

Statistics is Everywhere

Regression

Fitting a linear model in R

Add the regression line to
the scatter plot using
`geom_abline()`

Transforming data

How do outliers affect the
line of best fit?

Counfounding

L06: Intro to Linear Regression

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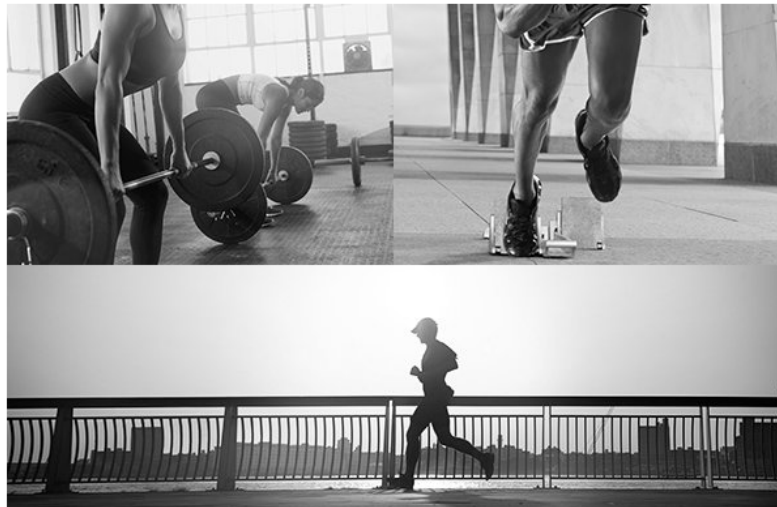
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Exercise and the Brain

PHYS ED

Which Type of Exercise Is Best for the Brain?

BY GRETCHEN REYNOLDS FEBRUARY 17, 2016 5:45 AM 509



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- ▶ from [The New York Times](#), February 2016:

“Some forms of exercise may be much more effective than others at bulking up the brain, according to a remarkable new study in rats. For the first time, scientists compared head-to-head the neurological impacts of different types of exercise: running, weight training and high-intensity interval training. The surprising results suggest that going hard may not be the best option for long-term brain health”

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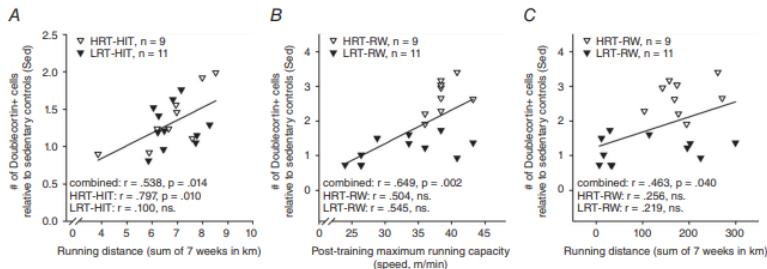
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► from *The Journal of Physiology*



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“More evidence that exercise can have beneficial effects on cognitive performance” December 10, 2024 Heard on Morning Edition (NPR)

“There’s no doubt that exercise is good for your heart, your muscles and your bones. Regular physical activity can also reduce the risk of cognitive decline. The latest study included men and women aged 50 up to 83 years old who agreed to wear fitness trackers and to complete a bunch of cognitive tests, including memory tests.”

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Bloomberg et al. Int J Behav Nutr Phys Act (2024) 21:133

76 British Adults ages 50 to 83 without evidence of cognitive issues or dementia.

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Table 2 Associations of previous day physical activity and sedentary behaviour with cognitive performance ($N = 76$)

	Physical activity and sedentary behaviour					
	β (95% CI)					
	MVPA	<i>p</i> -value	LPA	<i>p</i> -value	SB	<i>p</i> -value
Attention	0.02 (-0.10, 0.14)	0.74	-0.02 (-0.08, 0.04)	0.56	-0.01 (-0.06, 0.04)	0.70
Episodic memory	0.15 (0.01, 0.29)	0.03	-0.03 (-0.11, 0.05)	0.45	-0.04 (-0.10, 0.02)	0.17
Working memory	0.16 (0.03, 0.28)	0.01	0.02 (-0.05, 0.08)	0.61	-0.05 (-0.09, 0.00)	0.03
Psychomotor speed	0.09 (0.00, 0.18)	0.06	0.01 (-0.04, 0.05)	0.82	0.00 (-0.03, 0.03)	0.96
Executive function	0.07 (-0.07, 0.20)	0.35	0.01 (-0.05, 0.08)	0.72	-0.03 (-0.08, 0.02)	0.18
Processing Speed	0.03 (-0.08, 0.14)	0.58	0.00 (-0.05, 0.05)	0.87	0.01 (-0.03, 0.05)	0.66

Coefficients correspond to change in cognitive performance per 30-min increase in physical activity or sedentary behaviour on the previous day

Units are standard deviations. Adjusted for age, gender, education, mobility limitations, self-rated health, depressive symptoms, employment status, BMI, time-of-day of cognitive test taking, round of cognitive testing, weekend or weekday, habitual physical activity and sedentary behaviour, habitual sleep, previous cognitive score, sleep parameters at time $t - 2$, and physical activity parameters at time t

