Intro to Econometrics: Recitation 8

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Roadmap for today

- Chapter 6 quick review:
 - Asymptotic order notation
 - ► The basic toolkit: CMT, delta method and limit theorems
- BH Chapter 7:
 - Asymptotic properties in the LP/LR model

Deterministic sequences

Consider the claim:

"The sequence $a_n = n^2$ diverges to infinity an order of magnitude faster than $b_n = n$ "

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- Note: we're assuming for simplicity that $b_n > 0$

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 - **4** $x_n \in O(x_n)$ but the same doesn't hold for $o(x_n)$
 - **3** If $x_n \in O(An^2 + Bn + C)$, then $x_n \in O(n^2)$; moreover, $x_n \in o(n^3)$
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Deterministic sequences: results

Lemma

For any positive sequences x_n , y_n , z_n and w_n ,

• If $x_n \in o(z_n)$ and $y_n \in o(w_n)$, then

$$x_n y_n \in o(z_n w_n)$$

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and the same holds for O

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Deterministic sequences: results

Proof:

- ▶ First part of (1) follows from limit arithmetic
- ▶ First part of (2) is a consequence of the fact that $a_n \to 0$ and b_n bounded implies $a_n b_n \to 0$
- ▶ Second part of (1) and (2): use

$$\frac{x_{n} + y_{n}}{z_{n} + w_{n}} = \frac{x_{n}}{z_{n}} \frac{z_{n}}{z_{n} + w_{n}} + \frac{y_{n}}{w_{n}} \frac{w_{n}}{z_{n} + w_{n}}$$

the ratio $z_n/(z_n+w_n)\in[0,1]$ is bounded

Sequences of random variables

- Analogs: let X_n be a sequence of random variables and a_n a positive deterministic sequence
 - ▶ $X_n \in O_p(a_n)$ iff $\frac{X_n}{a_n}$ is *stochastically* bounded
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For every $\delta > 0$ there exists M_{δ} such that

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ullet Result: lemma goes through with o and O replaced with o_p and O_p

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A sequence (X_n) that converges in distribution is necessarily $O_p(1)$

- Proof. Let Y be st $X_n \stackrel{d}{\to} Y$. Denote F_Y and F_n the CDF of Y and X_n , respectively. Fix $\delta > 0$.
 - ▶ Let M_δ be such that

$$F_Y(M_\delta) - F_Y(-M_\delta) = \Pr\{Y \in (-M_\delta, M_\delta]\} > 1 - \frac{\delta}{2}$$

▶ In addition, we can choose M_δ st F_Y continuous at M_δ and $-M_\delta$ (why?)

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- Convergence in distribution implies $F_n(M_\delta) \to F_Y(M_\delta)$ and $F_n(-M_\delta) \to F_Y(-M_\delta)$

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- That implies

$$\Pr\{|X_n| \le M_{\delta}\} \ge F_n(M_{\delta}) - F_n(-M_{\delta}) > 1 - \delta$$



The asymptotics toolkit

- You need to know the following by heart:
 - General results
 - **★** Continuous mapping theorem
 - ★ Slutzky's theorem
 - ★ Delta method
 - Results about averages
 - ★ Law of large numbers
 - * Central Limit Theorem

Continuous mapping theorem

- ullet Let F:U o V have a set of discontinuities U_0
- If $X_n \stackrel{d}{\to} X$ or $X_n \stackrel{p}{\to} X$, then $F(X_n) \to F(X)$ as long as

$$\Pr\{X\in \textit{U}_0\}=0$$

Slutzky's theorem

- If $X_n \stackrel{d}{\to} X$ and $Y_n \stackrel{p}{\to} c$ where c is non-random, then

 - $2 X_n Y_n \stackrel{d}{\to} cX$
- These results follow from CMT, plus the fact that

$$(X_n,Y_n)\stackrel{d}{\to} (X,c)$$

provided that c is non-random.



