Dynamic Random Testing of Web Services: A Methodology and Evaluation

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Abstract—In recent years, Service Oriented Architecture (SOA) has been increasingly adopted to develop distributed applications in the context of the Internet. To develop reliable SOA-based applications, an important issue is how to ensure the quality of web services. In this paper, we propose a dynamic random testing (DRT) technique for web services that is an improvement over the widely-practiced random testing (RT) and partition testing (PT). We examine key issues when adapting DRT to the context of SOA, including a framework, guidelines for parameter settings, and a prototype for such an adaptation. Empirical studies are reported where DRT is used to test three real-life web services, and mutation analysis is employed to measure the effectiveness. Our experimental results show that, compared with the two baseline techniques, RT and Random Partition Testing (RPT), DRT demonstrates higher fault-detection effectiveness with a lower test case selection overhead. Furthermore, the theoretical guidelines of parameter setting for DRT are confirmed to be effective. The proposed DRT and the prototype provide an effective and efficient approach for testing web services.

Index Terms—Software Testing, Random Testing, Dynamic Random Testing, Web Service, Service Oriented Architecture.

1 Introduction

C ERVICE oriented architecture (SOA) [2] defines a loose-Iy coupled, standards-based, service-oriented application development paradigm in the context of the Internet. Within SOA, three key roles are defined: service providers (who develop and own services); service requestors (who consume or invoke services); and a service registry (that registers services from providers and returns services to requestors). Applications are built upon services that present functionalities through publishing their interfaces in appropriate repositories, abstracting away from the underlying implementation. Published interfaces may be searched by other services or users, and then invoked. Web services are the realization of SOA based on open standards and infrastructures [3]. Ensuring the reliability of SOA-based applications can become critical when such applications implement important business processes.

Software testing is a practical method for ensuring the quality and reliability of software. However, some SOA features can pose challenges for the testing of web services [4], [5]. For instance, service requestors often do not have access to the source code of web services which are published and owned by another organization, and, consequently, it is not

possible to use white-box testing techniques. Testers may, therefore, naturally turn to black-box testing techniques.

Random Testing (RT) [6] is one of the most widely-practiced black-box testing techniques. Because test cases in RT are randomly selected from the input domain (which refers to the set of all possible inputs of the software under test), it can be easy to implement. Nevertheless, because RT does not make use of any information about the software under test (SUT), or the test history, it may be inefficient in some situations. In recent years, many efforts have been made to improve to RT in different ways [7]–[9]. Adaptive random testing (ART) [8], [10]–[12], for example, has been proposed to improve RT by attempting to have a more diverse distribution of test cases in the input domain.

In contrast to RT, partition testing (PT) attempts to generate test cases in a more "systematic" way, aiming to use fewer test cases to reveal more faults. When conducting PT, the input domain of the SUT is divided into disjoint partitions, with test cases then selected from each and every one. Each partition is expected to have a certain degree of homogeneity—test cases in the same partition should have similar software execution behavior. Ideally, a partition should also be homogeneous in fault detection: If one input can reveal a fault, then all other inputs in the same partition should also be able to reveal a fault.

RT and PT are based on different intuitions, and each have their own advantages and disadvantages. Because it is likely that they can be complementary to each other, detecting different faults, it is intuitively appealing to investigate the their integration. Accordingly, Cai et al. [7] have proposed the random partition testing (RPT) strategy. In RPT, the input domain is first divided into m partitions, s_1, s_2, \ldots, s_m , where each s_i is allocated a probability p_i of selection. A partition s_i is randomly selected according to the testing profile $\{\langle s_1, p_1 \rangle, \langle s_2, p_2 \rangle, \ldots, \langle s_m, p_m \rangle\}$, where

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 $p_1 + p_2 + \cdots + p_m = 1$. A concrete test case is then randomly selected from the chosen s_i .

In traditional RPT testing, the partitions and corresponding test profiles remain constant throughout testing, which may not be the best strategy. Independent researchers from various areas have observed that fault-revealing inputs tend to cluster into "continuous regions" [13], [14]—there is similarity in the execution behavior of neighboring software inputs. Based on software cybernetics, Cai et al. proposed dynamic random testing (DRT) [7], which aims to improve on both RT and RPT. Unlike the original RPT, where the values of p_i are fixed, DRT attempts to dynamically change the values: If a test case from a partition s_i reveals a fault, the corresponding p_i will be increased by a constant ε ; otherwise, it is decreased by ε .

In practice, web services have usually been tested by the service providers, and simple or easy-to-test faults have been removed, meaning that the remaining faults are normally hard to detect. For ensuring a higher reliability of the web services, a simple RT strategy may not be an appropriate technique [15], especially when the scale is large, or there are some stubborn faults. Studies have shown that DRT can improve on RT in term of fault detection effectiveness [16]–[18].

In this paper, we present a dynamic random testing (DRT) approach for web services, as an enhanced version of RT that is an adaptation of DRT to the context of SOA. We examine key issues of such an adaption, and conduct empirical studies to evaluate the feasibility and effectiveness of the proposed DRT, with the experimental results showing DRT outperforms RT in terms of fault detection efficiency. The contributions of this work include:

- We develop an effective and efficient testing technique for web services. This includes a DRT framework that addresses key issues for testing web services, and a prototype that partly automates the framework.
- We evaluate the performance of DRT through a series of empirical studies on three real web services. These studies show that DRT has significantly higher faultdetection efficiency than RT and RPT. That is, to detect a given number of faults, DRT uses less time and fewer test cases than RT and RPT.
- We provide guidelines for the DRT parameter settings, supported by theoretical analysis, and validated by the empirical studies.

The rest of this paper is organized as follows. Section 2 introduces the underlying concepts for DRT, web services and mutation analysis. Section 3 presents the DRT framework for web services, guidelines for its parameter settings, and a prototype that partially automates DRT. Section 4 describes an empirical study where the proposed DRT is used to test three real-life web services, the results of which are summarized in Section 5. Section 6 discusses related work and Section 7 concludes the paper.

2 BACKGROUND

In this section, we present some of the underlying concepts for DRT, web services, and mutation analysis.

2.1 Dynamic Random Testing (DRT)

DRT combines RT and PT [31], with the goal of benefitting from the advantages of both. Given a test suite TS classified into m partitions (denoted s_1, s_2, \ldots, s_m), suppose that a test case from s_i ($i=1,2,\ldots,m$) is selected and executed. If this test case reveals a fault, $\forall j=1,2,\ldots,m$ and $j\neq i$, we then set

$$p'_{j} = \begin{cases} p_{j} - \frac{\varepsilon}{m-1} & \text{if } p_{j} \ge \frac{\varepsilon}{m-1} \\ 0 & \text{if } p_{j} < \frac{\varepsilon}{m-1} \end{cases}, \tag{1}$$

where ε is a probability adjusting factor, and then

$$p_i' = 1 - \sum_{\substack{j=1\\j \neq i}}^{m} p_j'. \tag{2}$$

Alternatively, if the test case does not reveal a fault, we set

$$p_i' = \begin{cases} p_i - \varepsilon & \text{if } p_i \ge \varepsilon \\ 0 & \text{if } p_i < \varepsilon \end{cases}, \tag{3}$$

and then for $\forall j = 1, 2, \dots, m$ and $j \neq i$, we set

$$p'_{j} = \begin{cases} p_{j} + \frac{\varepsilon}{m-1} & \text{if } p_{i} \ge \varepsilon \\ p_{j} + \frac{p'_{i}}{m-1} & \text{if } p_{i} < \varepsilon \end{cases}$$
 (4)

The detailed DRT algorithm is given in Algorithm 1. In DRT, the first test case is taken from a partition that has been randomly selected according to the initial probability profile $\{p_1, p_2, \ldots, p_m\}$ (Lines 2 and 3 in Algorithm 1). After each test case execution, the testing profile $\{\langle s_1, p_1 \rangle, \langle s_2, p_2 \rangle, \ldots, \langle s_m, p_m \rangle\}$ is updated by changing the values of p_i : If a fault is revealed, Formulas 1 and 2 are used; otherwise, Formulas 3 and 4 are used. The updated testing profile is then used to guide the random selection of the next test case (Line 8). This process is repeated until a termination condition is satisfied (Line 1). Examples of possible termination conditions include: "testing resources have been exhausted"; "a certain number of test cases have been executed"; and "a certain number of faults have been detected".

Algorithm 1 DRT

Input: $\varepsilon, p_1, p_2, \ldots, p_m$

- 1: while termination condition is not satisfied
- 2: Select a partition s_i according to the testing profile $\{\langle s_1, p_1 \rangle, \langle s_2, p_2 \rangle, \dots, \langle s_m, p_m \rangle\}$.
- 3: Select a test case t from s_i .
- 4: Test the software using t.
- 5: **if** a fault is revealed by *t*
- 6: Update p_j $(j = 1, 2, ..., m \text{ and } j \neq i)$ and p_i according to Formulas 1 and 2.
- 7: **els**e
- 8: Update p_j $(j=1,2,\ldots,m \text{ and } j\neq i)$ and p_i according to Formulas 3 and 4.
- 9: **end_if**
- 10: end_while

As can be seen from Formulas 1 to 4, updating the testing profile involves m simple calculations, thus requiring a constant time. Furthermore, the selection of partition s_i , and subsequent selection and execution of the test case, all also involve a constant time overhead. The execution time for one iteration of DRT is thus a constant, and therefore the overall time complexity for DRT to select n test cases is $O(m \cdot n)$.

