## An Introduction to Large Sample Asymptotics [Hansen, chapter 6]

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#### **Outline**

Introduction

Limit Superior and Limit Inferior

Basic Theory of Stochastic Convergence

Laws of Large Numbers and Central Limit Theorems

**Delta Method** 

### Limit Superior and Limit Inferior of a Sequence of Real Numbers

► The limit superior and limit inferior of a sequence  $x_n$  of real numbers are defined as

$$\limsup_{n\to\infty} x_n = \inf_{n\geq 1} \sup_{k\geq n} x_k,$$
  
$$\liminf_{n\to\infty} x_n = \sup_{n\geq 1} \inf_{k\geq n} x_k.$$

- ▶ The limit superior always exists as a real number if  $x_n$  is bounded from above. Likewise, the limit inferior exists as a real number whenever  $x_n$  is bounded from below.
- Note that we always have

$$\liminf_{n\to\infty} x_n \leq \limsup_{n\to\infty} x_n.$$

▶ The limit  $\lim_{n\to\infty} x_n$  exists if and only if

$$\liminf_{n\to\infty} x_n = \limsup_{n\to\infty} x_n.$$

▶ We have that  $\liminf_{n\to\infty} x_n = -\limsup_{n\to\infty} x_n$ .

## Limit Superior and Limit Inferior of a Sequence of Sets

▶ The limit superior and limit inferior of a sequence  $E_n$  of sets are defined as

$$\limsup_{n\to\infty} E_n = \bigcap_{n\geq 1}^{\infty} \cup_{k\geq n}^{\infty} E_k,$$
$$\liminf_{n\to\infty} x_n = \bigcup_{n\geq 1}^{\infty} \bigcap_{k\geq n}^{\infty} E_k.$$

- ▶ A point belongs to  $\limsup_{n\to\infty} E_n$  if and only if it belongs to infinitely many terms of the sequence  $\{E_n, n \geq 1\}$ . Hence, sometimes we write  $\limsup_{n\to\infty} E_n = E_n$  i.o. ("infinitely often").
- ▶ A point belongs to  $\liminf_{n\to\infty} E_n$  if and only if it belongs to all terms of the sequence from a certain term on.
- ▶ We have that  $\liminf_{n\to\infty} E_n = (\limsup_{n\to\infty} E_n^c)^c$ .
- We also have that

$$P(\limsup_{n\to\infty} E_n) = \lim_{n\to\infty} P(\cup_{k\geq n}^{\infty} E_k),$$
  
$$P(\liminf_{n\to\infty} x_n) = \lim_{n\to\infty} P(\cap_{k\geq n}^{\infty} E_k).$$

#### References

- The following materials on stochastic convergence are closely based on Sections 2.1 and 2.4 of A. W. van der Vaart (1998), Asymptotic Statistics, Cambridge University Press.
- Other useful references are:
  - P. Billingsley (1995), Probability and Measure, 3rd ed., Wiley.
  - ► K.L. Chung (2001), A course in Probability Theory, 3rd ed., Academic Press.

### Convergence in Distribution

A sequence of random vectors X<sub>n</sub> is said to converge in distribution to a random vector X if

$$P(X_n \leq x) \rightarrow P(X \leq x)$$

for every x at which the limit distribution function  $x \mapsto P(X \le x)$  is continuous.

- Alternatively, they are called weak convergence and convergence in law.
- ▶ Weak convergence is denoted by  $X_n \rightarrow_d X$ .
- Note that the convergence only depends on the induced laws (but not on the probability spaces on which they are defined). Hence, sometimes we write  $X_n \to_d L$  if X has distribution L (e.g.  $X_n \to_d N(0,1)$ ).

### Convergence in Distribution: Example

- ▶ If  $X_1, X_2, ...$  are iid uniform(0,1) and  $X_{(n)} = \max_{1 \le i \le n} X_i$ .
- Note that for every t > 0,

$$P\left(X_{(n)} \leq 1 - \frac{t}{n}\right) = \left[1 - \frac{t}{n}\right]^n \to e^{-t},$$

equivalently

$$P(n(1-X_{(n)}) \leq t) \to 1-e^{-t},$$

which implies that  $n(1 - X_{(n)})$  converges in distribution to an exponential(1) random variable.

