

Applied Data Science

Machine Learning Lecture 1

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June 10, 2024**

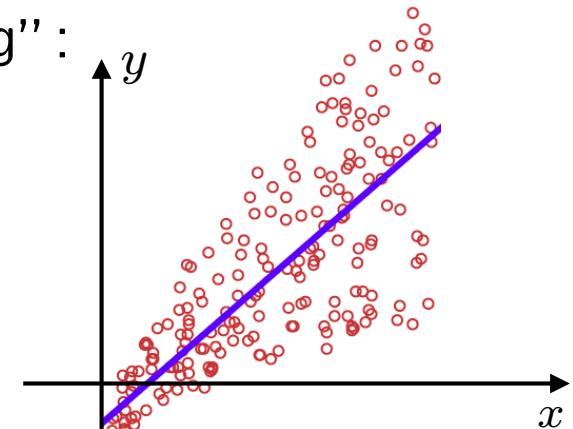
Overview of this week/module

- Central methods in Machine Learning

- we will only discuss “supervised learning”: learn from labeled examples

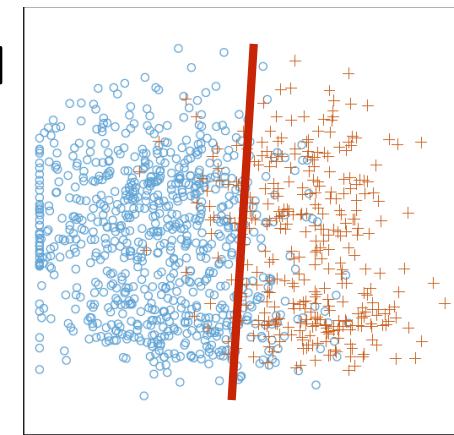
- Predict the value of an unobserved y

Regression (linear)



- Predict the type/color of a new individual

Classification



- **Assessment**

How good is our method, our model, and our prediction?

Testing, validation

Today's agenda

- Regression
 - formulation
 - solution
 - interpretation
 - (classical) performance assessment
- Further topics (next session)
 - what can go wrong
 - using nonlinear features of the data
 - overfitting and regularization
 - ridge regression
 - sparse regression and lasso
 - more on performance assessment
 - cross-validation
 - bootstrap

MACHINE LEARNING AND STATISTICS

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A conceptual big picture



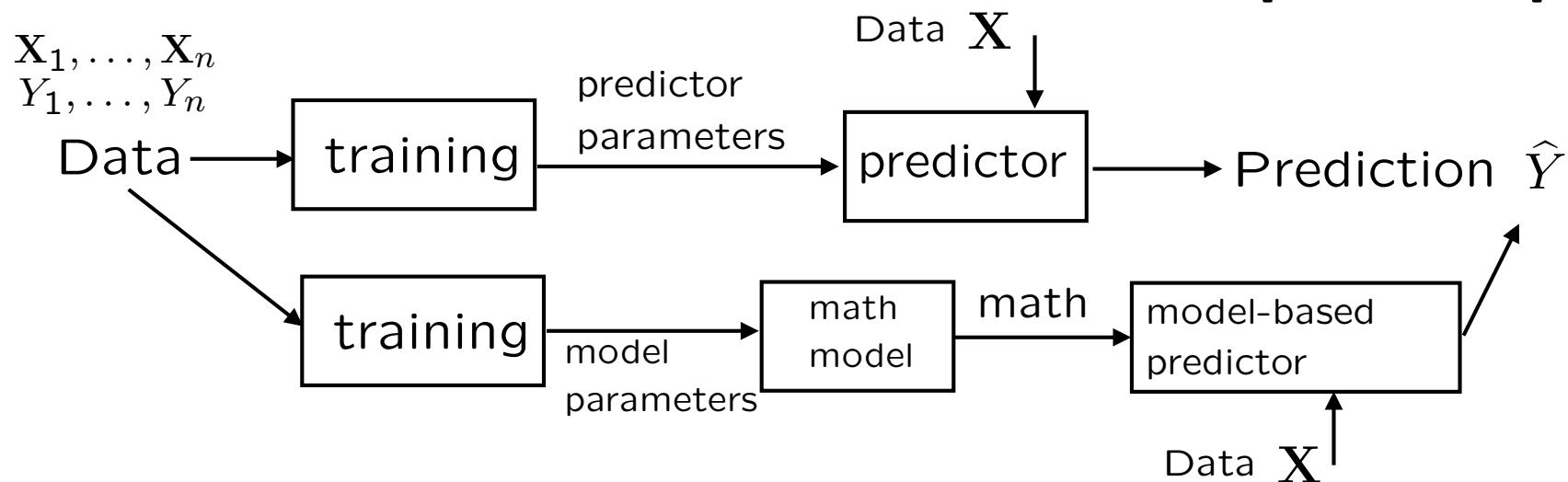
Model?	
X_1	Y_1
:	:
:	:
X_n	Y_n
X	$Y?$

X : symptoms, test results, etc.
 Y : state of health

- “Predict” Y based on X

Y : sick or not (binary) [classification]

Y : life expectancy (any real number)
[regression]



- **Understand**

- build a model, a theory, a narrative, a mechanism

“All models are wrong, some are useful” (George E.P. Box)

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Some language

- predict Y from $\mathbf{X} = (X_1, \dots, X_m)$
- X_i : covariates, independent variables, features
- Y : response, dependent variable, target

