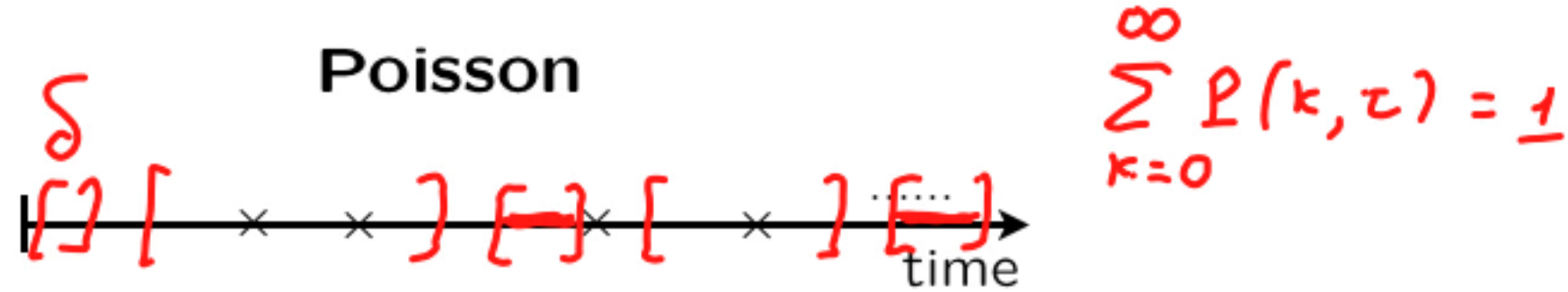


LECTURE 22: The Poisson process

- Definition of the Poisson process
 - applications
- Distribution of number of arrivals
- The time of the k th arrival
- Memorylessness
- Distribution of interarrival times

Definition of the Poisson process



- Numbers of arrivals in disjoint time intervals are **independent**

$P(k, \tau)$ = Prob. of k arrivals in interval of duration τ

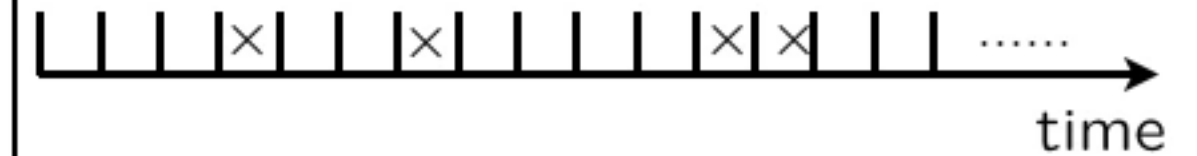
- Small interval probabilities:**

For VERY small δ :

$$P(k, \delta) \approx \begin{cases} 1 - \lambda\delta & \text{if } k = 0 \\ \lambda\delta & \text{if } k = 1 \\ 0 & \text{if } k > 1 \end{cases} \quad P(k, \delta) = \begin{cases} 1 - \lambda\delta + O(\delta^2) & \text{if } k = 0 \\ \lambda\delta + O(\delta^2) & \text{if } k = 1 \\ 0 + O(\delta^2) & \text{if } k > 1 \end{cases}$$

$$\frac{O(\delta^2)}{\delta} \xrightarrow{\delta \rightarrow 0} 0$$

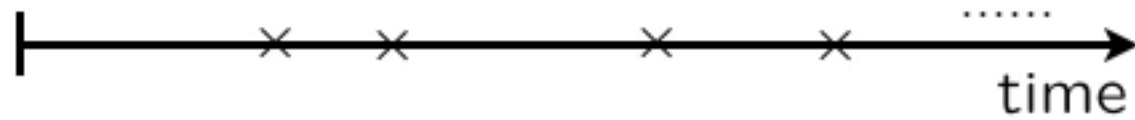
Bernoulli



- Independence
- Time homogeneity:**
Constant p at each slot

λ : "arrival rate"

Applications of the Poisson process



- Deaths from horse kicks in the Prussian army (1898)
- Particle emissions and radioactive decay
- Photon arrivals from a weak source
- Financial market shocks
- Placement of phone calls, service requests, etc. •



Siméon Denis Poisson
(1781-1840)