### Properties of Weak Convergence

- (Portmanteau Theorem) The following statements are equivalent.
  - $\rightarrow$   $X_n \rightarrow_d X_n$ .
  - ►  $Ef(X_n) \rightarrow Ef(X)$  for all bounded, continuous real-valued functions f.
  - ▶  $Ef(X_n) \rightarrow Ef(X)$  for all bounded, Lipschitz continuous real-valued functions f.
  - ▶  $\liminf Ef(X_n) \ge Ef(X)$  for all nonnegative, continuous functions f.
  - ▶  $\liminf P(X_n \in G) \ge P(X \in G)$  for every open set G.
  - ▶  $\limsup P(X_n \in F) \le P(X \in F)$  for every closed set F.
  - ▶  $P(X_n \in B) \rightarrow P(X \in B)$  for all Borel sets B with  $P(X \in \delta B) = 0$ , where  $\delta B$  is the boundary of B.
- See Van der Vaart (1998) for the proof.

### Convergence in Probability

A sequence of random vectors X<sub>n</sub> is said to converge in probability to a random vector X if for every € > 0,

$$P(\|X_n - X\| > \epsilon) \to 0, \tag{1}$$

where  $\|\cdot\|$  is the Euclidean norm.

An alternative expression to (1) is

$$P(\|X_n - X\| \le \epsilon) \to 1 \tag{2}$$

for every  $\epsilon > 0$ .

► Convergence in probability is denoted by  $X_n \rightarrow_p X$ . Equivalently,  $||X_n - X|| \rightarrow_p 0$ .



### Almost Sure Convergence

A sequence of random vectors X<sub>n</sub> is said to converge almost surely (a.s.) to a random vector X if for every ε > 0,

$$\lim_{m\to\infty} P(\|X_n - X\| > \epsilon \text{ for some } n \ge m) = 0.$$
 (3)

An alternative expression to (3) is

$$\lim_{m \to \infty} P(\|X_n - X\| \le \epsilon \text{ for all } n \ge m) = 1$$
 (4)

for every  $\epsilon > 0$ .

▶ Almost sure is denoted by  $X_n \rightarrow_{a.s.} X$ .

### Almost Sure Convergence (Cont.)

- Almost sure convergence is similar to pointwise convergence of a sequence of functions, except that the convergence need not occur on a set with probability 0 (hence, the "almost sure" or sometimes we say "almost everywhere").
- ▶ Given the definition, it is obvious that  $X_n \rightarrow_{a.s.} X$  implies that  $X_n \rightarrow_p X$ . (The converse is not true)
- Note that convergence in probability and almost sure convergence only make sense if X<sub>n</sub> and X are defined on the same probability space for each n (recall that this was not necessary for convergence in distribution).

### **Continuous Mapping Theorems**

- Let  $g: \mathbf{R}^k \mapsto \mathbf{R}^m$  be continuous at every point of a set C such that  $P(X \in C) = 1$ .
  - (i) If  $X_n \to_d X$ , then  $g(X_n) \to_d g(X)$ .
  - (ii) If  $X_n \to_p X$ , then  $g(X_n) \to_p g(X)$ .
  - (iii) If  $X_n \to_{a.s.} X$ , then  $g(X_n) \to_{a.s.} g(X)$ .
- This is a simple, but very useful result.

### Continuous Mapping Theorems: Proof

- Proof: part (iii) is obvious.
- ▶ To prove part (ii), fix arbitrary  $\epsilon > 0$ . Then for any  $\delta > 0$ , note that

$$\begin{split} & P(\|g(X_n) - g(X)\| > \epsilon) \\ & = P(\|g(X_n) - g(X)\| > \epsilon \text{ and } \|X_n - X\| < \delta) \\ & + P(\|g(X_n) - g(X)\| > \epsilon \text{ and } \|X_n - X\| \ge \delta) \\ & \le P(\|g(X_n) - g(X)\| > \epsilon \text{ and } \|X_n - X\| < \delta) \\ & + P(\|X_n - X\| \ge \delta). \end{split}$$

- ▶ The second term above converges to zero as  $n \to \infty$  for every fixed  $\delta > 0$  since  $X_n \to_p X$ .
- ▶ The first term converges to zero as  $\delta \downarrow 0$ . (In fact, this is a loose statement. A more rigorous proof requires more careful treatment here).
- Therefore, we proved part (ii).

### Continuous Mapping Theorems: Proof (Cont.)

▶ To prove part (i), note that the event  $\{g(X_n) \in F\}$  is identical to the event  $\{X_n \in g^{-1}(F)\}$ . For every closed set F,

$$g^{-1}(F) \subset \overline{g^{-1}(F)} \subset g^{-1}(F) \cup C^c$$
.