Notation key

- vectors: boldface
scalars: normal font

- \mathbf{X}_2 : second data record
- X_2 : second component of a vector \mathbf{X}

$$\mathbf{X} = \begin{bmatrix} X_1 \\ X_2 \\ X_3 \end{bmatrix} \quad \mathbf{X}^T = [X_1 \ X_2 \ X_3]$$

$$\mathbf{X}^T \mathbf{Y} = X_1 Y_1 + X_2 Y_2 + X_3 Y_3$$

- “star” for true quantity, e.g., θ^*
- “hat” for estimates, e.g., $\widehat{\Theta}$

The overall field

Statistics

Machine learning

- **Data Science:** Extracting useful information from data
- Need a language: probability
- Build on two centuries of statistical knowledge

(LINEAR) REGRESSION

formulation

solution

interpretation

An example: Advertising and Sales

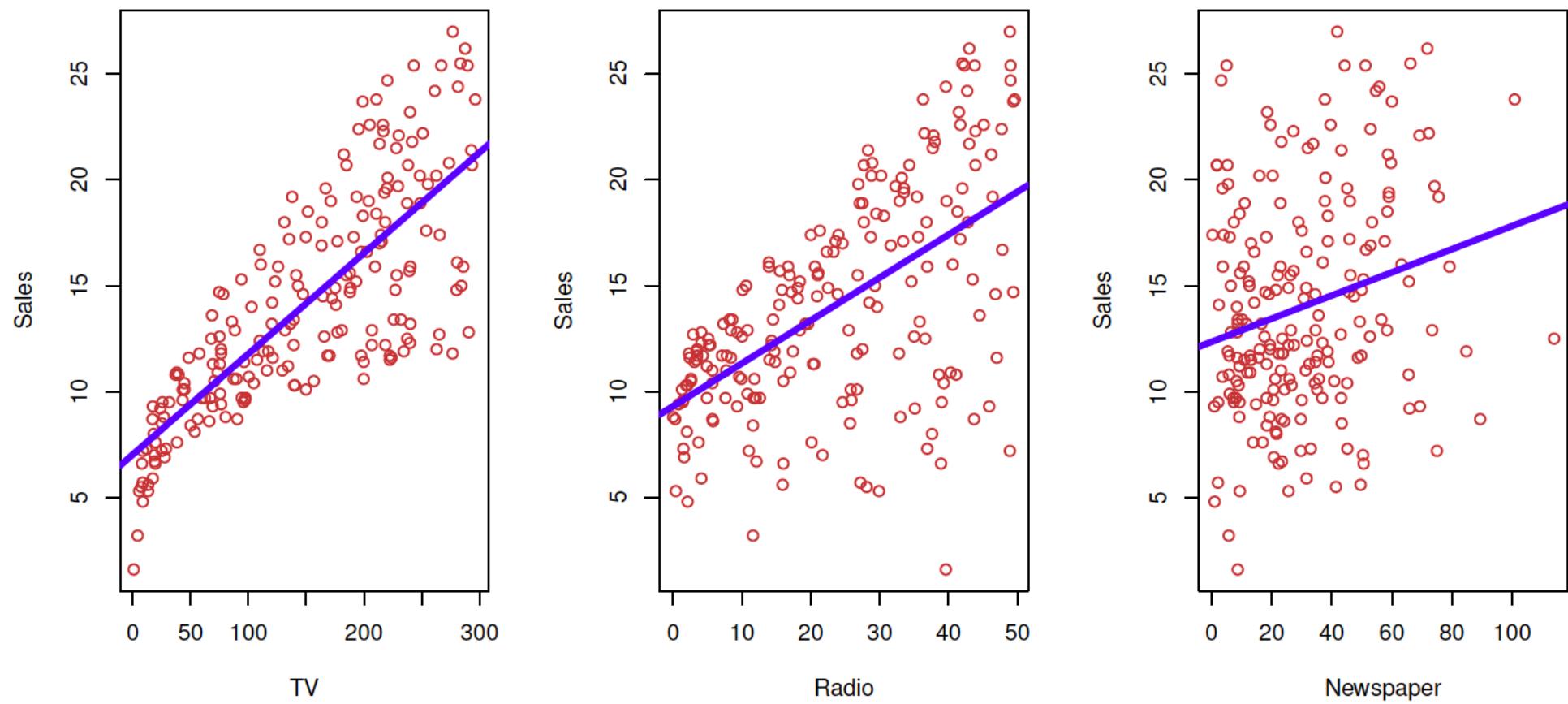
- Data across 200 Markets
 - Spending for TV, Radio, NewsPaper
 - Sales

	TV	Radio	Newspaper	Sales
0	230.1	37.8	69.2	22.1
1	44.5	39.3	45.1	10.4
2	17.2	45.9	69.3	9.3
3	151.5	41.3	58.5	18.5
4	180.8	10.8	58.4	12.9
...
195	38.2	3.7	13.8	7.6
196	94.2	4.9	8.1	9.7
197	177.0	9.3	6.4	12.8
198	283.6	42.0	66.2	25.5
199	232.1	8.6	8.7	13.4

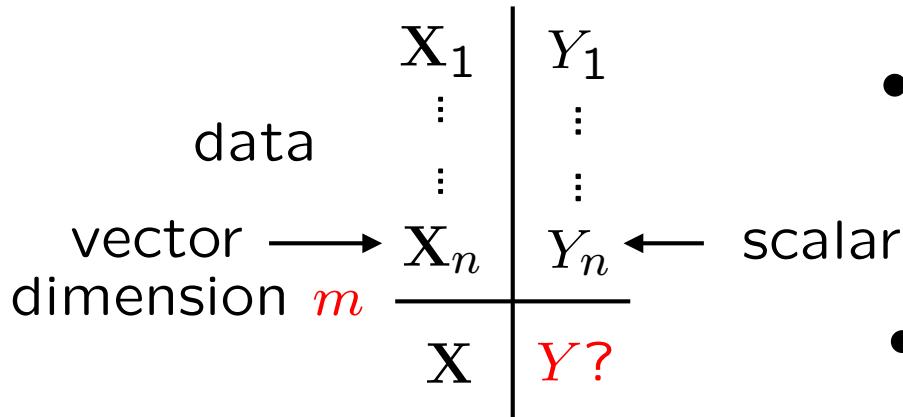
200 rows × 4 columns

- Questions
 - Is there a relation between Advertising Channel Budgets and Sales?
 - If yes, can we “predict” Sales given the Channel Budgets?

