

# A Look at Likelihood Ratio Tests for Variance Components in Mixed Effects Models

---

Peter Norwood

April 2019

Department of Statistics, North Carolina State University

# Table of contents

1. Introduction

2. Methods

3. Results

4. Discussion

5. Application

# Introduction

---

$$y = X\beta + Z\alpha + e$$

$$\alpha \sim N(\mu, \sigma_1^2 V_1)$$

$$e \sim N(0, \sigma^2 V)$$

- Treat some covariates as fixed values, others as random variables
- Natural Question: is there evidence of a random effect?
- How can we test whether or not there is a random effect?

# Main Question

$$H_0 : \sigma_1^2 = 0$$

$$H_a : \sigma_1^2 > 0$$

- Consider the Likelihood Ratio Test
- Asymptotically, the test statistic *should* converge to  $\chi_v^2$
- One major problem

$$\lambda(\mathbf{x}) = \frac{L(\hat{\theta}_0|\mathbf{x})}{L(\hat{\theta}|\mathbf{x})}$$

- The likelihood ratio test is based on the maximum likelihood estimation, which is restricted to the parameter space
- Parameter space:  $\sigma_1^2 \geq 0$
- $H_0 : \sigma_1^2 = 0$
- Since we are testing on the boundary, what does the test statistic converge to? (Hint: not  $\chi_v^2$ )

## p-values in Mixed Models

*"Perhaps I can try again to explain why I don't quote p-values or, more to the point, why I do not take the "obviously correct" approach of attempting to reproduce the results provided by SAS. Let me just say that, although there are those who feel that the purpose of the R Project - indeed the purpose of any statistical computing whatsoever - is to reproduce the p-values provided by SAS, I am not a member of that group." - Doug Bates, 2006*



# p-values in Mixed Models

*"Perhaps I can try again to explain why I don't quote p-values or, more to the point, why I do not take the "obviously correct" approach of attempting to reproduce the results provided by SAS. Let me just say that, although there are those who feel that the purpose of the R Project - indeed the purpose of any statistical computing whatsoever - is to reproduce the p-values provided by SAS, I am not a member of that group." - Doug Bates, 2006*





# Research Questions

1. Under the null hypothesis, what does the sampling distribution of the likelihood ratio test statistic (LRT) look like? What known distributions can we compare it to?
2. When comparing the LRT to known distributions, what is the type I error rate of the test? How often are we claiming significance when none actually exists?
3. For the different distributions, how powerful is the likelihood ratio test? How much signal is needed to be confident the likelihood ratio test will pick it up?

# Methods

---

# Random Intercept, Random Slope Model

$$y_{ij} = \beta_1 x_i + \alpha_j + \gamma_{i(j)} x_i + e_{ij}$$

$$i = 0, \dots, 9, x_i = i$$

$$j = 1, \dots, b$$

$$\alpha_j \sim N(0, \sigma_1^2), \gamma_{i(j)} \sim N(0, \sigma_2^2)$$

$$e_{ij} \sim N(0, \sigma^2)$$

We do not specify or restrict any correlation structure between  $\alpha_j$  and  $\gamma_{i(j)}$ . We assume  $\alpha_j$  and  $\gamma_{i(j)}$  are independent of  $e_{ij}$ . This is based off the *sleepstudy* dataset.

# Split Plot Model

$$y_{ijk} = \mu + \alpha_i + \gamma_{k(i)} + \beta_j + (\alpha\beta)_{ij} + e_{ijk}$$

$$i = 1, \dots, 3, j = 1, \dots, 4, k = 1, \dots, b$$

$$\gamma_{k(i)} \sim N(0, \sigma_1^2)$$

$$e_{ijk} \sim N(0, \sigma^2)$$

We assume  $e_{ijk}$  is uncorrected with  $\gamma_{k(i)}$ . Based of the *corn* dataset.

