

The rule of conditional probability is valid in quantum theory

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In a recent manuscript, Gelman & Yao (2020) claim that “the usual rules of conditional probability fail in the quantum realm” and purport to support such statement with an example. Such a statement is false. I would like to recall some literature that shows why it is false, and to sum up the fallacy underlying their example.

Let me point out at the outset that the rule of conditional probability and the other two rules (sum and negation) are in fact routinely used in quantum theory, with full validity, especially in problems of state “retrodiction” and measurement reconstruction (Jones 1991; Slater 1995; de Muynck 2002 chs 7, 8; Barnett et al. 2003; Ziman et al. 2006; D’Ariano et al. 2004; see Månsson et al. 2006 § 1 for many further references), for example to infer the state of a quantum laser given its output through different optical apparatus (Leonhardt 1997).

Similar incorrect claims with similar examples have appeared before in the quantum literature. Bernard O. Koopman (of the Pitman-Koopman theorem for sufficient statistics, Koopman 1936) discussed the falsity of such claims already in 1957. The Introduction in his work is very clear:

Ever since the advent of modern quantum mechanics in the late 1920’s, the idea has been prevalent that the classical laws of probability cease, in some sense, to be valid in the new theory. More or less explicit statements to this effect have been made in large number and by many of the most eminent workers in the new physics [. . .]. Some authors have even gone farther and stated that the formal structure of logic must be altered to conform to the terms of reference of quantum physics [. . .].

Such a thesis is surprising, to say the least, to anyone holding more or less conventional views regarding the positions of logic, probability, and experimental science: many of us have been apt – perhaps too naively – to assume that experiments can lead to conclusions only when worked up by means of logic and probability, whose laws seem to be on a different level from those of physical science.

The primary object of this presentation is to show that the thesis in question is entirely without validity and is the product of a confused view of the laws of probability.

A claim similar to Gelman & Yao's, with a similar supporting example, was made in a work by Brukner & Zeilinger (2001), somewhat famous in the quantum community; although their focus was on an alleged inconsistency of some properties of the Shannon entropy in quantum theory.

The fallacy in the reasoning of Brukner & Zeilinger and Gelman & Yao rests in the neglect of the experimental setup, leading to an incorrect calculation of conditional probabilities. Such fallacy was exposed and discussed at length by Porta Mana (2004) through a step-by-step analysis and calculation. This work also showed, through simple examples (*ibid.* § IV), that the same incorrect statements can be obtained *with completely classical systems*, such as drawing from an urn, if the setup is neglected. Here is a simple example.

Consider an urn with one Blue and one Red ball. There are two possible drawing setups:

D_a With replacement for blue, without replacement for red. That is, if blue is drawn, it is put back before the next draw (and the urn is shaken); if red is drawn, it is thrown away before the next draw.

D_b With replacement for red, without replacement for blue.

The two setups are obviously mutually exclusive.

We can easily calculate what is the unconditional probability for blue at the second draw in the setup D_a :

$$P(B_2 | D_a) = \frac{3}{4}, \quad (1)$$

and the conditional probabilities for blue at the second draw, conditional on the first draw, in the setup D_b :

$$P(B_2 | B_1 \wedge D_b) = 0 \quad P(B_2 | R_1 \wedge D_b) = \frac{1}{2}. \quad (2)$$

We find that

$$P(B_2 | D_a) \neq P(B_2 | B_1 \wedge D_b) P(B_1 | D_b) + P(B_2 | R_1 \wedge D_b) P(R_1 | D_b). \quad (3)$$

This inequality is no surprising. And it does not contradict the rule of conditional probability, because that rule is supposed to be used within the same probability space. You can call the inequality above “interference” if you like; for further examples see Kirkpatrick (2003a; 2003b) and Porta Mana (2004 § IV).

Note that, had we considered a drawing setup D_c with replacement for both colours, and a setup D_d without replacement for either colour, we would have found

$$P(B_2 | D_c) = P(B_2 | B_1 \wedge D_d) P(B_1 | D_d) + P(B_2 | R_1 \wedge D_d) P(R_1 | D_d). \quad (4)$$

The equality above, however, is *not* an expression of the conditional-probability rule, because the probability spaces are different. It is simply a peculiar equality contingent on the two specific setups. The probability calculus handles correctly situations such as (3) or (4).

