Midterm 1: Math 6266

Peter Williams

Section 1.1

Exercise 1. Consider the linear regression model with mean zero, uncorrelated, heteroscedastic noise:

$$Y_i = X_i^\intercal \theta + \varepsilon_i,\ for\ i = 1, .., n,\ E\varepsilon_i = 0,\ cov(\varepsilon_i, \varepsilon_j) = \begin{cases} \sigma_i^2, & \text{if } i = j \\ 0, & i \neq j \end{cases} \tag{1}$$

Find expressions for the LSE and response estimator in this model

To set up the problem, take $W^{-1} = diag\{\sigma_1^2, ..., \sigma_n^2\}$, $W = diag\{\frac{1}{\sigma_1^2}, ..., \frac{1}{\sigma_n^2}\}, W^{1/2} = diag\{\sqrt{\frac{1}{\sigma_1^2}}, ..., \sqrt{\frac{1}{\sigma_n^2}}\}$, with $W^\intercal = W$, and $W^{1/2}W^{1/2} = W$, since they are diagonal matrices. Also we will use $w_i = \frac{1}{\sigma_i^2} = W_{ii}$.

Under heteroscedastic noise assumptions, we define the least squares estimator, denoted $\hat{\theta}$, as:

$$\hat{\theta} = \underset{\theta}{argmin} \sum_{i=1}^{n} w_i (Y_i - X_i^\intercal \theta)^2 = \underset{\theta}{argmin} \sum_{i=1}^{n} (\sqrt{w_i} Y_i - \sqrt{w_i} X_i^\intercal \theta)^2 = \underset{\theta}{argmin} ||W^{1/2}Y - W^{1/2}X^\intercal \theta||^2$$

$G(\theta) = ||W^{1/2}Y - W^{1/2}X^\intercal\theta||^2 = (W^{1/2}Y - W^{1/2}X^\intercal\theta)^\intercal(W^{1/2}Y - W^{1/2}X^\intercal\theta) = Y^\intercal WY - 2\theta^\intercal XWY + \theta^\intercal XWX^\intercal\theta$

with gradient,

$$\nabla G(\theta) = -2XWY + 2XWX^\intercal\theta$$

Setting this expression equal to zero leads to estimator $\hat{\theta} = (XWX^\intercal)^{-1}XWY$, which leads to response estimator $\hat{Y} = X^\intercal\hat{\theta} = X^\intercal(XWX^\intercal)^{-1}XWY$.

Exercise 2. Assume that $\varepsilon_i \sim N(0, \sigma_i^2)$ in the previous problem. What is known about the distribution of $\hat{\theta}$ and $\hat{Y}$?

For $\hat{\theta}$, we have,

$$E[\hat{\theta}] = E[(XWX^\intercal)^{-1}XWY] = E[(XWX^\intercal)^{-1}XW(X^\intercal\theta^* + \varepsilon)] = E[\theta^*] + E[(XWX^\intercal)^{-1}XW\varepsilon] = \theta^*$$

indicating that $\hat{\theta}$ is unbiased. Further $\hat{\theta}$ is normally distributed, since is a linear transformation of $\varepsilon \sim N(0, W^{-1})$. Further we have,

$$Var(\hat{\theta}) = Var((XWX^\intercal)^{-1}XWY) = Var((XWX^\intercal)^{-1}XW(X^\intercal\theta^* + \varepsilon)) = Var((XWX^\intercal)^{-1}XW\varepsilon)) = ...$$

$$= (XWX^\intercal)^{-1}XWVar(\varepsilon)W^\intercal X^\intercal(XWX^\intercal)^{-1} = (XWX^\intercal)^{-1}XWX^\intercal(XWX^\intercal)^{-1} = (XWX^\intercal)^{-1} = Var(\hat{\theta})$$