Visualize!



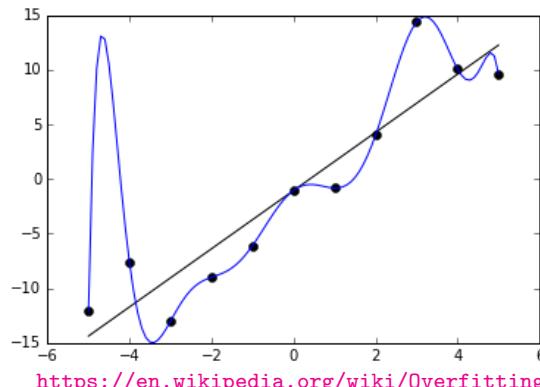
Regression



- **Regressor/predictor:** $\hat{Y} = g(\mathbf{X})$
- “Learn” a “good” g from the data

objective: $\mathbb{E}\left[(g(\mathbf{X}) - Y)^2\right]$
(risk)

proxy: $\frac{1}{n} \sum_{i=1}^n (g(\mathbf{X}_i) - Y_i)^2$ “empirical risk minimization”



- Restrict to limited class of predictors

<https://en.wikipedia.org/wiki/Overfitting>

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Linear regression

data	X_1 ⋮ X_n	Y_1 ⋮ Y_n
vector dimension m	\mathbf{X}	$Y?$

$$\frac{1}{n} \sum_{i=1}^n (g(\mathbf{X}_i) - Y_i)^2$$

sum of over data points

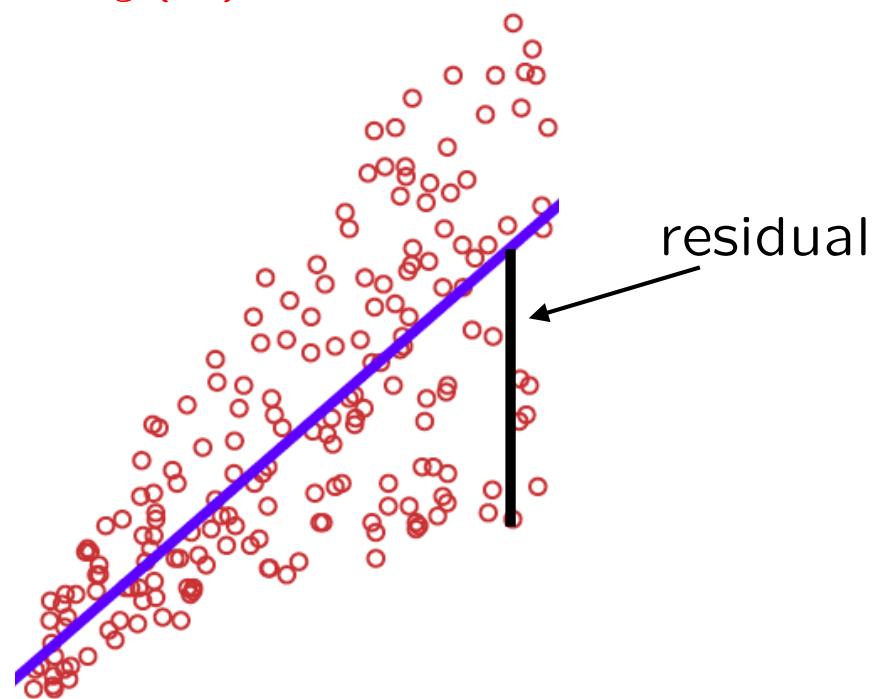
$$\min_{\theta} \sum_{i=1}^n (\theta^T \mathbf{X}_i - Y_i)^2$$

- Restrict to limited class of predictors:

$$\hat{Y} = \theta_0 + \theta_1 X_1 + \cdots + \theta_m X_m$$

let $\mathbf{X} = (1, X_1, \dots, X_m)$
 $\theta = (\theta_0, \theta_1, \dots, \theta_m)$

$$\hat{Y} = g(\mathbf{X}) = \theta^T \mathbf{X}$$



- ordinary least squares (OLS)

Solution to the regression problem

$$\min_{\theta} \sum_{i=1}^n (\theta^T \mathbf{X}_i - Y_i)^2 \quad \begin{array}{l} n \text{ data points} \\ \mathbf{X}_i \text{ and } \theta \text{ have dimension } m+1 \end{array}$$

$$n \begin{bmatrix} \cdots \mathbf{X}_1^T \cdots \\ \vdots \\ \cdots \mathbf{X}_n^T \cdots \\ m+1 \end{bmatrix} \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} = [\mathbb{X} \mid \mathbf{Y}]$$

- Formulas:

$$\hat{\theta} = (\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X}^T \mathbf{Y}$$

