

# Lecture 26: Inference about a population proportion

## Chapter 19

Corinne Riddell (Instructor: Tomer Altman)

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# Recap

- ▶ So far we've learned the  $z$ -test and the  $t$ -test that apply when the variable of interest is continuous
- ▶ We applied these tests to one-sample (e.g.,  $H_0 : \mu = 8$ ) and two-sample settings (e.g.,  $H_0 : \mu_1 = \mu_2$ )
- ▶ Today, we will generalize these procedures to binary data, for which we estimate a proportion  $\hat{p}$  from a sample and use that as our best guess of the underlying population parameter  $p$
- ▶ Notation:  $\bar{x}$  is to  $\mu$  as  $\hat{p}$  is to  $p$

# Agenda

- ▶ Confidence interval for a proportion
- ▶ Sample size estimates for a proportion
- ▶ Hypothesis tests for a proportion

## Recall the sampling distribution for $\hat{p}$

The sampling distribution for  $\hat{p}$  is centered on  $p$  with a standard error of  $\sqrt{\frac{p(1-p)}{n}}$

If we follow the same format for the CI from previous chapters we would get:

$$\hat{p} \pm z^* \sqrt{\frac{\hat{p}(1 - \hat{p})}{n}}$$

This is what is known as the **large sample confidence interval for a population proportion**

But...

- ▶ This confidence interval can perform poorly, meaning that if we repeated the confidence interval 100 times (based on 100 random samples), the coverage could be what is termed
  - ▶ overly conservative (e.g., coverage significantly above 95%), or
  - ▶ anti-conservative, or permissive, or “poor coverage”
    - ▶ Fewer than 95 out of 100 of the 95% confidence intervals, on average, would contain the true value for the proportion  $p$
  - ▶ For the Wald approximation, for different combinations of  $p$  and  $n$ , it can exhibit **both** problems!
- ▶ To overcome this, we will modify how we calculate the confidence interval slightly using what is known as the “plus four” method

## Introducing: the “Plus Four” Method

- ▶ If you add two imaginary successes and two failures to the data set (increasing the sample size by four imaginary trials), the interval can have better performance

- ▶ Let

$$\tilde{p} = \frac{\text{number of successes} + 2}{n+4}$$

- ▶ Let  $SE = \sqrt{\frac{\tilde{p}(1-\tilde{p})}{n+4}}$

- ▶ Then the CI is:

$$\tilde{p} \pm z^* \sqrt{\frac{\tilde{p}(1-\tilde{p})}{n+4}}$$

- ▶ This is called the **“Plus Four” Method**
- ▶ Note we use  $z^*$  rather than  $t^*$ . This is because the standard error of the sampling distribution is completely determined by  $p$  and  $n$ , we don't need to estimate a second parameter
- ▶ In addition, for smaller samples (when one cannot rely on the CLT), one can rely on alternative methods to get inference that do not rely on Normality of the  $\hat{p}$
- ▶ Use this method when  $n \geq 10$

# Why does the “Plus Four” Method work?

- ▶ It is a simplification of a more complex method known as the Wilson Score Interval
- ▶ You don't need to know why it works, just that it is better to use this “plus four” trick if you're making a confidence interval for a proportion by hand
- ▶ Note: if the number of successes and failures is relatively large, then the “plus four” method will converge with the large sample CI method

## Two methods so far...

We have so far introduced the large sample method to calculate the CI for  $p$ :

$$\hat{p} \pm z^* \sqrt{\frac{\hat{p}(1 - \hat{p})}{n}}$$

And the plus four method to calculate the CI for  $p$ :

- ▶  $\tilde{p} = \frac{\text{number of successes} + 2}{n+4}$
- ▶ Let  $SE = \sqrt{\frac{\tilde{p}(1-\tilde{p})}{n+4}}$
- ▶ Then the CI is:

$$\tilde{p} \pm z^* \sqrt{\frac{\tilde{p}(1 - \tilde{p})}{n + 4}}$$

We are going to talk about two more methods.



## What does R use?

- ▶ R has two functions to calculate confidence intervals for proportions
- ▶ The first function is `prop.test` (analogous to `t.test`) to calculate confidence intervals and hypothesis tests for binomial proportions
- ▶ This function uses the **Wilson score** method, specifically, the “Wilson score interval with a continuity correction”.
  - ▶ Thus, you don't need to “plus 4” the proportion or standard error
  - ▶ It will do it for you
  - ▶ It does an even better job because of the continuity correction
  - ▶ You do not need to know how to calculate the Wilson score method by hand
  - ▶ You only need to know how to use R to perform this method.

