

LECTURE X NOTES

1. Relevant Probability Background.

1.1. *Convergence in probability.* Last class we went through the epsilon, delta definition. A more intuitive definition is: A sequence of r.v.'s $X_n \xrightarrow{p} X$ if

$$P(|X_n - X| > \delta) \rightarrow 0 \text{ as } n \rightarrow \infty \text{ for any } \delta > 0$$

1.2. *Convergence in distribution (weak convergence).* There are many different ways to define convergence in distribution. One useful definition is a convergence in expectation: A sequence of r.v.'s $X_n \xrightarrow{d} X$ if

$$E[g(X_n)] \rightarrow E[g(X)] \text{ for any bounded and continuous } g$$

When $X_n \in \mathbb{R}$, this is equivalent to the CDF definition:

$$F_n(t) \rightarrow F(t) \text{ at all } t \text{ s.t. } F \text{ is continuous}$$

This definition has difficulties in the multivariate case in which there are many ways to define CDF's.

1.3. *Strength of convergence.* Convergence in probability is stronger than convergence in distribution, meaning that $X_n \xrightarrow{p} X \implies X_n \xrightarrow{d} X$, but the other implication does not hold. One notable case of the reverse implication is when a sequence of r.v.'s converge to a constant: $X_n \xrightarrow{d} C$.

1.4. *Notation.*

1.4.1. *Big O: Stochastic boundedness:* .

$$X_n = O_p(1) \text{ if } P(X_n > M) < \delta \text{ for any } \delta$$

In words, we can make the probability that this sequence goes above some bound M arbitrarily small by increasing n far enough.

1.4.2. *Small o: convergence in probability.*

$$X_n = o_p(1) \text{ if } X_n \xrightarrow{p} 0$$

Furthermore,

$$X_n \sim O_p(B_n) \text{ if } \frac{X_n}{B_n} \sim O_p(1)$$

and,

$$X_n \sim o_p(B_n) \text{ if } \frac{X_n}{B_n} \sim o_p(1)$$

1.5. *Continuous Mapping Theorem.* Let X_n be a sequence of r.v.'s:

1. $X_n \xrightarrow{p} X \implies g(X_n) \xrightarrow{p} g(X)$
2. $X_n \xrightarrow{d} X \implies g(X_n) \xrightarrow{d} g(X)$, for any continuous function g .

1.6. *Slutsky's Theorem.* First you consider the joint convergence of two sequences of r.v., then apply continuous mapping theorem.

Let $X_n \xrightarrow{d} X$ and $Y_n \xrightarrow{d} c$. Then:

1. $X_n + Y_n \xrightarrow{d} X + c$
2. $X_n Y_n \xrightarrow{d} cX$

1.7. *Thm. Delta Method.* First is a description of the theorem, and then a more applicable corollary.

Let a_n be a nonneg sequence of scalars, X_n be a nonneg sequence of r.v., and c be a constant $\in \mathbb{R}$

If:

$$a_n(X_n - c) \xrightarrow{d} Z$$

Then for any differentiable function g :

$$a_n(g(X_n) - g(c)) \xrightarrow{d} \nabla g(c)^T Z$$

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1.7.1. *Corollary.* Let X_n be asymptotically normal:

$$\sqrt{n}(X_n - c) \xrightarrow{d} N(0, \Sigma)$$

$$\sqrt{n}(g(X_n) - g(c)) \xrightarrow{d} \nabla g(c)^T Z \text{ where } Z \text{ is } N(0, \Sigma)$$

$$\text{or } \xrightarrow{d} Z \text{ where } Z \text{ is } N(0, \nabla g(c)^T \Sigma \nabla g(c))$$

2. Large Sample Theory.

2.1. Standard IID Setup.

1. $\hat{\theta}_n = \delta(X_1, \dots, X_n)$ is an estimator for the true θ^* . Because $\hat{\theta}_n$, depends on samples, it is random. δ is a fixed decision rule that does not depend on number of samples. (An example of δ would be "compute the MLE from the sample").
2. $X_i \sim IIDF_{\theta^*}$

2.2. *Consistency.* Consistency is an elementary criteria for the correctness of an estimator. Technically, consistency applies to a sequence of estimators (indexed by the sample number n), though some people say consistent estimators for short-hand. Informally, consistency means that the sequence of estimators converges to the right thing; it is asymptotically unbiased. Most statistics deals with only consistent estimators, though there are notable exceptions (which ones?)

