

The background of the slide features a large, faint watermark of the University of Basel seal. The seal is circular and contains a central figure, likely a saint or scholar, seated and holding a book. Above the figure are three smaller figures in niches. The entire seal is surrounded by a Latin inscription in a circular border.

Stochastic Processes

17. Examples of Markov processes

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Generator of a semigroup

- $\mathcal{X} = (X_t)_{t \in I}$ be a time-homogeneous Markov process with operator semigroup $(T_t)_{t \in I}$. The *generator* of \mathcal{X} is defined as

$$(Gf)(x) = \lim_{t \rightarrow 0} \frac{\mathbf{E}_x[f(X_t) - f(x)]}{t} = \lim_{t \rightarrow 0} \frac{1}{t}((T_t f)(x) - f(x)),$$

for all f for which the limit value exists.

$$\left(f(X_t) - \int_0^t (Gf)(X_s) ds \right)_{t \in I}$$

is a martingale for all $f \in \mathcal{D}(G)$.

Example: ODEs

- ▶ $\mathcal{X} = (X_t)_{t \geq 0}$ solution of the ODE

$$\frac{d}{dt}X_t = g(X_t)$$

where g is Lipschitz. Then, the generator of \mathcal{X} is

$$\begin{aligned}(G^{\mathcal{X}}f)(x) &= \lim_{t \rightarrow 0} \frac{1}{t}(f(X_t) - f(x)) = \left. \frac{d}{dt}(f(X_t)) \right|_{t=0} \\ &= \sum_{i=1}^d \frac{\partial f}{\partial x_i}(g(x)) \cdot g_i(x) = (\nabla f)(g(x)) \cdot g(x).\end{aligned}$$

Example: PPP

- ▶ Let $\mathcal{X} \sim \text{PPP}(\lambda)$. Recall

$$Gf(x) = \lambda(f(x+1) - f(x)).$$

Hence,

$$X_t - \int_0^t \lambda(X_s + 1 - X_s) ds = X_t - \lambda t,$$

$$X_t^2 - \int_0^t \lambda(X_s + 1)^2 - X_s^2 ds = X_t^2 - \int_0^t \lambda(2X_s + 1) ds$$

are martingales.

Example: BM

- Let $\mathcal{X} \sim \text{BM}$. Recall

$$G^{\mathcal{X}} f(x) = \frac{1}{2} f''(x).$$

Hence,

$$\left(X_t - \frac{1}{2} \int_0^t id''(X_s) ds \right)_{t \geq 0} = (X_t)_{t \geq 0},$$

$$\left(X_t^2 - \frac{1}{2} \int_0^t (id^2)''(X_s) ds \right)_{t \geq 0} = (X_t^2 - t)_{t \geq 0},$$

$$\left(\exp \left(\mu X_t - \frac{1}{2} \int_0^t \frac{(e^{\mu \cdot})''(X_s)}{e^{\mu X_s}} ds \right) \right)_{t \geq 0} = \left(\exp \left(\mu X_t - \frac{1}{2} \mu^2 t \right) \right)_{t \geq 0}$$

are martingales.

Example: Markov jump processes

- ▶ Given $X_s = x$, jump after $\exp(\lambda(x))$ distributed time according to $\mu(x, \cdot)$.
- ▶ If λ is bounded, probability of > 1 jump by time t is $\mathcal{O}(t^2)$, so for $f \in \mathcal{C}_b(E)$

$$\begin{aligned}(Gf)(x) &= \lim_{t \rightarrow 0} \frac{\mathbf{E}_x[f(X_t) - f(x)]}{t} \\&= \lim_{t \rightarrow 0} \frac{1}{t} \left((e^{-\lambda(x)t} - 1)f(x) + \lambda(x)t e^{-\lambda(x)t} \int \mu(x, dy) f(y) \right) \\&= \lambda(x) \int \mu(x, dy) (f(y) - f(x)) dy.\end{aligned}$$

- ▶ If $\lambda(x, dy) = \lambda(x)\mu(x, dy)$ is continuous, \mathcal{X} is Feller.

Example: Markov jump processes

- Assume E is discrete. Using $f(y) = 1_{y=x}$, the master equation

$$\begin{aligned}\frac{d}{dt}\mathbf{P}(X_t = x) &= \frac{d}{dt}\mathbf{E}[f(X_t)] = \mathbf{E}[(Gf)(X_t)] \\ &= \mathbf{E}\left[\sum_{y \in E} \lambda(X_t, y)(1_{y=x} - 1_{X_t=x})\right] \\ &= \sum_{z \in E} \mathbf{P}(X_t = z) \sum_{y \in E} \lambda(z, y)(1_{x=y} - 1_{x=z}) \\ &= \sum_{z \in E} \lambda(z, x)\mathbf{P}(X_t = z) - \lambda(x, z)\mathbf{P}(X_t = x).\end{aligned}$$

is a linear system of ODEs.