Abbreviations: CI, confidence interval; MVPA, moderate-vigorous physical activity; LPA, light physical activity; SB, sedentary behaviour

Table 3 Associations between previous day sleep characteristics and cognitive performance ($N = 76$)

	Sleep					
	β (95% CI)					
	Sleep duration ≥ 6 hours ^a	<i>p</i> -value	REM sleep ^b	<i>p</i> -value	SWS ^b	<i>p</i> -value
Attention	0.39 (0.00, 0.79)	0.05	0.13 (0.00, 0.25)	0.04	0.11 (-0.02, 0.23)	0.10
Episodic memory	0.60 (0.16, 1.03)	0.008	0.02 (-0.11, 0.16)	0.74	0.17 (0.05, 0.30)	0.008
Working memory	-0.16 (-0.59, 0.26)	0.45	-0.04 (-0.17, 0.09)	0.54	-0.03 (-0.18, 0.12)	0.67
Psychomotor speed	0.34 (0.04, 0.65)	0.03	0.05 (-0.04, 0.14)	0.26	-0.03 (-0.13, 0.06)	0.50
Executive function	0.23 (-0.21, 0.66)	0.31	0.03 (-0.10, 0.16)	0.67	0.03 (-0.12, 0.19)	0.67
Processing Speed	-0.13 (-0.46, 0.20)	0.45	0.05 (-0.05, 0.15)	0.31	0.01 (-0.09, 0.11)	0.88

Abbreviations: CI confidence interval, REM rapid eye movement, SWS slow wave sleep

^a Reference: overnight sleep duration < 6 h

^b Coefficient shown corresponds to change in cognitive performance per 30-min increase in sleep stage on the previous day

Units are standard deviations. Adjusted for age, gender, education, mobility limitations, self-rated health, depressive symptoms, employment status, BMI, time-of-day of cognitive test taking, round of cognitive testing, weekend or weekday, habitual physical activity and sedentary behaviour, habitual sleep, previous cognitive score, physical activity parameters at time $t - 1$, and physical activity parameters at time t

- ▶ Introduce linear regression
 - ▶ How do we find the line of best fit?
 - ▶ What is the slope?
 - ▶ What is the intercept?
 - ▶ What is the R squared?
- ▶ Using R to run a linear regression and add a regression line to a scatter plot
- ▶ How do we transform data that do not look linear to make a line?
- ▶ How do outliers influence our line of best fit?
- ▶ Some Important cautions
 - ▶ Association is not causation
 - ▶ Do not extrapolate beyond your data
 - ▶ Always consider potential **confounders** in your interpretation
 - ▶ Confirm the shape of your data visually

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Regression

What is a regression line?

- ▶ A straight line that is **fitted** to data to minimize the distance between the data and the fitted line.
- ▶ It is often called the **line of best fit**.
- ▶ It is also called the **least-squares regression line** (sometimes referred to as *ordinary least squares* or *ols*) this is because mathematically, the criteria for choosing this line is based on the sum of squares of the vertical distances from the line. We choose the line that minimizes this sum.

What is a regression line?

Once we have calculated this line, the line of best fit can be used to describe the relationship between the explanatory and response variables.

- ▶ Can you fit a line of best fit for non-linear relationships?
- ▶ Very important to visualize the relationship first. Why?

Equation of the line of best fit

The line of best fit can be represented by the equation for a line:

$$y = a + bx$$

where a is the **intercept** and b is the **slope**.

This equation encodes a lot of useful information

In earlier math classes you may have seen this expressed as:

$$y = mx + b$$

Equation of the line of best fit: the intercept

$$y = a + bx$$

If $x = 0$, the equation says that $y = a$, which is why a is known as the intercept.

Note: Is the value of the intercept always meaningful?

Equation of the line of best fit: the slope

$$y = a + bx$$

b is known as the slope because an increase from x to $x + 1$ is associated with an increase in y by the amount b .

The slope is closely related to the correlation coefficient:

$$b = r \frac{S_y}{S_x}$$

If the correlation coefficient is negative what will be the sign of the b ?

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The r^2 value or R squared, is the fraction of the variation in the values of y that is explained by the regression of y on x

In a regression where every observation fell exactly on the regression line, the value of r^2 would be 1.

In a linear regression with only one x the r^2 is the square of the correlation coefficient.

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Fitting a linear model in R

Fitting a linear model in R

Code template:

```
lm(formula = y ~ x, data = your_dataset)
```

- ▶ `lm()` is the function for a linear model.
- ▶ The first argument that `lm()` wants is a formula `y ~ x`.
 - ▶ `y` is the response variable from your dataset
 - ▶ `x` is the explanatory variable
 - ▶ be careful with the order of `x` and `y`! It is opposite from the default order in `ggplot`

```
ggplot(data,aes(x=your_x, y=your_y))
```

- ▶ The second argument sent to `lm()` is the data set.
 - ▶ the default order of declaring the data as the second argument in `lm()` is different from the `ggplot2` and `dplyr` functions

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Why the package broom?