## Random Intercept, Random Slope

- Vary subjects: 2, 5, 10, 20, 50, 100, 200, 500, 1000
- Vary signal  $\sigma_1 = \sigma_2$ : 0 to 25 by 5

## Split Plot

- Vary replicates: 1, 2, 4, 8, 16, 32, 64, 128, 256
- Vary signal  $\sigma_1$ : 0 to 2.5 by 5

10,000 simulated datasets for each model under each sample size, signal combination

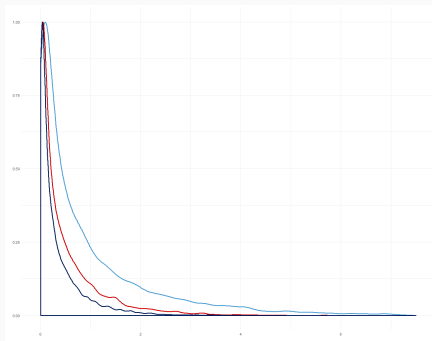
- *lme4* package does not give p-values or LRTs for random effects
- *lmerTest* gives p-values for fixed and random effects and the LRT for random effects

```
fit <- lmerTest::lmer(reaction ~ day + (day|subject),  
                     data=dat)
```

```
fit <- lmerTest::lmer(yield~pesticide*method +  
                     (1|field:pesticide), data=dat)
```

# Distributions

- Basic chi-square distribution:  $\chi_v^2$
  - Custom chi-square mixture distribution:  $p\chi_0^2 + (1-p)\chi_v^2$
  - 50/50 chi-square mixture distribution:  $.5\chi_0^2 + .5\chi_v^2$
- Zhang, Daowen and Lin, Xihong (2008)