2.2 Web Services

A web service is a platform-independent, loosely coupled, self-contained, programmable, web-enabled application that can be described, published, discovered, coordinated and configured using XML artifacts for the purpose of developing distributed interoperable applications [2]. A web service consists of a description (usually specified in WSDL) and implementation (that can be written in any programming language). Web services present their functionalities through published interfaces, and are usually deployed in a service container. Invocation of a web service requires analysis of the input message in its WSDL, test data generation based on its input parameters, and wrapping of test data in a SOAP message.

A web service is a basic component of SOA software, and, accordingly, the reliability of such SOA software depends heavily on the quality of the component web services. While testing is an obvious potential activity to help assuring the quality of web services, due to the unique features of SOA, web service testing can be more challenging than traditional software testing [5]. Some of these features include:

- Lack of access to service implementation: Normally, web service owners will not make source code of their web services accessible. Typically, service users only have access to the service interface defined in a WSDL file, which means that white-box testing approaches are not possible.
- Incomplete documentation or specification: A service provider may only offer an incomplete or inaccurate description of a service's functional and nonfunctional behavior. This makes it difficult to decide whether or not a test passes, especially when details about behavior or restrictions on implementations are missing [19].
- Lack of control: Unlike traditional software testing where testers can control the execution of software under test, there is usually no opportunity to intervene in the execution of the web service under test, which is often deployed in a remote service container.
- Side effects caused by testing: A large number of tests may introduce an additional communication load, and hence impact on the performance of the web service under test. This suggests that the number of tests should be kept as low as possible [20].

Although RT is a widely-used software testing method, some of its characteristics may make it inefficient for testing web services. This study explores the application of DRT to web services with an aim of improving on the fault detection efficiency of RT.

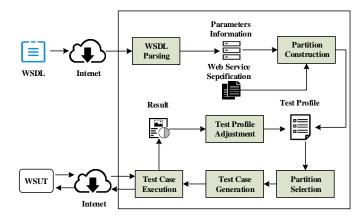


Fig. 1. DRT for web services framework

2.3 Mutation Analysis

Mutation analysis [12], [21]–[23] has been widely used to assess the adequacy of test suites and the effectiveness of testing techniques. Mutation operators are used to seed various faults into the program under test, and thus generate a set of variants, called mutants. If a test case causes a mutant to behave differently to the program under test (for example, by giving different output for the same input), then we say that this test case "kills" the mutant, and thus detects the injected fault. The mutation score (MS) is used to measure how thoroughly a test suite "kills" the mutants. The MS is defined as:

$$MS(p,ts) = \frac{N_k}{N_m - N_e} \tag{5}$$

where p is the program being mutated; ts is the test suite under evaluation; N_k is the number of mutants killed; N_m is the total number of mutants; and N_e is the number of equivalent mutants (mutants whose behavior is always the same as that of p). It has been highlighted that, compared with manually seeded faults, automatically generated mutants can be more similar to real-life faults, and thus the mutant score is a good indicator of the effectiveness of a testing technique [24]. In this paper, we use mutation analysis to evaluate the effectiveness of our proposed DRT for web services.

3 DRT FOR WEB SERVICES

In this section, we describe a framework for applying DRT to web services, discuss guidelines for DRT's parameter settings, and present a prototype that partially automates DRT for web services.

3.1 Framework

Considering the principle of DRT and the features of web services, we propose a DRT for web services framework, as illustrated in Figure 1. In the figure, the DRT components are inside the box, and the web services under test are located outside. Interactions between DRT components and the web services are depicted in the framework. We next discuss the individual framework components.

- 1 WSDL Parsing. Web services are composed of services and the relevant WSDL documents. By parsing the WSDL document, we can get the input information for each operation in the services. This includes the number of parameters, their names and types, and any additional requirements that they may have.
- 2 Partition Construction. Partition testing (PT) refers to a class of testing techniques that break the input domain into a number of partitions [25]. Because DRT is a black-box testing technique, combining RT and PT, the PT approaches used are at the specification level. Various approaches and principles for achieving convenient and effective partitions have been discussed in the literature [25]-[28]. The input domain of the web service under test (WSUT) can be partitioned based on the WSUT specifications and the parsed parameters. Once partitioned, testers can assign probability distributions to the partitions as an initial testing profile. This initial testing profile can be assigned in different ways, including using a uniform probability distribution, or one that sets probabilities according to the importance of the partition: For example, a partition within which faults were previously detected should be given higher priority.
- 3 *Partition Selection*. DRT randomly selects a partition according to the testing profile.
- 4 Test Case Generation. Given the selected partition s_i , a test case is then randomly and independently generated within s_i . Because the WSDL document has been parsed, generation of this test case is non difficult, and can be automated.
- 5 *Test Case Execution*. The relevant DRT component receives the generated test case, converts it into an input message, invokes the web service(s) through the SOAP protocol, and intercepts the test results (from the output message).
- 6 Test Profile Adjustment. Upon completion of each test, its pass or fail status is determined by comparing the actual and expected results (with the test passing if both are the same). The pass or fail status is then used to adjust the (partition) probability distribution accordingly. Situations where determination of the test outcome status is not possible (i.e. in the presence of the oracle problem [29]—[31]) may potentially be addressed using metamorphic testing [32].

Generally speaking, DRT test case generation is both in accordance with the probability distribution (for selection of the relevant partition), and with the principles of RT, taking advantage of the ease of RT and the effectiveness of PT. Furthermore, many of the DRT for web services framework components can be automated. To make DRT for web services more efficient and practical, we developed a prototype that will be described in Section 3.3.

3.2 Guidelines for Parameter Setting

Our previous work [1] found that the DRT performance can be influenced by the number of partitions and the parameter ε . We next explore these impacts through a theoretical analysis, which, to be mathematically tractable, has the following assumptions:

1 The failure rate θ_i of each partition s_i ($i=1,2,\ldots,m$, and m>1) is unknown, but can be estimated.

- 2 Each failure rate θ_i ($i=1,2,\ldots,m$, and m>1) remains unchanged throughout the testing process (faults are not removed after their detection).
- 3 Test cases are selected with replacement, which means that same test cases may be selected more than once.

A principle of the DRT strategy is to increase the selection probabilities (by amount ε) of partitions with larger failure rates. In addition to the impact of the parameter ε , the number of partitions also influences the speed of updating the testing profile (Formulas 1 to 4). Therefore, for a given number of partitions, we are interested in investigating what values of ε yield the best DRT performance.

Letting θ_M denote the maximum failure rate, and s_M denote partitions with that failure rate, then p_i^n denotes the probability of executing the n^{th} test case from partition s_i . As testing proceeds, the probability p_M of partition s_M being selected is expected to increase:

$$p_M^{n+1} > p_M^n \tag{6}$$

Initially, the testing profile is $\{\langle s_1,p_1^0\rangle,\langle s_2,p_2^0\rangle,\ldots,\langle s_m,p_m^0\rangle\}$, which, after n test cases have been executed, is then updated to $\{\langle s_1,p_1^n\rangle,\langle s_2,p_2^n\rangle,\ldots,\langle s_m,p_m^n\rangle\}$. During the testing process, p_i^n is increased or decreased by the value ε , which is relatively small (set to 0.05 in previous studies [16], [18]). Because the initial p_i^0 is larger than ε , and the adjustment of p_i is relatively small (Formulas 1 to 4), the following two situations are rare, and thus not considered here: $p_i < \varepsilon/(m-1)$ or $p_i < \varepsilon$ $(i=1,2,\ldots,m)$.