Then by the portmanteau theorem,

$$\limsup P(g(X_n) \in F) = \limsup P(X_n \in g^{-1}(F))$$

$$\leq \limsup P(X_n \in \overline{g^{-1}(F)})$$

$$\leq P(X \in \overline{g^{-1}(F)}).$$

Furthermore, since  $P(X \in C) = 1$ ,

$$P(X \in \overline{g^{-1}(F)}) = P(X \in g^{-1}(F)) = P(g(X) \in F).$$

Then the desired result follows from the portmanteau theorem again.



### Almost-Sure Representation

- Convergence in distribution does not imply convergence in probability or almost surely.
- ▶ However, for a given sequence  $X_n \to_d X$ , we can always find a sequence  $\tilde{X}_n \to_{a.s.} \tilde{X}$  such that  $X_n =_d \tilde{X}_n$  and  $X =_d \tilde{X}$  for each n.
- ► Here,  $X =_d Y$  means X and Y have the same marginal distributions.

### Skorohod-Dudley-Wichura theorem

- ▶ (Skorohod-Dudley-Wichura theorem) Suppose that  $X_n \rightarrow_d X_0$ . Then there exists a probability space such that
  - $\tilde{X}_n$  and  $\tilde{X}_0$  defined on this probability space;
  - $\tilde{X}_n$  has a probability distribution that is the same as that of  $X_n$  for each  $n \geq 0$ ;
  - $ilde{X}_n 
    ightharpoonup a.s. ilde{X}_0.$
- ▶ In view of Skorohod-Dudley-Wichura theorem, it is obvious that the continuous mapping theorem holds for all three modes of convergence. See Theorems 25.6 and 25.7 of Billingsley (1995, 3rd ed.).

- (i)  $X_n \rightarrow_{a.s.} X$  implies that  $X_n \rightarrow_p X$ .
- Proof: this follows from the definitions of the modes of convergence. Note that

$$P(\|X_n - X\| > \epsilon) \le P(\bigcup_{m \ge n} \|X_m - X\| > \epsilon)$$

$$= P(\|X_m - X\| > \epsilon \text{ for some } m \ge n)$$

$$\to 0$$

as  $n \to \infty$ .

- (ii)  $X_n \rightarrow_p X$  implies that  $X_n \rightarrow_d X$ .
  - Proof:
    - Note that for every f with range [0, 1] and Lipschitz norm at most 1 and every ∈ > 0,

$$|Ef(X_n) - Ef(X)| \le \epsilon E1\{||X_n - X|| \le \epsilon\} + 2E1\{||X_n - X|| > \epsilon\}.$$

- ▶ The second term converges to zero as  $n \to \infty$ .
- ▶ The first term can be arbitrarily small by choosing a sufficiently small  $\epsilon$ .
- ▶ Thus,  $Ef(X_n)$  converges to Ef(X) for all bounded, Lipschitz continuous functions f.
- Now the result follows from the portmanteau theorem.

- (iii)  $X_n \rightarrow_p c$  for a constant c if and only if  $X_n \rightarrow_d c$ .
  - ▶ Proof: it suffices to show that  $X_n \rightarrow_d c$  implies that  $X_n \rightarrow_p c$ .
  - ▶ To show this, let  $B(c, \epsilon)$  be a open ball of radius  $\epsilon$  around c. Then

$$P(||X_n - c|| \ge \epsilon) = P(X_n \in B(c, \epsilon)^c).$$

Note that

$$\limsup P(X_n \in B(c,\epsilon)^c) \le P(c \in B(c,\epsilon)^c) = 0$$

by the portmanteau theorem.

- (iv) If  $X_n \rightarrow_d X$  and  $||X_n Y_n|| \rightarrow_p 0$ , then  $Y_n \rightarrow_d X$ .
  - Proof: this can be proved as in the proof of part (ii).

- (v) If  $X_n \to_p X$  and  $Y_n \to_p Y$ , then  $(X_n, Y_n) \to_p (X, Y)$ .
  - Proof: this follows from

$$||(X_n, Y_n) - (X, Y)|| \le ||X_n - X|| + ||Y_n - Y||.$$

This states that convergence in probability of a sequence of random vectors is equivalent to convergence of each component separately.

- (vi) If  $X_n \to_d X$  and  $Y_n \to_p c$  for a constant c, then  $(X_n, Y_n) \to_d (X, c)$ .
  - Proof: We can prove this by applying part (iv). First of all,

$$||(X_n, Y_n) - (X_n, c)|| = ||Y_n - c|| \rightarrow_{p} 0.$$

- Now it suffices to show that  $(X_n, c) \to_d (X, c)$ . For every bounded, continuous function  $(x, y) \mapsto f(x, y)$ , the function  $x \mapsto f(x, c)$  is continuous and bounded.
- ▶ Hence,  $Ef(X_n, c) \rightarrow Ef(X, c)$  if  $X_n \rightarrow X$  (by the portmanteau theorem).
- Then the theorem follows again from the portmanteau theorem.