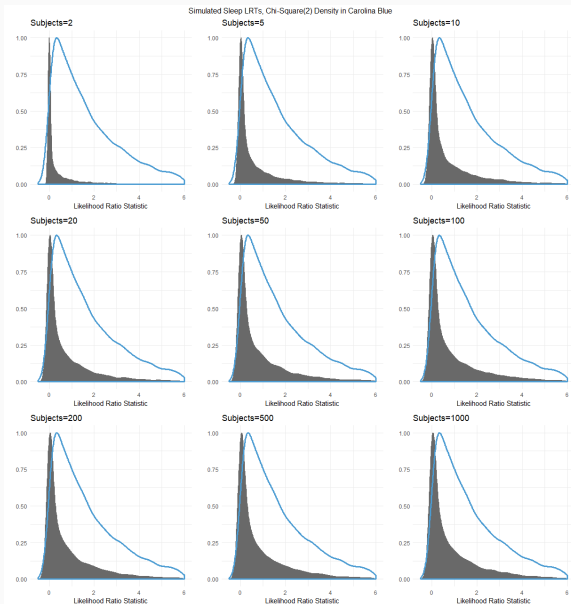


## Results

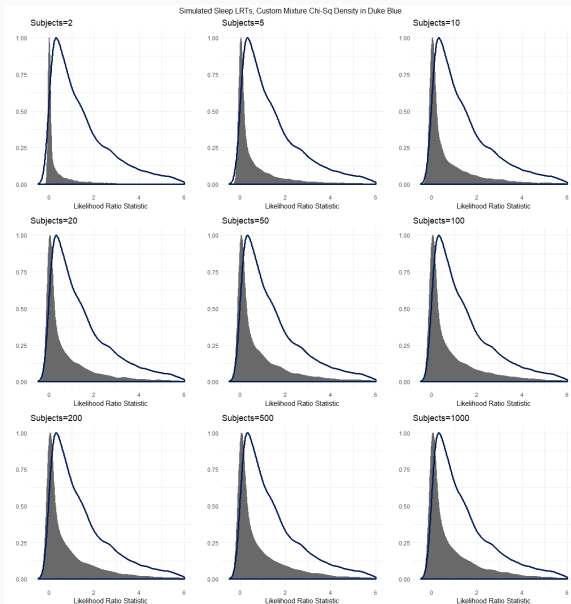
---



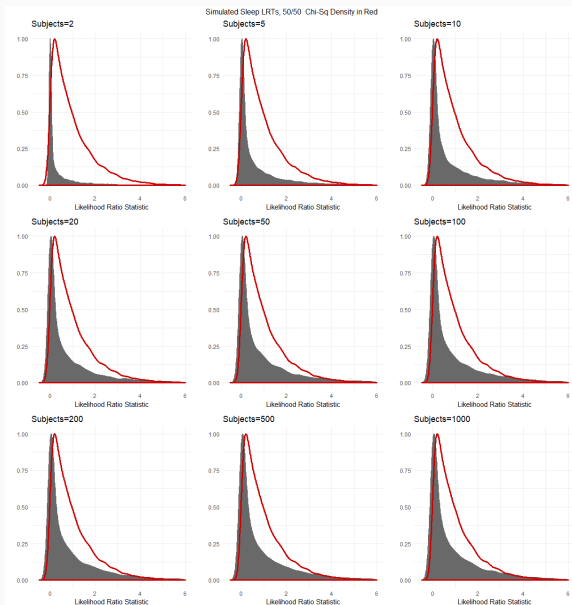
# Random Intercept, Random Slope LRT Distribution



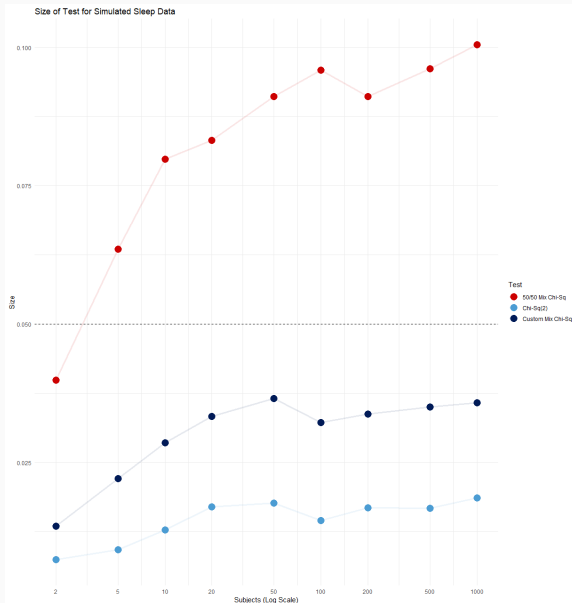
# Random Intercept, Random Slope LRT Distribution



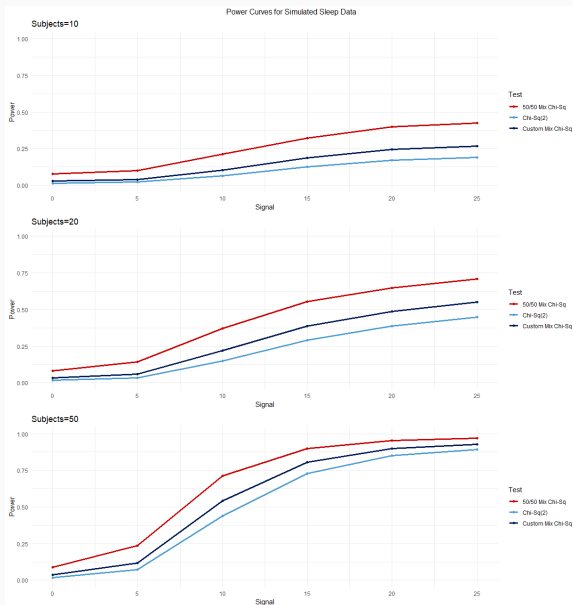
# Random Intercept, Random Slope LRT Distribution



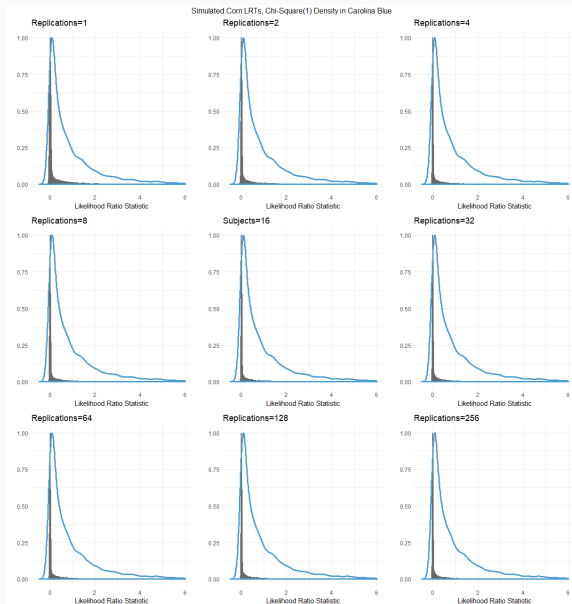
# Random Intercept, Random Slope Type I Error



