



Measurement of the branching fraction for the decay $B_s^0 \rightarrow D_s K \pi \pi$

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Abstract

We present the measurement of the branching fraction of decay $B_s^0 \rightarrow D_s K \pi \pi$ using the complete 3 fb^{-1} of data, collected during Run 1 of the LHC. The branching fraction is measured relative to the decay $B_s^0 \rightarrow D_s \pi \pi \pi$, from which we obtain

$$\frac{\mathcal{B}(B_s^0 \rightarrow D_s K \pi \pi)}{\mathcal{B}(B_s^0 \rightarrow D_s \pi \pi \pi)} = xx \pm xx \pm xx$$

The $B_s^0 \rightarrow D_s K \pi \pi$ decay can be further used to measure the weak CKM phase γ in a time-dependent analysis of the B_s^0 and \bar{B}_s^0 decay rates.

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₁ **0 To do list**

- ₂ 1. Efficiencies for 2012. Need MC12 for normalization channel?
- ₃ 2. Check BDT performance on data
- ₄ 3. Systematics

1 Introduction

The weak phase γ is the least well known angle of the CKM unitary triangle. A key channel to measure γ is the time-dependent analysis of $B_s^0 \rightarrow D_s K$ decays [REF HERE]. The measurement of γ presented in this note uses $B_s^0 \rightarrow D_s K \pi \pi$ decays, where the $K \pi \pi$ subsystem is dominated by excited kaon states, such as the $K_1(1270)$ and $K_1(1400)$ resonances. It is complementary to the above mentioned analysis of $B_s^0 \rightarrow D_s K$, making use of a fully charged final state, where every track is detected in the vertex locator. To account for the non-constant strong phase across the Dalitz plot, one can either develop a time-dependent amplitude model or select a suitable phase-space region and introduce a coherence factor as additional hadronic parameter to the fit. This analysis is based on the first observation of the $B_s^0 \rightarrow D_s K \pi \pi$ decay presented in [1] and [2], where its branching ratio is measured relative to $B_s^0 \rightarrow D_s \pi \pi \pi$. The branching ratio measurement is updated, exploiting the full Run 1 data sample, corresponding to 3 fb^{-1} of integrated luminosity.

2 Data samples

We use the full Run 1 sample from Stripping 21, consisting of 3 fb^{-1} of data, collected in the years 2011 and 2012 at a center of mass energies of 7 TeV and 8 TeV, respectively. The selected B_s^0 -candidates are required to pass the L0 Hadron trigger on signal (TOS) or the L0 Global trigger independent of signal (TIS). Events that pass the L0 stage are further required to pass the HLT1 TrackAllL0 trigger on signal (TOS). All remaining candidates have to pass either the 2, 3 or 4-body topological trigger (TOS) of the HLT2 stage. For the presented analysis the B02DKPiPiD2HHHPIDBeauty2CharmLine is used to preselect signal $B_s^0 \rightarrow D_s K \pi \pi$ candidates. A summary of the cuts employed by this stripping line can be found in Table 2.1. In this table and throughout the note, we abbreviate $B_s^0 \rightarrow D_s X_s (\rightarrow K \pi \pi)$ and $B_s^0 \rightarrow D_s X_d (\rightarrow \pi \pi \pi)$, identifying $X_s \rightarrow K \pi \pi$ and $X_d \rightarrow \pi \pi \pi$ as the various resonances through which the decays proceed.

3 Simulated samples

The simulated (MC) samples are generated using Pythia 8. Some more blabla here ! In order to use our MC samples during the BDT training, described in Chapter 4, and the calculation of efficiencies (Chapter 7), we have to make sure that the $B_s^0 \rightarrow D_s K \pi \pi$ decay is modelled correctly by the simulation. To check this we compare distributions of observables, which we use during the multivariate selection stage, as well as some key event observables. The compared distributions need to be generated by signal decays only, therefore we truth match all particles in the monte carlo samples. Signal distributions of observables in data are obtained using the sWeight technique [3]: We perform a fit of a