We will pull in a new package here: `library(broom)` and apply the `tidy()` function as follows: `tidy(your_lm)`

- ▶ `broom` has functions that make the output from the linear model look clean
- ▶ `tidy` is a function from the `broom` package that tidies up the output

Example: Manatee deaths and powerboat purchases

Let's apply the `lm()` function. Recall the manatee example from our last lecture that examined the relationship between the number of registered powerboats and the number of manatee deaths in Florida between 1977 and 2016.

Recall that the relationship appeared linear when we examined the scatter plot:

```
library(ggplot2)
mana_death1<-ggplot(mana_data, aes(x = powerboats, y = deaths)) +
  geom_point() +
  theme_minimal(base_size = 15)
```

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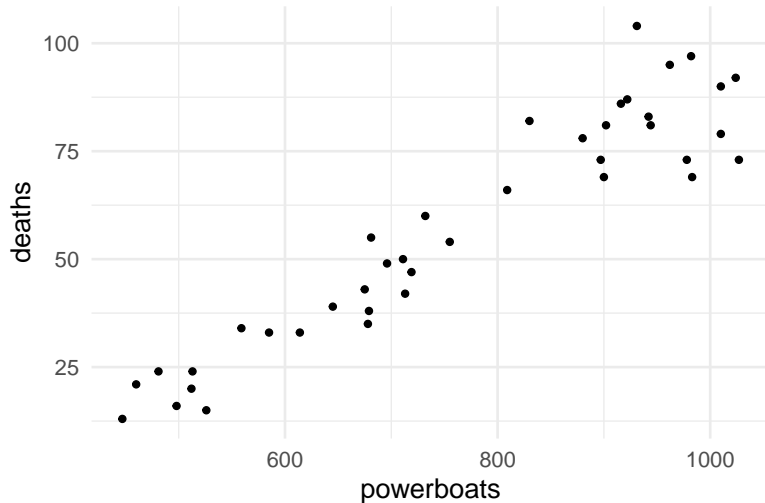
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Manatee deaths and powerboat purchases

mana_death1



lm() of manatee deaths and powerboat purchases

Calculate the line of best fit:

```
mana_lm <- lm(deaths ~ powerboats, mana_data)
library(broom)
tidy(mana_lm)
```

```
## # A tibble: 2 x 5
##   term          estimate std.error statistic  p.value
##   <chr>          <dbl>     <dbl>     <dbl>    <dbl>
## 1 (Intercept)   -46.8      6.03      -7.75 2.43e- 9
## 2 powerboats    0.136     0.00764    17.8 5.21e-20
```

Only pay attention to the term and estimate columns for now.

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lm() of manatee deaths and powerboat purchases

Interpret the model output

```
## # A tibble: 2 x 5
##   term          estimate std.error statistic  p.value
##   <chr>          <dbl>    <dbl>    <dbl>    <dbl>
## 1 (Intercept)   -46.8      6.03     -7.75 2.43e- 9
## 2 powerboats    0.136    0.00764    17.8 5.21e-20
```

- ▶ Intercept: The predicted number of deaths if there were no powerboats. But the prediction is negative. Why?
- ▶ Powerboats: This is the slope. What does the estimated slope for powerboats mean?

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Interpreting the slope

```
## # A tibble: 2 x 5
##   term          estimate std.error statistic  p.value
##   <chr>          <dbl>    <dbl>    <dbl>    <dbl>
## 1 (Intercept)   -46.8      6.03     -7.75 2.43e- 9
## 2 powerboats    0.136     0.00764    17.8 5.21e-20
```

- ▶ A one unit change in the number of powerboats registered (X 1,000) is associated with an increase of manatee deaths of 0.1358. That is, an increase in the number of powerboats registered by 1,000 is association with 0.1358 more manatee deaths.
- ▶ If powerboat registered increased by 100,000 how many more manatee deaths are expected?

Change units

```
mana_data_units<-mana_data%>%mutate(actual_powerboats = powerboats * 1000)
mana_lm_units <- lm(deaths ~ actual_powerboats, mana_data_units)
tidy(mana_lm_units)
```

```
## # A tibble: 2 x 5
```

##	term	estimate	std.error	statistic	p.value
##	<chr>	<dbl>	<dbl>	<dbl>	<dbl>
## 1	(Intercept)	-46.8	6.03	-7.75	2.43e- 9
## 2	actual_powerboats	0.000136	0.00000764	17.8	5.21e-20

What happened to the slope? To the intercept?

Getting the R-squared from your model

When we run a linear model, the r-squared is also calculated. Here is how to see the r-squared for the manatee data:

```
library(broom)  
glance(mana_lm)
```

```
## # A tibble: 1 x 12  
##   r.squared adj.r.squared sigma statistic  p.value    df logLik   AIC    BIC  
##   <dbl>      <dbl> <dbl>      <dbl>    <dbl> <dbl> <dbl> <dbl> <dbl>  
## 1    0.893      0.890  8.82      316. 5.21e-20     1  -143.  292.  297.  
## # i 3 more variables: deviance <dbl>, df.residual <int>, nobs <int>
```

Focus on:

- ▶ Column called `r.squared` values only.
- ▶ Interpretation of r-squared: The fraction of the variation in the values of y that is explained by the line of best fit.