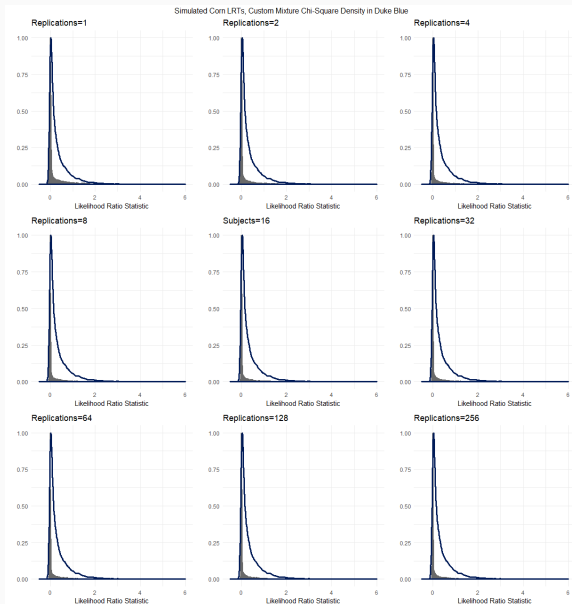
# Random Intercept, Random Slope Power



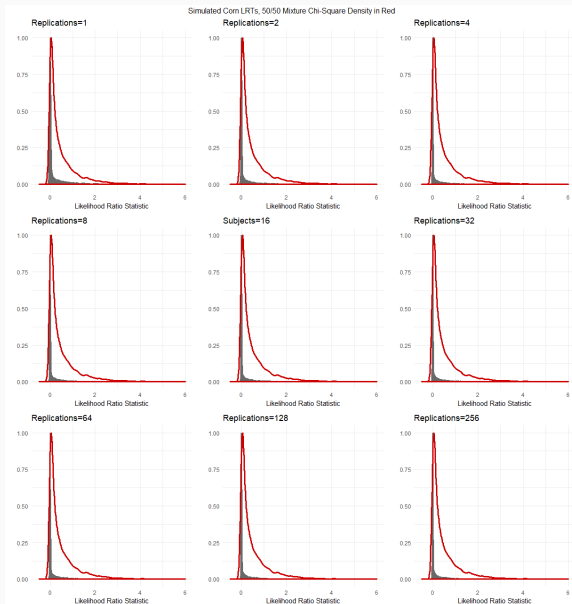
# Split Plot LRT Distribution



# Split Plot LRT Distribution

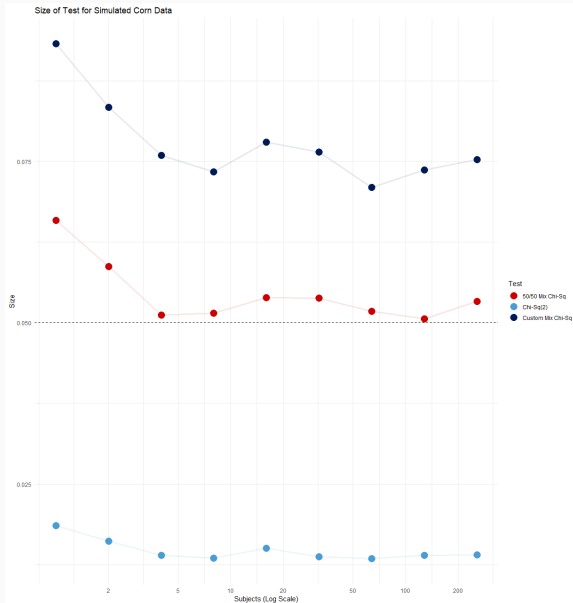


# Split Plot LRT Distribution

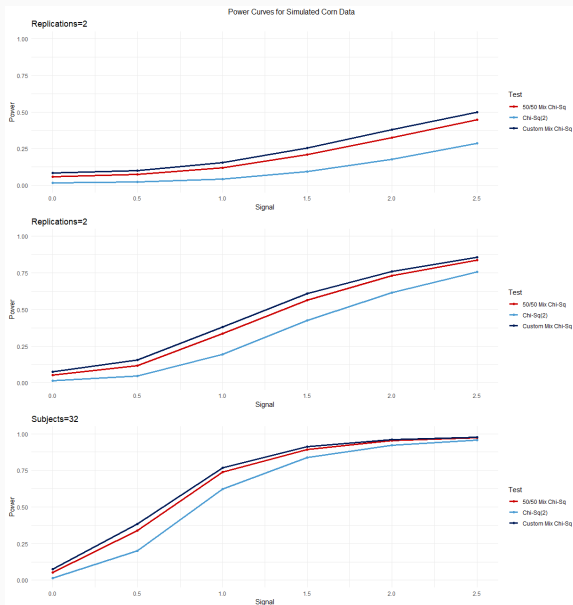




# Split Plot Type I Error



# Split Plot Power



## Discussion

---

# Chi-Square Distribution

- Does not pass the eye test
- Conservative for both models
- Lower power compared to the mixture distributions

- Visually fits better than chi-squared
- Slightly conservative for random intercept, random slope
- Liberal for split plot model
- Lower power than 50/50 for random intercept, random slope, slightly better for split plot

- Visually fits better than chi-squared
- Liberal for random intercept, random slope
- Close to correct size for split plot model
- Better power than chi-square, better or close to custom mixture

# Main Takeaways

- No one distribution is the perfect fit
- When  $p > .5$ , 50/50 looked solid. When  $p < .5$ , 50/50 was too liberal.
- If p-values are necessary consider if type I error or type II error is more grave.
- Best to look at p-values from both distributions. If they both render the same decision, we can be comfortable with the conclusion.
- Bootstrap confidence intervals and other distribution-free evaluations can be better if the answer isn't clear from these two distributions.

- Same study, more models
- Similar study, new tests. Zhang, Daowen and Lin, Xihong (2008) shows Score tests have appropriate size and solid power