Variable	Stripping Cut
Track χ^2/nDoF	< 3
Track p	$> 1000 \text{ MeV}/c$
Track p_T	$> 100 \text{ MeV}/c$
Track IP χ^2	> 4
D_s Daughter p_T	$\Sigma_{i=1}^3 p_i > 1800 \text{ MeV}/c$
D_s Daughter DOCA	0.5 mm
D_s mass m_{D_s}	within $\pm 40 \text{ MeV}/c^2$ of PDG value
D_s Vertex χ^2/nDoF	< 10
D_s min FD χ^2	> 36
X_d Daughter p_T	$> 2 \text{ GeV}/c$
$X_{s,d}$ Daughter DOCA	0.4 mm
$X_{s,d}$ Daughter p_T	$\Sigma_{i=1}^3 p_{t,i} > 1250 \text{ MeV}/c$
$X_{s,d}$ Vertex χ^2/nDoF	< 8
$X_{s,d}$ min FD χ^2/nDoF	> 16
$X_{s,d}$ DIRA	> 0.98
$X_{s,d}$ $\Delta\rho$ (vertex displacement perpendicular to z-axis)	$> 0.1 \text{ mm}$
$X_{s,d}$ ΔZ (vertex displacement along z-axis)	$> 2.0 \text{ mm}$
B_s^0 DIRA	> 0.98
B_s^0 min IP χ^2	> 25
B_s^0 Vertex χ^2/nDoF	< 10
B_s^0 $\tau_{B_s^0}$	$> 0.2 \text{ ps}$
K DLL $_{K\pi}$	> -5
π DLL $_{K\pi}$	< 10

Table 2.1: Summary of the stripping selections for $B_s^0 \rightarrow D_s K \pi \pi$ decays.

gaussian signal model and an exponential background to the invariant mass distribution of $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates (our normalization channel). Using the weights generated from this fit, we weight the distributions of data observables in $B_s^0 \rightarrow D_s K \pi \pi$ and obtain the corresponding signal distributions.

Figure 3.1 shows the distribution of the number of tracks per event and the distribution of the maximum ghost probability over all tracks, in MC and data.

In both cases, the distributions differ significantly. Therefore, we re-weight the MC samples using those two variables. All distributions of observables used in the BDT training, before and after the re-weighting procedure, are shown in the Appendix A.1.

4 Selection

A twofold approach is used to isolate the $B_s^0 \rightarrow D_s K \pi \pi$ candidates from data passing the stripping line. First, further one-dimensional cuts are applied to reduce the level of combinatorial background and to veto some specific physical background. After that, a

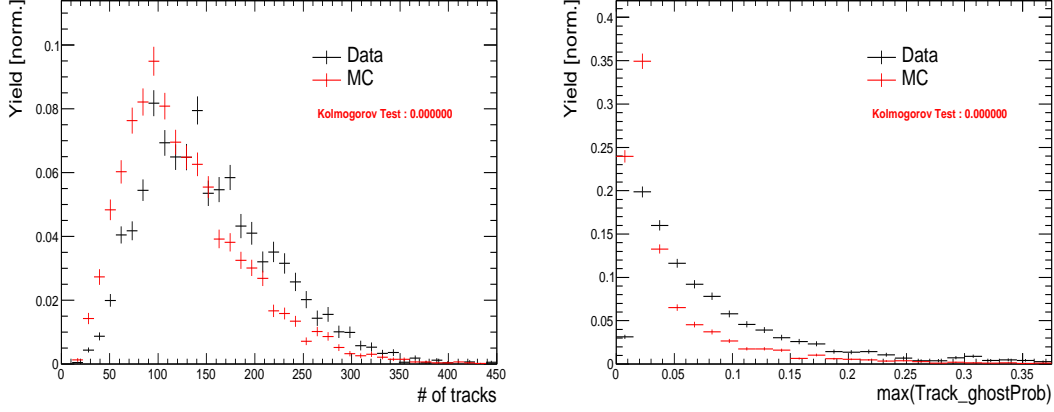


Figure 3.1: Comparison between the distribution of (left) the number of tracks and (right) the maximum ghost probability over all tracks, in (black) data and (red) simulation.

multivariate classifier is trained which combines the information of several input variables, including their correlation, into one powerful discriminator between signal and combinatorial background.