Delta method

ullet If μ is non-random,

$$\sqrt{n}(X_n-\mu)\stackrel{d}{\to} Z$$

and g is differentiable at μ , then

$$\sqrt{n}(g(X_n)-g(\mu))\stackrel{d}{\rightarrow} g'(\mu)Z$$

- Remarks.
 - ullet No assumption about normality is needed for Z.
 - The theorem goes through in the multivariate case
 - ★ Example: $F(\gamma_1, \gamma_2) = \gamma_1 \gamma_2$

• Let $\{X_i\}$ be sampled iid from some distribution. Denote

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$$

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• Central limit theorem. If $\{X_i\}$ are sampled iid from distribution with finite mean μ and variance σ^2 ,

$$\sqrt{n}\left(\bar{X}_n-\mu\right)\stackrel{d}{
ightarrow}Z\sim N(0,\sigma^2)$$



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• Hence (with finite second moments) the convergence of \bar{X}_n to μ is no slower than $n^{-1/2}$ goes to 0.

OLS asymptotics: the linear projection model

- Consider the following assumptions:
 - (1) (y_i, x_i) sampled iid
 - (2) $(y_i^2) < \infty$
 - (2') $(y_i^4) < \infty$
 - (3) $(\|x_i\|^2) < \infty$
 - (3') $(\|x_i\|^4) < \infty$
 - (4) $[x_i x_i']$ positive definite
- Notation:

$$[x_i x_i'] =: Q_{xx}$$

$$[x_iy_i] =: Q_{xy}$$

OLS asymptotics: the linear projection model

- Results:
 - Under (1), (2), (3) and (4), OLS is consistent for the population projection coefficient:

$$\hat{\beta} = \left(\sum_{i=1}^{n} x_i x_i'\right)^{-1} \sum_{i=1}^{n} x_i y_i \stackrel{p}{\rightarrow} \beta := Q_{xx}^{-1} Q_{xy}$$

② Under (1), (2'), (3') and (4), OLS is asymptotically normal:

$$\sqrt{n}\left[\left(\sum_{i=1}^{n}x_{i}x_{i}'\right)^{-1}\sum_{i=1}^{n}x_{i}y_{i}-\beta\right]\stackrel{d}{\to}N\left(0,Q_{xx}^{-1}\Omega Q_{xx}^{-1}\right)$$

- \star β is again the population projection coefficient
- ★ $\Omega = [u_i^2 x_i x_i']$
- * $u_i := y_i x_i' \beta$
- Notation:

$$V_{\beta} = Q_{xx}^{-1} \Omega Q_{xx}^{-1}$$



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 - (6) $\mathbf{V}[y_i|x_i] \equiv \sigma^2$
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- If (6) is assumed, we have the homoskedastic linear regression model

• Finite sample variance of OLS:

$$\hat{\beta} = (X'X)^{-1}X'Y \implies \mathbf{V}(\hat{\beta}|X) = (X'X)^{-1}X'\mathbf{V}(Y|X)X(X'X)^{-1}$$

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It is true however that

$$nV_{\hat{\beta}} \stackrel{p}{\rightarrow} V_{\beta}$$

(why?)



OLS asymptotics: estimating the limiting variance

We know that the limiting variance of OLS in the LPM is

$$V_{\beta} = Q_{xx}^{-1} \Omega Q_{xx}^{-1}$$

• A valid estimator for $\Omega = [u_i^2 x_i x_i']$ is

$$\hat{\Omega} = \frac{1}{n} \sum_{i=1}^{n} \hat{u}_i^2 x_i x_i'$$

ullet One estimator for V_eta is

$$\hat{V}_{\beta}^{\mathsf{HC0}} = \hat{Q}_{\mathsf{xx}}^{-1} \hat{\Omega} \hat{Q}_{\mathsf{xx}}^{-1}$$

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Note that for fixed x.

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That motivates the approximation

$$\hat{m}(x) \approx N\left(m(x), \frac{1}{n}x'\hat{V}_{\beta}x\right)$$



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Forecast error is

$$\tilde{e}_{n+1} = y_{n+1} - x'\hat{\beta} = x'\beta + e_i - x'\hat{\beta} = e_{n+1} + x'(\beta - \hat{\beta})$$

- ▶ The true regression has an error e_{n+1}
- ▶ In addition to that, there is an estimation noise coming from $\hat{\beta}$ being different than β

• Finite sample variance of e_{n+1} :

$$[\hat{e}_{n+1}^2|x_{n+1}] = \sigma^2(x_{n+1}) + x'_{n+1}V_{\hat{\beta}}x_{n+1}$$

• Can we build a confidence interval for the forecast?