Math details:

$$\min_{\theta} H(\theta) \quad \text{quadratic in } \theta$$

$$\text{optimality conditions: } \nabla H(\theta) = 0 \quad \frac{\partial H}{\partial \theta_j} = 0, \quad j = 0, 1, \dots, m$$

linear system of $m+1$ equations

Results for our example

$$n = 200$$

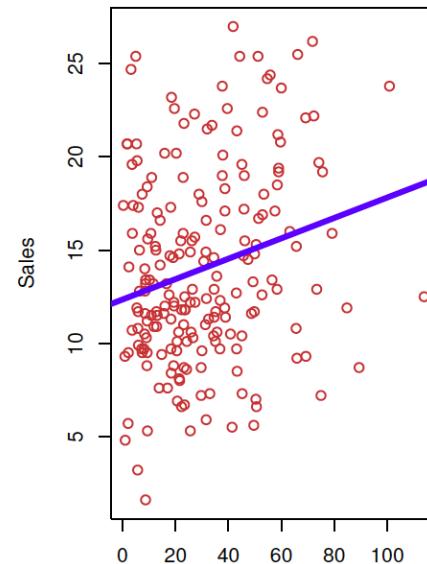
$$m + 1 = 4$$

$$\hat{\theta} = \begin{bmatrix} 2.94 \\ 0.046 \\ 0.19 \\ -0.001 \end{bmatrix}$$

$$\widehat{\text{Sales}} = 2.94 + 0.046 \cdot (\text{TV}) + 0.19 \cdot (\text{Radio}) - 0.001 \cdot (\text{NewsP})$$

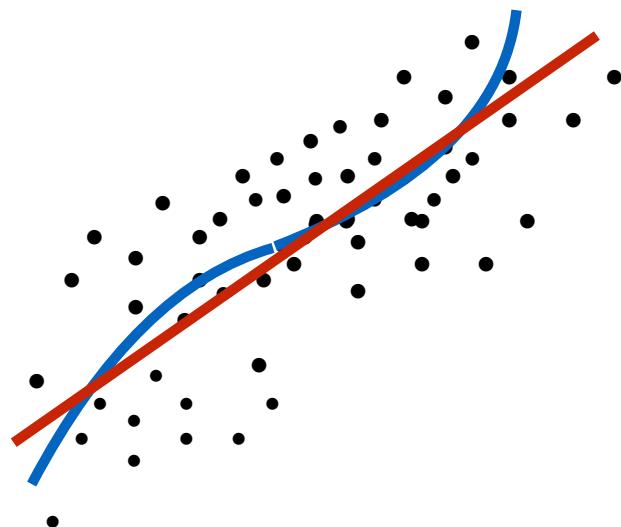
- Compare with **simple** linear regression

$$\widehat{\text{Sales}} = 12.35 + 0.055 \cdot (\text{NewsP})$$

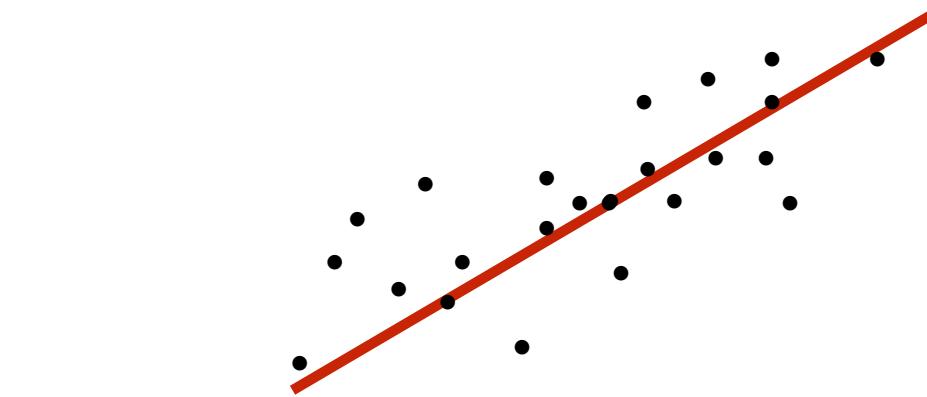
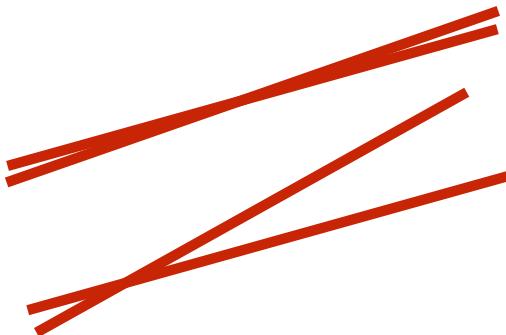


Interpretation and justification: empirical risk minimization

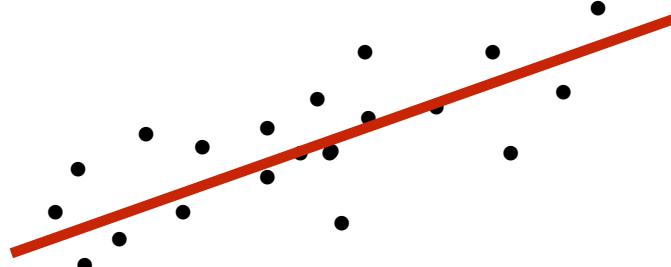
- Large true population



- true relation may be complex
- interested in best linear predictor



- Finite sample: find best linear fit
 $n \rightarrow \infty$: recover “population best”
(as long as samples are drawn representatively)



- Another finite sample: different results
how much variation do we expect?