4.1 Cut-based selection

In order to minimize the contribution of combinatorial background to our samples, we apply the following cuts to the b-hadron:

- (i) $\text{DIRA} > 0.99994$
- (ii) $\min \text{IP } \chi^2 < 20$ to any PV
- (iii) $\text{FD } \chi^2 > 100$ to any PV
- (iv) $\text{Vertex } \chi^2/\text{nDoF} < 8$
- (v) $(Z_{D_s} - Z_{B_s^0}) > 0$, where Z_M is the z-component of the position \vec{x} of the decay vertex for the B_s^0/D_s meson

Additionally, we veto various physical backgrounds, which have either the same final state as our signal decay, or can contribute via a single miss-identification of $K \rightarrow \pi$ or $K \rightarrow p$:

- $B_s^0 \rightarrow D_s^+ D_s^- : |M(K\pi\pi) - m_{D_s}| > 20 \text{ MeV}/c^2$
- $B_s^0 \rightarrow D_s K K \pi : \pi^- \text{ DLL}_{K\pi} < 5$

- 72 • $B^0 \rightarrow D^+(\rightarrow K^+\pi^-\pi^+)K\pi\pi$: possible with single miss-ID of $K^+ \rightarrow \pi^+$, vetoed by
 73 changing mass hypothesis and recompute $|M(K^+\pi^-\pi^+) - m_{D^+}| > 20 \text{ MeV}/c^2$, or
 74 the K^+ has to fulfill $\text{DLL}_{K\pi} > 10$
- 75 • $\Lambda_b^0 \rightarrow \Lambda_c^+(\rightarrow pK^-\pi^+)K\pi\pi$: possible with single miss-ID of $K^+ \rightarrow p$, vetoed by
 76 changing mass hypothesis and recompute $M(pK^-\pi^+) - m_{\Lambda_c^+} > 15 \text{ MeV}/c^2$, or the
 77 K^+ has to fulfill $\text{DLL}_{Kp} > 0$

78 All signal candidates for the branching ratio measurement are reconstructed via the
 79 $D_s \rightarrow K^+K^-\pi^+$ channel. This decay can either proceed via the narrow ϕ resonance, the
 80 broader K^{*0} resonance, or non-resonant. Depending on the decay process being resonant
 81 or not, we apply additional PID requirements:

- 82 1. resonant case:
 - 83 (a) $D_s^+ \rightarrow \phi\pi^+$, with $|M(K^+K^-) - m_\phi| < 20 \text{ MeV}/c^2$: no additional requirements
 - 84 (b) $D_s^+ \rightarrow \bar{K}^{*0}K^+$, with $|M(K^-\pi^+) - m_{K^{*0}}| < 75 \text{ MeV}/c^2$: $\text{DLL}_{K\pi} > 0$ for kaons
- 85 2. non-resonant case: $\text{DLL}_{K\pi} > 5$ for kaons

86 4.2 Multivariate stage

87 We use TMVA [4] to train a multivariate discriminator, which is used to further improve
 88 the signal to background ratio. The 17 variables used for the training are:

- 89 • $\text{max}(\text{ghostProb})$ over all tracks
- 90 • $\text{cone}(p_T)$ asymmetry of every track
- 91 • $\text{min}(\text{IP}\chi^2)$ over the X_s daughters
- 92 • $\text{max}(\text{DOCA})$ over all pairs of X_s daughters
- 93 • $\text{min}(\text{IP}\chi^2)$ over the D_s daughters
- 94 • D_s DIRA
- 95 • D_s FD significance
- 96 • $\text{max}(\cos(D_s h_i))$, where $\cos(D_s h_i)$ is the cosine of the angle between the D_s and
 97 another track i in the plane transverse to the beam
- 98 • B_s^0 $\text{IP}\chi^2$, $\text{FD}\chi^2$ and Vertex χ^2

Various classifiers were investigated in order to select the most efficient discriminator. As the result a boosted decision tree with gradient boost (BDTG) is chosen as nominal classifier. We use truth-matched Monte Carlo (MC) as signal input. Those simulated signal candidates are required to pass the same trigger, stripping and preselection requirements, that were used to select the data samples. For the background we use events from the high mass sideband ($m_{B_s^0 \text{ candidate}} > 5600 \text{ MeV}/c^2$) of our data samples.