## What does R use?

- ▶ There is a fourth often-used method to compute confidence intervals for proportions called the **Clopper Pearson method**
  - ▶ Also known as the **“Exact method”**
  - ▶ Implemented with the second R function, `binom.test()`
- ▶ The exact method is statistically conservative, meaning that it gives better coverage than it suggests
  - ▶ A 95% CI computed under this method includes the true proportion in the interval more than 95% of the time

- ▶ It can be thought of as an inversion of hypothesis testing
  - ▶ It finds the set of all values of the parameter  $p_0$ , where we consider the null value in a two-sided hypothesis test, would lead to a  $p$ -value  $p \geq 0.05$
  - ▶ It calculates the  $p$ -value using the binomial distribution function
  - ▶ As a reminder, this was the cumulative probability function for binomial (Lec 16):

$$P(X \leq x) = \sum_{i=0}^x \binom{n}{i} p_0^i (1-p_0)^{n-i}$$

## Example applying all the methods

Suppose that 500 elderly individuals suffered hip fractures, of which 100 died within a year of their fracture. Compute the 95% CI for the proportion who died using:

- ▶ the large sample method
- ▶ the “plus four” method (by hand)
- ▶ the Wilson Score method (using `prop.test`)
- ▶ the Clopper Pearson Exact method (using `binom.test`)

## Example of large sample method to calculate the CI for a proportion (by hand)

```
p.hat <- 100/500 # estimate proportion  
se <- sqrt(p.hat*(1-p.hat)/500) # standard error  
p.hat - 1.96*se # Lower confidence bound
```

```
## [1] 0.1649385
```

```
p.hat + 1.96*se # Upper confidence bound
```

```
## [1] 0.2350615
```

- ▶ Our estimate for the proportion is  $\hat{p} = 20\%$
- ▶ Using the large sample method, the 95% confidence interval is 16.5% to 23.5%
- ▶ Remember, this method has poor coverage, meaning that fewer than 95 of the 100 intervals we would make would contain the true value  $p$  on average

## Example using the “plus four” method to calculate the CI for a proportion (by hand)

```
p.tilde <- (100 + 2)/(500 + 4)
se <- sqrt(p.tilde * (1 - p.tilde)/(500 + 4)) # standard error
p.tilde - 1.96 * se # Lower confidence bound
```

```
## [1] 0.1673039
```

```
p.tilde + 1.96 * se # Upper confidence bound
```

```
## [1] 0.237458
```

Using the plus 4 method, the confidence interval is 16.7% to 23.7%.

## Example using the Wilson Score method to calculate the CI for a proportion (using R)

```
prop.test(x = 100, n = 500, conf.level = 0.95)
```

```
##  
## 1-sample proportions test with continuity correction  
##  
## data: 100 out of 500, null probability 0.5  
## X-squared = 178.8, df = 1, p-value < 2.2e-16  
## alternative hypothesis: true p is not equal to 0.5  
## 95 percent confidence interval:  
## 0.1663581 0.2383462  
## sample estimates:  
## p  
## 0.2
```

- ▶ The 95% confidence interval using the Wilson Score method is 16.6% to 23.8%
- ▶ Note that the `prop.test` function is also conducting a two-sided hypothesis test (where  $H_0 : p = 0.5$  unless

## Example using the Clopper Pearson “Exact” method to calculate the CI for a proportion (using R)

```
binom.test(x = 100, n = 500, conf.level = 0.95)
```

► The 95% confidence

interval using the exact

binomial test is 16.6% to 23.8%

► Note that this interval is wider than the one made

with the large sample method

► It has larger coverage (contains the true value more often) which necessitates a wider interval

► Note that the `binom.test` function is also conducting a two-sided hypothesis test (where  $H_0 : p_0 = 0.5$ , unless otherwise specified)

```
##  
## Exact binomial test  
##  
## data: 100 and 500  
## number of successes = 100, number of trials = 500, p-value  
## alternative hypothesis: true probability of success is not  
## 95 percent confidence interval: 0.1658001 0.2377918  
## sample estimates:  
## probability of success  
## 0.2
```

## Summary of the confidence intervals across the methods

| Method            | 95% Confidence Interval | R Function              |
|-------------------|-------------------------|-------------------------|
| Large sample      | 16.5% to 23.5%          | by hand                 |
| Plus four         | 16.7% to 23.7%          | by hand                 |
| Wilson Score*     | 16.6% to 23.8%          | <code>prop.test</code>  |
| Clopper Pearson** | 16.6% to 23.8%          | <code>binom.test</code> |

\*with continuity correction

\*\*also known as the exact method

- ▶ Only the large sample method is symmetric around  $\hat{p} = 20\%$ . This is okay. Symmetric confidence intervals are applicable to Normal sampling distributions, which might or might not describe the distribution of  $\hat{p}$ .
- ▶ Non-symmetric CIs make sense because  $p$  is bounded between 0 and 1. For example, if  $p$  is very small, say 0.012, you would not want a CI that has a lower bound which is negative, this would not make sense.