For $\hat{Y}$ we have,

$E[\hat{Y}] = E[X^\intercal(XWX^\intercal)^{-1}XWY] = E[X^\intercal(XWX^\intercal)^{-1}XW(X^\intercal\theta^* + \varepsilon)] = E[X^\intercal\theta^* + X^\intercal(XWX^\intercal)^{-1}XW\varepsilon] = E[X^\intercal\theta^*] = Y$

and,

$Var[\hat{Y}] = Var[X^\intercal(XWX^\intercal)^{-1}XWY] = Var[X^\intercal(XWX^\intercal)^{-1}XW(X^\intercal\theta^* + \varepsilon)] = Var[X^\intercal\theta^* + X^\intercal(XWX^\intercal)^{-1}XW\varepsilon] = ...$

$$... = Var[X^\intercal(XWX^\intercal)^{-1}XW\varepsilon] = X^\intercal(XWX^\intercal)^{-1}XW\ Var(\varepsilon)\ W^\intercal X^\intercal(XWX^\intercal)^{-1}X = ...$$

$$= X^\intercal(XWX^\intercal)^{-1}XWX^\intercal(XWX^\intercal)^{-1}X = X^\intercal(XWX^\intercal)^{-1}X = Var[\hat{Y}]$$

Now suppose additionally that $\sigma_i^2 \equiv \sigma^2 > 0$. What can be said about distribution of the estimator $\hat{\sigma^2}$?

With $\sigma_i^2 \equiv \sigma^2 > 0$, we have $\hat{\sigma^2} = \frac{||Y - X^\intercal \hat{\theta}||^2}{n-p} = \frac{||\hat{\varepsilon}||^2}{n-p}$. Further denote, $||\hat{\varepsilon}|| = ||Y - \hat{Y}|| = ||Y - \Pi Y|| = ||(I_n - \Pi)Y||$, also noting that $(I_n - \Pi)X^\intercal = X^\intercal - \Pi X^\intercal = X^\intercal - X^\intercal(XX^\intercal)^{-1}XX^\intercal = X^\intercal - X^\intercal = 0$.

Then we have,

$$(n-p)E[\hat{\sigma^2}] = E||Y - X^\intercal \hat{\theta}||^2 = E||\hat{\varepsilon}||^2 = E[tr(\hat{\varepsilon}\hat{\varepsilon}^\intercal)] = E[tr((I_n - \Pi)YY^\intercal(I_n - \Pi))] = ...$$

$$= E[tr((I_n - \Pi)(X^\intercal\theta^* + \varepsilon)(X^\intercal\theta^* + \varepsilon)^\intercal(I_n - \Pi))] = E[tr((I_n - \Pi)\varepsilon\varepsilon^\intercal(I_n - \Pi))] = tr((I_n - \Pi)E[\varepsilon\varepsilon^\intercal]) = ...$$

Using the cylic property of the trace operator, the property that $(I_n - \Pi)(I_n - \Pi) = (I_n - \Pi)$, and the expectation $E[\varepsilon\varepsilon^\intercal] = \sigma^2 I_n$, leading to

$$... = \sigma^2 tr(I_n - \Pi) = \sigma^2(n-p) = (n-p)E[\hat{\sigma}^2]$$

Looking further at the distribution of $||Y - X^\intercal\hat{\theta}||^2 = \hat{\varepsilon}^\intercal\hat{\varepsilon}$, we have

$$\hat{\varepsilon}^\intercal\hat{\varepsilon} = ((I_n - \Pi)Y)^\intercal((I_n - \Pi)Y) = Y^\intercal(I_n - \Pi)Y = (X^\intercal\theta^* + \varepsilon)^\intercal(I_n - \Pi)(X^\intercal\theta^* + \varepsilon) = \varepsilon^\intercal(I_n - \Pi)\varepsilon$$

Since we know that $\varepsilon \sim N(0, \sigma^2 I_n)$, and further $\frac{\varepsilon^\intercal\varepsilon}{\sigma^2} \sim \chi^2(n)$, $(\frac{\varepsilon}{\sigma})^\intercal(I_n - \Pi)(\frac{\varepsilon}{\sigma}) \sim \chi^2(n-p)$, since we know from earlier that $(I_n - \Pi)$, is idempotent, with rank equal to $tr(I_n - \Pi) = tr(I_n) - tr(\Pi) = n - p$.