### Joint vs. Marginal Convergence in Distribution

- ▶ In general,  $X_n \rightarrow_d X$  and  $Y_n \rightarrow_d Y$  would not imply that  $(X_n, Y_n) \rightarrow_d (X, Y)$ .
- ▶ Convergence in the distribution of the sequence  $(X_n, Y_n)$  is stronger than convergence of  $X_n$  and  $Y_n$  separately.
- ▶ The implication of part (vi) is that if  $X_n \rightarrow_d X$  and  $Y_n \rightarrow_d c$  for a constant c, then  $(X_n, Y_n) \rightarrow_d (X, c)$ .
- ▶ It follows from the continuous mapping theorem that if  $X_n \rightarrow_d X$  and  $Y_n \rightarrow_d c$  for a constant c, then

$$g(X_n, Y_n) \rightarrow_d g(X, c)$$

for every continuous map g.

#### Slutsky's Lemma

- As special cases to the previous result, we have the following result.
- ▶ (Slutsky's Lemma) If  $X_n \rightarrow_d X$  and  $Y_n \rightarrow_d c$  for a constant c, then
  - $\rightarrow$   $X_n + Y_n \rightarrow_d X + c$ .
  - $ightharpoonup Y_n X_n \rightarrow_d c X.$
  - $Y_n^{-1}X_n \rightarrow_d c^{-1}X$  provided that  $c \neq 0$ .

## Convergence in $L_p$

A sequence of random vectors X<sub>n</sub> is said to converge in L<sub>p</sub> to a random vector X if

$$\lim_{n\to\infty} E(\|X_n-X\|^p)\to 0.$$

- ▶ Convergence in  $L_p$  is denoted by  $X_n \rightarrow_{L_p} X$ .
- ► (Theorem) If  $X_n$  converges in  $L_p$ , then it converges to 0 in probability.
- Proof: By Chebyshev's inequality,

$$P(\|X_n - X\| \ge \epsilon) = P(\|X_n - X\|^{\rho} \ge \epsilon^{\rho}) \le \frac{E(\|X_n - X\|^{\rho})}{\epsilon^{\rho}},$$

which proves the theorem if we let  $n \to \infty$ .

# When is convergence in $L_p$ implied by convergence in probability?

- ▶ First, assume that X belongs to  $L_p$  (an obvious condition to think about the convergence in  $L_p$ ).
- ▶ Suppose that  $X_n$  is dominated by some Y that belongs to  $L_p$ , that is  $||X_n|| \le Y$  a.s. with  $EY^p < \infty$ .
- Note that

$$E(\|X_{n} - X\|^{p}) \leq E\left[1(\|X_{n} - X\| < \epsilon) \|X_{n} - X\|^{p}\right]$$

$$+ E\left[1(\|X_{n} - X\| \ge \epsilon) \|X_{n} - X\|^{p}\right]$$

$$\leq \epsilon^{p} + E\left[1(\|X_{n} - X\| \ge \epsilon)(Y + \|X\|)^{p}\right].$$

▶ Since Y + ||X|| is in  $L_p$ , we have shown that  $X_n \to_p X$  implies that  $X_n \to_{L_p} X$  by letting first  $n \to 0$  and then  $\epsilon \to 0$ .

#### Borel-Cantelli Lemma

▶ For any sequence of events  $\{E_n : n = 1,...\}$ , if

$$\sum_{n=1}^{\infty} P(E_n) < \infty,$$

then

$$P(\limsup_{n\to\infty}E_n)=0.$$

Proof: the lemma follows from the fact that

$$P(\limsup_{n\to\infty} E_n) = \lim_{n\to\infty} P(\cup_{k\geq n}^{\infty} E_k) \leq \lim_{n\to\infty} \sum_{k=n}^{\infty} P(E_k).$$

▶ Thus, to show that  $X_n \rightarrow_{a.s.} X$ , it suffices to show that there exists an  $\epsilon > 0$  satisfying

$$\sum_{n=1}^{\infty} P(\|X_n - X\| > \epsilon) < \infty.$$

## Lévy's Continuity Theorem

- ▶ Let  $X_n$  and X be random vectors in  $\mathbf{R}^k$ . Then  $X_n \to_d X$  if and only if  $Ee^{it'X_n} \to Ee^{it'X}$  for every  $t \in \mathbf{R}^k$ .
- ▶ Moreover, if  $Ee^{it'X_n}$  converges to a function  $\phi(t)$  that is continuous at zero, then  $\phi$  is the characteristic function of a random vector X and  $X_n \rightarrow_d X$ .