The Poisson PMF for the number of arrivals



- N_τ : arrivals in $[0, \tau]$ $P(k, \tau) = P(N_\tau = k)$

$n = \tau/\delta$ intervals/slots of length δ *← small*

$P(\text{some slot contains two or more arrivals})$

$$\leq \sum_i P(\text{slot } i \text{ has } \geq 2 \text{ arrivals})$$

$$= \frac{\tau}{\delta} O(\delta^2) \xrightarrow{\delta \rightarrow 0} 0$$

$P(k \text{ arrivals in Poisson}) \approx P(k \text{ slots have arrival})$

$N_\tau \approx \text{binomial} \quad p = \lambda\delta + O(\delta^2)$

$np = \lambda\tau + O(\delta) \approx \lambda\tau$

Bernoulli

$$p_S(k) = \frac{n!}{(n-k)!k!} \cdot p^k (1-p)^{n-k},$$

$$k = 0, \dots, n$$

$$\lambda = np \quad n \rightarrow \infty \quad p \rightarrow 0$$

For fixed $k = 0, 1, \dots$,

$$p_S(k) \rightarrow \frac{\lambda^k}{k!} e^{-\lambda},$$

$$P(k, \tau) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}, \quad k = 0, 1, \dots$$

Mean and variance of the number of arrivals

$$P(k, \tau) = \mathbf{P}(N_\tau = k) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}, \quad k = 0, 1, \dots$$

$$\mathbf{E}[N_\tau] = \sum_{k=0}^{\infty} k \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!} = \dots = \lambda\tau$$

$$N_\tau \approx \text{Binomial}(n, p)$$

$$n = \tau/\delta, \quad p = \lambda\delta + O(\delta^2)$$

$$\mathbf{E}[N_\tau] \approx np \approx \lambda\tau$$

$$\text{var}(N_\tau) \approx np(1-p) \approx \lambda\tau$$

$$\mathbf{E}[N_\tau] = \lambda\tau$$

$$\text{var}(N_\tau) = \lambda\tau$$

$$\lambda = \frac{\mathbf{E}[N_\tau]}{\tau}$$

Example

- You get email according to a Poisson process, at a rate of $\lambda = 5$ messages per hour.

$$E[N_\tau] = \lambda\tau$$

$$\text{var}(N_\tau) = \lambda\tau$$

- Mean and variance of mails received during a day = $5 \cdot 24$
- $P(\text{one new message in the next hour}) = P(1,1) = 5e^{-5}$

$$P(k, \tau) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}, \quad k = 0, 1, \dots$$

- $P(\text{exactly two messages during each of the next three hours}) =$

$$\begin{array}{|c|c|c|} \hline 2 & 2 & 2 \\ \hline \end{array} \quad \left(P(2,1) \right)^3 = \left(\frac{5^2 e^{-5}}{2} \right)^3$$

The time T_1 until the first arrival



$$P(k, \tau) = \frac{(\lambda \tau)^k e^{-\lambda \tau}}{k!}, \quad k = 0, 1, \dots$$

- Find the CDF: $P(T_1 \leq t) =$

$$= 1 - P(T_1 > t) = 1 - P(0, t) = 1 - e^{-\lambda t}$$

$$f_{T_1}(t) = \lambda e^{-\lambda t}, \quad \text{for } t \geq 0$$

Exponential(λ)

Memorylessness: conditioned on $T_1 > t$,
the PDF of $T_1 - t$ is again exponential

The time Y_k of the k th arrival

$$P(k, \tau) = \frac{(\lambda \tau)^k e^{-\lambda \tau}}{k!}, \quad k = 0, 1, \dots$$