In fact, strictly speaking it is wrong to use the expression ‘ B_2 ’ for all these setups, because ‘ B_2 ’ in the one setup denotes a different statement (or random variable) than in another. Just like “it rains (on 2020-07-14T09:00+0200 in Trondheim)” is different from “it rains (on 2019-01-20T18:00+0200 in Rome)”. I should have used different symbols. The explicit presence of ‘ $D...$ ’, which represents given information, luckily avoided any ambiguities. But if in our formulae we omit the notation of the setup *and* we use the same notation for actually different statements or random variables, then we’re in for trouble and for incorrect applications of the probability rules.

The example above stresses the importance of the probability space – even in “non-quantum” situations. But it is not meant as a parallel of Gelman & Yao’s (2020 § 2) example. In fact their example has important differences and their analysis of it is incorrect from the point of view of quantum theory. Here are the main points:

- (i) It *does* matter whether many photons are sent at once, or one at a time, as well as their wavelength, temporal spread, and so on. These details lead to different probabilities distributions of detection at the screen¹. More precisely, the spatio-temporal dependence of the optic-field operator (which defines the quantum system: van Enk 2003) must be specified.

All these details are not “latent variables”: they are the initial and boundary conditions that define the physical system. They correspond to the different drawing setups in the urn example above

¹ e.g. Mandel & Wolf 1965; Morgan & Mandel 1966; Paul 1982; Jacobson et al. 1995; and textbooks such as Loudon 2000; Mandel & Wolf 2008; Scully & Zubairy 2001; Bacher & Ralph 2004; Walls & Milburn 1994.

– we would not call the specification “drawing with replacement” a latent variable. The rules of probability apply seamlessly in each case.

It is also possible to consider situations where part of the setup, such as slit width or presence or absence of detectors, is unknown. This is similar, in the urn example above, to not knowing whether D_a or D_b applies. In this case one can make inferences by giving the probability for each setup, e.g. $P(D_a)$, and applying the conditional-probability rule. The same rule is indeed applied in the analogous quantum situation (e.g. Ziman et al. 2006). Again, no violations of the probability rules in the quantum realm.

- (ii) Owing to the point above, it is important not to conflate the *probability* for single-photon detections and the *frequency* distribution of a long-run of such detections, as Gelman & Yao (2020 p. 1) instead do. For example, in some setups we can have a detection probability density $p(y_1)$ for the first photon, and a *different* density for the second photon $p(y_2 | y_1)$, conditional on the detection of the first – both being different from the long-run joint density of detections $f(y)$ (see e.g. the phenomena of higher-order coherence and bunching in the references above¹). The rules of the probability calculus also apply in such situations. We can infer, for example, the position of the first photon detection given the second from $p(y_1 | y_2) \propto p(y_2 | y_1) p(y_1)$.
- (iii) The long-run frequency distribution for screen detection in the setup with one slit open (denoted ‘ $p_1(y)$ ’ by Gelman & Yao), and the one conditional on slit detection in the setup with detectors at the slits (‘ $p_4(y | x)$ ’) are a priori unrelated, since they belong to different setups. In some situations (but not always, see points (i), (ii) above) they experimentally turn out to be equal. But this was not required by the probability calculus. Such equality is similar to eq. (4) between the with-replacement and without-replacement urn setups, which is not an a priori rule of the probability calculus, and in fact does not hold for the setups in eq. (3).
- (iv) In the setup with one or both slits open, the quantity “detection at slit x ” *does not exist*, because no such measurement is included in the setup. So in these cases the idea of a marginal probability does not even apply.

Of course one is free to introduce such a quantity within some hidden-variable theory (which goes beyond the quantum realm), such as Bohmian mechanics (Berndl et al. 1995; Dürr et al. 2004; Valentini & Westman 2005; this theory includes as hidden variables the positions of the particles and the spatial configuration of a so-called pilot wave). But then note, first, that such theory gives experimental predictions in complete agreement with standard quantum theory; and second, that the rules of probability theory are again used without violations in this theory (e.g. the relation between joint and conditional distributions of the particles' positions, and so on).

I may add that the idea and parlance of “photons passing through slits” are used today only out of tradition; maybe a little poetically. The technical parlance, as routinely used in quantum-optics labs for example (Leonhardt 1997; Bachor & Ralph 2004), has a different underlying picture. The basic “system” in a quantum-optics experiment are not photons, but the modes of the field-configuration operator (note that this is not yet Quantum ElectroDynamics). “Photon numbers” denote the discrete outcomes of a specific energy-measurement operator; “photon states” denote specific states of the field operators. As another example, “entanglement” is strictly speaking not among photons, but among modes of the field operator (van Enk 2003). Several quantum physicists indeed oppose the idea and parlance of “photons”, owing to the confusion it leads to. Lamb (of the Lamb shift, Lamb & Retherford 1947) wrote in 1995:

the author does not like the use of the word “photon”, which dates from 1926. In his view, there is no such thing as a photon. Only a comedy of errors and historical accidents led to its popularity among physicists and optical scientists.

