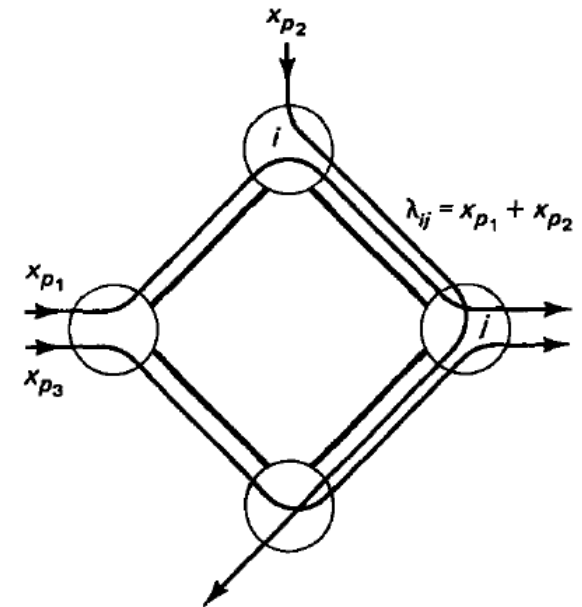
- ♦ Let
 - » x_p = arrival rate of packets along path p
 - » λ_{ij} = arrival rate of packets to link (i,j)
 - » μ_{ij} = service rate on link (i,j)
- ♦ Link queues \rightarrow independent M/M/1 queues

$$\lambda_{ij} = \sum_{\text{all } p \text{ traversing link}(i,j)} x_p \quad \rho_{ij} = \frac{\lambda_{ij}}{\mu_{ij}} \quad N_{ij} = \frac{\rho_{ij}}{1 - \rho_{ij}}$$

- ♦ And
 - » N = Average number of packets in network
 - » T – Average packet delay in network

$$N = \sum_{i,j} N_{ij} \quad \lambda = \sum_{\text{all paths } p} x_p = \text{total external arrival rate}$$

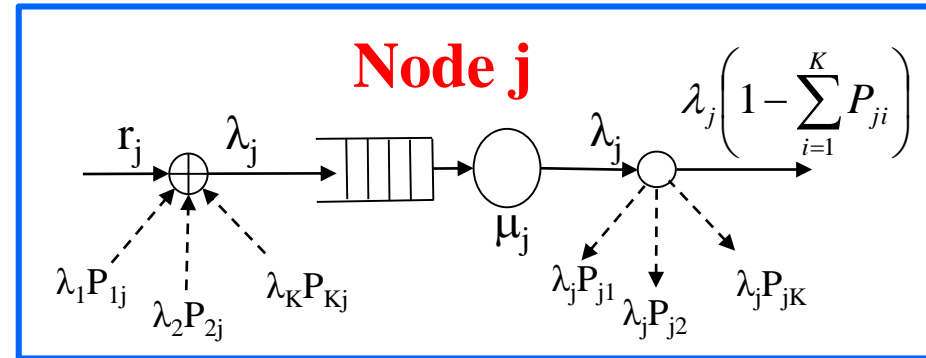
$$T = \frac{N}{\lambda}$$



Jackson Networks

- ◆ Arrival rate at node j

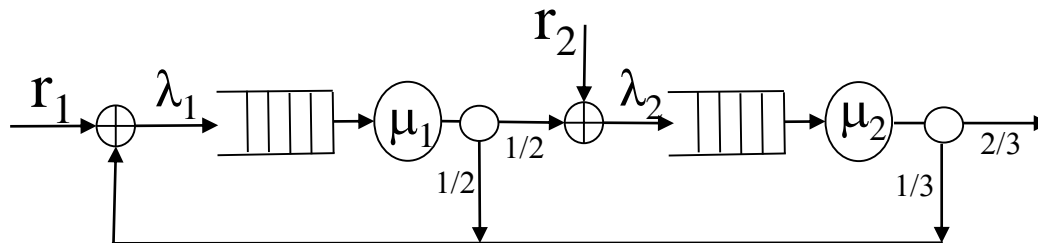
$$\lambda_j = r_j + \sum_{i=1}^K \lambda_i P_{ij} \quad , j = 1, 2, \dots, K$$