As part of exploring the relationship between p_i^{n+1} and p_i^n , we calculate the conditional probability, $p(i|\delta)$, of the following four situations (denoted $\delta_1, \delta_2, \delta_3$, and δ_4 , respectively):

Situation 1 (δ_1): If $t_n \notin s_i$ and a fault is detected by t_n , then $p(i|\delta_1)$ is calculated according to Formula 1:

$$p(i|\delta_1) = \sum_{i \neq j} \theta_j (p_i^n - \frac{\varepsilon}{m-1}).$$

Situation 2 (δ_2): If $t_n \in s_i$ and a fault is detected by t_n , then $p(i|\delta_2)$ is calculated according to Formula 2:

$$p(i|\delta_2) = \theta_i(p_i^n + \varepsilon).$$

Situation 3 (δ_3): If $t_n \in s_i$ and no fault is detected by t_n , then $p(i|\delta_3)$ is calculated according to Formula 3:

$$p(i|\delta_3) = (1 - \theta_i)(p_i^n - \varepsilon).$$

Situation 4 (δ_4): If $t_n \notin s_i$ and no fault is detected by t_n , then $p(i|\delta_4)$ is calculated according to Formula 4:

$$p(i|\delta_4) = \sum_{i \neq j} (1 - \theta_j)(p_i^n + \frac{\varepsilon}{m-1}).$$

Therefore, p_i^{n+1} for all cases together is:

$$\begin{split} p_i^{n+1} &= p_i^n \theta_i(p_i^n + \varepsilon) + p_i^n (1 - \theta_i)(p_i^n - \varepsilon) \\ &+ \sum_{j \neq i} p_j^n \theta_j(p_i^n - \frac{\varepsilon}{m-1}) \\ &+ \sum_{j \neq i} p_j^n (1 - \theta_j)(p_i^n + \frac{\varepsilon}{m-1}) \\ &= (p_i^n)^2 \theta_i + p_i^n \theta_i \varepsilon + (p_i^n)^2 - p_i^n \varepsilon - (p_i^n)^2 \theta_i + p_i^n \theta_i \varepsilon \\ &+ (p_i^n - \frac{\varepsilon}{m-1}) \sum_{j \neq i} p_j^n \theta_j + (p_i^n + \frac{\varepsilon}{m-1}) \sum_{j \neq i} p_j^n \\ &- (p_i^n + \frac{\varepsilon}{m-1}) \sum_{j \neq i} p_j^n \theta_j \\ &= (p_i^n)^2 + 2p_i^n \theta_i \varepsilon - p_i^n \varepsilon + (p_i^n - \frac{\varepsilon}{m-1} - p_i^n) \\ &- \frac{\varepsilon}{m-1} \sum_{j \neq i} p_j^n \theta_j + (p_i^n + \frac{\varepsilon}{m-1})(1 - p_i^n) \\ &= p_i^n + (p_i^n)^2 - (p_i^n)^2 + 2p_i^n \theta_i \varepsilon - p_i^n \varepsilon + \frac{\varepsilon}{m-1} - \\ &\frac{\varepsilon}{m-1} p_i^n - \frac{2\varepsilon}{m-1} \sum_{j \neq i} p_j^n \theta_j \\ &= p_i^n + \frac{\varepsilon}{m-1} (2p_i^n \theta_i m - p_i^n m - 2p_i^n \theta_i + 1) \\ &- \frac{2\varepsilon}{m-1} \sum_{j \neq i} p_j^n \theta_j \\ &= p_i^n + Y_i^n, \end{split}$$

where

$$Y_i^n = \frac{\varepsilon}{m-1} (2p_i^n \theta_i m - p_i^n m - 2p_i^n \theta_i + 1) - \frac{2\varepsilon}{m-1} \sum_{j \neq i} p_j^n \theta_j.$$
(8)

From Formula 8, we have:

$$\begin{split} Y_M^n - Y_i^n &= \frac{\varepsilon}{m-1} (2p_M^n \theta_M m - p_M^n m - 2p_M^n \theta_M + 1) \\ &- \frac{2\varepsilon}{m-1} \sum_{j \neq M} p_j^n \theta_j - \frac{\varepsilon}{m-1} (2p_i^n \theta_i m - p_i^n m \\ &- 2p_i^n \theta_i + 1) + \frac{2\varepsilon}{m-1} \sum_{j \neq i} p_j^n \theta_j \\ &= \frac{\varepsilon}{m-1} (2m(p_M^n \theta_M - p_i^n \theta_i) - m(p_M^n - p_i^n) - \\ &2(p_M^n \theta_M - p_i^n \theta_i)) - \sum_{j \neq M} p_j^n \theta_j + \sum_{j \neq i} p_j^n \theta_j \\ &= \frac{2\varepsilon}{m-1} (m(p_M^n \theta_M - p_i^n \theta_i) - \frac{m(p_M^n - p_i^n)}{2} - \\ &(p_M^n \theta_M - p_i^n \theta_i)) + \frac{2\varepsilon}{m-1} (p_M^n \theta_M - p_i^n \theta_i) \\ &= \frac{2\varepsilon}{m-1} (m(p_M^n \theta_M - p_i^n \theta_i) - \frac{m(p_M^n - p_i^n)}{2}). \end{split}$$

Before presenting the final guidelines, we need the following lemma.

Lemma 1. If
$$p_i^n - p_M^n > 2(p_i^n \theta_i - p_M^n \theta_M)$$
, then $p_M^{n+1} > p_M^n$.

Proof: The condition $p_i^n - p_M^n > 2(p_i^n \theta_i - p_M^n \theta_M)$ can be equivalently expressed as:

$$\frac{p_M^n - p_i^n}{2} < p_M^n \theta_M - p_i^n \theta_i. \tag{10}$$

From Formula 10, $(p_M^n \theta_M - p_i^n \theta_i) - \frac{p_M^n - p_i^n}{2} > 0$, and because $0 < \varepsilon < 1$, and m > 1, therefore:

$$\frac{2m\varepsilon}{m-1}((p_M^n\theta_M - p_i^n\theta_i) - \frac{p_M^n - p_i^n}{2}) > 0.$$
 (11)

Furthermore:

$$\frac{2\varepsilon}{m-1}(m(p_M^n\theta_M - p_i^n\theta_i) - \frac{m(p_M^n - p_i^n)}{2}) > 0.$$
 (12)

According to Formulas 12 and 9, if $p_i^n - p_M^n > 2(p_i n \theta_i - p_M^n)$

According to Formulas 12 and 2, if $p_i = p_M > 2 \sqrt{n} + 1 = 1 = 1 = 1$, therefore $p_M^n \theta_M$), then $Y_M^n - Y_i^n > 0$.

Also, because $\sum_{i=1}^m p_i^{n+1} = 1$, and $\sum_{i=1}^m p_i^n = 1$, therefore $Y_M^n > 0$, and thus $\sum_{i=1}^m Y_i^n = 0$.

According to Formula 7, $p_M^{n+1} = p_M^n + Y_M^n$. Because $Y_M^n > 0$, therefore $p_M^{n+1} > p_M^n$.

Accordingly, we can now present the following theorem that states a sufficient condition for achieving $p_M^{n+1} > p_M^n$.

that states a sufficient condition for achieving $p_M^{n+1} > p_M^n$.

Theorem 1. For failure rate $\theta_{min} = min\{\theta_1,\ldots,\theta_m\}$, $\theta_M > \theta_{min}$, if $0 < \theta_{min} < \frac{1}{2}$, the following condition is sufficient to guarantee that $p_M^{n+1} > p_M^n$:

$$\frac{2m\theta_{min}^2}{1 - 2\theta_{min}} < \varepsilon < \frac{(m-1)m\theta_{min}}{2(m+1)}.$$
 (13)

Proof: In order to guarantee $p_M^{n+1} > p_M^n$, we consider the following three situations (where $i \in \{1, 2, \dots, m\}$ and $i \neq m$

Situation 1 ($p_i^n = p_M^n$): Because $\theta_i < \theta_M$, therefore $(p_i^n \theta_i - p_M^n \theta_M) < 0.$

Therefore, $(p_i^n-p_M^n)>2(p_i^n\theta_i-p_M^n\theta_M)$. According to Lemma 1, we have $p_M^{n+1}>p_M^n$. Situation 2 $(p_i^n > p_M^n)$: According to Formula 13, we have the following:

$$\varepsilon > \frac{2m\theta_{min}^2}{1 - 2\theta_{min}}$$

Because

$$\frac{2m\theta_{min}^2}{1 - 2\theta_{min}} = \frac{\theta_{min}}{1/2m\theta_{min} - 1/m},$$

we have the following:

$$\varepsilon > \frac{\theta_{min}}{1/2m\theta_{min} - 1/m}.$$

Because $\theta_{min} < 1/2$, therefore $1/2m\theta_{min} - 1/m > 0$ and $\varepsilon(1/2m\theta-1/m)>\theta_{min}$, which gives $\varepsilon/2m\theta_{min}>$ $\theta_{min} + \varepsilon/m$.

Because $\varepsilon > 0$, and m > 1, therefore

$$\frac{1}{2\theta_{min}} > \frac{(\theta_{min} + \varepsilon/m)}{(\varepsilon/m)}.$$

 $(1/2\theta_{min})(p_i^n - p_M^n) > (p_i^n - p_M^n)(\theta_{min} + \varepsilon/m)/(\varepsilon/m)$ as $p_i^n > p_M^n$, and

$$p_i^n - p_M^n > 2\theta_{min}(p_i^n - p_M^n)\frac{\theta_{min} + \varepsilon/m}{\varepsilon/m}.$$

Because $(\theta_{min} + \varepsilon/m)/(\varepsilon/m) > 1$, therefore

$$2\theta_{min}(p_i^n - p_M^n)\frac{\theta_{min} + \varepsilon/m}{\varepsilon/m} > 2\theta_{min}(p_i^n - p_M^n).$$

Because $\theta_{min} < \theta_M$, therefore

$$2\theta_{min}(p_i^n - p_M^n) > 2(p_i^n \theta_{min} - p_M^n \theta_M).$$

Thus,

$$p_i^n - p_M^n > 2(p_i^n \theta_{min} - p_M^n \theta_M).$$

According to Lemma 1, we have $p_M^{n+1}>p_M^n$. Situation 3 ($p_i^n< p_M^n$): For this proof, we make the assumption that $\frac{1}{2}<\theta_M<1$.

