

Midterm 1: Math 6266

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Section 1.1

Exercise 1. Consider the linear regression model with mean zero, uncorrelated, heteroscedastic noise:

$$Y_i = X_i^\top \theta + \varepsilon_i, \text{ for } i = 1, \dots, n, \quad E\varepsilon_i = 0, \quad \text{cov}(\varepsilon_i, \varepsilon_j) = \begin{cases} \sigma_i^2, & \text{if } i = j \\ 0, & \text{if } i \neq j \end{cases} \quad (1)$$

Find expressions for the LSE and response estimator in this model

Under heteroscedastic noise assumptions, the LSE estimator, denoted $\hat{\theta}_{OLS}$, is:

$$\hat{\theta}_{OLS} = \underset{\theta}{\operatorname{argmin}} \|Y - X^\top \theta\|^2 = \underset{\theta}{\operatorname{argmin}} G(\theta)$$

,

$$\|Y - X^\top \theta\|^2 = G(\theta) = (Y - X^\top \theta)^\top (Y - X^\top \theta) = Y^\top Y - 2\theta^\top X Y + \theta^\top X X^\top \theta$$

with gradient,

$$\nabla G(\theta) = -2XY + 2XX^\top \theta$$

Setting this expression equal to zero leads to estimator $\hat{\theta} = \hat{\theta}_{OLS} = (XX^\top)^{-1}XY$, which leads to response estimator $\hat{Y} = X^\top \hat{\theta} = X^\top (XX^\top)^{-1}XY$.

Exercise 2. Assume that $\varepsilon_i \sim N(0, \sigma_i^2)$ in the previous problem. What is known about the distribution of $\hat{\theta}$ and \hat{Y} ? Denote $n \times n$ matrix $D = \operatorname{diag}\{\sigma_1^2, \sigma_2^2, \dots, \sigma_n^2\} = \operatorname{Var}(\varepsilon)$.

For $\hat{\theta}$, we have,

$$E[\hat{\theta}] = E[(XX^\top)^{-1}XY] = E[(XX^\top)^{-1}X(X^\top \theta^* + \varepsilon)] = E[\theta^*] + E[\varepsilon] = \theta^*$$

indicating that $\hat{\theta}$ is unbiased despite the presence of heteroscedastic noise. Further $\hat{\theta}$ is normally distributed, since is a linear transformation of $\varepsilon \sim N(0, D)$. Further we have,

$$\begin{aligned} \operatorname{Var}(\hat{\theta}) &= \operatorname{Var}((XX^\top)^{-1}XY) = \operatorname{Var}((XX^\top)^{-1}X(X^\top \theta^* + \varepsilon)) = \operatorname{Var}((XX^\top)^{-1}X\varepsilon) = \\ &= (XX^\top)^{-1}X \operatorname{Var}(\varepsilon) X^\top (XX^\top)^{-1} = (XX^\top)^{-1}X D X^\top (XX^\top)^{-1} = \operatorname{Var}(\hat{\theta}) \end{aligned}$$

For \hat{Y} we have,

$$E[\hat{Y}] = E[X^\top (XX^\top)^{-1}XY] = E[X^\top (XX^\top)^{-1}X(X^\top \theta^* + \varepsilon)] = E[X^\top \theta^* + X^\top (XX^\top)^{-1}X\varepsilon] = E[X^\top \theta^*] = Y$$

and,

$$\begin{aligned} \operatorname{Var}[\hat{Y}] &= \operatorname{Var}[X^\top (XX^\top)^{-1}XY] = \operatorname{Var}[X^\top (XX^\top)^{-1}X(X^\top \theta^* + \varepsilon)] = \operatorname{Var}[X^\top \theta^* + X^\top (XX^\top)^{-1}X\varepsilon] = \dots \\ &= \operatorname{Var}[X^\top (XX^\top)^{-1}X\varepsilon] = X^\top (XX^\top)^{-1}X \operatorname{Var}(\varepsilon) X^\top (XX^\top)^{-1}X = \Pi D \Pi^\top \end{aligned}$$

where $\Pi = X^\top (XX^\top)^{-1}X = \Pi^\top$, and $D = \operatorname{diag}\{\sigma_1^2, \sigma_2^2, \dots, \sigma_n^2\}$.