- ◆ Independent routing of packets

- » When a packet leaves node i it comes to node j with probability P_{ij}
- » Packets can loop inside network
- » Packet leaves the system at node j with probability

$$P = 1 - \sum_{i=1}^K P_{ji}$$

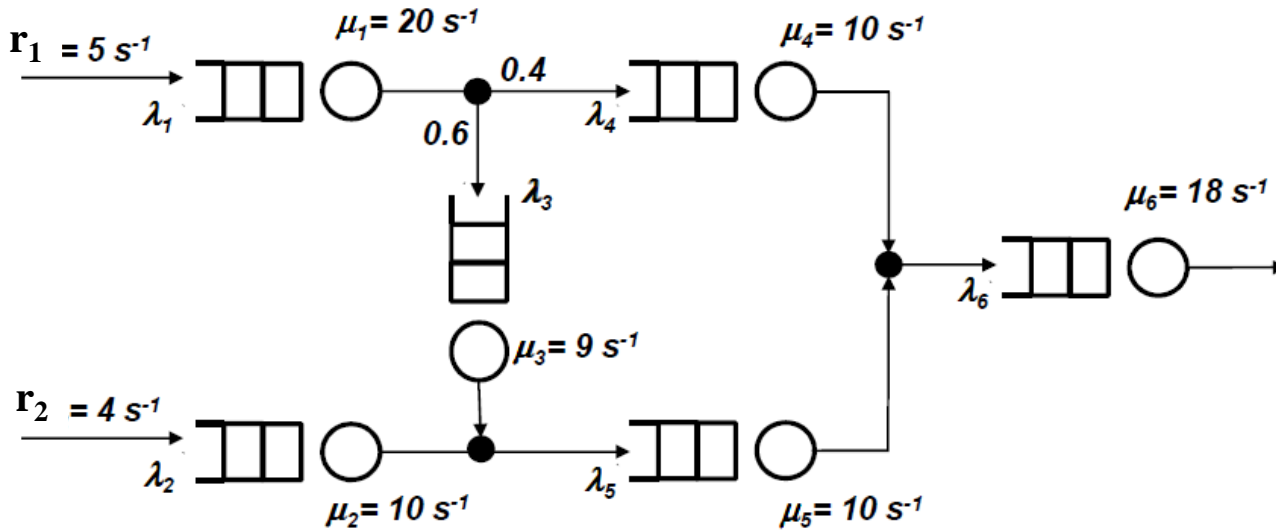


Jackson Networks

- ♦ Let the state of the system be defined by $\vec{n} = (n_1, n_2, \dots, n_K)$
 n_j – number of clients in Q_j
- ♦ Jackson's theorem: $P(\vec{n}) = \prod_{j=1}^K P_j(n_j) = \prod_{j=1}^K \rho_j^{n_j} (1 - \rho_j)$, where $\rho_j = \frac{\lambda_j}{\mu_j}$
 - » State of Q_j (n_j) is independent $\left(\prod_{j=1}^K \right)$ of state of other queues
 - » Similar to independent M/M/1 queues!
 - » Similar to Kleinrock's independence
- ♦ Again, by Little's theorem

$$N_j = \frac{\rho_j}{1 - \rho_j} \quad N = \sum_{j=1}^K N_j \quad \lambda = \sum_{j=1}^K r_j \quad T = \frac{N}{\lambda}$$

Jackson Network - Example



$$\lambda = \sum_{i=1}^6 r_i = 9 \text{ s}^{-1}$$

$$N = \sum_{i=1}^6 N_i = 5.08$$

$$T = \frac{N}{\lambda} = \frac{5.08}{9} = 0.56 \text{ s}$$

Queue i	\mathbf{r}_i (s^{-1})	λ_i (s^{-1})	μ_i (s^{-1})	$\rho_i = \lambda_i/\mu_i$	$\mathbf{N}_i = \rho_i/(1 - \rho_i)$
1	5	5	20	0.25	0.33
2	4	4	10	0.40	0.67
3	-	3	9	0.33	0.50
4	-	2	10	0.20	0.25
5	-	7	10	0.70	2.33
6	-	9	18	0.50	1

Homework

1. Review slides
2. Read *Bertsekas&Gallager*
 - » Sections 3.1, 3.2, 3.3, 3.5, 3.6, 3.8
3. Answer questions at moodle