

DISCRIMINATIVE PROBABILISTIC FRAMEWORK FOR GENERALIZED MULTI-INSTANCE LEARNING

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①. Compute α . Recall that $\alpha_i(k) \triangleq p(n_{bi} = k | \mathbf{X}_b, \mathbf{w}')$. Consequently, $\alpha_{i+1}(k)$ can be computed by marginalizing $p(n_{b(i+1)} = k, n_{bi} = l, y_{b(i+1)} = c | \mathbf{X}_b, \mathbf{w}')$ over $y_{b(i+1)}$ and n_{bi} as follows

$$\begin{aligned} \alpha_{i+1}(k) &= p(n_{b(i+1)} = k | \mathbf{X}_b, \mathbf{w}') \\ &= \sum_{l=0}^{n_b} \sum_{c=0}^1 p(n_{b(i+1)} = k, n_{bi} = l, y_{b(i+1)} = c | \mathbf{X}_b, \mathbf{w}') \end{aligned}$$

Using conditional rule and the independent assumption among instance labels we obtain

$$\begin{aligned} \alpha_{i+1}(k) &= \sum_{l=0}^{n_b} \sum_{c=0}^1 p(n_{b(i+1)} = k | n_{bi} = l, y_{b(i+1)} = c) \\ &\quad p(n_{bi} = l | \mathbf{X}_b, \mathbf{w}') p(y_{b(i+1)} = c | \mathbf{x}_{b(i+1)}, \mathbf{w}') \end{aligned}$$

Replace $p(n_{b(i+1)} = k | n_{bi} = l, y_{b(i+1)} = c) = \mathbb{I}[k = l + c]$ we obtain the result for $\alpha_{i+1}(k)$ as follows

$$\begin{aligned} \alpha_{i+1}(k) &= p(y_{b(i+1)} = 1 | \mathbf{x}_{b(i+1)}, \mathbf{w}') \alpha_i(k-1) \\ &\quad + p(y_{b(i+1)} = 0 | \mathbf{x}_{b(i+1)}, \mathbf{w}') \alpha_i(k). \end{aligned} \quad (1)$$

The intuition behind above equation is that: there are two cases that there are k positive instances in the first $(i+1)$ instances. First, the $(i+1)$ th instance is positive and there are $k-1$ positive instances in the first i instances. Second, the $(i+1)$ th instance is negative and there are k positive instances in the first i instances.

②. Compute β . Recall that $\beta_i(k) \triangleq p(Y_b | n_{bi} = k, \mathbf{X}_b, \mathbf{w}')$. Consequently, $\beta_i(k)$ can be computed by marginalizing $p(Y_b | n_{bi} = k, \mathbf{X}_b, \mathbf{w}')$ over $y_{b(i+1)}$ and $n_{b(i+1)}$ as follows

$$\begin{aligned} \beta_i(k) &= p(Y_b | n_{bi} = k, \mathbf{X}_b, \mathbf{w}') \\ &= \sum_{l=0}^{n_b} \sum_{c=0}^1 p(Y_b, n_{b(i+1)} = l, y_{b(i+1)} = c | n_{bi} = k, \mathbf{X}_b, \mathbf{w}') \end{aligned}$$

Using conditional rule and the independent assumption among instance labels we obtain

$$\begin{aligned} \beta_i(k) &= \sum_{l=0}^{n_b} \sum_{c=0}^1 p(Y_b | n_{b(i+1)} = l, \mathbf{X}_b, \mathbf{w}') \\ &\quad p(n_{b(i+1)} = l | y_{b(i+1)} = c, n_{bi} = k) p(y_{b(i+1)} = c | \mathbf{x}_{b(i+1)}, \mathbf{w}') \end{aligned}$$

Replace $p(n_{b(i+1)} = l | y_{b(i+1)} = c, n_{bi} = k) = 1$ if $l = k + c$ we obtain the result for $\beta_i(k)$ as follows

$$\begin{aligned} \beta_i(k) &= p(y_{b(i+1)} = 1 | \mathbf{x}_{b(i+1)}, \mathbf{w}') \beta_{i+1}(k+1) \\ &\quad + p(y_{b(i+1)} = 0 | \mathbf{x}_{b(i+1)}, \mathbf{w}') \beta_{i+1}(k). \end{aligned} \quad (2)$$

③. Compute $p(y_{bi} = 1 | Y_b, \mathbf{X}_b, \mathbf{w}', \mathbf{v}')$. Using conditional probability we obtain

$$\begin{aligned} p(y_{bi} = 1 | Y_b, \mathbf{X}_b, \mathbf{w}', \mathbf{v}') &= \frac{p(y_{bi} = 1, Y_b | \mathbf{X}_b, \mathbf{w}', \mathbf{v}')}{p(y_{bi} = 0, Y_b | \mathbf{X}_b, \mathbf{w}', \mathbf{v}') + p(y_{bi} = 1, Y_b | \mathbf{X}_b, \mathbf{w}', \mathbf{v}')} \end{aligned} \quad (3)$$

