

STA 504 Mathematical Statistics with Calculus Review

Midterm Exam #2

11/19/2022

Please Print: _____
(First Name) (Last Name)

Instructions

- This is an open-book test. Textbook and notes can be used. However, you must complete this exam independently. All forms of collaborations are NOT allowed.
- You may use a calculator for the exam.
- Please show your detailed work to earn full credit.
- Partial credit will be granted to the key steps that reflect your correct reasoning even if your numerical answer is incorrect.

Problem 1.

Consider two discrete random variables X and Y whose values are r and s respectively and suppose that the probability of the event $\{X = r\} \cap \{Y = s\}$ is given by:

$$f(s, t) = \begin{cases} \frac{r+s}{48}, & 0 \leq r, s \leq 3 \\ 0, & \text{elsewhere} \end{cases}$$

The above probability distribution can be tabulated in the following

		Y				
		$s \rightarrow$				
		0	1	2	3	
X	0	$\frac{0}{48}$	$\frac{1}{48}$	$\frac{2}{48}$	$\frac{3}{48}$	$\frac{6}{48}$
	1	$\frac{1}{48}$	$\frac{2}{48}$	$\frac{3}{48}$	$\frac{4}{48}$	$\frac{10}{48}$
	2	$\frac{2}{48}$	$\frac{3}{48}$	$\frac{4}{48}$	$\frac{5}{48}$	$\frac{14}{48}$
	3	$\frac{3}{48}$	$\frac{4}{48}$	$\frac{5}{48}$	$\frac{6}{48}$	$\frac{18}{48}$
		$\frac{6}{48}$	$\frac{10}{48}$	$\frac{14}{48}$	$\frac{18}{48}$	
		$P(Y = s) \rightarrow$				

$P(X = r)$
 \downarrow

Find the expectation of

- Are X and Y independent?

Solution. Since $P[X = 0, Y = 0] = \frac{0}{48} \neq \left(\frac{6}{48}\right) \times \left(\frac{6}{48}\right) = P[X = 0] \times P[Y = 0]$, X and Y are dependent.

- $E[X + Y]$

Solution. Since the two marginal distributions are identical, $E[X + Y] = E[X] + E[Y] = 2E[X]$.

We only find $E[X] = 0 \times \left(\frac{6}{48}\right) + 1 \times \left(\frac{10}{48}\right) + 2 \times \left(\frac{14}{48}\right) + 3 \times \left(\frac{18}{48}\right) = \frac{92}{48}$.

Therefore, $E[X + Y] = 2 \times \left(\frac{92}{48}\right) = \frac{23}{6}$

- $E[XY]$

Solution. $E[XY] = (1 \times 1) \left(\frac{2}{48}\right) + (1 \times 2) \left(\frac{3}{48}\right) + (1 \times 3) \left(\frac{4}{48}\right) + (2 \times 1) \left(\frac{3}{48}\right) + (2 \times 2) \left(\frac{4}{48}\right) + (2 \times 3) \left(\frac{5}{48}\right) + (3 \times 1) \left(\frac{4}{48}\right) + (3 \times 2) \left(\frac{5}{48}\right) + (3 \times 3) \left(\frac{6}{48}\right) = \frac{168}{48} = 3.5$.

- $COV(X, Y)$

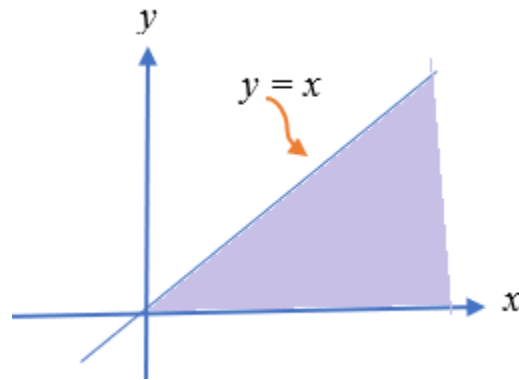
Solution. $COV(X, Y) = E[XY] - E[X]E[Y] = \frac{168}{48} - \frac{92}{48} \times \frac{92}{48} = -\frac{25}{144}$.

Problem 2.

Let X be the total time that a customer spends at a bank, and Y the time she spends waiting in line. Assume that X and Y have joint density

$$f(x, y) = \begin{cases} \lambda^2 e^{-\lambda x}, & 0 \leq y \leq x < \infty \\ 0, & \text{elsewhere} \end{cases}$$

Sketch the domain or related regions whenever needed.



1. Find the marginal density functions of X and Y .

Solution: $f(x) = \int_0^x f(x, y) dy = \int_0^x \lambda^2 e^{-\lambda x} dy = \lambda^2 x e^{-\lambda x}$
 $f(y) = \int_y^\infty f(x, y) dx = \int_y^\infty \lambda^2 e^{-\lambda x} dx = -\lambda \times e^{-\lambda x} \Big|_y^\infty = \lambda e^{-\lambda y}$

2. Are X and Y independent?

Solution: Since $f(x, y) = \lambda^2 e^{-\lambda x} \neq \lambda^2 x e^{-\lambda x} \times \lambda e^{-\lambda y}$, random variables X and Y are dependent.

