

Intro to Econometrics: Recitation 2

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Review

Random variables - *univariate* case

$$(\Omega, \mathcal{F}, \mathbf{P})$$

► $X : \Omega \rightarrow \mathbf{R}$

► CDF:

$$F_X(x) = \mathbf{P}(\{\omega : X(\omega) \leq x\})$$

► Completely characterizes $\mathbf{P}\{X \in B\}$ for $B \subset \mathbf{R}$

► Absolutely continuous:

$$F_X(x) = \int_{-\infty}^x f_X(x) dx$$

Review

Random variables - *multivariate* case

$$(\Omega, \mathcal{F}, \mathbf{P})$$

- ▶ $X : \Omega \rightarrow \mathbf{R}^S$ where $X(\omega) = (X_1(\omega), \dots, X_S(\omega))'$
- ▶ CDF:

$$F_X(x_1, \dots, x_S) = \mathbf{P}(\{\omega : X_1(\omega) \leq x_1, \dots, X_S(\omega) \leq x_S\})$$

- ▶ Completely characterizes $\mathbf{P}\{X \in B\}$ for $B \subset \mathbf{R}^S$
- ▶ Absolutely continuous:

$$F_X(x_1, \dots, x_S) = \int_{-\infty}^{x_1} \cdots \int_{-\infty}^{x_S} f_X(x_1, \dots, x_S) dx_S \cdots dx_1$$

Review

Random variables - *multivariate* case

► Result: if $F : \mathbf{R} \rightarrow [0, 1]$ is

1. Increasing
2. Right-continuous
3. Satisfies $\lim_{x \rightarrow \infty} F(x) = 1 - \lim_{x \rightarrow -\infty} F(x) = 1$

Then it is the CDF of some random variable $X : \Omega \rightarrow \mathbf{R}$

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► Can you think of (or prove?) an S -dimensional analog of the statement above?

Review

Random variables - *multivariate* case

- ▶ If $F : \mathbf{R}^2 \rightarrow [0, 1]$ is
 1. Increasing
 2. “Continuous from above”
 3. Has the following limits:
 - 3.1 $\lim_{x_1 \rightarrow -\infty} F(x_1, x_2) = 0$ for all x_2
 - 3.2 $\lim_{x_2 \rightarrow -\infty} F(x_1, x_2) = 0$ for all x_1
 - 3.3 $\lim_{x_1 \rightarrow \infty} \lim_{x_2 \rightarrow \infty} F(x_1, x_2) = 1$

Then F is the CDF of a random variable $X : \Omega \rightarrow \mathbf{R}^2$
(Durrett, sec 2.9)

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 - 3.3 $\lim_{x_1 \rightarrow \infty} \lim_{x_2 \rightarrow \infty} F(x_1, x_2) = 1$
 4. Satisfies, for $x_1^* \geq x_1$ and $x_2^* \geq x_2$,

$$F(x_1^*, x_2^*) - F(x_1^*, x_2) - F(x_1, x_2^*) + F(x_1, x_2) \geq 0$$

Then F is the CDF of a random variable $X : \Omega \rightarrow \mathbf{R}^2$

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Review

Marginals

- ▶ Marginal with respect to coordinate s , $F_s : \mathbf{R} \rightarrow [0, 1]$

$$F_s(x) = \mathbf{P}(\{\omega : X_s(\omega) \leq x\})$$

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$$F_s(x) = \mathbf{P}(\{\omega : X_s(\omega) \leq x\})$$

- ▶ How do you obtain it?

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Marginals

- ▶ Marginal with respect to coordinate s , $F_s : \mathbf{R} \rightarrow [0, 1]$

$$F_s(x) = \mathbf{P}(\{\omega : X_s(\omega) \leq x\})$$

- ▶ How do you obtain it?
- ▶ Just take limits. Suppose $S = 2$ and we want to recover first coordinate:

$$F_1(x_1) = \lim_{x_2 \rightarrow \infty} F(x_1, x_2)$$

Proof?

Review

Marginals

- ▶ How do you recover a marginal pdf? Suppose $X : \Omega \rightarrow \mathbf{R}^2$ has pdf $f(x_1, x_2)$:

$$f_1(x_1) = \int_{-\infty}^{\infty} f(x_1, x_2) dx_2$$

- ▶ Proof?

Review

Digression: marginals don't determine joints

- ▶ A very useful counterexample:
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- ▶ A very useful counterexample:
 - ▶ Let $X \sim N(0, 1)$
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- ▶ Define $Y = WX$. The marginal is the same as X :

$$\begin{aligned} F_Y(y) &= \mathbf{P}(WX \leq y) = \frac{1}{2}\mathbf{P}(X \leq y) + \frac{1}{2}\mathbf{P}(-X \leq y) \\ &= F_X(y) \end{aligned}$$

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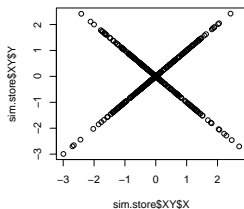
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- ▶ (X, Y) is not multivariate normal. Why?
- ▶ $X + Y$ has a mass at zero!