Interpretation and justification: maximum likelihood

- independent variables \mathbb{X} are somehow fixed; then \mathbb{Y} is observed
 - For any candidate θ , how probable would it be to observe the Y s that were actually observed?
Likelihood: $\mathbb{P}(\mathbb{Y} | \mathbb{X}; \theta)$
 - **Maximum likelihood method:** $\max_{\theta} \mathbb{P}(\mathbb{Y} | \mathbb{X}; \theta)$
- Illustrate for $m = 1$. **Assume:**
 - **structural model:** $Y_i = \theta_0^* + \theta_1^* X_i + W_i$
 - conditioned on all the X_i : all the W_i are **independent** and $\text{Normal}(0, \sigma^2)$

Maximizing the likelihood function = minimizing the empirical risk

$$Y_i : \text{Normal}(\theta_0^* + \theta_1^* X_i, \sigma^2) \quad \mathbb{P}(\mathbb{Y} | \mathbb{X}; \theta) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{(Y_i - \theta_0 - \theta_1 X_i)^2}{2\sigma^2} \right\}$$

$$\log \mathbb{P}(\mathbb{Y} | \mathbb{X}; \theta) = (\text{constant}) - \frac{1}{2\sigma^2} \sum_{i=1}^n (Y_i - \theta_0 - \theta_1 X_i)^2$$

Summary of the two interpretations

- (X, Y) (data, and new examples) come from some distribution
 - we learn best **linear predictor**

versus

- The world is linear; we know the structure of the relation
 - we learn the **coefficients of the structural relation**

PERFORMANCE ASSESSMENT

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R^2 (R-squared)

- Prediction if no regression:

$$\bar{Y} = \frac{1}{n} Y_i$$

- Total sum of squares:

$$TSS = \sum_{i=1}^n (Y_i - \bar{Y})^2$$

"initial" variation in Y

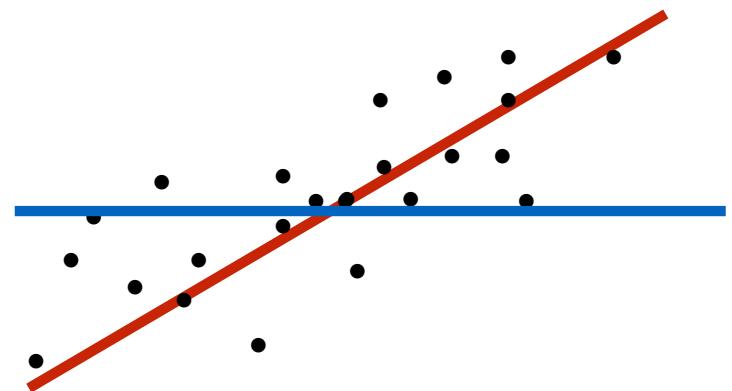
- Residual sum of squares: $RSS = \sum_{i=1}^n (Y_i - \hat{\theta}^T \mathbf{X}_i)^2$

unexplained variation in Y , after taking into account X

- $R^2 = 1 - \frac{RSS}{TSS}$ fraction of variation in Y that has been explained

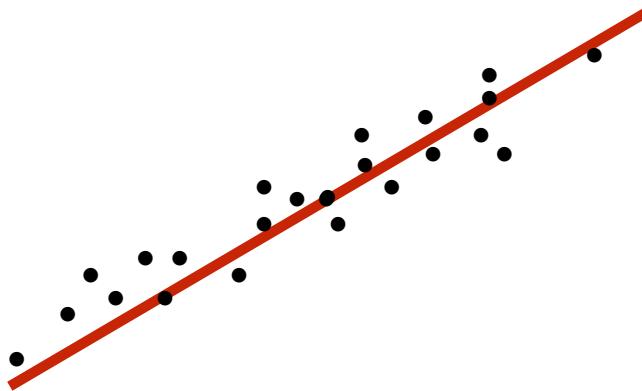
$$0 \leq R^2 \leq 1 \quad \text{high } R^2 \text{ is preferred}$$

in simple regression R^2 is an estimate of
the squared correlation coefficient between X and Y

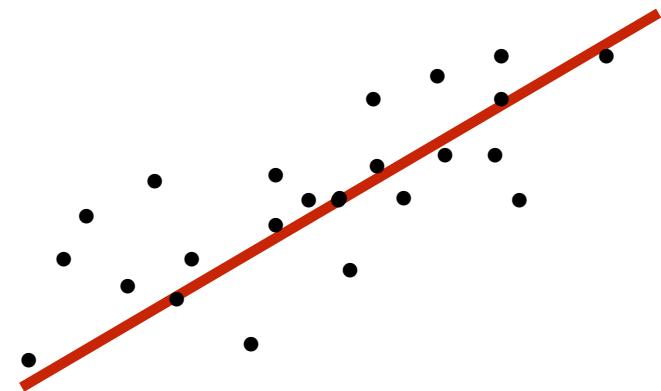


R^2 illustration

- Higher R^2



- Lower R^2



R^2 for our example

- $\widehat{\text{Sales}} = 2.94 + 0.046 \cdot (\text{TV}) + 0.19 \cdot (\text{Radio}) - 0.001 \cdot (\text{NewsP})$
 $R^2 = 0.897$ All the budgets together explain a lot
- $\widehat{\text{Sales}} = 12.35 + 0.055 \cdot (\text{NewsP})$
 $R^2 = 0.05$ Newspaper budget explains little
For TV alone: $R^2 = 0.61$
For Radio alone: $R^2 = 0.33$
- More variables: R^2 can only go up (or stay the same)
 - but this may be a mirage
 - adjusted R^2 : $1 - \frac{\text{RSS}/(n - m - 1)}{\text{TSS}/(n - 1)}$ 0.897 → 0.896