Section 1.3

Exercise 4. Let $A \in R^{n\times n}$ be a matrix (corresponding to a linear map in R^n). Show that A preserves length for all $x \in R^n$ iff it preserves the inner product. I.e. one needs to show the following:

$||Ax|| = ||x|| \ \forall \ x \in R^n \iff (Ax)^\intercal(Ay) \ \forall \ x, y \in R^n$.

Take,

$$||x|| = \sqrt{x \cdot x} = \sqrt{x^\intercal x} \implies ||Ax|| = \sqrt{Ax \cdot Ax} = \sqrt{x^\intercal A^\intercal Ax} \implies$$

,

$$A^\intercal A = I_n = A^{-1}, \ A^\intercal = A^{-1}, ||Ax|| = ||x||$$

this implies A is an orthogonal matrix, and further,

$$(Ax)^\intercal(Ay) = ||AxAy||^2 = x^\intercal A^\intercal Ay = x^\intercal y = ||xy||^2$$

Exercise 5. (a) Let $x_0 \in R^n$ be some fixed vector, find a projection map on the subspace $span(x_0)$. Compare your result with matrix Π (from section 1.3) for the case of $p = 1$.

Let $x = span(x_0) = span(x_1, x_2, .., x_n)$, denote the subspace of interest, and $x_1, x_2, ...$ are basis vectors and $y = (y_1, y_2, ..., y_n)^\intercal$. The projection map is,

$$Proj_x(y) = \frac{y \cdot x}{y \cdot y}x = \sum_{i=1}^{n} \frac{y_i \cdot \ x_i}{y_i \cdot y_i}x_i$$

For the case $p = 1$, and $\Pi = X^\intercal(XX^\intercal)^{-1}X, X^\intercal \in R^n$, we have,

$$\Pi y = \hat{y} = X^\intercal(XX^\intercal)^{-1}Xy = X^\intercal\frac{Xy}{XX^\intercal} = \frac{\sum_i^n x_i y_i}{\sum_i^n x_i^2}(x_1, x_2, ..., x_n)^\intercal = \frac{< X \cdot y >}{< y \cdot y >}X^\intercal = Proj_X(y)$$

(b) Prove part 3) of Lemma 1.1 for an arbitrary orthogonal projection in R^n. Show $\forall h \in R^n$, $||h||^2 = ||\Pi h||^2 + ||h - \Pi h||^2$.

Using the fact that $(I_n - \Pi)^\intercal (I_n - \Pi) = I_n - 2\Pi + \Pi = I_n - \Pi$, we have,

$$||h||^2 = ||\Pi h||^2 + ||h - \Pi h||^2 = h^\intercal \Pi^\intercal \Pi h + h^\intercal (I_n - \Pi)^\intercal (I_n - \Pi) h = h^\intercal \Pi h + h^\intercal (I_n - \Pi) h = h^\intercal I_n h + h^\intercal \Pi h - h^\intercal \Pi h = ||h||^2$$

Exercise 6. Let L_1, L_2 be some subspaces in R^n, and $L_2 \subseteq L_1 \subseteq R^n$. Let P_{L_1}, P_{L_2} denote orthogonal projections on these subspaces. Prove the following properties:
(a) $P_{L_2} - P_{L_1}$ is an orthogonal projection,
Denote L_1 as a subset of R^n with orthonormal basis $span\{u_1, u_2, ..., u_p\}$, and L_2 with basis $span\{u_1, u_2, ..., u_{p-k}\} \subseteq span\{u_1, ..., u_p\}$. For a vector $x \in R^n$, we have an orthogonal projection onto L_1 and L_2 denoted as follows:

$$P_{L_1}(x) = \sum_{i=1}^{p} (x \cdot u_i) u_i, \ P_{L_2}(x) = \sum_{i=1}^{p-k} (x \cdot u_i) u_i$$

The difference of these projections is then:

$$P_{L_2}(x) - P_{L_1}(x) = (P_{L_2} - P_{L_1}) x = \sum_{i=1}^{p-k} (x \cdot u_i) u_i - \sum_{i=1}^{p} (x \cdot u_i) u_i = (-1) \cdot \sum_{i=p-k+1}^{p} (x \cdot u_i) u_i$$

which is an orthogonal projection onto the subspace, defined as $span\{u_{p-k+1}, u_{p-k+2}, ..., u_p\} \subseteq span\{u_1, ..., u_p\}$.