Because we have

$$\varepsilon < \frac{(m-1)m\theta_{min}}{2(m+1)}$$

and

$$\frac{(m-1)m\theta_{min}}{2(m+1)} = \frac{2m - (m+1)}{2(m+1)}m\theta_{min},$$

thus

$$\varepsilon<(\frac{m}{m+1}-\frac{1}{2})m\theta_{min}.$$

Obviously, $\varepsilon/m < (m/(m+1)-1/2)\theta_{min}$ as m>1. Furthermore, we have

$$-\frac{\varepsilon}{m} > (\frac{1}{2} - \frac{m}{m+1})\theta_{min}$$

and

$$\frac{m\theta_{min}}{m+1} - \frac{\varepsilon}{m} + \frac{2\varepsilon}{m} > \frac{\theta_{min}}{2} + \frac{2\varepsilon}{m}$$

which means that

$$\frac{m\theta_{min}}{m+1} + \frac{\varepsilon}{m} > \frac{1}{2}(\theta_{min} + \frac{4\varepsilon}{m}).$$

It follows that

$$(m\theta_{min}/(m+1) + \varepsilon/m)/(4\varepsilon/m + \theta_{min}) > 1/2$$

for any $m>1, \varepsilon>0$, and $0<\theta_{min}<1$.

Because $\frac{1}{2} < \theta_M < 1$, therefore $(m\theta_{min}/(m+1) + \varepsilon/m)/(4\varepsilon/m + \theta_{min}) > 1/2\theta_M$.

Thus, we have

$$2(p_M^n - p_i^n)\theta_M \frac{\frac{\varepsilon}{m} + \frac{m\theta_{min}}{m+1}}{\frac{4\varepsilon}{m} + \theta_{min}} > p_M^n - p_i^n$$

as $p_M^n > p_i^n$.

Because $\varepsilon/m < 4\varepsilon/m$, and $m\theta_{min}/(m+1) < \theta_{min}$, therefore

$$\frac{\frac{\varepsilon}{m} + \frac{m\theta_{min}}{m+1}}{\frac{4\varepsilon}{m} + \theta_{min}} < 1$$

and

$$2(p_M^n-p_i^n)\theta_M>2(p_M^n-p_i^n)\theta_M\frac{\frac{\varepsilon}{m}+\frac{m\theta_{min}}{m+1}}{\frac{4\varepsilon}{m}+\theta_{min}}$$

Hence we have

$$2(p_M^n - p_i^n)\theta_M > p_M^n - p_i^n,$$



Fig. 2. Prototype interface

which can be equivalently expressed as

$$p_i^n - p_M^n > 2(p_i^n - p_M^n)\theta_M.$$

Because $\theta_{min} < \theta_M$, therefore $2(p_i^n - p_M^n)\theta_M > 2(p_i^n\theta_{min} - p_M^n\theta_M)$, and thus

$$p_i^n - p_M^n > 2(p_i^n \theta_{min} - p_M^n \theta_M).$$

According to Lemma 1, we have $p_M^{n+1}>p_M^n$. $\hfill\Box$ In summary, when $\frac{1}{2}<\theta_M<1$, there is always an interval E:

$$\varepsilon \in \left(\frac{2m\theta_{min}^2}{1 - 2\theta_{min}}, \frac{(m-1)m\theta_{min}}{2(m+1)}\right) \tag{14}$$

where $\theta_{min} \leq \theta_i, i \in \{1, 2, ..., m\}$, and $\theta_i \neq 0$, which can guarantee $p_M^{n+1} > p_M^n$.

From the proof above, it is clear that the value of θ_M affects the upper bound (E_{upper}) of E. When $\theta_{min} < \theta_M < \frac{1}{2}$, the value of E_{upper} should close to the lower bound of E. In practice, we should set

$$\varepsilon \approx \frac{2m\theta_{min}^2}{1 - 2\theta_{min}}.$$
 (15)

3.3 Prototype

Figure 2 shows a screenshot of a prototype tool that partially automates DRT for web services. To start, testers input the address of the web service being tested (the URL of the WSDL), and press the Parse button to analyze the input and output formats. Next, an operation is selected from the operation list (in the bottom left). The tool provides two options for the partitions and test suites: either to manually specify the partitions (and test cases); or to upload the predefined partitions and test suites. Before beginning testing (by pressing the Test button), testers must set the maximum number of tests (Test Repetition Limit). During the testing, if a failure is detected before having executed the maximum number of tests, then the tool suspends testing and asks for the tester's instruction. Testers can choose to remove defects and continue testing, or to stop testing. When all tests have completed, the test report is summarized and output in a file.

4 EMPIRICAL STUDY

We conducted a series of empirical studies to evaluate the performance of DRT.

TABLE 1 Studied Web Services

Web service	LOC	Number of mutants
ACMS	116	3
CUBS	131	11
PBS	129	4

4.1 Research Questions

In our experiments, we focused on addressing the following three research questions:

RQ1 How effective is DRT at detecting web service faults?

Fault-detection effectiveness is a key criterion for evaluating the performance of a testing technique. In our study, we chose three commonly used, real-life web services as subject programs, and applied mutation analysis to evaluate the effectiveness.

RQ2 How do the number of partitions and the DRT parameter ε impact on the failure detection effectiveness and efficiency of DRT?

In our earlier work [1], we found that the DRT parameter ε had a significant effect on DRT efficiency, and that the optimal value of the parameter could be related to the number of partitions. The relationship between ε and the number of partitions is examined through theoretical analysis, and verified through the empirical studies. What is the actual test case generation overhead

RQ3 What is the actual test case generation overhead when using the DRT strategy?

In Section 2.1, we have showed that DRT only requires linear time to generate test case. We wish to validate this theoretical finding through empirical examination of the actual test case

generation and execution.

4.2 Subject Web Services

We selected three real-life web services as the subject programs for our study: Aviation Consignment Management Service (ACMS), China Unicom billing service (CUBS), and Parking billing service (PBS). We used mutation analysis to generate a total of 1563 mutants. After removing equivalent mutants, we then also removed mutants that were too easily detected — deleting mutants that could be detected with less than 20 randomly generated test cases. Table 1 summarizes the basic information of the used web services and their mutants. A detailed description of each web service is given in the following.

4.2.1 Aviation Consignment Management Service (ACMC)

ACMS helps airline companies check the allowance (weight) of free baggage, and the cost of additional baggage. Based on the destination, flights are categorised as either domestic or international. For international flights, the baggage allowance is greater if the passenger is a student (30kg), otherwise it is 20kg. Each aircraft offers three cabins classes from which to choose (economy, business, and first), with passengers in different classes having different allowances.

TABLE 5 Plan C

	Plan details	Month charge (CNY)					
	Tan details	$option_1$	$option_2$	$option_3$			
i.	Free calls (min)	260	380	550			
Basic	Free data (MB)	40	60	80			
<u> </u>	Free data (MB)	Domestic (including video calls)					
g	Incoming calls (CNY/min)	0.25	0.20	0.15			
Extra	Data (CNY/KB)		0.0003				
Щ	Video calls (CNY/min)		0.60				

The detailed price rules are summarized in Table 2, where $price_0$ means economy class fare.

4.2.2 China Unicom Billing Service (CUBS)

CUBS provides an interface through which customers can know how much they need to pay according to cell-phone plans, calls, and data usage. The details of several cell-phone plans are summarized in Tables 3, 4, and 5.

4.2.3 Parking Billing Service (PBS)

Consider a parking billing service that accepts the parking details for a vehicle, including the vehicle type, day of the week, discount coupon, and hours of parking. This service rounds up the parking duration to the next full hour, and then calculates the parking fee for according to the hourly rates in Table 6. If a discount voucher is presented, a 50% discount off the parking fee is applied.

To facilitate better parking management, at the time of parking, customers may provide an estimation of parking duration, in terms of three different time ranges ((0.0, 2.0], (2.0, 4.0], and (4.0, 24.0]). If the estimation and actual parked hours fall into the same time range, then the customer will receive a 40% discount; but if they are different ranges, then a 20% markup is applied. A customer may choose to either use a discount coupon, or provide an estimation of parking duration, but may not do both. (Obviously, a customer may also choose to neither provide an estimation, nor use a discount coupon.) No vehicles are allowed to remain parked for two consecutive days on a continuous basis.

4.3 Variables

4.3.1 Independent Variables

The independent variable in this study is the testing technique, DRT. RPT and RT were used as baseline techniques for comparison.