How noisy/reliable are my estimates of θ^*

- Assume structural model



(If not, need to resort to simulation/bootstrap methods)

next
session

$$Y_i = (\theta^*)^T \mathbf{X}_i + W_i$$

W_i : independent,
zero mean, variance σ^2

$\widehat{\Theta}$ is a random variable
(depends on random data)

$$\widehat{\Theta} = (\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X}^T \mathbf{Y}$$

$$\mathbb{E}[(\widehat{\Theta}_j - \theta_j^*)^2] = (\mathbb{E}[\widehat{\Theta}_j] - \theta_j^*)^2 + \text{var}(\widehat{\Theta}_j)$$

$$\mathbb{E}[X^2] = (\mathbb{E}[X])^2 + \text{var}(X)$$

bias

variance

$$\mathbb{E}[\widehat{\Theta}_j] = \theta_j^*$$

(OLS is unbiased)

- Hence focus on the variance of $\widehat{\Theta}_j$

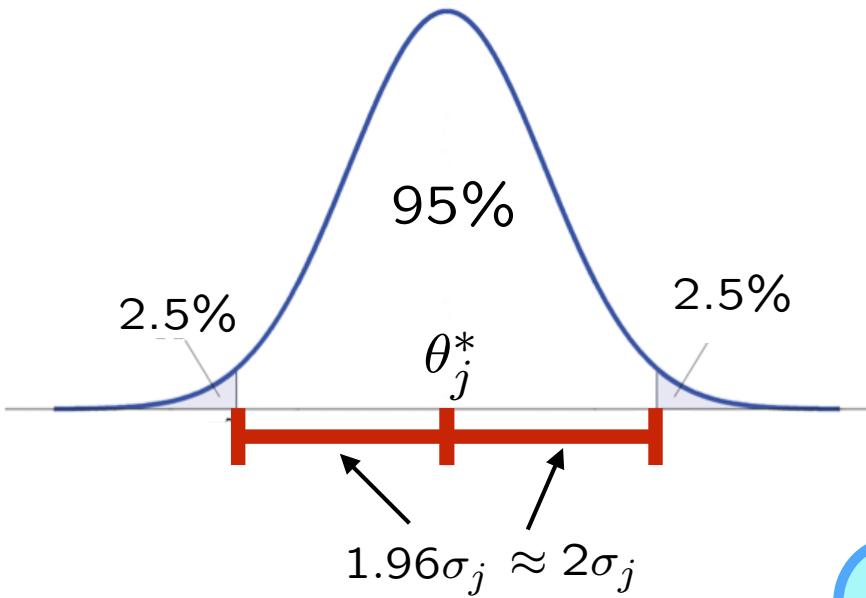


The distribution of $\widehat{\Theta}$

(given, deterministic, \mathbb{X})

- Each $\widehat{\Theta}_j$ is normal

$$\widehat{\Theta}_j \sim \mathcal{N}(\theta_j^*, \sigma_j^2)$$



$$\sigma_j = \sqrt{\text{var}(\widehat{\Theta}_j)} = \text{se}(\widehat{\Theta}_j)$$

standard error

$$Y_i = (\theta^*)^T \mathbb{X}_i + W_i$$

W_i : independent,
zero mean, variance σ^2

$$\widehat{\Theta} = (\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X}^T \mathbf{Y}$$

- approximately: large n , central limit theorem
- exactly: if W_i are normal

- Agenda:
calculate/approximate standard error
use it (confidence intervals, hypothesis testing)

Standard error calculation

- There is a formula
- Software implements it (approximately)

The covariance matrix of $\widehat{\Theta}$ (given, deterministic, \mathbb{X})

$$Y_i = (\theta^*)^T \mathbf{X}_i + W_i$$

W_i : independent,
zero mean, variance σ^2

$$\widehat{\Theta} = (\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X}^T \mathbf{Y}$$

diagonal entries: $\text{Var}(\widehat{\Theta}_j)$

off-diagonal entries: $\text{Cov}(\widehat{\Theta}_i, \widehat{\Theta}_j)$

	const	TV	Radio	Newspaper
const	9.72867479E-02	-2.65727337E-04	-1.11548946E-03	-5.91021239E-04
TV	-2.65727337E-04	1.9457371E-06	-4.47039463E-07	-3.26595026E-07
Radio	-1.11548946E-03	-4.47039463E-07	7.41533504E-05	-1.78006245E-05
Newspaper	-5.91021239E-04	-3.26595026E-07	-1.78006245E-05	3.44687543E-05

dimensions $(m + 1) \times (m + 1)$

formula: $\sigma^2 (\mathbb{X}^T \mathbb{X})^{-1}$

use $\hat{\sigma}^2$

- Estimate σ^2 by

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (Y_i - \hat{\theta}^T \mathbf{X}_i)^2$$