Wald (1994) warns:

standard treatments of quantum field theory in flat spacetime rely heavily on Poincaré symmetry (usually entering the analysis implicitly via plane-wave expansions) and interpret the theory primarily in terms of a notion of “particles”. Neither Poincaré (or other) symmetry nor a useful notion of “particles” exists in a general, curved spacetime, so a number of the familiar tools and concepts of field theory must be “unlearned” in order to have a clear grasp of quantum field theory in curved spacetime. [p. ix] [...] the notion of “particles” plays no fundamental role either in the formulation or interpretation of the theory. [p. 2]

It may be useful to give a summary of how modern quantum theory works. A quantum system is defined by its sets of possible states and possible measurements. A state ρ is represented by an Hermitean, unit-trace matrix ρ (which satisfies additional constraints), called ‘density matrix’. States traditionally represented by kets, $|\psi\rangle$, are just special cases of density matrices. A measurement setup M is represented by a set of Hermitean matrices $\{\mathbf{M}_r\}$ adding up to the identity matrix, of the same order as the density matrices. They are called ‘positive-operator-valued measures’. Traditional von Neumann projection operators $\{|\phi_r\rangle\langle\phi_r|\}$ are just special cases. Each matrix \mathbf{M}_r is associated with an outcome r of the measurement. These outcomes are mutually exclusive. An outcome can actually represent a combination of simpler outcomes, $r \equiv (x, y, z, \dots)$, such as two intensities at two detectors.

The probability of observing outcome $r \equiv (x, y, \dots)$ given the measurement setup M and the state S is encoded in the trace-product of the respective matrices:

$$p(x, y, \dots | M \wedge S) \equiv \text{tr}(\mathbf{M}_{x,y,\dots} \rho) , \quad (5)$$

forming a probability distribution. Such probabilities come from repeated measurement experiences in the same experimental conditions – we could invoke de Finetti’s theorem here, and some quantum physicists indeed do².

Once the probability distribution above is given, we can use the full-fledged probability calculus. We can for example sum (or integrate) over the detector outcomes y, \dots , obtaining the marginal probabilities for the detector outcomes x :

$$p(x | M \wedge S) \equiv \text{tr}[(\sum_{y,\dots} \mathbf{M}_{x,y,\dots}) \rho] , \quad (6)$$

obviously represented by the matrices $\{\sum_{y,\dots} \mathbf{M}_{x,y,\dots}\}$.

obviously encoded by the matrix

Once the probabilities ab

First, in the experiments with only one slit open or both slits open, the outcome space is $\{\text{‘no event’}\} \cup \mathbf{R}$, because either an emulsion is produced at some point on the screen, or none is produced.

After all, we do not expect that marginal probabilities from, say, a drawing-without-replacement urn setup should be obtainable from

² Caves et al. 2002; van Enk & Fuchs 2002; Fuchs et al. 2004.

the joint distribution of a drawing-with-replacement setup, or of Pólya drawings. These setups are mutually exclusive. A random variable of one of them is not the same random variable of another. If you thought that you could consistently combine probabilities from such different urn-drawing setups and find that you actually cannot, well, too bad for you. The probability calculus, in fact, makes clear at the outset that the probabilities of these setups cannot generally be combined. It is thus somewhat funny that one ends up blaming the probability calculus, which makes a clear distinction, for one's neglect of that distinction.

Likewise, measurement setups in quantum theory – and in many classical-physics situations – are generally mutually exclusive.

This kind of incorrect claims about

I refer to the work just cited for the full analysis and counterexamples. Here I summarize the basic fallacy with a simpler counterexample.

The basic fallacy is the confusion of the probability conditional with a temporal ordering.

Take the conditional probability $P(A | B)$, where the statement or event A refers to a time t_2 , and B to a time t_1 that precedes t_2 . This probability is related to the reverse conditional $P(B | A)$ by

$$P(A | B) P(B) = P(B | A) P(A) = P(A \wedge B) . \quad (7)$$

It goes without saying that *the statements or events A and B must be the same on both sides of each equation*. In particular, in the conditional $P(B | A)$ the statement B still refers to the time t_1 , and A to the time t_2 , with t_1 preceding t_2 . We can represent these times explicitly and rewrite (7) as

$$P(A_2 | B_1) P(B_1) = P(B_1 | A_2) P(A_2) = P(A_2 \wedge B_1) . \quad (8)$$

The fallacy is to think that, in calculating $P(B | A)$, we should now ensure that B refers to time t_2 , and A to time t_1 , *swapping the times*. But these would be *different* events, not the original events. In symbols, we would be calculating $P(B_2 | A_1)$, which is different from $P(B_1 | A_2)$, to which the conditional-probability rule (7) refers.

