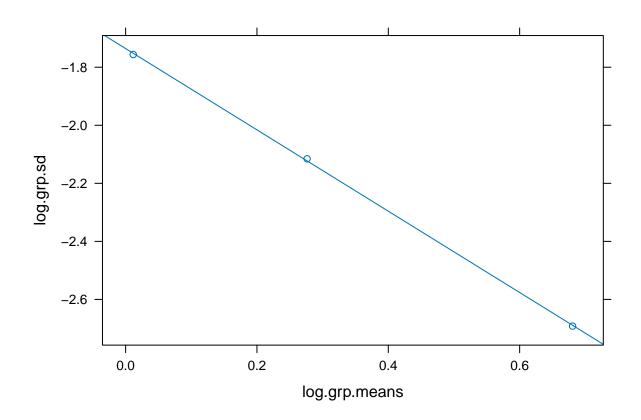
pset6

Peter

2023-10-20

```
library(Stat2Data)
library(mosaic)
## Registered S3 method overwritten by 'mosaic':
##
     method
                                       from
##
     fortify.SpatialPolygonsDataFrame ggplot2
##
## The 'mosaic' package masks several functions from core packages in order to add
## additional features. The original behavior of these functions should not be affected by this.
##
## Attaching package: 'mosaic'
## The following objects are masked from 'package:dplyr':
##
##
       count, do, tally
## The following object is masked from 'package:Matrix':
##
##
       mean
## The following object is masked from 'package:ggplot2':
##
##
       stat
## The following objects are masked from 'package:stats':
##
##
       binom.test, cor, cor.test, cov, fivenum, IQR, median, prop.test,
##
       quantile, sd, t.test, var
## The following objects are masked from 'package:base':
##
##
       max, mean, min, prod, range, sample, sum
#5.44c:
data("Olives")
log.grp.means = log(mean(Fenthion~Time, data = Olives))
```

```
log.grp.sd = log(sd(Fenthion~Time, data = Olives))
xyplot(log.grp.sd~log.grp.means, type=c("p","r"))
```



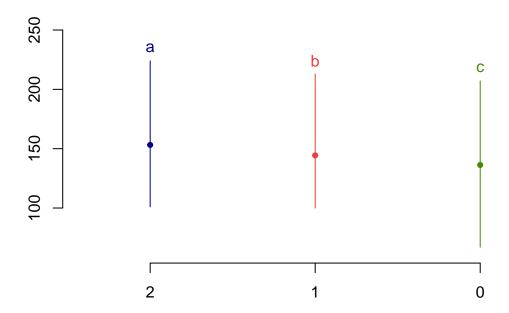
```
(trnsline = lm(log.grp.sd~log.grp.means))
```

```
##
## Call:
## lm(formula = log.grp.sd ~ log.grp.means)
##
## Coefficients:
## (Intercept) log.grp.means
## -1.736 -1.401

#if 1 - slope =
#1 => no change to y
#.5 => sqrt y
#0 => log y
#-1 => 1/y
#The diagnostic plot suggests that we do not do any transformations on Y. 1 - -1.401 = 2.401 which is g
```

```
#5.56:
data("Blood1")
#5.56a: Two sample t tests are used to determine if two population means are equal, in this case we hav
#5.56b:
?Blood1 #Overwt O=normal, 1=overweight, or 2=obese
BPcalc <- lm(SystolicBP~Overwt, data = Blood1) #factor method for more than one numeric
anova(BPcalc)
## Analysis of Variance Table
## Response: SystolicBP
              Df Sum Sq Mean Sq F value
## Overwt
             1 27788 27788.5 38.093 1.399e-09 ***
## Residuals 498 363286 729.5
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
#With f = 38.093, df = (1,498), and a pual <0.05, we can conclude that at least one alpha is != 0.
#Yes, there is a significant difference in blood pressure based on weight.
#5.56c:
Overwt1 <- factor(Blood1$Overwt)</pre>
BPcalc1 <- aov(SystolicBP ~ Overwt1, data = Blood1)</pre>
library(asbio)
## Loading required package: tcltk
library(agricolae)
LSDout <- LSD.test(BPcalc1, "Overwt1", group = TRUE, p.adj = "none")
plot(LSDout)
```