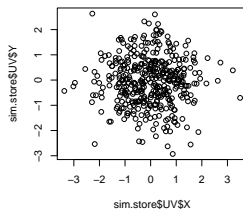
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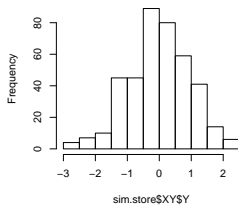
Scatterplot of (X,Y) simulations



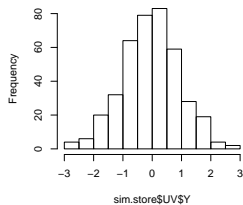
Scatterplot of (U,V) simulations



Histogram of Y



Histogram of V



Review

Moments of multivariate RVs

- ▶ Focus on the case when there is a pdf
- ▶ “Definition”

$$\mathbf{E}g(X) = \int_{\mathbf{R}^S} g(x)f_X(x)dx$$

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- ▶ Second moment:

$$V(X) = \mathbf{E} [(X - \mu_X)(X - \mu_X)']$$

- ▶ When is $V(X)$ finite?

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- ▶ Second moment:

$$V(X) = \mathbf{E} [(X - \mu_X)(X - \mu_X)']$$

- ▶ When is $V(X)$ finite?
- ▶ Covariance btw X and Y :

$$\text{Cov}(X, Y) = \mathbf{E} [(X - \mu_X)(Y - \mu_Y)']$$

Review

Moment generating functions of multivariate RVs

► MGF:

$$m_X(\mathbf{t}) = \mathbf{E} \left[e^{\mathbf{t}'X} \right] = \mathbf{E} \left[e^{\sum_{i=1}^S t_i X_i} \right]$$

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- ▶ Result: suppose X and Y have a moment generating function, and

$$m_X(\mathbf{t}) = m_Y(\mathbf{t})$$

for all \mathbf{t} . Then $F_X(\mathbf{t}) = F_Y(\mathbf{t})$ for all \mathbf{t} .

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- ▶ Result (stronger): suppose that, for all $\mathbf{t} \in \mathbf{R}^S$, $\alpha \in \mathbf{R}$,

$$\mathbf{P}\{\mathbf{t}'X \leq \alpha\} = \mathbf{P}\{\mathbf{t}'Y \leq \alpha\}$$

then $F_X(z) = F_Y(z)$ for all $z \in \mathbf{R}^S$

PS2: Projections, conditioning, linear predictors

Projections

Let $(V, \langle \cdot, \cdot \rangle)$ be a vector space with an inner product.

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- Orthogonal projection of v into (closed) $W \subseteq V$:

$$v - \text{Proj}_W(v) \perp w$$

for all $w \in W$

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Theorem (Projection in a Hilbert Space)

Let $W \subset V$ be a closed vector subspace of V .

For any $v \in V$, the distance minimization problem

$$\min_{w \in W} \|v - w\|$$

has a unique solution $w^ \in W$. Moreover, $w^* = \text{Proj}_W(v)$.*

PS2: Projections, conditioning, linear predictors

Projections

What if W has a finite basis?

$$W = \text{span}\{w_1, \dots, w_K\}$$

► Orthogonal projection of v into W is

$$\text{Proj}_W(v) = \sum_{i=1}^K \frac{\langle w_i, v \rangle}{\langle w_i, w_i \rangle} w_i$$

Using this result in the pset is fair game

PS2: Projections, conditioning, linear predictors

Projections

Space $V = \{X : \Omega \rightarrow \mathbf{R}^S : \mathbf{E}\|X\|^2 < \infty\}$ is a Hilbert space with

$$\langle X, Y \rangle = \mathbf{E}XY$$

PS2: Projections, conditioning, linear predictors

Projections

Space $V = \{X : \Omega \rightarrow \mathbf{R}^S : \mathbf{E}\|X\|^2 < \infty\}$ is a Hilbert space with

$$\langle X, Y \rangle = \mathbf{E}XY$$

- Fix variables X, Y in V and consider the subspace

$$W = \{Z : \Omega \rightarrow \mathbf{R} : Z = \alpha + \beta(X - \mu_X)\}$$

(Is there a finite basis for W ?)

PS2: Projections, conditioning, linear predictors

Projections

The problem

$$\min_{(\alpha, \beta)} [Y - \alpha - \beta(X - \mu_X)]^2$$

is equivalent to some norm minimization problem involving Y , X and W .

What is it?