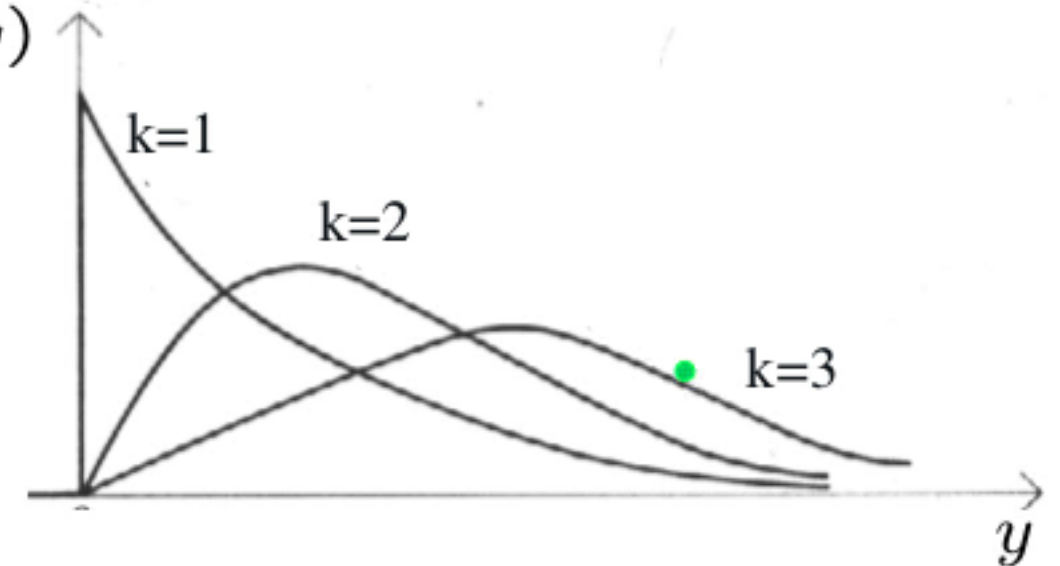
- Can derive its PDF by first finding the CDF
- More intuitive argument:

$$P(Y_k \leq \gamma) = \sum_{n=k}^{\infty} P(n, \gamma)$$

$$\begin{aligned} f_{Y_k}(y) \delta &\approx P(y \leq Y_k \leq y + \delta) = \\ &\approx P(k-1, y) \lambda \delta \\ &\quad + P(k-2, y) O(\delta^2) \\ &\quad + P(k-3, y) O(\delta^2) \end{aligned}$$

$$\frac{(\lambda y)^{k-1} e^{-\lambda y}}{(k-1)!} \lambda$$

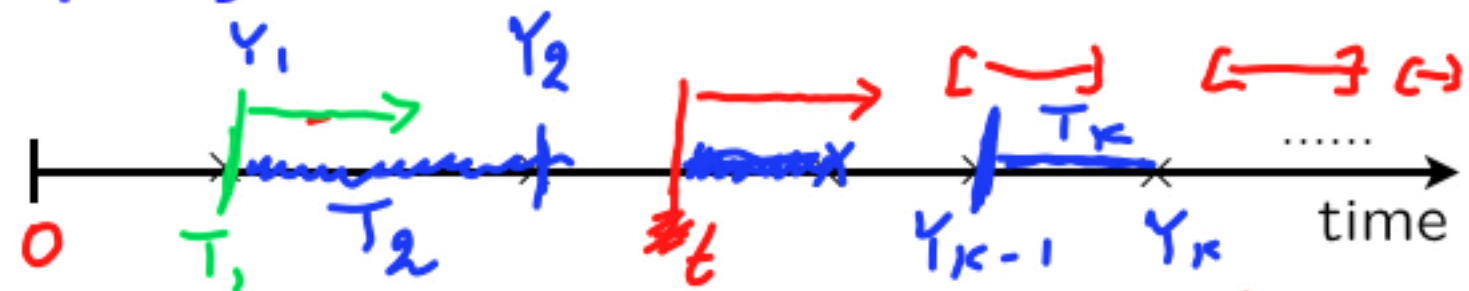
Erlang distribution: $f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}, \quad y \geq 0$
 order k



Memorylessness and the fresh-start property

- Analogous to the properties for the Bernoulli process
 - plausible, given the relation between the two processes
 - use intuitive reasoning
 - can be proved rigorously

Memorylessness and the fresh-start property



- If we start watching at time t ,

we see Poisson process, independent of the history until time t

- time until next arrival: $\text{Exp}(\lambda)$, independent of past

- If we start watching at time T_1 , $T_1 = 3$

we see Poisson process, independent of the history until time T_1

- hence: time between first and second arrival, $T_2 = Y_2 - Y_1$ is: $\text{Exp}(\lambda)$
- similarly for all $T_k = Y_k - Y_{k-1}$, $k \geq 2$ incl. of T_1