Groups and Range

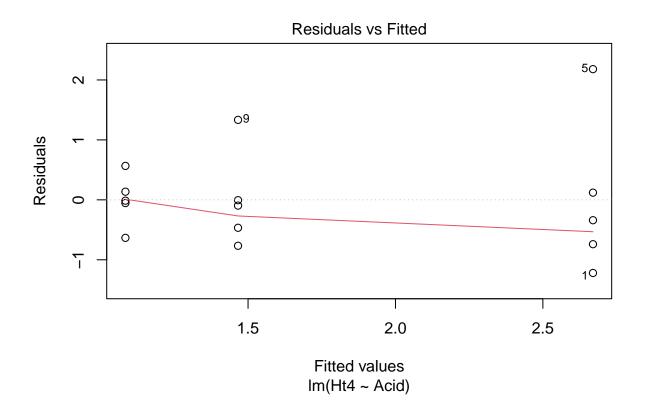


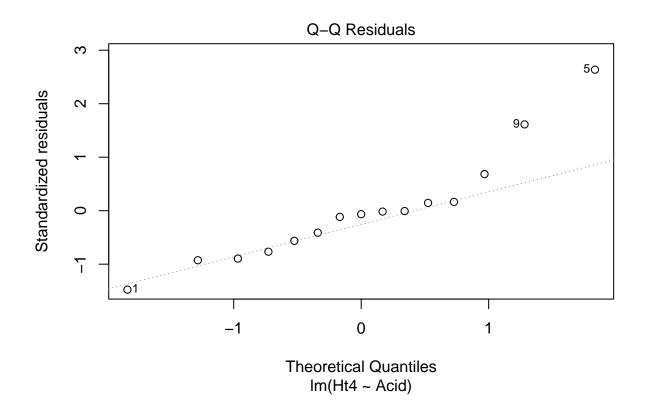
LSDout\$groups

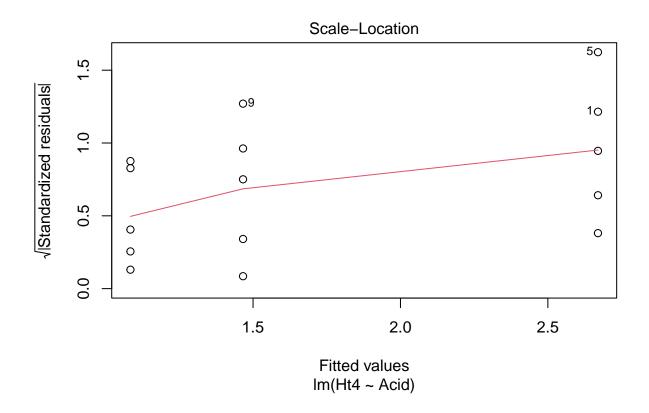
```
##
     SystolicBP groups
## 2
     153.1814
      144.3670
## 1
                     b
## 0
      136.3155
#6.2:
#6.2a: Response: Height of Tomato Plants
#6.2b: Factors: Type of Fertilizer (Factor 1), Amount of Water (Factor 2)
#6.2c: Factor 1: Experimental, 3 levels (Regular Fertilizer, Organic Fertilizer, No Fertilizer)
#Factor 2: Experimental, 2 levels (Normal Watering, Reduced Watering)
#6.2d: Complete Block Design: No
#a. Degrees of Freedom (DF) for Factor A:
\#DF\ Factor\ A = (3 - 1) = 2
#b. Degrees of Freedom (DF) for Factor B:
\#DF\ Factor\ B = (3 - 1) = 2
#c. Degrees of Freedom (DF) for Error:
\#DF\ Error = (1 - (2 + 2)) = (1 - 4) = -3
```

```
#If your dataset has one facotr of interest and one blocking factor, and you carry
#out a two-way ANOVA, your SSE will be too SMALL (small/large) and you will be MORE
#(more, less) likely to detect real differences than would a one-way ANOVA.
#6.29a: You could measure each rat twice, one time for each hormone injection (leptin and insulin), and
#like we did in AP Stats:
#Select male rats and female rats.
#Randomly assign each rat to one of two time slots.
#Time 1, inject hormones (e.g., leptin) in all rats.
#Time 2, inject hormones (e.g., insulin) in all rats.
#Measure amount eaten in Time 1 and Time 2 for each rat.
#6.29b:
#Reasons to not use blocks:
#takes much longer to do both treatments to all the rats, can't run simultaneously,
#and also uses more reasources.
#blocking minimizes variability, which may not be a good thing in this case.
#6.36a:
data("Alfalfa")
favstats(Ht4~Row, data = Alfalfa)
    Row min
                Q1 median
##
                             Q3 max mean
                                                 sd n missing
## 1 a 1.00 1.015 1.03 1.240 1.45 1.16 0.2515949 3
## 2 b 0.70 0.960 1.22 2.005 2.79 1.57 1.0880717 3
                                                            0
                    1.37 1.650 1.93 1.25 0.7472617 3
     c 0.45 0.910
                                                            0
                                                            0
## 4 d 1.65 1.990 2.33 2.565 2.80 2.26 0.5781868 3
## 5 e 1.07 1.265 1.46 3.155 4.85 2.46 2.0789661 3
#mean a: 1.16
#mean b: 1.57
#mean c: 1.25
#mean d: 2.26
#mean e: 2.46
favstats(Ht4 ~ 1, data = Alfalfa)
## 1 min Q1 median Q3 max mean
                                            sd n missing
## 1 1 0.45 1.05 1.45 2.13 4.85 1.74 1.105396 15
#mean for all 15: 1.74
#SD for all 15: 1.105396
#6.36b:
out = aov(Ht4~Acid + Row, data=Alfalfa)
anova(out)
```