Now suppose additionally that $\sigma_i^2 \equiv \sigma^2 > 0$. What can be said about distribution of the estimator $\hat{\sigma}^2$?

With $\sigma_i^2 \equiv \sigma^2 > 0$, we have $\hat{\sigma}^2 = \frac{\|Y - X^\top \hat{\theta}\|^2}{n-p} = \frac{\|\hat{\varepsilon}\|^2}{n-p}$. Further denote, $\|\hat{\varepsilon}\| = \|Y - \hat{Y}\| = \|Y - \Pi Y\| = \|(I_n - \Pi)Y\|$, also noting that $(I_n - \Pi)X^\top = X^\top - \Pi X^\top = X^\top - X^\top (XX^\top)^{-1}XX^\top = X^\top - X^\top = 0$.

Then we have,

$$(n-p)E[\hat{\sigma}^2] = E[\|Y - X^\top \hat{\theta}\|^2] = E[\|\hat{\varepsilon}\|^2] = E[\operatorname{tr}(\hat{\varepsilon}\hat{\varepsilon}^\top)] = E[\operatorname{tr}((I_n - \Pi)YY^\top(I_n - \Pi))] = \dots$$

,

$$\dots = E[\text{tr}((I_n - \Pi)(X^\top \theta^* + \varepsilon)(X^\top \theta^* + \varepsilon)^\top (I_n - \Pi))] = E[\text{tr}((I_n - \Pi)\varepsilon\varepsilon^\top (I_n - \Pi))] = \text{tr}((I_n - \Pi)E[\varepsilon\varepsilon^\top]) = \dots$$

Using the cyclic property of the trace operator, the property that $(I_n - \Pi)(I_n - \Pi) = (I_n - \Pi)$, and the expectation $E[\varepsilon\varepsilon^\top] = \sigma^2 I_n$, which leads to

$$\dots = \sigma^2 \text{tr}(I_n - \Pi) = \sigma^2(n - p) = (n - p)E[\hat{\sigma}^2]$$

Looking further at the distribution of $\|Y - X^\top \hat{\theta}\|^2 = \hat{\varepsilon}^\top \hat{\varepsilon}$, we have $\hat{\varepsilon}^\top \hat{\varepsilon} = ((I_n - \Pi)Y)^\top ((I_n - \Pi)Y) = Y^\top (I_n - \Pi)Y = (X^\top \theta^* + \varepsilon)^\top (I_n - \Pi)(X^\top \theta^* + \varepsilon) = \varepsilon^\top (I_n - \Pi)\varepsilon$.

Since we know that $\varepsilon \sim N(0, \sigma^2 I_n)$, and further $\frac{\varepsilon^\top \varepsilon}{\sigma^2} \sim \chi^2(n)$, $(\frac{\varepsilon}{\sigma})^\top (I_n - \Pi)(\frac{\varepsilon}{\sigma}) \sim \chi^2(n - p)$, since we know from earlier that $(I_n - \Pi)$, is idempotent, with rank equal to $\text{tr}(I_n - \Pi) = \text{tr}(I_n) - \text{tr}(\Pi) = n - p$.

Exercise 3. Consider the linear regression model from exercise 1. Suppose, that the target of estimation is $h^\top \theta$ for some determinate non-zero vector $h \in R^p$. Find expression for the LSE of $h^\top \theta$. Is this estimate optimal in sense of Gauss-Markov theorem, i.e. does it have the smallest variance among all linear unbiased estimators?