2.2.1. Example: Bernoulli Trials. Setup:

$$X_i \sim \text{Bern}(p^*) \text{ IID}$$

We know the MLE:

$$\hat{p}_{mle} = \delta(X_1, \dots, X_n) = \frac{1}{n} \sum X_i$$

By weak law of large numbers (WLLN), we know that the sample mean converges to the expected value of the r.v.:

$$\frac{1}{n} \sum Z_i \xrightarrow{p} p^*$$

2.2.2. Example: Sample Max of Uniform. Setup:

$$X_i \sim U(0, \theta^*) \text{ IID}$$

MLE:

$$\hat{\theta}_{mle} = \max(X_i)$$

$$|\hat{\theta}_{mle} - \theta^*| = \theta^* - \max(X_i)$$

We want to show, convergence in probability.

$$P(\theta^* - \max(X_i) > \delta) \rightarrow 0$$

$$\begin{aligned}
&= P(\cap (X_i < \theta - \delta)) \\
&= \prod P(X_i < \theta - \delta) \\
&= \left(\frac{\theta - \delta}{\theta}\right)^n \rightarrow 0
\end{aligned}$$

as $n \rightarrow \infty$

2.3. Uniform Law of Large Numbers (ULLN). A set of functions satisfies ULLN if

$$\sup_{g \in G} \left| \frac{1}{n} \sum_{i=1}^n g(X_i - E(g(X))) \right| \xrightarrow{P} 0$$

If G is finite: $G = \{g_1, g_2, \dots, g_m\}$

Suppose $g_i, i \in [m]$ obeys LLN for each i . Then $G = \{g_1, g_2, \dots, g_m\}$ obey a ULLN:

$$P \left(\sup_{g \in G} \left| \frac{1}{n} \sum_{i=1}^n g(X_i - E(g(X))) \right| > \epsilon \right) \leq P \left(\bigcup_{i=1}^m \left\{ \left| \frac{1}{n} \sum_{i=1}^n g_i(X_i - E(g_i(X))) \right| > \epsilon \right\} \right) \leq \sum_{i=1}^m P \left(\left| \frac{1}{n} \sum_{i=1}^n g_i(X_i - E(g_i(X))) \right| > \epsilon \right) \rightarrow 0$$

2.3.1. Consistency of MLE.

$$\hat{\theta}_{MLE} := \arg \max_{\theta \in \Theta} \frac{1}{n} \sum_{i \in [n]} l_{X_i}(\theta)$$

Proof Sketch:

1. Show $\theta^* = \arg \max_{\theta} E[l_{X_i}(\theta)]$
2. Show $\frac{1}{n} \sum_{i=1}^n l_{X_i}(\theta) \xrightarrow{\text{uniformly}} E_{\theta^*}[l_{X_i}(\theta)]$

2.3.2. Theorem (Consistency of MLE). Let $X_i \stackrel{IID}{\sim} f_{\theta^*}$

Assume:

1. $f_{\theta^*}(X) = f_{\theta}(X)$ if $\theta_* = \theta$
2. Uniform LLN $\sup_{\theta \in \Theta} \left| \frac{1}{n} \sum_{i=1}^n \log(f_{\theta}(X_i)) - E_{\theta^*}[\log(f_{\theta}(X_i))] \right| \xrightarrow{P} 0$

Note:

1. θ_* is the unique maximizer of $E_{\theta^*}[l_{X_i}(\theta)]$
2. $E_{\theta^*}[l_{X_i}(\hat{\theta}_{MLE})] \xrightarrow{P} E_{\theta^*}[l_{X_1}(\theta^*)]$

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