3. Find out the mean service time: $E[T] = E[X - Y]$.

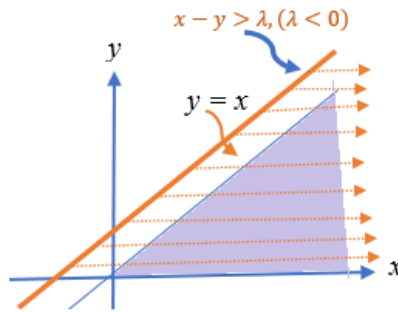
Solution: $E[T] = \iint_D (x - y) f(x, y) dA = \int_0^\infty \int_0^x (x - y) \lambda^2 e^{-\lambda x} dy dx$
 $= \int_0^\infty \int_0^x x \lambda^2 e^{-\lambda x} dy dx - \int_0^\infty \int_0^x y \lambda^2 e^{-\lambda x} dy dx = \int_0^\infty \lambda^2 x^2 e^{-\lambda x} dx - 0.5 \int_0^\infty \lambda^2 x^2 e^{-\lambda x} dx$
 $= 0.5 \int_0^\infty \lambda^2 x^2 e^{-\lambda x} dx = -0.5 \lambda [x^2 e^{-\lambda x} \Big|_0^\infty - \int_0^\infty 2x e^{-\lambda x} dx] = \int_0^\infty \lambda x e^{-\lambda x} dx$
 $= - \int_0^\infty x d e^{-\lambda x} = -x e^{-\lambda x} \Big|_0^\infty + \int_0^\infty e^{-\lambda x} dx = -\frac{1}{\lambda} \int_0^\infty e^{-\lambda x} d(-\lambda x) = \frac{1}{\lambda}$

4. Find the probability $P[T > \lambda]$

Solution: We calculate the probability in two cases

Case 1. $\lambda \leq 0$

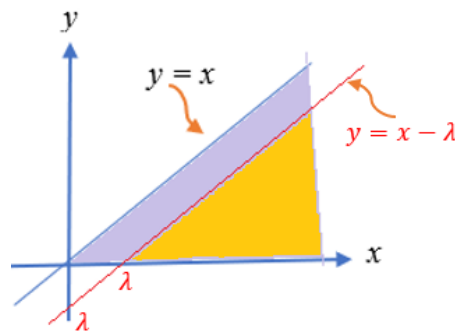
$T > \lambda$ implies $X - Y > \lambda$. This defines a region that contains the domain as a sub-region. The following figure depicts the region defined by $T > \lambda$ and the domain. You can test location of the defined region by taking a special value at $\lambda = -1$.



Since $T > \lambda$ contain the domain, therefore, $P[T > \lambda] = 1$.

Case 2. $\lambda < 0$

The yellow region in the following figure is defined by $T > \lambda$.



$$\begin{aligned}
 P[T > \lambda] &= P[X - Y > \lambda] = \iint_{\text{orange}} f(x, y) dA \\
 &= \int_{\lambda}^{\infty} \left[\int_0^{x-\lambda} \lambda^2 e^{-\lambda x} dy \right] dx = \int_{\lambda}^{\infty} \lambda^2 e^{-\lambda x} \left[\int_0^{x-\lambda} dy \right] dx \\
 &= \int_{\lambda}^{\infty} (x - \lambda) \lambda^2 e^{-\lambda x} dx = \int_{\lambda}^{\infty} \lambda^2 x e^{-\lambda x} dx - \int_{\lambda}^{\infty} \lambda^3 e^{-\lambda x} dx
 \end{aligned}$$

Note that

$$\begin{aligned}
 \int_{\lambda}^{\infty} \lambda^2 x e^{-\lambda x} dx &= - \int_{\lambda}^{\infty} \lambda x e^{-\lambda x} d(-\lambda x) = - \int_{\lambda}^{\infty} \lambda x d e^{-\lambda x} \\
 &= -\lambda x e^{-\lambda x} \Big|_{\lambda}^{\infty} + \int_{\lambda}^{\infty} e^{-\lambda x} d\lambda x = \lambda^2 e^{-\lambda^2} - e^{-\lambda x} \Big|_{\lambda}^{\infty} = \lambda^2 e^{-\lambda^2} + e^{-\lambda^2}
 \end{aligned}$$

and

$$\int_{\lambda}^{\infty} \lambda^3 e^{-\lambda x} dx = -\lambda^2 \int_{\lambda}^{\infty} e^{-\lambda x} d(-\lambda x) = -\lambda^2 e^{-\lambda x} \Big|_{\lambda}^{\infty} = \lambda^2 e^{-\lambda^2}$$

$$P[T > \lambda] = P[X - Y > \lambda] = \lambda^2 e^{-\lambda^2} + e^{-\lambda^2} - \lambda^2 e^{-\lambda^2} = e^{-\lambda^2}.$$

5. Find the correlation coefficient between X and Y .

Solution: Since $COV(X, Y) = E[XY] - E[X]E[Y]$. We first calculate the involved moments in the following using the marginal density functions derived in part 1 and the joint density as well.