4.3.2 Dependent Variables

The dependent variable for RQ1 is the metric for evaluating the fault-detection effectiveness. Several effectiveness metrics exist, including: the P-measure [33] (the probability of at least one fault being detected by a test suite); the E-measure [34] (the expected number of faults detected by a test suite); the F-measure [35] (the expected number of test case executions required to detect the first fault); and the T-measure [36] (the expected number of test cases required to detect all faults). Since the F- and T-measures have been widely used for evaluating the fault-detection efficiency and

TABLE 2
ACMC Baggage Allowance and Pricing Rules

		Domestic fligh	its	International flights				
	First class Business class Economy class			First class	Business class	Economy class		
Carry on (kg)	5	5	5	7	7	7		
Free checked-in (kg)	40 30 20			40	30	20/30		
Additional baggage pricing (kg)		$price_0 * 1.5\%$	ó		$price_0 * 1.5\%$	ó		

TABLE 3 Plan A

	Plan details		Month charge (CNY)										
	i iaii detaiis	$option_1$	$option_2$	$option_3$	$option_4$	$option_5$	$option_6$	$option_7$	$option_8$	$option_9$	$option_{10}$	$option_{11}$	
10.	Free calls (min)	50	50	240	320	420	510	700	900	1250	1950	3000	
Bas	Free data (MB)	150	300	300	400	500	650	750	950	1300	2000	3000	
^m	Free incoming calls		Domestic (including video calls)										
ia.	Incoming calls (CNY/min)	0.25	0.20					0.15					
1 X	Data (CNY/KB)						0.0003						
田	Video calls (CNY/min)						0.60						

TABLE 4 Plan B

	Plan details		Month charge (CNY)								
	i iaii detaiis	$option_1$	$option_2$	$option_3$	$option_4$	$option_5$	$option_6$				
.c	Free calls (min)	120	200	450	680	920	1180				
Basic	Free data (MB)	40	60	80	100	120	150				
m	Free incoming calls		Domestic (including video calls)								
Ġ	Incoming calls (CNY/min)	0.25	0.20		0.	15					
Extra	Data (CNY/KB)			0.0	0003						
H	Video calls (CNY/min)	0.60									

TABLE 6 Hourly Parking Rates

	Hourly parking rates									
Actual parking hours		Weekday		Saturday and sunday						
	Motorcycle	Car: 2-door coupe	Car: others	Motorcycle	Car: others					
(0.0, 2.0]	\$4.00	\$4.50	\$5.00	\$5.00	\$6.00	\$7.00				
(2.0, 4.0]	\$5.00	\$5.50	\$6.00	\$6.50	\$7.50	\$8.50				
(4.0, 24.0]	\$6.00	\$6.50	\$7.00	\$8.00	\$9.00	\$10.00				

effectiveness of DRT-related testing techniques [7], [9], [16]–[18], [36], they are also adopted in this study. We use F and T to represent the F-measure and the T-measure of a testing method. As shown in Algorithm 1, the testing process may not terminate after the detection of the first fault. Furthermore, because the fault detection information can lead to different probability profile adjustment mechanisms, it is also important to see what would happen after the first fault is revealed. For our study, therefore, we introduce the F2-measure [35], which is the number of additional test cases required to reveal the second fault after detection of the first fault. We use F2 to represent the F2-measure of a testing method, and $SD_{measure}$ to represent the standard deviation of metrics (where measure can be F, F2, or T).

An obvious metric for RQ3 is the time required to detect faults. Corresponding to the T-measure, in this study we used T-time, the time required to detect all faults. F-time and F2-time denote the time required to detect the first fault, and the additional time needed to detect the second fault (after detecting the first), respectively.

For each of these metrics, smaller values indicate a better performance.

4.4 Experimental Settings

4.4.1 Partitioning

In our study, we set the partitions by making use of a decision table (DT) [37]. A DT presents a large amount of complex decisions in a simple, straightforward manner, representing a set of decision rules under all exclusive conditional scenarios in a pre-defined problem. Typically, a DT consists of four parts:

- 1) The upper-left part lists the conditions denoted C_i $(i=1,\ldots,n,$ where n is the number of conditions in the pre-defined problem, and $n\geq 1$). Each condition C_i contains a set of possible options $O_{i,q}\in CO_i=\{O_{i,1},\ldots,O_{i,t_i}\}$, where t_i is the number of possible options for C_i , and $q=\{1,\ldots,t_i\}$.
- 2) The upper-right part shows the condition space, which is a Cartesian product of all the CO_i ($SP(C) = CO_1 \times CO_2 \times \ldots \times CO_n$). Each element in the SP(C) is a condition entry (CE) with the ordered n-tuple.
- 3) The lower-left part shows all possible actions, represented A_j ($j=1,\ldots,m$, where m is the number of possible actions and $m \geq 1$). Similar to CO_i , an action A_j contains a set of possible options $O'_{i,p} \in$

TABLE 7 Numbers of Partitions per Scheme for Three Subject Programs

Web service	Scheme 1	Scheme 2
ACMS	24	7
CUBS	20	3
PBS	18	3

 $AO_j = \{O'_{j,1}, \dots, O'_{j,k_j}\}$, where k_j is the number of alternatives for A_j , and $p = \{1, \dots, k_j\}$.

4) The lower-right part shows the action space SP(A), which is also a Cartesian product of all the AO_j ($SP(A) = AO_1 \times AO_2 \times \ldots \times AO_m$). Similar to CE, each element in the SP(A) is an action entry (AE) with the ordered m-tuple.

A DT *rule* is composed of a *CE* and its corresponding *AE*. With the constructed DT, it is possible to obtain partition schemes with different granularities. For fine-grain partition schemes, each *CE* of a DT *rule* corresponds to a partition; while for coarse-grained schemes, a partition corresponds to the union of a group of partitions of which all *CE* of DT *rules* have the same *AE*. Accordingly, we obtained two partition schemes for each subject web service: the numbers of partitions within these partition schemes are summarized in Table 7.

4.4.2 Initial Test Profile

Because test cases may be generated randomly during the test process, a feasible method is to use a uniform probability distribution as the initial testing profile. On the other hand, testers may also use past experience to guide a different probability distribution as the initial profile.

4.4.3 Constants

In the experiments, we were interested in exploring the relationship between the number of partitions and the DRT strategy parameter ε , and therefore selected a set of parameter values: $\varepsilon \in \{1.0E\text{-}05, 5.0E\text{-}05, 1.0E\text{-}04, 5.0E\text{-}04, 1.0E\text{-}03, 5.0E\text{-}03, 1.0E\text{-}02, 5.0E\text{-}02, 1.0E\text{-}01, 2E\text{-}01, 3E\text{-}01, 4E\text{-}01, 5E\text{-}01\}.$ It should be noted that $\varepsilon = 5E\text{-}01$ is already a large value. Consider the following scenario. For PBS, when the test is carried out under partition scheme 2, if $\varepsilon = 7.5E\text{-}01$ and a uniform probability distribution is used as the testing profile (that is, $p_i = 1/3$) , then suppose that the first test case belonging to c_1 is executed and does not reveal any faults, then, according to Formula 3, the value of p_1 would become 0. It is important, therefore, that the initial value of ε should not be set too large.

4.5 Experimental Environment

Our experiments were conducted on a virtual machine running the Ubuntu 11.06 64-bit operating system, with two CPUs, and a memory of 2GB. The test scripts were written in Java. To ensure statistically reliable values [38] of the metrics (F-measure, F2-measure, T-measure, F-time, F2-time, and T-time), each testing session was repeated 30 times with 30 different seeds, and the average value calculated.

4.6 Threats To Validity

4.6.1 Internal Validity

A threat to internal validity is related to the implementations of the testing techniques, which involved a moderate amount of programming work. However, our code was cross-checked by different individuals, and we are confident that all techniques were correctly implemented.

4.6.2 External Validity

The possible threat to external validity is related to the subject programs and seeded faults under evaluation. Although the three subject web services are not very complex, they do implement real-life business scenarios of diverse application domains. Furthermore, 18 distinct faults were used to evaluate the performance. These faults cover different types of mutation operators and require an average of more than 20 randomly generated test cases to be detected. Although we have tried to improve the generalisability of the findings by applying different partitioning granularities, and 13 kinds of parameters, we cannot be certain as to whether or not similar results would be observed in other types of web services.

4.6.3 Construct Validity

The metrics used in our study are simple in concept and straightforward to apply, and hence there should be little threat to the construct validity.

4.6.4 Conclusion Validity

As has been reported for empirical studies in the field of software engineering [38], at least 30 observations are necessary to ensure the statistical significance of results. Accordingly, we have run a sufficient number of trials to ensure the reliability of our experimental results. Furthermore, as will be discussed in Section 5, further statistical tests were also conducted to confirm their significance.

5 EXPERIMENTAL RESULTS

5.1 RQ1: Fault Detection Effectiveness

The F-, F2-, and T-measure results are summarized in Tables 8 to 10, and their distributions for each program are displayed using boxplots in Figures 3 to 5. In each boxplot, the upper and lower bounds of the box represent the third and first quartiles of the metric, respectively; the middle line represents the median value; the upper and lower whiskers mark, respectively, the largest and smallest data within the range of $\pm 1.5 \times IQR$ (where IQR is the interquartile range); outliers beyond the IQR are denoted with hollow circles; and each solid circle represents the mean value of the metric.