First, we obtain $p(y_{bi} = 1, Y_b | \mathbf{X}_b, \mathbf{w}', \mathbf{v}')$ by marginalizing over n_{bi} and $n_{b(i-1)}$ as follows

$$\begin{aligned} p(y_{bi} = 1, Y_b | \mathbf{X}_b, \mathbf{w}', \mathbf{v}') &= \sum_{l=0}^{n_b} \sum_{k=0}^{n_b} p(y_{bi} = 1, Y_b, n_{bi} = l, n_{b(i-1)} = k | \mathbf{X}_b, \mathbf{w}', \mathbf{v}') \end{aligned}$$

Using the conditional rule and the independent assumption among instance labels we obtain

$$\begin{aligned} p(y_{bi} = 1, Y_b | \mathbf{X}_b, \mathbf{w}', \mathbf{v}') &= \sum_{l=0}^{n_b} \sum_{k=0}^{n_b} p(Y_b | n_{bi} = l, \mathbf{X}_b, \mathbf{w}') p(n_{bi} = l | n_{b(i-1)} = k, y_{bi} = 1) \\ &\quad p(n_{b(i-1)} = k | \mathbf{X}_b, \mathbf{w}') p(y_{bi} = 1 | \mathbf{x}_{bi}, \mathbf{w}'). \end{aligned}$$

Using the fact that $p(n_{bi} = l | n_{b(i-1)} = k, y_{bi} = 1) = \mathbb{I}[l = k + 1]$, we obtain

$$\begin{aligned} p(y_{bi} = 1, Y_b | \mathbf{X}_b, \mathbf{w}', \mathbf{v}') &= \sum_{k=0}^{n_b} p(Y_b | n_{bi} = k+1, \mathbf{X}_b, \mathbf{w}') \times \\ &\quad p(n_{b(i-1)} = k | \mathbf{X}_b, \mathbf{w}') p(y_{bi} = 1 | \mathbf{x}_{bi}, \mathbf{w}'). \end{aligned}$$

Using the definition of α and β we obtain

$$\begin{aligned} p(y_{bi} = 1, Y_b | \mathbf{X}_b, \mathbf{w}', \mathbf{v}') &= \sum_{k=0}^{n_b} \beta_i(k+1) \alpha_{i-1}(k) p(y_{bi} = 1 | \mathbf{x}_{bi}, \mathbf{w}'). \end{aligned} \quad (4)$$

Using similar proof technique, we obtain

$$\begin{aligned} p(y_{bi} = 0, Y_b | \mathbf{X}_b, \mathbf{w}', \mathbf{v}') &= \sum_{k=0}^{n_b} \beta_i(k) \alpha_{i-1}(k) p(y_{bi} = 0 | \mathbf{x}_{bi}, \mathbf{w}'). \end{aligned} \quad (5)$$

Then, we obtain the instance membership probability $p(y_{bi} = 1|Y_b, \mathbf{X}_b, \mathbf{w}', \mathbf{v}')$ as follows

$$p(y_{bi} = 1|Y_b, \mathbf{X}_b, \mathbf{w}', \mathbf{v}') = \frac{\pi_1}{\pi_1 + \pi_0} \quad (6)$$

where

$$\begin{aligned} \pi_1 &= \sum_{k=0}^{n_b} \alpha_{i-1}(k) \beta_i(k+1) p(y_{bi} = 1|\mathbf{x}_{bi}, \mathbf{w}'), \\ \pi_0 &= \sum_{k=0}^{n_b} \alpha_{i-1}(k) \beta_i(k) p(y_{bi} = 0|\mathbf{x}_{bi}, \mathbf{w}'). \end{aligned}$$

④. Compute $p(n_{bi} = n|Y_b, \mathbf{X}_b, \mathbf{w}', \mathbf{v}')$. Using conditional probability we obtain

$$p(n_{bi} = n|Y_b, \mathbf{X}_b, \mathbf{w}', \mathbf{v}') = \frac{p(n_{bi} = n, Y_b|\mathbf{X}_b, \mathbf{w}', \mathbf{v}')}{\sum_{k=0}^{n_b} p(n_{bi} = k, Y_b|\mathbf{X}_b, \mathbf{w}', \mathbf{v}')} ,$$

where $p(n_{bi} = k, Y_b|\mathbf{X}_b, \mathbf{w}', \mathbf{v}')$ is computed by using the conditional rule

$$\begin{aligned} p(n_{bi} = k, Y_b|\mathbf{X}_b, \mathbf{w}', \mathbf{v}') &= p(Y_b|n_{bi} = k, \mathbf{v}') \\ &\times p(n_{bi} = k|\mathbf{X}_b, \mathbf{w}'). \end{aligned}$$

Using the definition of α and β we obtain the probability of the number of positive instances $p(n_{bi} = n|Y_b, \mathbf{X}_b, \mathbf{w}', \mathbf{v}')$ as follows

$$p(n_{bi} = n|Y_b, \mathbf{X}_b, \mathbf{w}', \mathbf{v}') = \frac{\alpha_i(n) \beta_i(n)}{\sum_{k=0}^{n_b} \alpha_i(k) \beta_i(k)}. \quad (7)$$