(b) $||PL2x|| \leq ||PL1x|| \ \forall x \in R^n$,

We have $||P_{L_2} x|| = || \sum_{i=1}^{p-k} (x \cdot u_i) u_i||$ and $||P_{L_1} x|| = || \sum_{i=1}^{p} (x \cdot u_i) u_i||$. For $k < p$, we have

$$||P_{L_1} x - P_{L_2}|| = || \sum_{i=p-k+1}^{p} (x \cdot u_i) u_i|| \geq 0 \ ,$$

and

$$||P_{L_2} x|| \leq ||P_{L_1}|| = ||P_{L_1} x - P_{L_2} x + P_{L_2} x|| \leq ||P_{L_1} x - P_{L_2} x|| \ ||P_{L_2} x||$$

(c) $PL2 \cdot PL1 = PL2$
We can denote $P_{L_1}(x) = \sum_{i=1}^{p} (x \cdot u_i) u_i = U U^\intercal x$, where matrix $U_{n \times p}$ consists of orthnormal vectors $[u_1, ..., u_p]$, and denote

$$P_{L_2}(x) = \sum_{i=1}^{p-k} (x \cdot u_i) u_i = V V^\intercal x$$

where matrix $V_{n \times (p-k)}$ consists of orthnormal vectors $[u_1, ..., u_{p-k}]$. So the product $P_{L_2} P_{L_1}$ can be written

$$P_{L_2} P_{L_1} = V V^\intercal U U^\intercal$$

Since the first $p - k$ column vectors of V and U are the same, and orthonormal, the inner product $V^\intercal U$ generates a $(p - k) \times p$ block matrix of the form $\left[\begin{array}{cc} I_{p-k} & 0 \end{array} \right]$ where 0 is a $k \times k$ matrix of zeroes. We then have

$$P_{L_2} P_{L_1} = V V^\intercal U U^\intercal = V \left[\begin{array}{cc} I_{p-k} & 0 \end{array} \right] U^\intercal = V V^\intercal = P_{L_2}$$

Section 2.1

Exercise 8. Let $X \sim N(0, I_n)$, $Q = X^\intercal X$. Suppose that Q is decomposed into the sum of two quadratic forms: $Q = Q1 + Q2$, where $Qi = X^\intercal A_i X$, $i = 1, 2$ for some symmetric matrices $A1, A2$ with $rank(A1) = n1$ and $rank(A2) = n2$. Show that if $n1 + n2 = n$, then $Q1$ and $Q2$ are independent and $Q_i \sim \chi^2(n_i) for$ $i = 1, 2$.

First note that $X^\intercal X \sim \chi^2(n)$, since $X^\intercal X = \sum_{i=1}^n x_i^2$, which is the sum of iid squared normal random variables with variance 1.

Since $A1$ is a symmetric matrix, we can diagonalize it, $A_1 = U^\intercal \Lambda U$. We know the rank of A_1 is n_1. This implies that $U^\intercal A_1 U = \Lambda = diag\{\Lambda_1, ..., \Lambda_{n_1}, ..., \Lambda_n\}$, has n_1 non-zero, positive eigenvalues, and n_2 eigenvalues that equal zero.

Using the orthogonal matrix U from the decomposition of A_1, we set $X = UY$, so that $X^\intercal X = Y^\intercal U^\intercal UY = Y^\intercal I_n Y = Y^\intercal Y$. So $Q = X^\intercal X = Y^\intercal Y = \sum_{i=1}^n Y_i^2$.