It can observed from the tables and figures that, in general, DRT is the best performer, followed by RPT. We also conducted statistical testing to verify the significance of this observation, using the Holm-Bonferroni method [35] (with p-value equal to 0.05) to determine which pairs of testing techniques had significant differences. The statistical data are shown in Tables 11 to 13, where each cell gives the number of scenarios where the technique above (in the table) performed better than one to the left. Where the difference is significant, the number is underlined and in

TABLE 8 ACMS Results

Str	ategy	Partition scheme 1						Partition scheme 2					
3116	ategy	F	SD_F	F2	SD_{F2}	T	SD_T	F	SD_F	F2	SD_{F2}	T	SD_T
]	RT		10.34	14.90	24.64	41.80	27.13	13.30	10.34	14.90	24.64	41.80	27.13
R	PT.	12.04	11.13	12.26	19.08	35.31	25.95	9.03	8.24	15.13	16.08	34.29	29.53
	1.0E-5	11.42	10.16	11.93	18.62	34.93	22.46	8.73	9.88	13.28	18.83	36.19	28.96
	5.0E-5	12.42	10.63	12.05	20.27	36.90	25.39	8.65	9.30	13.53	19.07	35.76	29.86
	1.0E-4	11.34	10.65	11.66	21.27	35.19	26.01	7.36	8.49	13.09	17.89	33.00	27.94
	5.0E-4	12.16	12.13	11.50	19.36	34.19	26.18	7.80	8.37	13.45	17.36	33.59	27.67
	1.0E-3	11.46	11.10	11.39	19.01	34.86	23.17	7.68	8.11	15.07	19.85	33.65	29.95
	5.0E-3	11.02	9.67	12.24	18.83	32.92	20.75	7.47	8.65	15.40	18.81	35.14	29.01
DRT	1.0E-2	10.48	9.60	9.46	14.17	29.59	19.10	7.66	9.18	14.93	18.69	34.30	30.15
	5.0E-2	8.75	6.59	7.06	10.35	23.05	11.67	7.26	7.74	14.70	18.15	34.81	28.14
	1.0E-1	8.59	6.66	6.37	9.41	21.36	10.81	6.67	7.34	16.26	17.54	34.27	27.48
	2.0E-1	8.50	6.21	5.80	9.07	22.10	10.56	5.71	6.04	16.63	10.33	33.58	30.78
	3.0E-1	9.23	6.84	7.12	10.34	22.57	11.13	5.43	6.28	17.60	10.39	33.86	30.33
	4.0E-1	9.22	7.03	6.72	9.36	22.57	10.44	5.14	5.19	17.56	19.11	33.83	28.94
	5.0E-1	8.61	6.41	7.70	10.45	22.64	10.95	5.86	6.50	16.18	17.06	33.31	27.40

TABLE 9 CUBS Results

C+m	atomy			Partiti	on scheme	e 1				Partiti	on scheme	e 2	
Sur	ategy	F	SD_F	F2	SD_{F2}	T	SD_T	F	SD_F	F2	SD_{F2}	T	SD_T
RT		21.93	20.37	38.17	38.17	4203.07	3219.40	21.93	20.37	38.17	38.17	4203.07	3219.40
R	RPT		20.36	28.74	27.56	2590.38	1768.49	23.66	21.92	27.31	26.99	4195.72	2777.89
	1.0E-5	21.21	23.24	25.62	23.29	2720.70	2051.85	22.90	22.51	25.29	28.21	4106.81	2589.29
	5.0E-5	19.50	19.94	25.77	24.78	2503.45	1873.76	23.88	23.25	26.68	26.84	4130.03	2588.36
	1.0E-4	20.71	22.51	25.59	26.00	2516.91	1843.11	23.55	22.76	26.71	30.26	4196.01	2247.57
	5.0E-4	21.79	21.10	26.82	28.21	2519.39	1942.65	25.27	28.74	26.11	24.36	4190.61	2753.74
	1.0E-3	20.95	20.79	31.41	31.54	2532.84	1752.56	23.84	25.44	27.34	27.64	4291.41	2884.39
	5.0E-3	22.32	21.90	25.93	26.48	2535.97	1572.42	24.11	23.27	26.80	25.20	4218.74	2887.01
DRT	1.0E-2	22.47	21.55	26.01	23.85	2550.88	1873.01	23.46	25.01	26.74	26.43	4117.11	2798.92
	5.0E-2	21.61	20.04	27.66	29.12	2559.56	1777.16	24.01	24.32	26.52	27.04	4105.51	2570.57
	1.0E-1	21.72	21.71	28.31	28.91	2533.08	1774.39	23.30	24.45	26.07	27.91	4271.32	3011.37
	2.0E-1	21.71	21.83	28.68	32.75	2552.29	1879.60	23.55	25.20	28.25	30.31	4170.32	2796.61
	3.0E-1	22.82	21.65	26.68	31.28	2623.00	1770.36	23.40	25.52	27.35	27.84	4138.49	2594.70
	4.0E-1	23.34	24.02	27.32	27.42	2664.34	1886.21	23.07	24.08	29.18	30.39	4192.68	2706.73
	5.0E-1	22.18	22.32	27.14	28.54	2599.40	1640.05	23.30	23.63	26.98	27.71	4195.45	2535.00

TABLE 10 PBS Results

Ctm	akaarr	Partition scheme 1						Partition scheme 2					
311	ategy	F	SD_F	F2	SD_{F2}	T	SD_T	F	SD_F	F2	SD_{F2}	T	SD_T
	RT	20.17	16.32	16.50	13.20	252.80	191.54	20.17	16.32	16.50	13.20	252.80	191.54
R	RPT	17.71	16.78	16.72	25.12	178.91	147.96	16.49	15.76	15.45	21.76	182.66	130.43
	1.0E-5	18.07	19.08	14.27	21.70	176.18	146.49	15.24	15.06	15.31	23.97	163.51	135.48
	5.0E-5	18.23	18.34	15.39	24.35	176.93	143.93	15.13	15.35	14.30	23.57	160.82	116.59
	1.0E-4	16.39	17.42	13.96	21.04	171.61	142.79	15.14	15.05	13.96	22.31	159.93	121.90
	5.0E-4	16.75	17.56	12.60	21.98	165.94	140.33	14.95	12.66	14.98	24.52	166.18	125.20
	1.0E-3	17.96	18.93	15.40	22.22	164.56	138.74	15.72	16.03	16.30	24.41	166.27	128.63
	5.0E-3	16.93	17.23	14.34	22.34	160.70	117.02	15.54	12.79	15.45	23.70	170.14	129.61
DRT	1.0E-2	17.12	16.60	15.10	22.68	166.15	136.59	15.19	15.02	15.13	25.78	166.60	124.97
	5.0E-2	17.02	18.70	16.46	24.75	168.20	129.72	17.10	17.23	15.07	24.09	170.09	130.88
	1.0E-1	17.09	16.78	14.12	22.00	172.94	153.50	16.02	17.16	15.98	24.97	174.16	134.34
	2.0E-1	17.45	18.36	14.14	22.77	174.71	139.02	15.27	15.54	15.52	23.86	167.07	132.81
	3.0E-1	17.23	18.45	16.95	27.43	178.09	161.44	15.43	15.54	15.15	22.04	175.02	136.66
	4.0E-1	17.05	17.72	16.21	20.08	169.21	145.43	15.28	15.93	15.28	24.48	164.67	124.24
	5.0E-1	17.17	18.26	16.23	25.83	172.29	134.94	16.10	15.78	15.26	23.80	167.21	124.05

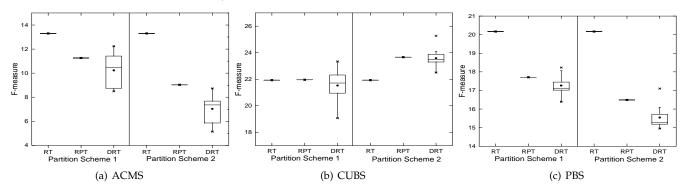


Fig. 3. F-measure boxplots for each program

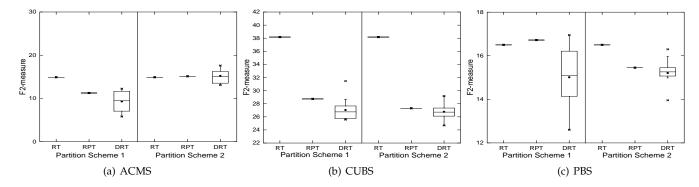


Fig. 4. F2-measure boxplots for each program

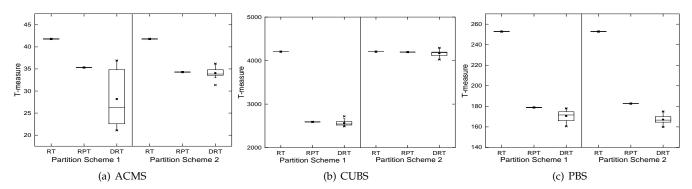


Fig. 5. T-measure boxplots for each program

bold face. For example, the $\overline{75}$ in the top right cell of Table 13 indicates that, of 78 scenarios (13 parameters \times two partition schemes \times three web services), DRT had lower F2-measure scores than RT for 75, with the fault-detection capabilities of these two techniques being significantly different.