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Correlation vs R Squared

```
library(dplyr)
mana_cor <- mana_data %>%
  summarize(corr_mana = cor(powerboats, deaths))
mana_cor
```

```
## # A tibble: 1 x 1
##   corr_mana
##   <dbl>
## 1      0.945
```

Correlation vs R Squared

```
glance(mana_lm)%>% pull(r.squared)
```

```
## [1] 0.8926573
```

```
#square the correlation coefficient  
.9448054^2
```

```
## [1] 0.8926572
```

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Add the regression line to the scatter plot using `geom_abline()`

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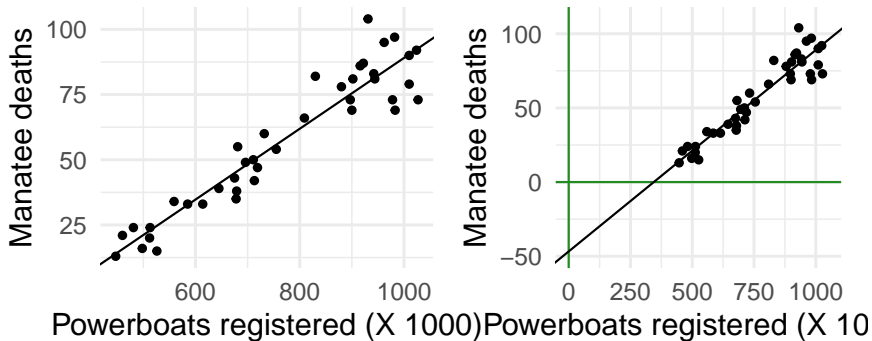
Confounding

We add a statement to our ggplot `geom_abline(intercept = your_intercept, slope = your_slope)`

so for our manatee data `geom_abline(intercept = -46.7520, slope = 0.1358)`

Note: by default, ggplot only shows the plotting region that corresponds to the range of data

Add the regression line to the scatter plot using `geom_abline()`



Add the regression line to the scatter plot using `geom_abline()`

- ▶ When we add the line, we can see the intercept estimate. It is where the line of best fit intersects the y axis. Should we interpret it?
 - ▶ It is far from the bulk of the data, there is no data near powerboats = 0
 - ▶ Interpretation would be [extrapolation](#), and is not supported by these data

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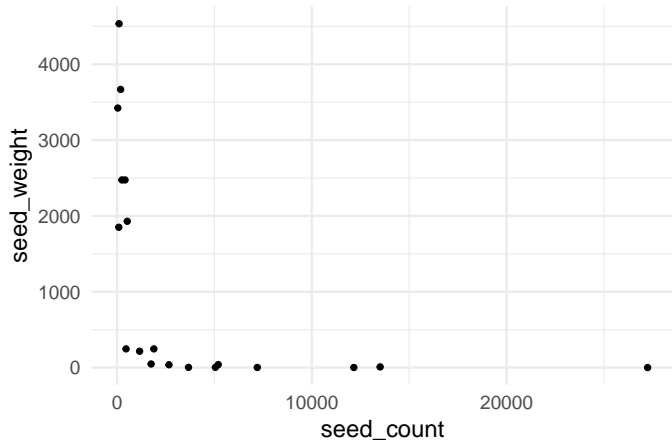
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- ▶ Sometimes, the data is transformed to another scale so that the relationship between the transformed x and y is linear
- ▶ Table 3.4 in B&M provides data on the mean number of seeds produced in a year by several common tree species and the mean weight (in milligrams) of the seeds produced.

Scatter plot of seed_weight vs. seed_count



- ▶ seed_count and seed_weight both vary widely
- ▶ Their relationship is not linear

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Investigate the relationship between their logged variables

- ▶ Add transformed variables to the dataset using `mutate()`.
- ▶ We add both log base e and log base 10 variables for illustration

```
library(dplyr)
seed_data <- seed_data %>% mutate(log_seed_count = log(seed_count),
                                   log_seed_weight = log(seed_weight),
                                   log_b10_count = log(seed_count, 10),
                                   log_b10_weight = log(seed_weight, 10))
```

Plot transformed data (log base e)

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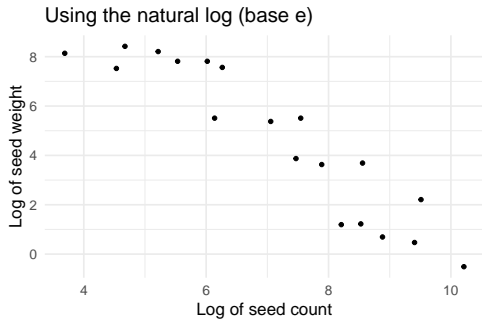
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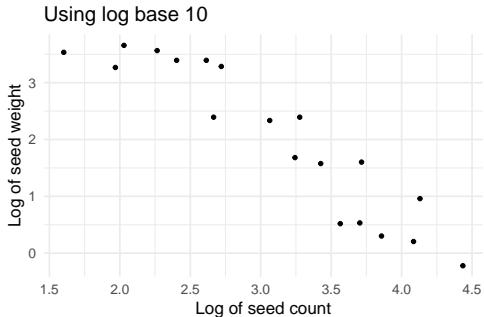
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How do outliers affect the
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Plot transformed data (log base 10)