Branching process in continuous time

- Individuals die at rate 1, and are replaced by $Z \sim \mu$ offspring.

$$Gf(x) = x \sum_{n=0}^{\infty} \mu(n)(f(x-1+n) - f(x)).$$

$$\begin{aligned} \text{For } f_r(x) = r^x: \quad Gf_r(x) &= xr^{x-1} \sum_{n=0}^{\infty} \mu(n)(r^n - r) \\ &= xr^{x-1}(g_{\mu}(r) - r) = (g_{\mu}(r) - r) \frac{d}{dr} f_r(x). \end{aligned}$$

So the function $u : (t, r) \mapsto \mathbf{E}_x[r^{X_t}]$ solves the equation

$$\frac{d}{dt} u(t, r) = (g_{\mu}(r) - r) \frac{d}{dr} u(t, r)$$

with the boundary conditions $u(0, r) = r^x$, $u(t, 1) = 1$.

Yule process

- This is the special case $\mu = \delta_2$, i.e. $g_\mu(r) = r^2$;

$$\frac{d}{dt}u(t, r) = -r(1 - r)\frac{d}{dr}u(t, r)$$

For $u(0, r) = r$, $u(t, 1) = 1$, the solution is

$$u(t, r) = \frac{e^{-t}r}{1 - r(1 - e^{-t})}.$$

The rhs is the generating function of $\text{geo}(e^{-t})$, i.e.

$X_t \sim \text{geo}(e^{-t})$. This can also be shown using the master equation

$$\frac{d}{dt}\mathbf{P}(X_t = x) = (x - 1)\mathbf{P}(X_t = x - 1) - x\mathbf{P}(X_t = x).$$

Extinction probability

- ▶ $T := \inf\{t : X_t = 0\}$ extinction time. Then,

$$\mathbf{P}_x(T < \infty) = \mathbf{P}_1(T < \infty)^x \text{ and } r := \mathbf{P}_1(T < \infty)$$

$$r = (1 - h)r + h \sum_{n=0}^{\infty} \mu(n)r^n + o(h),$$

so

$$r = g_{\mu}(r)$$

- ▶ For binary branching, $g_{\mu}(r) = p\delta_0 + (1 - p)\delta_2$ the only solution ≤ 1 is

$$r = \frac{p}{1 - p} \wedge 1.$$

Hitting times

- ▶ Let $E' \subseteq E$ and $T := T_{E'}$ hitting time of E' .

Goal: calculate $u : x \mapsto \mathbf{E}_x[T]$.

Since $u(x) = \mathbf{E}_x[T] = 0$ for $x \in E'$, with $\lambda(x) = \sum_y \lambda(x, y)$

$$\begin{aligned}\mathbf{E}_x[T] &= (1 - h\lambda(x))\mathbf{E}_x[T + h] + \sum_y \mathbf{E}_x[T | X_h = y] \cdot \mathbf{P}(X_h = y) \\ &= \mathbf{E}_x[T] + h(1 - \lambda(x))\mathbf{E}_x[T] + \sum_y \lambda(x, y)\mathbf{E}_y[T] + O(h^2) \\ &= \mathbf{E}_x[T] + h(1 + G\mathbf{E}_\bullet[T]) + O(h^2).\end{aligned}$$

Therefore, u must fulfill the equation

$$Gu(x) = -1, \quad x \notin E',$$

$$u(x) = 0, \quad x \in E'.$$

Birth-death processes

- If $E = \mathbb{Z}_+$ and $\lambda(x, y) = 0$ for $|x - y| > 1$, we have a birth-death process. Set

$$\lambda(n, n+1) =: \lambda_n, \quad \lambda(n, n-1) =: \mu_n,$$

$$Gf(n) = \lambda_n(f(n+1) - f(n)) + \mu_n(f(n-1) - f(n)).$$

$$\text{For } u(n) := \mathbf{E}_n[T_0], \quad u(0) = 0, \quad u(n) = \sum_{k=1}^n \frac{1}{\mu_k \pi_k} \sum_{j=k}^{\infty} \pi_j$$

with $\pi_1 = 1$ and $\pi_i = \prod_{j=2}^i \frac{\lambda_{j-1}}{\mu_j}$.. Indeed,

$$\begin{aligned} Gu(n) &= \lambda_n \frac{1}{\mu_{n+1} \pi_{n+1}} \sum_{j=n+1}^{\infty} \pi_j - \mu_n \frac{1}{\mu_n \pi_n} \sum_{j=n}^{\infty} \pi_j \\ &= \frac{1}{\pi_n} \sum_{j=n+1}^{\infty} \pi_j - \frac{1}{\pi_n} \sum_{j=n}^{\infty} \pi_j = -1. \end{aligned}$$