TABLE 11 Number of Scenarios Where the Technique on the Top Row Has a Lower F-measure Score Than That on the Left Column

	RT	RPT	DRT
RT	_	4	<u>60</u>
RPT	2	_	62
DRT	18	<u>16</u>	_

Tables 11 to 13 clearly show that the difference between each pair of testing techniques is always significantly different.

TABLE 12 Number of Scenarios Where the Technique on the Top Row Has a Lower F2-measure Score Than That on the Left Column

	RT	RPT	DRT
RT	_	4	<u>69</u>
RPT	2	_	<u>63</u>
DRT	9	15	_

TABLE 13 Number of Scenarios Where the Technique on the Top Row Has a Lower T-measure Score Than That on the Left Column

	RT	RPT	DRT
RT	_	6	<u>75</u>
RPT	0	_	<u>64</u>
DRT	3	<u>14</u>	_

TABLE 14
Theoretical Optimal Values of DRT Parameter

Web	Partition	θ_{min}	ε^*
service	scheme		
ACMS	1	5.452E-2	1.601E-1
ACIVIS	2	2.797E-3	1.102E-4
CUBS	1	1.193E-3	5.702E-5
CODS	2	1.397E-3	1.734E-5
PBS	1	1.760E-3	1.118E-4
1 03	2	1.492E-3	1.340E-5

5.2 RQ2: Relationship between Partition Number and ε

In 3.2, we analyzed the relationship between the number of partitions and the DRT strategy parameter ε . In this section, we show that our theoretical analysis provides useful guidance to testers to set the value of ε .

We used three web services to validate our theoretical analysis. Before starting the test, the failure rate θ_i of partition s_i was obtained by executing (k) test cases from s_i until revealing a fault, then $\theta_i = k/k_i$, where k_i is the total number of test cases in s_i . According to Formula 19, the theoretically optimal values of ε in each scenario for each web service are shown in Table 14, where ε^* denotes the theoretical value of ε . We ran a series of experiments with the parameters set according to those in Table 14: The F-, F2-, and T-measure results for each program are shown in Figure 6, where ε_1^* and ε_2^* denote the theoretical values of parameter ε in the two different partition schemes, respectively. For ease of presentation and understanding, we used $log_{100}(1.0E05 \times \varepsilon)$ for the horizontal axis in Figure 6. Apart from the DRT strategy parameter ε , all other experimental settings remained the same as in Section 5.1.

From Figure 6, we have the following observations:

- In most scenarios, the DRT strategy with theoretically optimum parameter value performs best. Furthermore, the DRT strategy performs better when the parameter values are near the theoretically optimum value than when not.
- From Figure 6 (a), it can be observed that the DRT strategy with larger parameter values performs better than with the theoretically optimum value, in terms of the F-measure. The main reason for this is that, for this scenario, the maximum failure rate $(\theta_M=4.781E-3)$ is large and the number of partitions is small: When the parameter value is large, the probability of selecting partitions with lower failure rates is quickly reduced, and the probability of selecting partitions with larger failure rates is quickly increased, according to Formulas 3 and 4.

5.3 RQ3: Selection Overhead

Tables 15 to 17 summarize the F-, F2-, and T-time results, respectively, and their distributions for each web service is shown in Figures 7 to 9. It can be observed from the figures that, in general, DRT had the best performance, and RPT just marginally outperforms RT.

As was done for the F-, F2-, and T-measure data, we used the Holm-Bonferroni method to check the difference between each pair of testing strategies in terms of F-time, F2-time, and T-time, as shown in Tables 18 to 20.

TABLE 15 F-time, F2-time and T-time for Web Service ACMS (in ms)

Strategy		Parti	tion sche	me 1	Partition scheme 2		
3116	negy	F-time	F-time F2-time		F-time	F2-time	T-time
	RT	0.43	0.29	0.85	0.43	0.29	0.85
R	.PT	0.57	0.31	1.08	0.33	0.45	0.78
	1.0E-5	0.47	0.28	0.96	0.38	0.41	0.95
	5.0E-5	0.35	0.21	0.78	0.30	0.27	0.71
	1.0E-4	0.33	0.19	0.66	0.24	0.28	0.65
	5.0E-4	0.30	0.17	0.60	0.27	0.25	0.67
	1.0E-3	0.29	0.19	0.61	0.22	0.24	0.60
	5.0E-3	0.28	0.23	0.62	0.21	0.31	0.64
DRT	1.0E-2	0.24	0.17	0.53	0.23	0.29	0.66
	5.0E-2	0.28	0.13	0.49	0.23	0.35	0.71
	1.0E-1	0.26	0.08	0.43	0.23	0.29	0.63
	2.0E-1	0.23	0.12	0.43	0.30	0.46	0.89
	3.0E-1	0.24	0.11	0.45	0.18	0.40	0.69
	4.0E-1	0.27	0.10	0.47	0.16	0.31	0.58
	5.0E-1	0.28	0.11	0.47	0.28	0.45	0.86

TABLE 16 F-time, F2-time and T-time for Web Service CUBS (in ms)

	Partition scheme 1 Partition scheme 2						2
Strategy					Partition scheme 2		
Julia	F-time		F2-time	T-time	F-time	F2-time	T-time
1	RT	0.82	1.14	34.69	0.82	1.14	34.69
R	PT	0.91	0.87	30.54	0.75	0.79	34.59
	1.0E-5	0.91	0.87	30.14	0.87	0.83	36.49
	5.0E-5	0.95	0.86	30.21	0.82	0.75	34.70
	1.0E-4	0.77	0.89	29.27	0.86	0.69	35.34
	5.0E-4	0.79	1.00	31.05	0.87	0.90	36.85
	1.0E-3	0.86	0.88	29.93	0.92	0.77	36.44
	5.0E-3	0.86	0.83	30.28	0.83	0.83	36.27
DRT	1.0E-2	0.95	0.88	29.33	0.84	0.72	35.23
	5.0E-2	0.84	0.82	30.56	0.88	0.72	37.07
	1.0E-1	0.88	0.93	29.45	0.82	0.76	37.14
	2.0E-1	0.83	0.83	30.16	0.86	0.73	36.79
	3.0E-1	0.95	0.86	27.81	0.84	0.84	36.89
	4.0E-1	0.84	0.84	28.23	0.86	0.82	35.22
	5.0E-1	0.87	0.84	27.91	0.90	0.86	38.21

In Table 20, four entries ("61" & "17" for DRT vs. RT, "57" vs. "21" for DRT vs. RPT) are in bold font and underline, meaning that, in terms of T-time, DRT was significantly better than RT, and DRT only marginally outperformed RPT. Similar observations can be made regarding the F-time and F2-time results. In other words, the additional computation incurred in DRT by updating the test profile is compensated

TABLE 17
F-time, F2-time and T-time for Web Service PBS (in ms)

Strategy		Parti	tion sche	me 1	Partition scheme 2		
Sur	ategy	F-time	F2-time	T-time	F-time	F2-time	T-time
1	RT	0.81	0.42	4.12	0.81	0.42	4.12
R	PT	0.85	0.52	3.83	0.66	0.35	2.98
	1.0E-5	0.76	0.48	3.85	1.05	0.51	3.44
	5.0E-5	0.73	0.35	3.02	0.91	0.52	3.79
	1.0E-4	0.54	0.36	2.97	0.49	0.34	2.26
	5.0E-4	0.68	0.34	3.20	0.50	0.30	2.31
	1.0E-3	0.62	0.38	2.87	0.52	0.38	2.42
	5.0E-3	0.60	0.38	2.99	0.43	0.51	2.40
DRT	1.0E-2	0.58	0.42	2.84	0.51	0.34	2.41
	5.0E-2	0.64	0.44	3.27	0.59	0.36	2.49
	1.0E-1	0.65	0.38	3.46	0.60	0.38	2.56
	2.0E-1	0.60	0.36	3.22	0.54	0.59	2.47
	3.0E-1	0.59	0.44	3.24	0.57	0.46	2.57
	4.0E-1	0.61	0.33	2.90	0.51	0.33	2.54
	5.0E-1	0.62	0.44	3.16	0.56	0.36	2.56

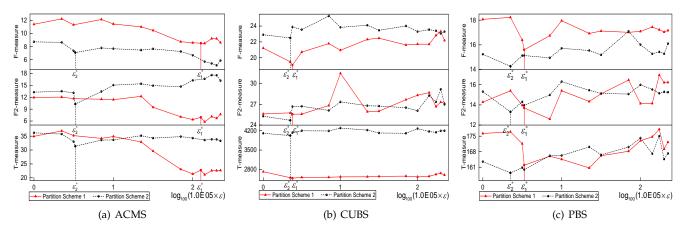


Fig. 6. Line charts of F-measure, F2-measure, and T-measure values for each program (for both the theoretically optimum parameter value, and other values