- ▶ You can use either base 10 or base e for class.
- ▶ The calculations using base e are easier

lm() on the log (base e) variables

```
seed_mod <- lm(log_seed_weight ~ log_seed_count, data = seed_data)
tidy(seed_mod)
```

```
## # A tibble: 2 x 5
```

##	term	estimate	std.error	statistic	p.value
##	<chr>	<dbl>	<dbl>	<dbl>	<dbl>
## 1	(Intercept)	15.5	1.08	14.3	6.37e-11
## 2	log_seed_count	-1.52	0.147	-10.4	9.28e- 9

```
glance(seed_mod) %>% pull(r.squared)
```

```
## [1] 0.8631177
```

- ▶ Interpret the intercept:
- ▶ Interpret the slope:

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lm() on the log (base 10) variables

```
seed_mod_b10 <- lm(log_b10_weight ~ log_b10_count, data = seed_data)
tidy(seed_mod_b10)
```

```
## # A tibble: 2 x 5
```

```
##   term                estimate std.error statistic  p.value
##   <chr>              <dbl>      <dbl>      <dbl>    <dbl>
## 1 (Intercept)         6.73        0.469      14.3 6.37e-11
## 2 log_b10_count     -1.52        0.147     -10.4 9.28e- 9
```

```
glance(seed_mod_b10) %>% pull(r.squared)
```

```
## [1] 0.8631177
```

- What is different from the log base e output?

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Predictions from `lm()` when using \log (base e) data

- ▶ What seed weight is predicted for a seed count of 2000?
- ▶ Worked calculation:

1. Write down the line of best fit:

$$\log_e(\text{seed.weight}) = 15.49130 - 1.522220 \times \log_e(\text{seed.count})$$

2. Plug in $\text{seed.count} = 2000$ into the line of best fit:

$$\log_e(\text{seed.weight}) = 15.49130 - 1.522220 \times \log_e(2000)$$

3. Solve for seed count by exponentiating both sides:

$$\text{seed.weight} = \exp(15.49130 - 1.522220 \times \log_e(2000))$$

(this uses the property that $e^{\log_e(x)} = x$)

$$\text{seed.weight} = 50.45$$

4. Interpret: Seeds are expected to weigh 50.45 for trees having a seed count of 2000.

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**How do outliers affect the
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How do outliers affect the line of best fit?

How do outliers affect the line of best fit?

To study this, we use data from the Organization for Economic Co-operation and Development (OECD). This dataset was downloaded from <http://dx.doi.org/10.1787/888932526084> and contains information on the health expenditure per capita and the GDP per capita for 40 countries.

```
library(readxl)

spending_dat <- read_xlsx("Ch04_Country-healthcare-spending.xlsx",
                          sheet = 2,
                          range = "A7:D47")
```

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Have a look

Next, we want to examine the imported data to see if it is how we expect:

```
head(spending_dat)
```

```
## # A tibble: 6 x 4
##   Country Country.code 'Health expenditure per capita' 'GDP per capita'
##   <chr>      <chr>                <dbl>                <dbl>
## 1 Australia AUS                3445                39409
## 2 Austria   AUT                4289                38823
## 3 Belgium   BEL                3946                36287
## 4 Brazil    BRA                 943                10427
## 5 Canada    CAN               4363                38230
## 6 Chile     CHL               1186                14131
```

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Rename() some variables to use a consistent naming style

If the variable name has spaces, we must use back ticks when referring to it:

```
library(dplyr)
spending_dat <- spending_dat %>%
  rename(country_code = Country.code,
         health_expenditure = `Health expenditure per capita`, # back ticks
         GDP = `GDP per capita`) # back ticks
```

Statistics is Everywhere

Regression

Fitting a linear model in R

Add the regression line to
the scatter plot using
geom_abline()

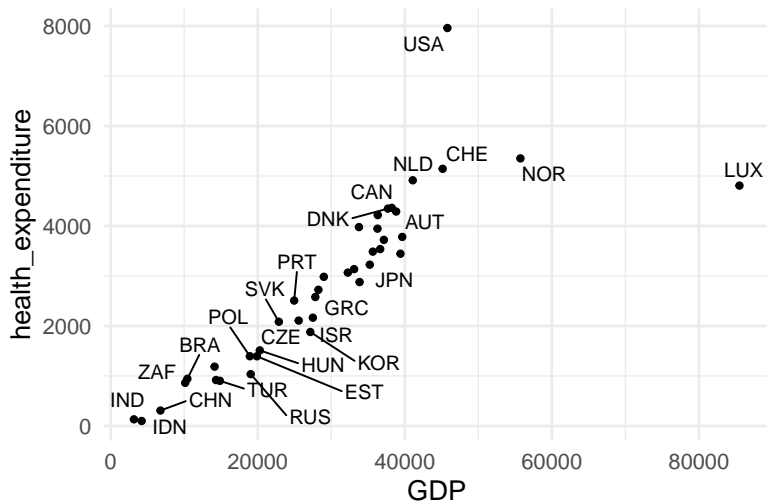
Transforming data

How do outliers affect the
line of best fit?