## Another example of the “plus four” method (by hand)

[We are including another example for you to read so you can practice working out the calculations by hand.]

A study examined a random sample of 75 SARS patients, of which 64 developed recurrent fever.

Therefore  $\hat{p} = 64/75 = 85.33\%$

Using the plus 4 method:  $\tilde{p} = \frac{64+2}{75+4} = 83.54\%$

$$SE = \sqrt{\frac{\tilde{p}(1 - \tilde{p})}{75 + 4}} = \sqrt{\frac{.8354 \times (1 - 0.8354)}{79}} = 0.04172$$

Thus the plus four 95% CI is:

$$\tilde{p} \pm 1.96 \times SE = 0.8354 \pm 0.04172 = 79.37\% \text{ to } 87.71\%$$

## How big should the sample be to estimate a proportion?

Suppose that you want to estimate a sample size for a proportion within a given margin of error. That is, you want to put a maximum bound on the width of the corresponding confidence interval.

Let  $m$  denote the desired margin of error. Then  $m = z^* \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$

We can solve this equation for  $n$ , but we also need to plug in a value for  $p$ . To do that we make a guess for  $p$  denoted by  $p^*$ .

$p^*$  is your best estimate for the underlying proportion. You might gather this estimate from a completed pilot study or based on previous studies published by someone else. If you have no best guess, you can use  $p^* = 0.5$ . This will produce the most conservative estimate of  $n$ . However if the true  $p$  is less than 0.3 or greater than 0.7, the sample size estimated may be much larger than you need.

## How big should the sample be to estimate a proportion?

Rearranging the formula on the last slide for  $n$ , we get:

$$m = z^* \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

$$\sqrt{n}m = z^* \sqrt{p(1-p)}$$

$$\sqrt{n} = \frac{z^*}{m} \sqrt{p(1-p)}$$

$$n = \left(\frac{z^*}{m}\right)^2 p^*(1-p^*)$$

This last formula is the one we will use to estimate the required sample size.

## Example of estimating sample size

Suppose after the general election, you were interested in estimating the number of STEM undergraduate students who voted. So you want to do a study to estimate this proportion. How many students should you include in your sample?

First you need to decide what margin of error you desire. Suppose it is 4 percentage points or  $m = 0.04$  for a 95% CI.

If you had no idea what proportion of STEM students voted then you let  $p^* = 0.5$  and solve for  $n$ :

$$n = \left(\frac{z^*}{m}\right)^2 p^*(1 - p^*) = \left(\frac{1.96}{0.04}\right)^2 \times 0.5 \times 0.5 = 600.25 = 601$$

This implies you would need to sample 601 students to get an estimate with a 95% confidence interval that is +/- 4 percentage points.

However, suppose you found a previous study that estimated the number of STEM students who voted to be 25%. Then what sample size would you need to detect this proportion?

## Example of estimating sample size

What if you want the width of the 95% confidence interval to be 6 percentage points. What would  $m$  be in this case?

## Example of estimating sample size

What if you want the width of the 95% confidence interval to be 6 percentage points. What would  $m$  be in this case?

The width of the 95% CI is equal to twice the margin of error. So if you want the width to be 0.06, then this is equivalent to saying you want a margin of error of 0.03.

# Hypothesis tests of a proportion

When you only have one sample what is the null hypothesis? You're interested in knowing whether there is evidence against the null hypothesis that the population proportion  $p$  is equal to some specified value  $p_0$ . That is:

$$H_0 : p = p_0$$

For example, you may want to test whether there is evidence against the null hypothesis that  $p = 0.25$ .

# Hypothesis tests of a proportion

Recall the sampling distribution for the proportion:

- ▶ Normally distributed
- ▶ Centered at  $p_0$  under the null hypothesis
- ▶ Has a standard error of  $\sqrt{\frac{p_0(1-p_0)}{n}}$



# Hypothesis tests of a proportion

The test statistic for the null hypothesis is:

$$z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$$

This is a  $z$ -test (not a  $t$ -test) so we compare to the standard Normal distribution and ask what is the probability of observing a  $z$  value of this magnitude (or more extreme).