slight downwards bias
negligible bias if $m \ll n$

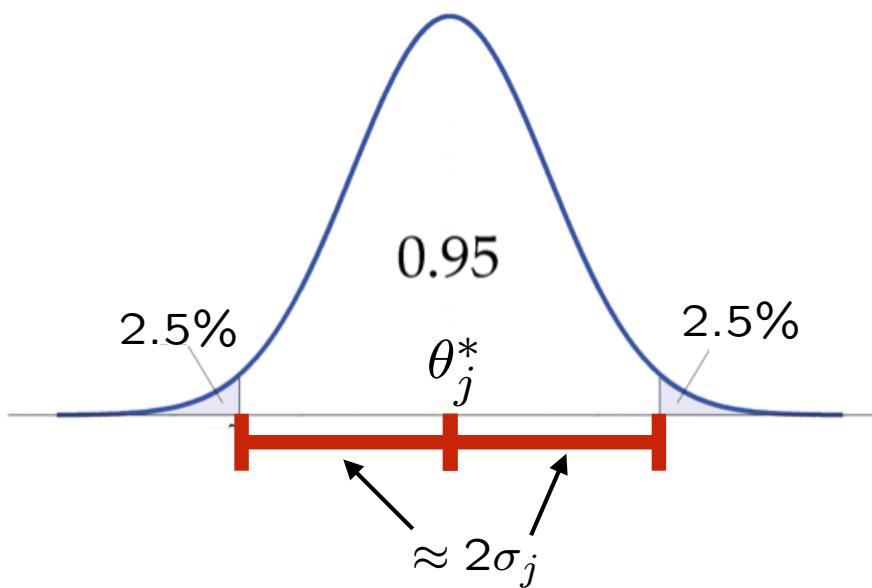
Why? For large samples, $\widehat{\Theta} \approx \theta^*$, and

$$\sigma^2 = \mathbb{E}[W_i^2] \approx \frac{1}{n} \sum_{i=1}^n (Y_i - (\theta^*)^T \mathbf{X}_i)^2 \approx \frac{1}{n} \sum_{i=1}^n (Y_i - \widehat{\Theta}^T \mathbf{X}_i)^2$$

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Confidence Interval (CI)



$$\widehat{\Theta}_j \sim \mathcal{N}(\theta_j^*, \sigma_j^2)$$

- With probability 95%: $|\text{error}| = |\widehat{\Theta}_j - \theta_j^*| \leq 2\sigma_j$

$$\theta_j^* \in [\widehat{\Theta}_j - 2\widehat{\sigma}_j, \widehat{\Theta}_j + 2\widehat{\sigma}_j]$$

95%-CI

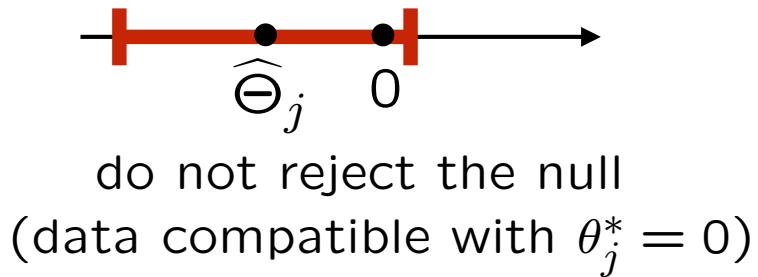
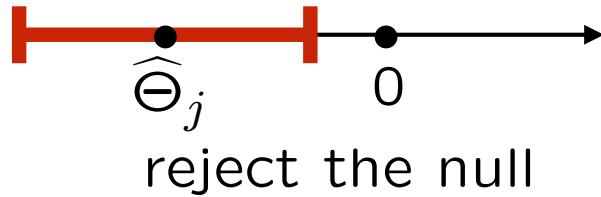
$$\mathbb{P}(\theta_j^* \in \text{CI}) \approx 0.95$$

- Needs careful ("frequentist") interpretation

Testing the hypothesis $\theta_j^* = 0$

- Are the data compatible with the **null hypothesis** $\theta_j^* = 0$?
(*j*th feature has “no effect”)

Wald test:

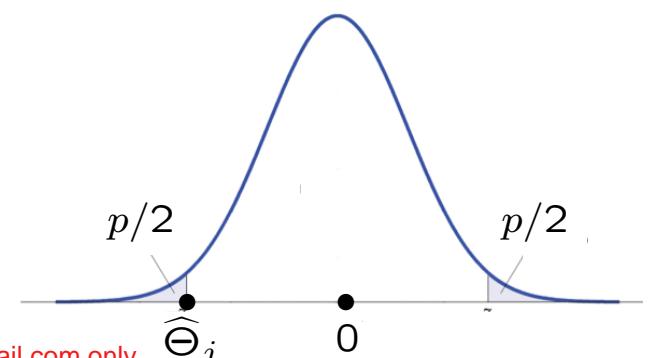


$$\begin{aligned} & \mathbb{P}(\text{reject} \mid \theta_j^* = 0) \quad (\text{false discovery rate}) \\ &= \mathbb{P}(\text{the CI “misses” } 0 \mid \theta_j^* = 0) \approx 5\% \end{aligned}$$

- how much of an outlier (under the null) do I see?

p-value: probability of seeing something at least as extreme as the observed $\hat{\theta}_j$, under $\theta_j^* = 0$

- reject if $p\text{-value} < 0.05$



Back to our example

	coef	std err	Confidence intervals	
			[0.025	0.975]
<hr/>				
Intercept	2.9389	0.312	2.324	3.554
TV	0.0458	0.001	0.043	0.049
Radio	0.1885	0.009	0.172	0.206
Newspaper	-0.0010	0.006	-0.013	0.011

- Wald test: Intercept, TV, Radio are “significant”
(reject the hypothesis that they are zero)
Newspaper: the hypothesis that $\theta_{\text{NewsP}}^* = 0$ “survives”
(not rejected)

Interpretation needs care

Scientists rise up against statistical significance,
Nature, 20 March 2019

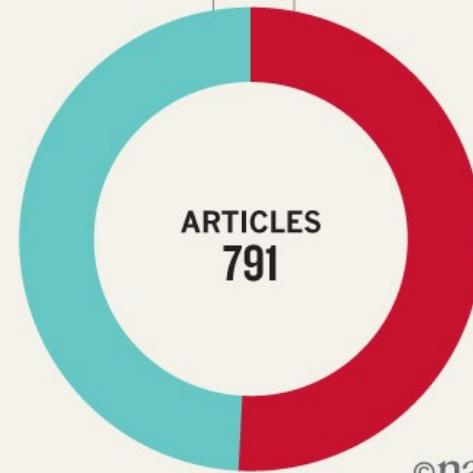
WRONG INTERPRETATIONS

An analysis of 791 articles across 5 journals* found that around half mistakenly assume non-significance means no effect.