$Y_k = T_1 + \dots + T_k$ is sum of i.i.d. exponentials

$$\mathbf{E}[Y_k] = k/\lambda \quad \text{var}(Y_k) = k/\lambda^2$$

- An equivalent definition
- A simulation method

Bernoulli/Poisson relation



$$n = \tau/\delta,$$

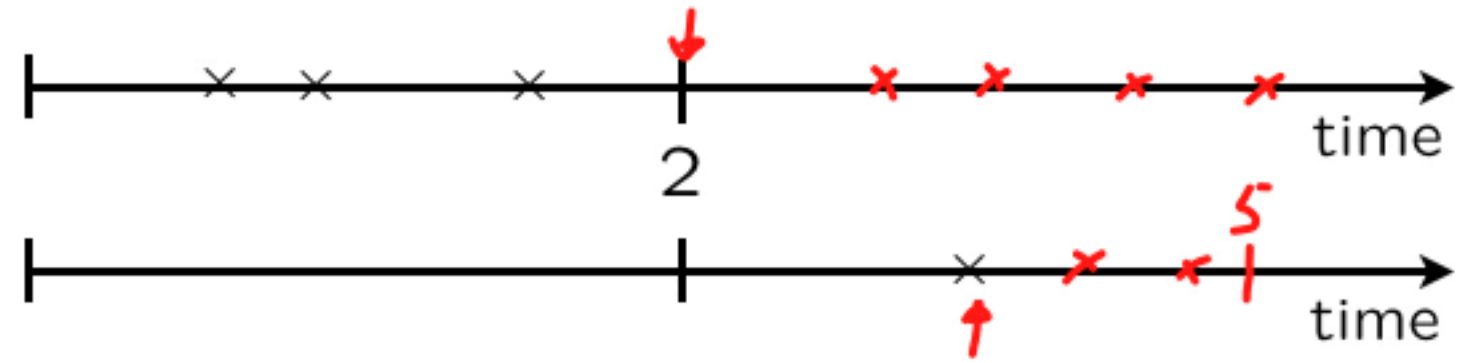
$$p = \lambda\delta$$

$$np = \lambda\tau$$

	POISSON	BERNOULLI
Times of Arrival	Continuous	Discrete
Arrival Rate	λ /unit time	p /per trial
PMF of # of Arrivals	<div>•</div> Poisson	Binomial
Interarrival Time Distr.	Exponential	Geometric
Time to k -th arrival	Erlang	Pascal

Example: Poisson fishing

- Fish are caught as a Poisson process, $\lambda = 0.6/\text{hour}$
 - fish for two hours;
 - if you caught at least one fish, stop
 - else continue until first fish is caught



$P(\text{fish for more than two hours}) = P(0, 2)$

$$P(\tau_1 > 2) = \int_2^{\infty} f_{\tau_1}(t) dt$$

$P(\text{fish for more than two and less than five hours}) =$

$$P(0, 2) (1 - P(0, 3))$$

$$P(2 < \tau_1 \leq 5) = \int_2^5 f_{\tau_1}(t) dt$$

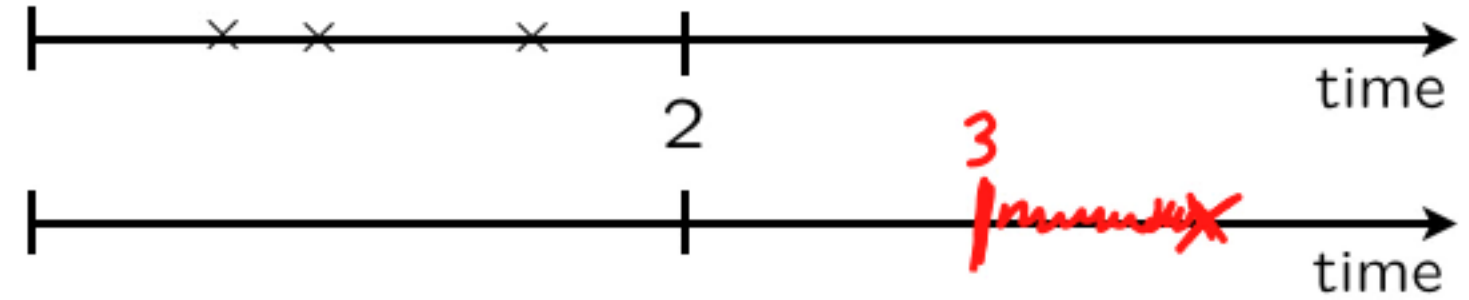
$$P(k, \tau) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}$$