### Laws of Large Numbers

- Let  $\bar{X}_n$  denote sample average of iid random vectors  $X_1, \ldots, X_n$ .
- ▶ Suppose that  $E ||X_1|| < \infty$ .
- ▶ Then  $\bar{X}_n \rightarrow_{a.s.} EX_1$  (the strong law of large numbers (SLLN)).
- ▶ Obviously, this implies that  $\bar{X}_n \rightarrow_p EX_1$  (the weak law of large numbers (WLLN)).
- ▶ The WLLN can be proved easily by Chebyshev's inequality, if we make a stronger assumption that  $E \|X_1\|^2 < \infty$  (prove this by yourself).

### Laws of Large Numbers (Cont.)

- We now prove WLLN using the characteristic function.
- Let  $X_1, \ldots, X_n$  be iid random variables with characteristic function  $\phi$ .
- ▶ Then  $\bar{X}_n \to_p \mu$  if  $\phi$  is differentiable at zero with  $\phi'(0) = i\mu$ .
- ▶ Proof: Since  $\phi(0) = 1$ ,  $\phi(t) = 1 + t\phi'(0) + o(t)$  as  $t \to 0$ .
- Note that

$$Ee^{it\bar{X}_n} = \prod_{i=1}^n Ee^{itX_i/n} = [\phi(t/n)]^n$$
$$= \left[1 + \frac{t}{n}\phi'(0) + o\left(n^{-1}\right)\right]^n \to e^{it\mu},$$

which is the characteristic function of the constant variable  $\mu$ .

▶ A sufficient condition for  $\phi(t)$  to be differentiable at zero is that  $E|X_1| < \infty$ . Also, in this case,  $\mu = EX_1$ .

#### Central Limit Theorem

- Let  $X_1, ..., X_n$  be iid random variables with  $EX_1 = 0$  and  $EX_1^2 = 1$ .
- ▶ Then  $n^{1/2}\bar{X}_n \to_d N(0,1)$ .
- ▶ Proof: Since  $EX_1^2 < \infty$ ,  $\phi$  is twice differentiable. Also, note that  $\phi'(0) = iEX_1 = 0$  and  $\phi''(0) = i^2EX_1^2 = -1$ . By Taylor expansion,  $\phi(t) = 1 (t^2/2) + o(t^2)$  as  $t \to 0$ .
- Now note that

$$\begin{aligned} Ee^{itn^{1/2}\bar{X}_n} &= \prod_{i=1}^n Ee^{itX_i/n^{1/2}} = [\phi(t/n^{1/2})]^n \\ &= \left[1 - \frac{t^2}{2n} + o\left(n^{-1}\right)\right]^n \to e^{-\frac{1}{2}t^2}, \end{aligned}$$

which is the characteristic function of N(0, 1).

#### Cramér-Wold Device

- ▶ (Cramér-Wold device)  $X_n \to_d X$  if and only if  $t'X_n \to t'X$  for all  $t \in \mathbb{R}^k$ .
- In other words, the Cramér-Wold device allows to reduce higher-dimensional problems to the one-dimensional case.
- Why? (prove this by yourself).

#### **Delta Method**

Let  $Y_n$  be a sequence of random variables that satisfies

$$\sqrt{n}(Y_n - \theta) \rightarrow_d N(0, \sigma^2).$$

- For a given function g and a specific value of  $\theta$ , suppose that  $g'(\theta)$  exists and is not zero.
- Then

$$\sqrt{n}(g(Y_n)-g(\theta)) \rightarrow_d N(0,\sigma^2[g'(\theta)]^2).$$

▶ Why?  $[g(Y_n) - g(\theta)] \approx g'(\theta)(Y_n - \theta)$  (prove this rigorously by yourself).

### Delta Method: Example

- Let  $X_1, \ldots, X_n$  be iid random variables with  $EX_1 = \mu \neq 0$  and  $EX_1^2 < \infty$ .
- Then

$$\sqrt{n}\left(\frac{1}{\bar{X}_n}-\frac{1}{\mu}\right)\to_{d}N\left(0,\mu^{-4}\mathrm{Var}(X_1)\right).$$