*Data taken from: P. Schatz et al. *Arch. Clin. Neuropsychol.* **20**, 1053–1059 (2005); F. Fidler et al. *Conserv. Biol.* **20**, 1539–1544 (2006); R. Hoekstra et al. *Psychon. Bull. Rev.* **13**, 1033–1037 (2006); F. Bernardi et al. *Eur. Sociol. Rev.* **33**, 1–15 (2017).

Appropriately interpreted
49%

Wrongly interpreted
51%



- Reject the null $\theta_j^* = 0$ (decide “there is an effect”): what we see is unlikely to have been generated by a model with $\theta_j^* = 0$
 - but also could be due to noise in the data; 5% “false discovery” prob.
- Do not reject the null $\theta_j^* = 0$ (“see no effect”): data do not provide compelling evidence that $\theta_j^* \neq 0$
 - no effect: θ_j^* is zero
 - small effect: θ_j^* is so close to zero that data cannot detect it
 - too few data: θ_j^* may be nonzero, but need more data to “see it”

Making new predictions

- After running the regression given some new \mathbf{X} , predict $\hat{Y} = \hat{\boldsymbol{\theta}}^T \mathbf{X}$
- Keep assuming structural model: $Y = (\boldsymbol{\theta}^*)^T \mathbf{X} + W$
- $\hat{\boldsymbol{\theta}}$ is unbiased estimate of $\boldsymbol{\theta}^*$
 $\Rightarrow \hat{\boldsymbol{\theta}}^T \mathbf{X}$ is unbiased estimate of $(\boldsymbol{\theta}^*)^T \mathbf{X}$ (and of Y)
- Two sources of error:
 - unavoidable, from W ; variance σ^2
 - variance of $(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}^*)^T \mathbf{X}$
 - Total prediction error variance:

$$\begin{array}{c|c}\mathbf{X}_1 & Y_1 \\ \vdots & \vdots \\ \mathbf{X}_n & Y_n \\ \hline \mathbf{X} & Y?\end{array}$$

$$\sigma^2 \mathbf{X}^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}$$

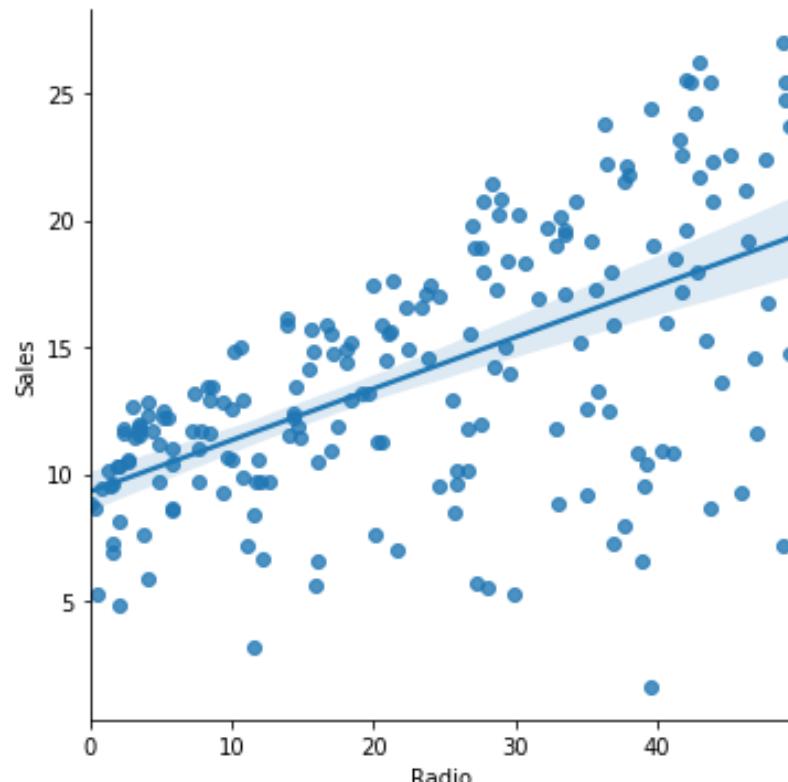
$$\sigma^2 + \sigma^2 \mathbf{X}^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}$$

Confidence bands

- 95% confidence interval about the value of $(\theta^*)^T \mathbf{X}$:

$$(\hat{\theta})^T \mathbf{X} \text{ plus or minus } 2 \cdot \hat{\sigma} \sqrt{\mathbf{X}^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}}$$

- confidence interval width changes with \mathbf{X}
- in simple regression, this gives a **confidence band**



Summary

- Linear regression
 - formulation
 - underlying assumptions
 - formulas
 - results: their interpretation and usage
- Two types of questions:
 $\widehat{\Theta} \approx \theta^*$? (modeling) $\widehat{Y} \approx Y$? (prediction)
- Still, many things can go wrong or be misinterpreted
- New issues when θ has high dimension
- Next session...

Some references

- James et al., An Introduction to Statistical Learning: with Applications in R
(accessible)
- Hastie et al., The Elements of Statistical Learning: Data Mining, Inference, and Prediction
(more comprehensive, and more advanced version of above)
- Wasserman, All of Statistics
(short, elegant, and more mathematical)