The distributions of the input variables for signal and background are shown in Fig. 4.1.

The relative importance of the input variables for the BDTG training is summarized in Table 4.1.

Variable	relative importance [%]
max_ghostProb	14.93
log_Bs_IPCHI2_OWNPV	10.91
log_DsDaughters_min_IPCHI2	10.67
K_plus_ptasy_1.00	9.60
Bs_ENDVERTEX_CHI2	9.38
K_minus_fromDs_ptasy_1.00	8.99
log_Ds_FDCHI2_ORIVX	8.78
log_XsDaughters_min_IPCHI2	7.23
K_plus_fromDs_ptasy_1.00	6.62
Xs_max_DOCA	4.13
log_Bs_DIRA	3.36
pi_minus_ptasy_1.00	1.63
pi_minus_fromDs_ptasy_1.00	1.46
cos(Ds h)	0.93
log_Bs_FDCHI2_OWNPV	0.69
pi_plus_ptasy_1.00	0.43
log_Ds_DIRA	0.27

Table 4.1: Summary of the relative importance of each variable in the training of the BDTG.

The BDTG output distribution for test and training samples is shown in Fig 4.2. No sign of overtraining is observed.

We determine the optimal cut value by maximizing the figure of merit $S/\sqrt{S+B}$ where S is the signal yield and B the background yield in the signal region, defined to be within $\pm 50 \text{ MeV}/c^2$ of the nominal B_s^0 mass. To avoid a bias in the determination of the branching fraction, we determine S and B using our normalization channel. All trigger, stripping and additional selections described in this and the previous chapters are applied to the $B_s^0 \rightarrow D_s \pi \pi \pi$ data samples. After that, we perform a simplified version of the fit to the invariant mass distribution of $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates described in Sec. 6. Here, a gaussian signal model and an exponential function to model combinatorial background is used. From this fit we can estimate the number of signal events in our normalization channel. Multiplying that number with the PDG branching fraction of $\frac{\mathcal{B}(B_s^0 \rightarrow D_s K \pi \pi)}{\mathcal{B}(B_s^0 \rightarrow D_s \pi \pi \pi)}$ and