—Start with this —By Gauss Markov, we know that a BLUE estimator has $\text{Var}(\theta_{OLS}) = \sigma^2 (XX^\top)^{-1}$. However in the case of heteroscedastic noise, we have $\text{Var}(\theta) = (XX^\top)^{-1} XDX^\top (XX^\top)^{-1}$, which must be greater than $\sigma^2 (XX^\top)^{-1}$. And so, in this case, our estimator is not BLUE. Study the same issue for the target $\eta = H^\top \theta$, where $H \in R^{q \times p}$ is some non-zero matrix with $q \leq p$.

Section 1.3

Exercise 4. Let $A \in R^{n \times n}$ be a matrix (corresponding to a linear map in R^n). Show that A preserves length for all $x \in R^n$ iff it preserves the inner product. I.e. one needs to show the following:

$$\|Ax\| = \|x\| \quad \forall x \in R^n \iff (Ax)^\top (Ay) = x^\top y \quad \forall x, y \in R^n.$$

Take,

$$\|x\| = \sqrt{x \cdot x} = \sqrt{x^\top x} \implies \|Ax\| = \sqrt{Ax \cdot Ax} = \sqrt{x^\top A^\top A x} \implies$$

,

$$A^\top A = I_n = A^{-1}, \quad A^\top = A^{-1}, \quad \|Ax\| = \|x\|$$

this implies A is an orthogonal matrix, and further,

$$(Ax)^\top (Ay) = \|AxAy\|^2 = x^\top A^\top A y = x^\top y = \|xy\|^2$$

Exercise 5. (a) Let $x_0 \in R^n$ be some fixed vector, find a projection map on the subspace $\text{span}(x_0)$. Compare your result with matrix Π (from section 1.3) for the case of $p = 1$.

(b) Prove part 3) of Lemma 1.1 for an arbitrary orthogonal projection in R^n .

Exercise 6. Let $L1, L2$ be some subspaces in R^n , and $L2 \subseteq L1 \subseteq R^n$. Let $PL1, PL2$ denote orthogonal projections on these subspaces. Prove the following properties:

(a) $PL2 - PL1$ is an orthogonal projection,

(b) $|PL2| \leq |PL1| \quad \forall x \in R^n$,

(c) $PL2 \cdot PL1 = PL2$

Section 2.1

Exercise 7. (a) Using the notation from section 2.1, consider $X \sim N(\mu, I_n)$ for some $\mu \in R^n$. Find $EQ(X)$ and $\text{Var}Q(X)$. (b) Generalize the results from part (a) to the case $X \sim N(\mu, \Sigma)$ for some positive-definite covariance matrix $\Sigma \in R^{n \times n}$.

Exercise 8. Let $X \sim N(0, I_n)$, $Q = XX$. Suppose that Q is decomposed into the sum of two quadratic forms: $Q = Q1 + Q2$, where $Qi = X^\top A_i X$, $i = 1, 2$ for some symmetric matrices $A1, A2$ with $\text{rank}(A1) = n1$ and $\text{rank}(A2) = n2$. Show that if $n1 + n2 = n$, then $Q1$ and $Q2$ are independent and $Qi \sim \chi^2(n_i)$ for $i = 1, 2$.

Section 2.2

Exercise 9. In the Gaussian linear regression model (3), consider the target of estimation $\eta = H^\top \theta^*$, where $H \in \mathbb{R}^{q \times p}$ is some non-zero matrix with $q \leq p$. Find an analogue of the quadratic form S_2 (from (4)) for the new target η^* , and prove for the new quadratic form statements similar to (e) from Theorem 2.1, and Corollary 2.1.2.

Exercise 10. (a) Consider model (3) for $p = 2$, $X_i = (1, x_i)^\top$, $\theta^* = (\theta_1^*, \theta_2^*)^\top$ (similarly to section 1.5). Write explicit expressions for the confidence sets for θ^* , θ_1^* , θ_2^* .

(b) Find a confidence interval for the expected response $E[Y_i]$ in the model in part (a).

Exercise 11. Find an elliptical confidence set for the expected response $E[Y]$ in model (3).

Exercise 12. Construct simultaneous confidence intervals (e.g., as in Corollary 2.2.1) for the expected responses $E[Y_1], \dots, E[Y_n]$ in model (3).