Counfoundings

Examine the relationship

Make a scatter plot of `health_expenditure` (our response variable) vs. each country's level of GDP:



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Confounding

Examine the relationship

Is the relationship linear? Which countries are outliers?

Fit a linear model to these data

```
lm(health_expenditure ~ GDP, data = spending_dat)
```

```
##  
## Call:  
## lm(formula = health_expenditure ~ GDP, data = spending_dat)  
##  
## Coefficients:  
## (Intercept)          GDP  
##    44.65623      0.09399
```

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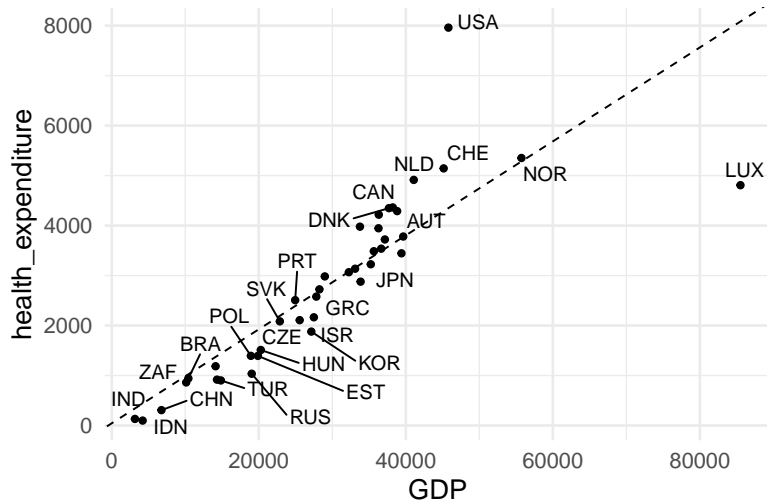
Counfoundings

Examine the relationship

Add the regression line to the graph:

```
GDP_withline<-ggplot(spending_dat, aes(x = GDP, y = health_expenditure)) +  
  geom_point() +  
  geom_text_repel(aes(label = country_code)) + # this adds the country code a  
  geom_abline(intercept = 44.65623, slope = 0.09399, lty = 2) +  
  theme_minimal(base_size = 15)
```

Examine the relationship



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Counfoundings

Examine the relationship without Luxembourg in the data

Let's see whether removing Luxembourg changes the fit of the line. We can remove Luxembourg using the `filter()` command from `dplyr`:

```
spending_dat_no_LUX <- spending_dat %>% filter(country_code != "LUX")  
  
lm(health_expenditure ~ GDP, data = spending_dat_no_LUX)
```

```
##  
## Call:  
## lm(formula = health_expenditure ~ GDP, data = spending_dat_no_LUX)  
##  
## Coefficients:  
## (Intercept)          GDP  
##   -785.1044         0.1264
```

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Transforming data

How do outliers affect the
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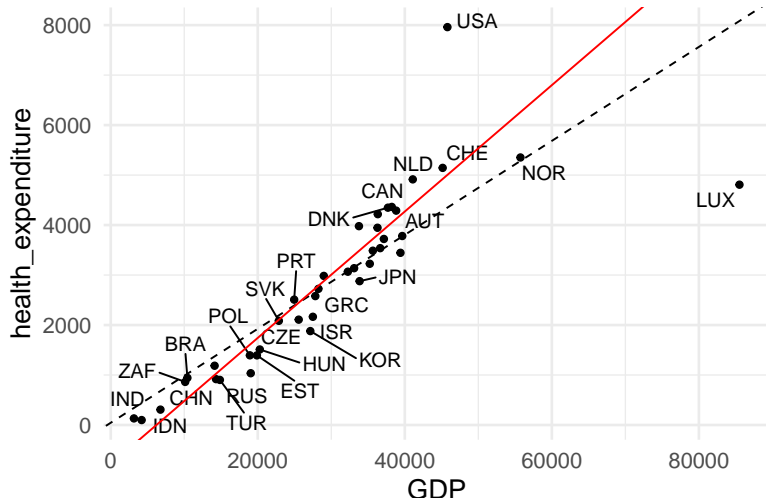
Confounding

Examine the relationship without Luxembourg in the data

```
GDP_nolux<-ggplot(spending_dat, aes(x = GDP, y = health_expenditure)) + geom_
  geom_text_repel(aes(label = country_code)) +
  geom_abline(intercept = 44.65623, slope = 0.09399, lty = 2) +
  geom_abline(intercept = -785.1044, slope = 0.1264, col = "red") +
  theme_minimal(base_size = 15)
```