Probability theory has nothing to say, a priori, about the relation between $P(B_1 | A_2)$ and $P(B_2 | A_1)$. They are two logically different situations. A simple example can illustrate this point.

Consider an urn with Red and Blue balls. The drawing scheme is with replacement for blue, and without replacement for red. That is, if a

Figure 1 “Equiprobable worlds” diagram

1st draw	R	R	B	B
	↓	↓	↓	↓
2nd draw	B	B	B	R

blue ball is drawn, it is put back in the urn (and the urn shaken) before the next draw. If a red ball is drawn, it is not put back before the next draw.

The urn initially has one blue and one red ball. The probabilities for the first draw are straightforward:

$$P(B_1) = \frac{1}{2} \quad P(R_1) = \frac{1}{2} , \quad (9)$$

as are the conditional probabilities for the second draw, conditional on the first:

$$P(B_2 | B_1) = \frac{1}{2} \quad P(R_2 | R_1) = 0 \quad (10a)$$

$$P(R_2 | B_1) = \frac{1}{2} \quad P(B_2 | R_1) = 1 . \quad (10b)$$

These probabilities can be visualized with the “equiprobable worlds” diagram of fig. 1. In half of the worlds the first draw yields red; in the other half, blue, according to (9). In all worlds with red first, the second must yield blue. In half of the worlds with blue first, the second yields blue; and in the other half, red; as for eqs (10).

Let us now calculate the conditional probabilities for the first draw conditional on the second (imagine the first draw was hidden from you, and upon seeing the second you’re asked to guess what the first was). Using the conditional-probability rule (8) in the form of Bayes’s theorem,

$$P(X_1 | Y_2) = \frac{P(Y_2 | X_1) P(X_1)}{\sum_X P(Y_2 | X_1) P(X_1)} , \quad (11)$$

we obtain

$$P(B_1 | B_2) = \frac{1}{3} \quad P(R_1 | R_2) = 0 \quad (12a)$$

$$P(R_1 | B_2) = \frac{2}{3} \quad P(B_1 | R_2) = 1 . \quad (12b)$$

These conditional probabilities are intuitively correct, as can be checked by simple enumeration of the possible cases in fig. 1. For example, among

all three worlds that have blue at the second draw, two of them have red at the first; hence $P(R_1 | B_2) = 2/3$. Also, if we have red at the second draw, then logically red cannot have been drawn at the first, otherwise there would not have been any red left; hence blue must have been drawn at the first: $P(B_1 | R_2) = 1$.

The following two relations between the conditional probabilities (10b) and (12b) are relevant to our discussion:

$$P(R_2 | B_1) \neq P(R_1 | B_2) \quad P(B_2 | R_1) = P(B_1 | R_2) . \quad (13)$$

As we see, the probability-calculus does not prescribe the equality nor the inequality between probabilities for events of similar kind but at different times (which therefore are *not* the same event). Any relations of this kind will depend on the specific situation. You can check, for example, that in a scheme of drawing with replacement for both blue and red, we would obtain two equalities in place of (13). In a scheme of Pólya draws for blue (blue is returned plus an additional blue) and replacement draws for red, instead, we would obtain two inequalities in place of (13).

Something analogous (see Porta Mana 2004 for an exact parallel) happens in inferences about quantum measurements. The conditional probability for the result of measurement R made at time t_2 , given the result of measurement B made at time t_1 , is not necessarily the same as that for measurement R made at time t_1 , given measurement B at time t_2 . The time ordering of measurements is extremely important in quantum theory (as it is in the urn example above). In particular, if our goal is to actually *retrodict* the result of a previous measurement or state from the information gained in a subsequent measurement, we must be very careful not to confuse the conditional probabilities. The times of the two measurements cannot be swapped.

It should be noted that the conditional-probability rule (7) is in fact routinely used in quantum theory in problems of state retrodiction and measurement reconstruction (Jones 1991; Slater 1995; de Muynck 2002 chs 7, 8; Ziman et al. 2006; D'Ariano et al. 2004; see Månsson et al. 2006 § 1 for many further references), for example to infer the state of a quantum laser given its output through different optical apparatus (Leonhardt 1997).

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- (“de X ” is listed under D, “van X ” under V, and so on, regardless of national conventions.)
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