We can write

$$Q = Q_1 + Q_2 = \sum_{i=1}^n Y_i^2 = Y^\intercal U^\intercal A_1 UY + Y^\intercal U^\intercal A_2 UY = Y^\intercal \Lambda Y + Y^\intercal U^\intercal A_2 UY = \sum_{i=1}^n \Lambda_i Y_i^2 + Y^\intercal U^\intercal A_2 UY$$

Since only n_1 eigenvalues in Λ are non-zero, we have

$$Q = \sum_{i=1}^{n_1} \Lambda_i Y_i^2 + \sum_{i=n_1+1}^{n} \Lambda_i Y_i^2 + Y^\intercal U^\intercal A_2 UY = Q = \sum_{i=1}^{n_1} \Lambda_i Y_i^2 + Y^\intercal U^\intercal A_2 UY$$

,

if we organize Λ in way such that the positive eigenvalues on the diagonal are present in the first n_1 diagonal elements. So we have $Q_1 = \sum_{i=1}^{n_1} \Lambda_i Y_i^2$

To solve for $Q_2 = X^\intercal X = Y^\intercal U^\intercal A_2 UY$, from above we have

$$Y^\intercal U^\intercal A_2 UY = Q - Q_1 = Q - \sum_{i=1}^{n_1} \Lambda_i Y_i^2 = \sum_{i=1}^{n_1} Y_i^2 + \sum_{i=n_1+1}^{n} Y_i^2 - \sum_{i=1}^{n_1} \Lambda_i Y_i^2 = \sum_{i=1}^{n_1} (1 - \Lambda_i) Y_i^2 + \sum_{i=n_1+1}^{n} Y_i^2$$

We know the rank of A_2 is $n_2 = n - n_1$. So the term $\sum_{i=1}^{n_1} (1 - \Lambda_i) Y_i^2$ must equal zero, implying that $\Lambda_1 = \Lambda_2 = ... = \Lambda_{n_1} = 1$. This also implies $Q = Q1 + Q2 = \sum_{i=1+1}^{n_1} Y_i^2 + \sum_{i=n_1+1}^{n} Y_i^2$.

Since each squared element $Y_i^2 = X_i^2 \sim \chi^2(1)$ in Q only occurs once in the summand, we can say that and $Q_1 = \sum_{i=1}^{n_1} Y_i^2 \sim \chi^2(n_1)$, and $Q_2 = \sum_{i=n_1+1}^{n} Y_i^2 \sim \chi^2(n_2)$, since $Q = Q_1 + Q_2 \sim \chi^2(n)$.

Section 2.2

Exercise 10. (a) Consider model (3) for $p = 2, X_i = (1, x_i)^\intercal, \theta^* = (\theta_1^*, \theta_2^*)^\intercal$ (similarly to section 1.5). Write explicit expressions for the confidence sets for $\theta^*, \theta_1^*, \theta_2^*$.

To set up explicit expression for the case above, we have:

$$XX^\intercal = \begin{bmatrix} 1 & ... & 1 \\ x_1 & ... & x_n \end{bmatrix} \begin{bmatrix} 1 & x_1 \\ ... & ... \\ 1 & x_n \end{bmatrix} = \begin{bmatrix} n & \sum_{i=1}^n x_i \\ \sum_{i=1}^n x_i & \sum_{i=1}^n x_i^2 \end{bmatrix}$$

and $det(XX^\intercal) = n \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i)^2 = n \sum_{i=1}^n (x_i - \bar{x})^2$, and

$$(XX^\intercal)^{-1} = \frac{n}{det(XX^\intercal)} \begin{bmatrix} \sum_{i=1}^n x_i^2 & -\bar{x} \\ -\bar{x} & 1 \end{bmatrix}$$

So we have

$$\hat{\theta} = (XX^\intercal)^{-1}XY = \frac{n}{det(XX^\intercal)}\begin{bmatrix} \sum_{i=1}^n x_i^2 & -\bar{x} \\ -\bar{x} & 1 \end{bmatrix}\begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n x_i y_i \end{bmatrix} = (\hat{\theta}_1, \hat{\theta}_2)^\intercal = \ ...$$

$$... \ = \frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2}\begin{bmatrix} \bar{y}\sum_i x_i^2 - \bar{x}\sum_i x_i y_i \\ \sum_i x_i y_i - n\bar{y}\bar{x} \end{bmatrix} = (\hat{\theta}_1, \hat{\theta}_2)^\intercal = \hat{\theta}$$

To find a confidence region for θ^*, using a mixture of matrix and summation notation, we use the property:

$$\frac{||(XX^\intercal)^{1/2}(\hat{\theta} - \theta^*)||^2}{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2}x_i)^2}\frac{n-2}{2} \sim F(2, n-2)$$

and denote $\hat{\sigma}^2 = \frac{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2}x_i)^2}{n-2}$. Where F denotes the F distribution with $df_1 = 2$, and $df_2 = n-2$.