Examine the relationship without Luxembourg in the data

```
## Warning: ggrepel: 15 unlabeled data points (too many overlaps).
## increasing max.overlaps
```



Consider

Statistics is Everywhere

Regression

Fitting a linear model in R

Add the regression line to
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Transforming data

How do outliers affect the
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Counfoundng

Examine the relationship without USA in the data

```
spending_dat_no_USA <- spending_dat %>% filter(country_code != "USA")  
  
lm(health_expenditure ~ GDP, data = spending_dat_no_USA)
```

```
##  
## Call:  
## lm(formula = health_expenditure ~ GDP, data = spending_dat_no_USA)  
##  
## Coefficients:  
## (Intercept)          GDP  
##   152.26274       0.08714
```

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Add the regression line to
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Transforming data

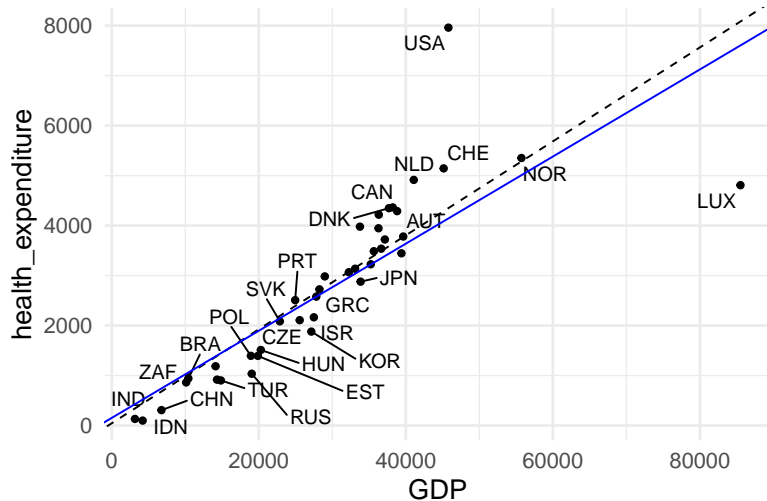
How do outliers affect the
line of best fit?

Counfounding

Examine the relationship without USA in the data

```
GDP_nousa<-ggplot(spending_dat, aes(x = GDP, y = health_expenditure)) +  
  geom_text_repel(aes(label = country_code)) +  
  geom_abline(intercept = 44.65623, slope = 0.09399, lty = 2) +  
  geom_abline(intercept = 152.26274, slope = 0.08714, col = "blue") +  
  theme_minimal(base_size = 15)
```


Examine the relationship without USA in the data



Examine the relationship without LUX or USA in the data

Let's write the code together to remove both the USA and LUX and see how it affects the fit:

```
spending_dat_no_USA_LUX <- spending_dat %>%  
  
  filter(country_code != "USA" & country_code != "LUX")  
  
#alternatively, you could have written:  
spending_dat_no_USA_LUX <- spending_dat %>%  
  
  filter(! country_code %in% c("USA", "LUX"))  
  
#pick the filter command that makes the most sense to you.
```

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Confounding

Examine the relationship without LUX or USA in the data

```
lm(health_expenditure ~ GDP, data = spending_dat_no_USA_LUX)
```

```
##
```

```
## Call:
```

```
## lm(formula = health_expenditure ~ GDP, data = spending_dat_no_USA_LUX)
```

```
##
```

```
## Coefficients:
```

```
## (Intercept)          GDP
```

```
##    -592.6973      0.1166
```

```
GDP_noluxnousa <- ggplot(spending_dat_no_USA_LUX, aes(x = GDP, y = health_expen  
  geom_text_repel(aes(label = country_code)) +  
  geom_abline(intercept = 44.65623, slope = 0.09399, lty = 2) +  
  geom_abline(intercept = -592.6973, slope = 0.1166, col = "green") +  
  theme_minimal(base_size = 15)
```

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Fitting a linear model in R

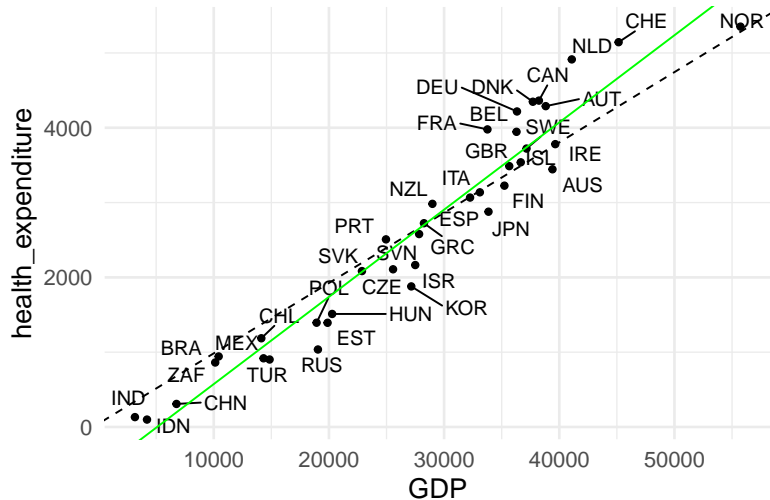
Add the regression line to
the scatter plot using
geom_abline()

Transforming data

How do outliers affect the
line of best fit?

