# midterm1WIP

Exercise 10. (a) Consider model (3) for  $p = 2, X_i = (1, x_i)^{\mathsf{T}}, \theta^* = (\theta_1^*, \theta_2^*)^{\mathsf{T}}$  (similarly to section 1.5). Write explicit expressions for the confidence sets for  $\theta^*, \theta_1^*, \theta_2^*$ .

To set up explicit expression for the case above, for parameter estimates we have:

$$XX^{\mathsf{T}} = \left[ \begin{array}{ccc} 1 & \dots & 1 \\ x_1 & \dots & x_n \end{array} \right] \left[ \begin{array}{ccc} 1 & x_1 \\ \dots & \dots \\ 1 & x_n \end{array} \right] = \left[ \begin{array}{ccc} n & \sum_{i=1}^n x_i \\ \sum_{i=1}^n x_i & \sum_{i=1}^n x_i^2 \end{array} \right]$$

and  $det(XX^{\intercal}) = n \sum_{i=1}^{n} x_i^2 - (\sum_{i=1}^{n} x_i)^2 = n \sum_{i=1}^{n} (x_i - \bar{x})^2$ , and

$$(XX^{\mathsf{T}})^{-1} = \frac{n}{\det(XX^{\mathsf{T}})} \begin{bmatrix} \sum_{i=1}^{n} x_i^2 & -\bar{x} \\ -\bar{x} & 1 \end{bmatrix}$$

So we have

$$\hat{\theta} = (XX^{\mathsf{T}})^{-1}XY = \frac{n}{\det(XX^{\mathsf{T}})} \begin{bmatrix} \sum_{i=1}^{n} x_i^2 & -\bar{x} \\ -\bar{x} & 1 \end{bmatrix} \begin{bmatrix} \sum_{i=1}^{n} y_i \\ \sum_{i=1}^{n} x_i y_i \end{bmatrix} = (\hat{\theta}_1, \hat{\theta}_2)^{\mathsf{T}} = \dots$$

$$\dots = \frac{1}{\sum_{i=1}^{n} (x_i - \bar{x})^2} \begin{bmatrix} \bar{y} \sum_{i} x_i^2 - \bar{x} \sum_{i} x_i y_i \\ \sum_{i} x_i y_i - n \bar{y} \bar{x} \end{bmatrix} = (\hat{\theta}_1, \hat{\theta}_2)^{\mathsf{T}} = \hat{\theta}$$

To find a confidence region for  $\theta^*$ , using a mixture of matrix and summation notation, we use the property:

$$\frac{||(XX^{\mathsf{T}})^{1/2}(\hat{\theta} - \theta^*)||^2}{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2}x_i)^2} \frac{n-2}{2} \sim F(2, n-2)$$

and denote  $\hat{\sigma}^2 = \frac{\sum_{i=1}^n (y_i - \hat{\theta_1} - \hat{\theta_2} x_i)^2}{n-2}$ . Where F denotes the F distribution with  $df_1 = 2$ , and  $df_2 = n - 2$ .

We can create a confidence interval for  $\theta^*$ , such that,  $qF_{\alpha}$  denotes the  $\alpha^{th}$  quantile for F(2, n-2).

$$P(\frac{||(XX^\intercal)^{1/2}(\hat{\theta}-\theta^*)||^2}{p\hat{\sigma}^2} < qF_{1-\alpha}) = 1 - \alpha = P((\hat{\theta}-\theta^*)^\intercal \left[ \begin{array}{cc} n & \sum_{i=1}^n x_i \\ \sum_{i=1}^n x_i & \sum_{i=1}^n x_i^2 \end{array} \right] (\hat{\theta}-\theta^*) < p\hat{\sigma}^2 qF_{1-\alpha})$$

(b) Find a confidence interval for the expected response  $E[Y_i]$  in the model in part (a).