$$E[N_\tau] = \lambda\tau$$

$$f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}$$

Example: Poisson fishing

- Fish are caught as a Poisson process, $\lambda = 0.6/\text{hour}$
 - fish for two hours;
 - if you caught at least one fish, stop
 - else continue until first fish is caught



$P(\text{catch at least two fish}) =$

$$\sum_{k=2}^{\infty} P(k, 2) = 1 - P(0, 2) - P(1, 2)$$
$$P(Y_2 \leq 2) = \int_0^2 f_{Y_2}(y) dy$$

$E[\text{future fishing time} \mid \text{already fished for three hours}] = \frac{1}{\lambda}$

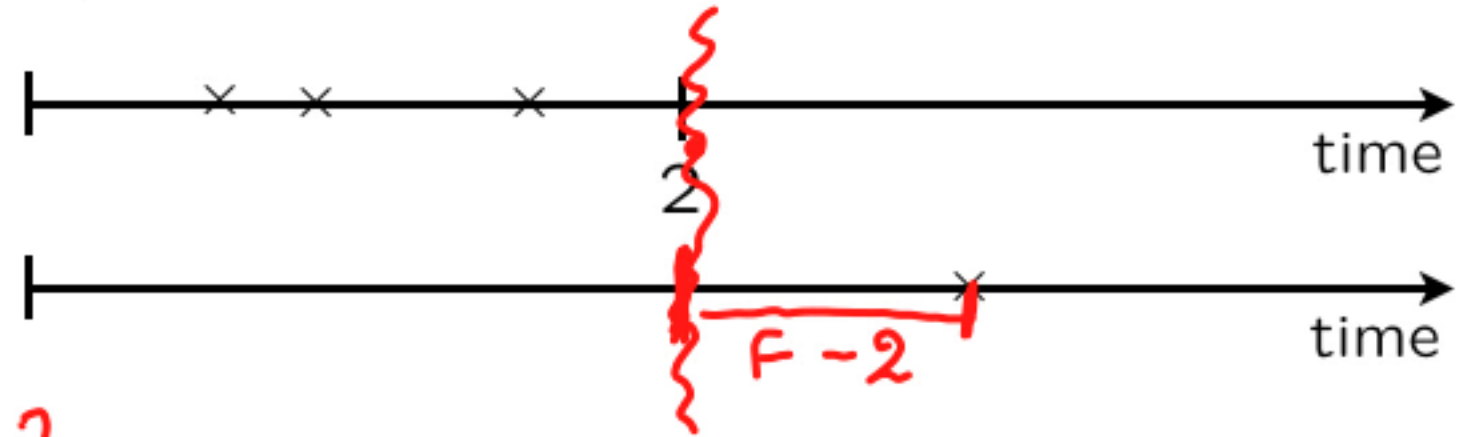
$$P(k, \tau) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}$$

$$E[N_\tau] = \lambda\tau$$

$$f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}$$

Example: Poisson fishing

- Fish are caught as a Poisson process, $\lambda = 0.6/\text{hour}$
 - fish for two hours;
 - if you caught at least one fish, stop
 - else continue until first fish is caught



$$\begin{aligned} E[\text{total fishing time}] &= E[F] = 2 + E[F - 2] \\ &= 2 + P(F = 2) \cdot 0 + P(F > 2) E[F - 2 | F > 2] \\ &= 2 + P(0, 2) \cdot 1/\lambda \end{aligned}$$

$$E[\text{number of fish}] = \lambda \tau + P(0, 2) \cdot 1$$

0.6×2

$$P(k, \tau) = \frac{(\lambda \tau)^k e^{-\lambda \tau}}{k!}$$

$$E[N_\tau] = \lambda \tau$$

$$f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}$$