Table 1: *Size and Power comparisons of the likelihood ratio tests and score tests for a single variance component based on 500 simulations under the logistic model (15)*

Method	Size		Power			
	$\psi = 0$	$\psi = 0.2$	$\psi = 0.4$	$\psi = 0.6$	$\psi = 0.8$	$\psi = 1.0$
LRT	0.034	0.370	0.790	0.922	0.990	1.000
Regular LRT	0.020	0.280	0.672	0.882	0.968	0.992
One-sided score test	0.054	0.416	0.834	0.938	0.996	1.000
Two-sided score test	0.050	0.336	0.736	0.910	0.980	0.998



## Problems Encountered

- Simulating the data correctly
- R packages
- Knowing when to stop/settle
- Computational Time

$$y_{ij} = 10 \text{ day}_j + \underbrace{\gamma_{0i}}_{\text{random intercept for each person}} + \gamma_{1j}(\text{day}_j | \text{sub}) \text{ret}_j + \epsilon_{ij}$$

$$\begin{array}{c|cccc}
 & S_1 & S_2 & S_3 & d \\
 \hline
 S_1 & 1 & 0 & 0 & 1 \\
 S_2 & 0 & 1 & 0 & 0 \\
 S_3 & 0 & 0 & 1 & 0 \\
 \hline
 \end{array}$$

# Application

---

The likelihood ratio statistic is 42.84.

Chi-Sq(2)	Custom Chi-Sq Mixture	50/50 Chi-Sq Mixture
4.99e-10	0	0

Even with the problems previously discussed, it is clear random variation is coming from the day nested within subject.

## Corn Dataset p-values

The likelihood ratio statistic is 7.30.

Chi-Sq(1)	Custom Chi-Sq Mixture	50/50 Chi-Sq Mixture
0.00691	0.00001	0.00014

Again, it is clear that the field nested within pesticide – the sub plot error – is introducing random variation.

# Thank you

- Dr. Hughes-Oliver
- Dr. Maity

# Questions?

1. Doug Bates. [r] lmer, p-values and all that.
2. Daowen Zhang and Xihong Lin. *Variance Component Testing in General-ized Linear Mixed Models for Longitudinal/Clustered Data and other RelatedTopics*, pages 19-35. 01 2008.