## Section 1.1

Exercise 3. Consider the linear regression model from exercise 1. Suppose, that the target of estimation is  $h^{\dagger}\theta$  for some determinate non-zero vector  $h \in R^p$ . Find expression for the LSE of  $h^{\dagger}\theta$ . Is this estimate optimal in sense of Gauss-Markov theorem, i.e. does it have the smallest variance among all linear unbiased estimators?

—Start with this —By Gauss Markov, we know that a BLUE estimator has  $Var(\theta_{OLS}) = \sigma^2(XX^{\intercal})^{-1}$ ). However in the case of heterscedastic noise, we have  $Var(\theta) = (XX^{\intercal})^{-1}XDX^{\intercal}(XX^{\intercal})^{-1}$ , which must be greater than  $\sigma^2(XX^{\intercal})^{-1}$ ). An so, in this case, our estimator is not BLUE. Study the same issue for the target  $\eta = H^{\intercal}\theta$ , where  $H \in \mathbb{R}^{q \times p}$  is some non-zero matrix with  $q \leq p$ .

#### Section 1.3

Exercise 6. Let L1, L2 be some subspaces in  $\mathbb{R}^n$ , and L2  $\subseteq$  L1  $\subseteq$   $\mathbb{R}^n$ . Let PL1, PL2 denote orthogonal projections on these subspaces. Prove the following properties:

- $(a)\ PL2-PL1\ is\ an\ orthogonal\ projection,$
- (b)  $|PL2| \le |PL1| \ \forall x \in \mathbb{R}^n$ ,
- (c)  $PL2 \cdot PL1 = PL2$

### Section 2.1

Exercise 7. (a) Using the notation from section 2.1, consider  $X \sim N(\mu, I_n)$  for some  $\mu \in \mathbb{R}^n$ . Find E(Q(X)) and Var(Q(X))

For  $Q(X) = \sum_{i} \sum_{j} a_{ij} X_i X_j = X^{\mathsf{T}} A X_i X_j \sim N(\mu, I_n)$ , we have, using the property of trace operator:

$$E(Q(X)) = tr(E(Q(X)) = E(tr(Q(X)) = E(tr(X^\intercal A X)) = E(tr(A X X^\intercal)) = tr(A E(X X^\intercal))$$

Since  $E(XX^{\intercal}) = I_n + \mu \mu^{\intercal}$ , we have,

$$tr(AE(XX^{\mathsf{T}})) = tr(A(I_n + \mu\mu^{\mathsf{T}})) = trA + tr(A\mu\mu^{\mathsf{T}}) = trA + \mu^{\mathsf{T}}A\mu$$

Var(Q(X)) =

(b) Generalize the results from part (a) to the case  $X \sim N(\mu, \Sigma)$  for some positive-definite covariance matrix  $\Sigma \in \mathbb{R}^{n \times n}$ . For  $X \sim N(\mu, \Sigma)$  we have,

$$E(Q(X)) = tr(AE(XX^{\mathsf{T}})) = tr(A(\Sigma + \mu\mu^{\mathsf{T}})) = tr(A\Sigma) + tr(A\mu\mu^{\mathsf{T}}) = tr(A\Sigma) + \mu^{\mathsf{T}}A\mu$$

Var(Q(X)) =

## Section 2.2

Exercise 9. In the Gaussian linear regression model 3, consider the target of estimation  $\eta = H^{\dagger}\theta^*$ , where  $H \in R^{q \times p}$  is some non-zero matrix with  $q \leq p$ . Find an analogue of the quadratic form S2 (from (4)) for the new target  $\eta^*$ , and prove for the new quadratic form statements similar to (e) from Theorem 2.1, and Corollary 2.1.2.

Exercise 11. Find an elliptical confidence set for the expected response E[Y] in model (3).

Exercise 12. Construct simultaneous confidence intervals (e.g., as in Corollary 2.2.1) for the expected responses  $E[Y_1], ..., E[Y_n]$  in model (3).