# Hypothesis tests of a proportion

One sided alternatives:

▶  $H_a : p > p_0$

▶  $H_a : p < p_0$

Two-sided alternative:

▶  $H_a : p \neq p_0$

When to use this test? Use this test when the expected number of successes and failures is  $\geq 10$ . That is, when  $np_0 \geq 10$  and  $n(1 - p_0) \geq 10$ .

## Example of a hypothesis test for a proportion

Consider an SRS of 200 patients undergoing treatment to alleviate side effects from a rigorous drug regimen at a particular hospital, where 33 patients experienced reduced or no side effects.

$$\hat{p} = 33/200 = 0.165 = 16.5\%$$

Suppose that historically, the rate of patients with little or no side effects is 10%. Does the new treatment increase the rate? That is:

$$H_0 : p = 0.10$$

$$H_a : p > 0.10$$

## Example of a hypothesis test for a proportion

Step 1: Calculate  $\hat{p} = 16.5\%$  from previous slide.

Step 2: Calculate the standard error of the sampling distribution for  $p$  under the null hypothesis:

$$SE = \sqrt{\frac{p_0(1-p_0)}{n}} = \sqrt{\frac{0.1(1-0.1)}{200}} = 0.0212132$$

Step 3: Calculate the z-test for the proportion:

$$z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}} = \frac{0.165 - 0.10}{0.0212132} = 3.06413$$

Step 4: Calculate the probability of seeing a z-score of this magnitude *or larger*:

```
pnorm(q = 3.06413, lower.tail = F)
```

```
## [1] 0.00109152
```

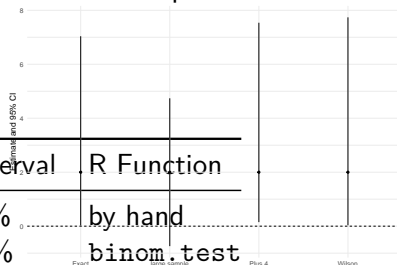
Step 5: Evaluate the evidence against the null hypothesis. Because the  $p$ -value is so small (0.1%), there is little chance of seeing a proportion equal to 16.5% or larger if the true proportion is actually 10%. This is the evidence in favor of the alternative hypothesis.

## Another Example

Suppose that there were 100 elderly individuals with falls observed, and 2 died. Here are the 95% CIs applying the four different methods:

| Method        | 95% Confidence Interval | R Function              |
|---------------|-------------------------|-------------------------|
| Large sample  | -0.74% to 4.74%         | by hand                 |
| Exact         | 0.024% to 7.04%         | <code>binom.test</code> |
| Wilson Score* | 0.034% to 7.74%         | <code>prop.test</code>  |
| Plus four     | 0.15% to 7.54%          | by hand                 |

We can graphically compare the CIs from the previous slide:



\*with continuity correction

## Elderly falls example

### Findings:

- ▶ Notice how different the intervals are, especially large sample vs. others
- ▶ Notice that the large sample lower bound is nonsensical (i.e., we can't have negative proportions!)
- ▶ The large sample CI differs from the others because the Normal approximation assumptions are not satisfied

## Code for elderly falls example

```
p.hat <- 2/100 # estimate proportion
se <- sqrt(p.hat*(1-p.hat)/100) # standard error
c(p.hat - 1.96*se, p.hat + 1.96*se) # CI
## [1] -0.00744 0.04744

binom.test(x = 2, n = 100, p = 0.5, conf.level = 0.95)
## data: 2 out of 100, null p = 0.5
## X-squared = 90.25, df = 1,
## alternative hypothesis: true
## 95 percent confidence interval:
## 0.003471713 0.077363988
## sample estimates:
## number of successes = 2, number of trials = 100, p-value
## alternative hypothesis: true, probability of success is not
## 0.02
## 95 percent confidence interval:
## 0.002431337 0.070383932
## sample estimates:
## probability of success
## 0.02
```

Check your understanding!



# Recap

- ▶ The binomial distribution has all of the same statistical “procedures”:
  - ▶ Confidence intervals
  - ▶ Sample size estimation
  - ▶ Hypothesis testing
- ▶ However, there are problems with the coverage of the CIs:
  - ▶ Subtle interactions between discrete binomial distribution and Normal-based approximations
  - ▶ Use “Plus 4” method when calculating “by hand” and sample is “small”
  - ▶ Use `prop.test` when using R functions
  - ▶ The coverage problem goes away as the sample size gets larger, and more of the Normality assumptions are met