Courtesy of

Examine the relationship without LUX or USA in the data



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Add the regression line to

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How do outliers affect the

line of best fit?

Counfoundings

Examine the relationship without LUX or USA in the data

What would happen if USA's point had actually been along the original line of best fit (say at $x = 80000$ and $y = 7500$) and we re-fit the line without USA's point?

Would USA have been an **outlier**? Would it be considered **influential**?

But, is it causal?

- ▶ Creating a scatter plot and a simple linear model is an important step in many analyses. It allows you to see the relationship between two quantitative variables and estimate the line of best fit.
- ▶ Sometimes these relationships will be used to make claims of causality.

Baldi & Moore emphasize that experiments are the best way to study causality. While this is often true, sophisticated causal methods have been developed for the analysis of observational data.

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Counfounding

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Counfounding

Your book talks about “lurking variables” which Baldi & Moore define as:

A variable that is not among the explanatory or response variables in a study and yet may influence the interpretation of relationships among those variables.

They also (pg 157) define confounding by saying:

Two variables (explanatory or lurking) are confounded when their effects on a response variable cannot be distinguished from each other.

I strongly disagree with this definition. We will use a different definition in this class.

Definition of Confounding

A relationship between your variable of interest (exposure, treatment) and your outcome of interest (disease status, health condition etc) is confounded when there is a variable that is associated with both the exposure and outcome, and is not on the causal pathway between the two.

Variables that are on the causal pathway are those that represent a way in which the exposure acts on the outcome. For example, poor cognitive function would be on the causal pathway between lack of sleep and trying to pay for groceries with your library card.

Discussion of Music example from Baldi & Moore

Example 4.7 “Nature, nurture, and lurking variables” presents an advertisement from the Michigan Symphony:

“Question: Which students scored 51 points higher in verbal skills and 39 points higher in math?

Answer: Students who had experience in music.”

Marketers often make leading statements that make their product or service sound appealing. The purpose of this ad was to have the target audience impute that music causes higher marks at school because there is an association between enrollment in music and higher marks. However, are students enrolled in music lessons otherwise the same as students not enrolled in music lessons? What else do you expect to differ between these groups of students?

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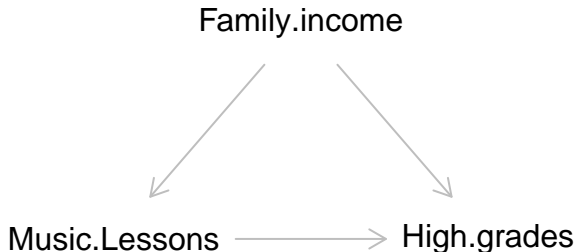
Transforming data

How do outliers affect the
line of best fit?

Counfoundings

Discussion of some examples from Baldi & Moore

We can encode these differences in a causal diagram. Here is a simple one to



demonstrate the concept:

The direction of the arrows from the “Family Income” node makes explicit that we believe family income to be a confounder of the relationship between taking music lessons and achieving higher grades. It means that not only do these children take music lessons, they also come from families with higher incomes, and higher incomes lead to higher grades in other ways. Of course, family income is not the only possible confounder. What are some others?

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Counfounding

In this course, we don't have time to go into methods that adjust for multiple variables or address how to control for confounding or other types of bias that limit causal interpretations.

However, know that causality can be studied using observational data and relies on clever study designs and oftentimes on advanced methods.

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Counfounding

We introduced some code today for running linear regressions in R

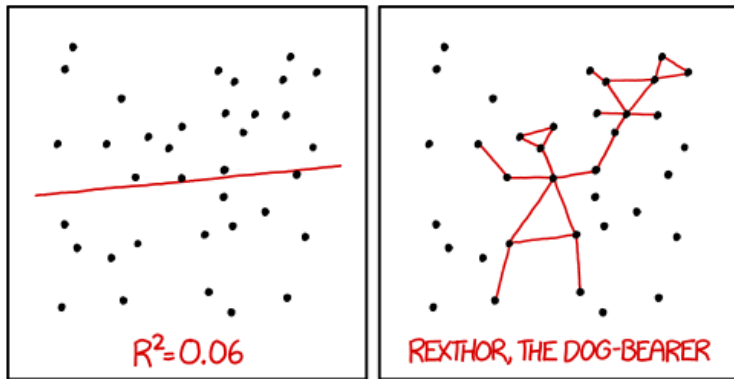
`-lm()` is the function that runs the linear model

`-tidy()` is a function in the `broom` library that cleans up the output from our linear model

`-glance()` is a function that gives us output related to the model fit - we used it to pull the R-squared value from our model

Comic Relief

From xkcd.com



I DON'T TRUST LINEAR REGRESSIONS WHEN IT'S HARDER
TO GUESS THE DIRECTION OF THE CORRELATION FROM THE
SCATTER PLOT THAN TO FIND NEW CONSTELLATIONS ON IT.