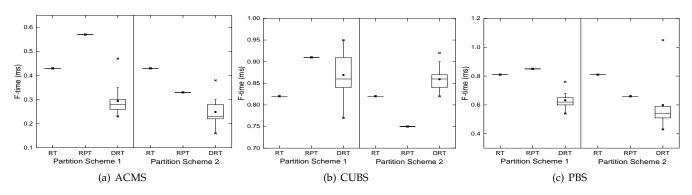


Fig. 7. F-time boxplots for each program

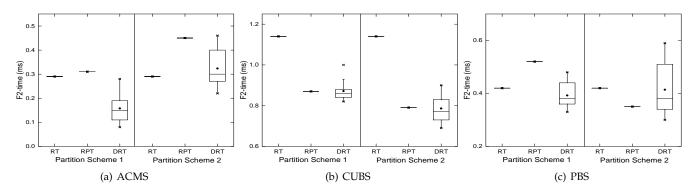


Fig. 8. F2-time boxplots for each program

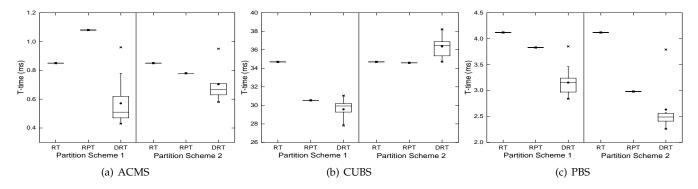


Fig. 9. T-time boxplots for each program

for in terms of test execution savings.

TABLE 18 Number of Scenarios Where The Technique on the Top Row Has a Lower F-time Than That on the Left Column

	RT	RPT	DRT
RT		5	<u>53</u>
RPT	1	_	<u>59</u>
DRT	<u>25</u>	<u>19</u>	_

TABLE 19 Number of Scenarios Where the Technique on the Top Row Has a Lower F2-time Than That on the Left Column

	RT	RPT	DRT
RT	_	3	<u>62</u>
RPT	3	_	57
DRT	16	21	_

TABLE 20 Number of Scenarios Where the Technique on the Top Row Has a Lower T-time Than That on the Left Column

	RT	RPT	DRT
RT	_	5	<u>61</u>
RPT	1	_	<u>57</u>
DRT	<u>17</u>	<u>21</u>	_

It can also be observed from Tables 18 to 20 that DRT only slightly outperformed RPT, but that DRT was significantly better than RT, especially in term of *T-time*.

In summary, the DRT strategy is considered the best testing technique across all six metrics, and RPT marginally outperformed RT.

6 RELATED WORK

In this section, we describe related work from two perspectives: related to testing techniques for web services; and related to improving RT and PT.

6.1 Testing Techniques for Web Services

In recent years, a lot of effort has been made to test web services [5], [15], [39], [40]. Test case generation, involving both the generation and selection of test cases, is core to testing web services, and model-based [41] and specificationbased [42] techniques are two common approaches. Before making services available on the Internet, testers can use model-based techniques to verify whether or not the behavior of the WSUT meets their requirements. In these techniques, test data can be generated from a data model that specifies the inputs to the software—this data model can be built before, or in parallel to, the software development process. Verification methods using technologies such as theorem-proving [43], models [44] and Petri-Nets [45] also exist. Sinha et al. [43], for example, used theorem-proving to generate test cases, making use of existing test generation methods based on extended finite state machine (EFSM) specifications. Paradkar et al. [44] proposed a model-based test data generation technique for semantic web services,

where test cases are generated using pre-defined faultmodels and the IOPE (Inputs, Outputs, Preconditions, Effects paradigm) information from semantic specifications. Dong et al. [46] proposed a web service testing technique based on fault coverage, in which a High-level Petri Net (HPN) is first constructed by analyzing the WSDL specification for each operation, and HPNs for all operations are merged into one for the service. Then, a UIO (Unique Input Output) sequence is generated based on the resulting HPN through graph transformation. Finally, a test case for each UIO sequence is generated to meet the fault-coverage requirements. When testing web services, because it is often only the service's specification that users can receive, specification-based testing is a natural choice. Typically, the web service specification is contained in the web service description language (WSDL) document, which provides information about the available operations and parameters. Many methods proposed for WSDL-based test data generation are based on the XML schema data type. Hanna and Munro proposed a framework that can be used to test the robustness of a web service [47]. Their framework analyzes the WSDL documents of web services to identify what faults could impact the robustness, facilitating the design of test cases to detect those faults.

The approaches listed above all aim to generate test cases without consideration of the impact of test case execution order on test efficiency. In contrast, Bertolino et al. [48] proposed using the category-partition method [49] with XML schemas to perform XML-based partition testing. Because PT aims to find subsets of all possible test cases to adequately test a system, it can help reduce the required number of test cases. Our proposed approach involved using software cybernetics with PT: In DRT, selection of a partition is done according to the testing profile, which is updated throughout the test process. An advantage of DRT is that partitions with larger failure rates have higher probabilities of selection. Zhu and Zhang [50] proposed a collaborative testing framework, where test tasks are completed using collaborating test services—a test service is a service assigned to perform a specific testing task. Our framework (Section 3.1) aims to find more faults in the WSUT, with the result of the current test case execution providing feedback to the control system so that the next test case selected has a greater chance to reveal faults.

Most existing web service testing techniques assume that the computed output for each test case is verifiable, something that is not always true in practice. The oracle problem [30], [31] refers to those situations where the test case output is not verifiable, and has meant that many testing techniques may not be applicable in some situations. To address the outstanding oracle problem for testing web services, a metamorphic testing [51], [52] technique has been proposed that not only alleviates the oracle problem, but also presents a feasible and efficient option for testing web services. Sun et al. proposed a metamorphic testing framework for web services [32] and conducted a case study that showed that up to 94.1% of seeded faults could be detected without the need for oracles.

6.2 Improving RT and PT

Based on the observation that failure-causing inputs tend to cluster into contiguous regions within the input domain [13], [14], significant work has been done aiming to improve RT [7], [8]. Adaptive random testing [8] is a family of advanced techniques based on random testing that aim to improve the failure-detection effectiveness by evenly spreading test cases throughout the input domain. One well-known ART approach, FSCS-ART, selects a next test input from the fixed-size candidate set of tests that is farthest from all previously executed tests [53]. Many other ART algorithms have also been proposed, including RRT [54], [55], DF-FSCS [22], and ARTsum [56], with their effectiveness examined and validated through simulations and experiments.

Adaptive testing (AT) [9], [57], [58] takes advantage of feedback information to control the execution process, and has been shown to outperforms RT and RPT in terms of the T-measure and the number of detected faults, which means that AT has higher efficiency and effectiveness than RT and RPT. However, AT may require a very long execution time in practice. To alleviate this, Cai et al. [7] proposed DRT, which uses testing information to dynamically adjust the testing profile. There are several things that can impact on DRT's test efficiency. Yang et al. [17] proposed A-DRT, which adjusts parameters during the testing process. Li et al. [18] developed O-DRT, which uses an objective function and a pre-defined parameter f which, when exceeded during the testing process (by the objective function), results in the test profile being adjusted to a theoretically optimal one. Lv et al. [16] introduced two parameters for adjusting the selection probability of different partitions: ε adjusts the probability of partitions where a test case detects a fault; and δ adjusts the probability of partitions where a test case does not detect a fault. (ε should be greater than δ). Furthermore, they recommended setting the ε and δ values to have an interval of ε/δ . However, it is unclear how to set ε and δ individually. In this study, we only considered one parameter (i.e. ε) in the DRT algorithm for adjusting the probability of different partitions, and provided detailed settings guidance.

7 CONCLUSION

In this paper, to address the challenges of testing SOA-based applications, we have presented a dynamic random testing (DRT) technique for web services. Our technique uses random testing to generate test cases, and selects test cases for execution from different partitions in accordance with a testing profile that is dynamically updated in response to the test data collected. In this way, the proposed test technique includes benefits form both random testing and partition testing.

We proposed a framework that examines key issues when applying DRT to testing web services, and developed a prototype to make the method feasible and effective. To help guide testers to correctly set the DRT parameters, we used a theoretical analysis to identify the relationships between the number of partitions (m) and the probability adjusting factor (ε) . Three real web services were used as experimental subjects to validate the feasibility and effectiveness of our approach. The results of the empirical study

show that, in general, DRT has better performance than both RT and RPT, in terms of the F-, F2-, and T-measures, and always outperforms when the ε settings follow our guidelines. In other words, our theoretical analysis can provide genuinely useful guidance when DRT is used.

In our future work, we plan to conduct experiments on more web services to further validate its effectiveness, and identify the limitations of our method.

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