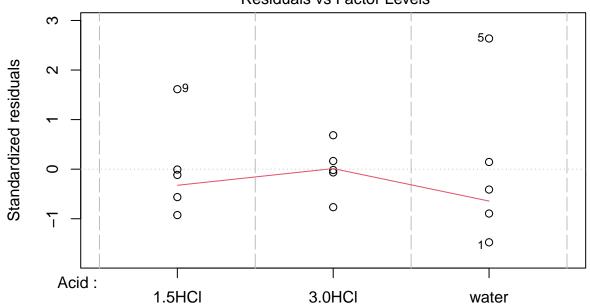
Analysis of Variance Table







Constant Leverage: Residuals vs Factor Levels



Factor Level Combinations

favstats(Ht4~Acid, data = Alfalfa)

```
## Acid min Q1 median Q3 max mean sd n missing
## 1 1.5HCl 0.70 1.00 1.37 1.46 2.80 1.466 0.8050963 5 0
## 2 3.0HCl 0.45 1.03 1.07 1.22 1.65 1.084 0.4312540 5 0
## 3 water 1.45 1.93 2.33 2.79 4.85 2.670 1.3151426 5
```

#Constant and additive: Yes the treatment are constant and additive based on the context.

#Errors zero-mean: According to the resids vs fits plot, yes the errors are 0 mean visually.

#Errors have same variance: This condition is not met, our lowest SD (.4312)*2 is #NOT greater than the largest SD (1.31). The errors do NOT have the same variance.

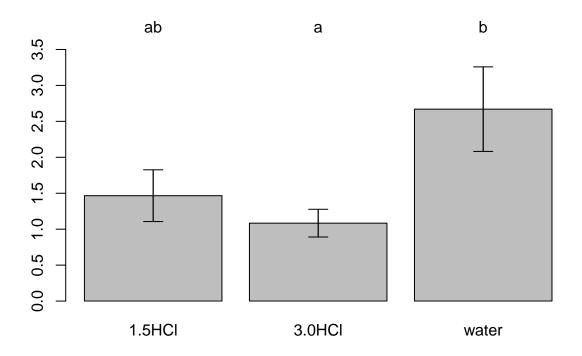
#Errors are normal: With so few observations it's hard to tell, but there is a very #concerning curve upwards on the right hand side. points 9 and 5 are especially concerning.

#errors are independent: The errors are independent. The collection of the data was #done in a good environment and no contamination could be found.

#6.36d: With an F=4.5132, df=(2, 8), p val < 0.05, there is significant evidence to #conclude that at least one alpha i is not 0, so there is a significant difference in #the average growth due to the treatments.

```
#6.36e: With an F=1.3776, df=(4,8), p val .32>0.05, I can not conclude that at #least one alpha i is not equal to 0, so I can't conclude that there is not a significant #difference in average growth due to distance from the window.
```

```
#6.37:
LSDalf <- aov(Ht4~Acid, data=Alfalfa)
library(asbio)
library(agricolae)
pairwise.t.test(Alfalfa$Ht4, Alfalfa$Acid, p.adj="none")
##
## Pairwise comparisons using t tests with pooled SD
## data: Alfalfa$Ht4 and Alfalfa$Acid
##
##
          1.5HCl 3.0HCl
## 3.0HCl 0.526 -
## water 0.062 0.019
## P value adjustment method: none
(LSDout = pairw.anova(Alfalfa$Ht4,Alfalfa$Acid,method="lsd"))
##
## 95% LSD confidence intervals
##
##
                         LSD
                              Diff
                                                Upper Decision Adj. p-value
                                       Lower
## mu1.5HCl-mu3.0HCl 1.27388 0.382 -0.89188 1.65588
                                                         FTR HO
                                                                     0.52584
                                                         FTR HO
## mu1.5HCl-muwater 1.27388 -1.204 -2.47788 0.06988
                                                                     0.06185
## mu3.0HCl-muwater 1.27388 -1.586 -2.85988 -0.31212 Reject HO
                                                                     0.01886
plot(LSDout)
```



```
##
## Bars are means. Errors are SEs.
##
## The population means of factor levels with the same letter are not
## significantly different at alpha = 0.05 using the Fisher LSD method.
plot(LSDout, type=2)
```

95% LSD confidence intervals