We can create a confidence interval for θ^*, such that, qF_α denotes the α^{th} quantile for $F(2, n-2)$.

$$P(\frac{||(XX^\intercal)^{1/2}(\hat{\theta} - \theta^*)||^2}{p\hat{\sigma}^2} < qF_{1-\alpha}) = 1 - \alpha = P((\hat{\theta} - \theta^*)^\intercal \begin{bmatrix} n & \sum_{i=1}^n x_i \\ \sum_{i=1}^n x_i & \sum_{i=1}^n x_i^2 \end{bmatrix} (\hat{\theta} - \theta^*) < p\hat{\sigma}^2 qF_{1-\alpha})$$

We know that $\frac{(XX^\intercal)^{1/2}(\hat{\theta} - \theta^*)}{\sigma} \sim N(0, I_p)$. We can then set up confidence intervals for θ_1^* and θ_2^*.

For θ_1^*, we can set up a T-statistic by taking the difference of the first parameter estimate and the true estimate and dividing it the corresponding standard error:

$$T_{1(n-2-1)} = \frac{\hat{\theta_1} - \theta_1^*}{\sqrt{\hat{\sigma^2}[(XX^\intercal)^{-1}]_{11}}} = \frac{\hat{\theta_1} - \theta_1^*}{\sqrt{\frac{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2}x_i)^2}{n-p}\frac{\sum_{i=1}^n x_i^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}}$$

Using T_1 we can set up a % $100(1-\alpha)$ confidence interval for $\hat{\theta_1^*}$ via:

$$\hat{\theta_1^*} \pm T_{1(n-3),\alpha/2}\sqrt{\frac{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2}x_i)^2}{n-p}\frac{\sum_{i=1}^n x_i^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$

For θ_2^* we have:

$$T_{2(n-3)} = \frac{\hat{\theta_2} - \theta_2^*}{\sqrt{\hat{\sigma^2}[(XX^\intercal)^{-1}]_{22}}} = \frac{\hat{\theta_2} - \theta_2^*}{\sqrt{\frac{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2}x_i)^2}{n-p}\frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2}}}$$

With T_2 we can set up a % $100(1-\alpha)$ confidence interval for $\hat{\theta_2^*}$ via:

$$\theta_2^* \pm T_{2(n-3),\alpha/2}\sqrt{\frac{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2}x_i)^2}{(n-p)\sum_{i=1}^n (x_i - \bar{x})^2}}$$

(b) Find a confidence interval for the expected response $E[Y_i]$ in the model in part (a). The variance of the expected response $var(\hat{Y}) = var(X^\intercal(XX^\intercal)^{-1}XY) = var(X^\intercal(XX^\intercal)^{-1}X(X^\intercal\theta^* + \varepsilon)) = var(X^\intercal(XX^\intercal)^{-1}X\varepsilon) = \sigma^2 X^\intercal(XX^\intercal)^{-1}X$. Using the standard error for $\hat{Y}$, we can set up up the following confidence interval for the expected response for the i^{th} record using a T-statistic:

$$T_{(n-3)} = \frac{\hat{y}_i - y_i}{\sqrt{\hat{\sigma^2} x_i^\intercal (XX^\intercal)^{-1} x_i}} = \frac{\hat{y}_i - y_i}{\sqrt{\frac{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2} x_i)^2}{n-2} x_i^\intercal (XX^\intercal)^{-1} x_i}}$$

With this statistic a % $100(1-\alpha)$ confidence interval for y_i can be created:

$$y_i \pm T_{n-3,\alpha/2} \sqrt{\frac{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2} x_i)^2}{n-2} x_i^\intercal \frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2} \begin{bmatrix} \sum_{i=1}^n x_i^2 & -\bar{x} \\ -\bar{x} & 1 \end{bmatrix} x_i}$$