$$\begin{aligned} E[X] &= \int_0^{\infty} x \lambda^2 x e^{-\lambda x} dx = -\lambda \int_0^{\infty} x^2 d e^{-\lambda x} = -\lambda x^2 e^{-\lambda x} \Big|_0^{\infty} + \lambda \int_0^{\infty} 2x e^{-\lambda x} dx \\ &= \frac{2}{\lambda} \int_0^{\infty} \lambda^2 x e^{-\lambda x} dx = \frac{2}{\lambda} \times 1 = \frac{2}{\lambda}. \end{aligned}$$

$$E[Y] = \int_0^{\infty} y \lambda e^{-\lambda y} dy = - \int_0^{\infty} y d e^{-\lambda y} = -y e^{-\lambda y} \Big|_0^{\infty} + \int_0^{\infty} e^{-\lambda y} dy = \frac{1}{\lambda}$$

$$\begin{aligned} E[XY] &= \iint xy f(x, y) dA = \int_0^{\infty} \int_0^x xy \lambda^2 e^{-\lambda x} dy dx = \int_0^{\infty} x \lambda^2 e^{-\lambda x} \times \frac{y^2}{2} \Big|_0^x dx \\ &= \frac{1}{2} \int_0^{\infty} \lambda^2 x^3 e^{-\lambda x} dx = -\frac{1}{2} \int_0^{\infty} \lambda x^3 d(e^{-\lambda x}) \\ &= -\frac{1}{2} \left[\lambda x^3 e^{-\lambda x} \Big|_0^{\infty} - \int_0^{\infty} 3\lambda x^2 e^{-\lambda x} dx \right] \\ &= \frac{3}{2\lambda} \int_0^{\infty} \lambda^2 x^2 e^{-\lambda x} dx = \frac{2}{\lambda} \times \frac{3}{2\lambda} = \frac{3}{\lambda^2}. \end{aligned}$$

$$COV[X, Y] = E[XY] - E[X]E[Y] = \frac{3}{\lambda^2} - \frac{2}{\lambda} \times \frac{1}{\lambda} = \frac{1}{\lambda^2}$$

To find the Pearson correlation coefficient between X and Y , we also need $V[X]$ and $V[Y]$. Next, we calculate the second moment of X and Y .

$$E[X^2] = \int_0^{\infty} x^2 \lambda^2 x e^{-\lambda x} dx = \int_0^{\infty} \lambda^2 x^3 e^{-\lambda x} dx = 2 \times \frac{3}{\lambda^2} = \frac{6}{\lambda^2}$$

$$E[Y^2] = \int_0^{\infty} y^2 \lambda e^{-\lambda y} dy = \int_0^{\infty} \lambda y^2 e^{-\lambda y} dy = \frac{1}{\lambda} \frac{2\lambda}{3} \left\{ \frac{3}{2\lambda} \int_0^{\infty} \lambda^2 y^2 e^{-\lambda y} dy \right\} = \frac{2}{3} \times \frac{3}{\lambda^2} = \frac{2}{\lambda^2}$$

We used part of the results in the calculation of $E[XY]$.

$$V[X] = \frac{6}{\lambda^2} - \left(\frac{2}{\lambda}\right)^2 = \frac{2}{\lambda^2},$$

$$V[Y] = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$$

The correlation coefficient

$$\rho = \frac{COV(X, Y)}{\sqrt{V[X]V[Y]}} = \frac{\frac{1}{\lambda^2}}{\sqrt{\frac{2}{\lambda^2} \times \frac{1}{\lambda^2}}} = \frac{\sqrt{2}}{2}$$

6. Find the variance of service time T .

Solution: We have calculated all expectations needed in the variance of T in the following

$$\begin{aligned} V[T] &= V[X - Y] = V[X] + V[Y] - 2COV(X, Y) \\ &= \frac{2}{\lambda^2} + \frac{1}{\lambda^2} - \frac{2}{\lambda^2} = \frac{1}{\lambda^2}. \end{aligned}$$

7. Find the conditional density function of $X|Y = y$.

Solution: By the definition, the conditional density function of $X|Y = y$ is given by

$$f(x|y) = \frac{f(x, y)}{f(y)} = \frac{\lambda^2 e^{-\lambda x}}{\lambda e^{-\lambda y}} = \lambda e^{-\lambda(x-y)}$$

where $0 \leq y \leq x < \infty$.

8. Given that waiting time $Y = \lambda$, what is $E[X|Y = \lambda]$?

Solution: $E[X|Y = \lambda] = \int_{\lambda}^{\infty} x \lambda e^{-\lambda(x-\lambda)} dx = - \int_{\lambda}^{\infty} x d e^{-\lambda(x-\lambda)}$

$$= -x \lambda e^{-\lambda(x-\lambda)} \Big|_{\lambda}^{\infty} + \int_{\lambda}^{\infty} e^{-\lambda(x-\lambda)} dx = \lambda - \frac{1}{\lambda} e^{-\lambda(x-\lambda)} \Big|_{\lambda}^{\infty} = \lambda + \frac{1}{\lambda}.$$