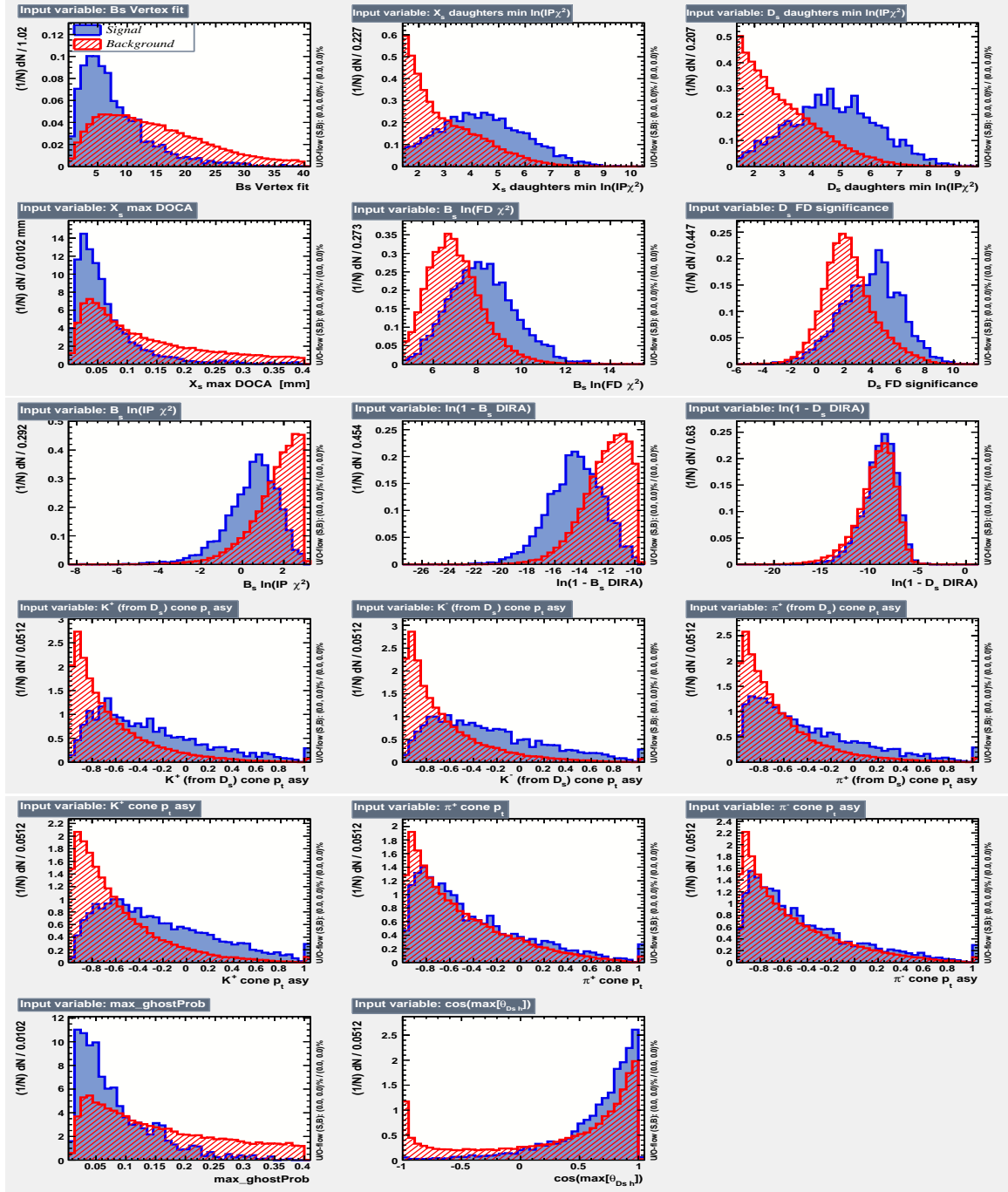


Figure 4.1: Distributions of the input variables used in the BDTG training. The background is shown as red hatched, while the signal is depicted solid blue.

120 the ratio of efficiencies discussed in Sec. 7 allows us to estimate the expected number of
 121 $B_s^0 \rightarrow D_s K \pi \pi$ signals. The number of background events can then be computed as

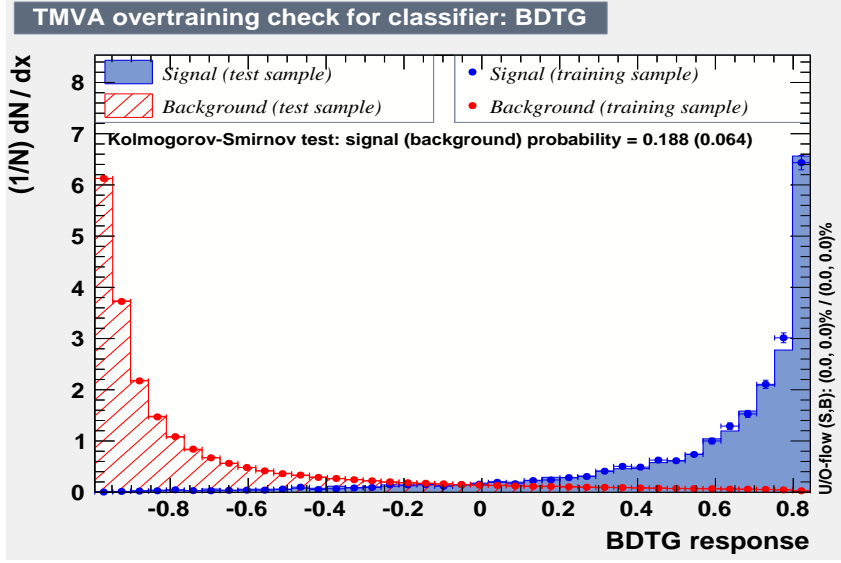


Figure 4.2: BDTG output classifier distribution for (blue) signal and (red) background. The response of an independent test sample is overlaid.

$$N_{bkg} = N_{all} - N_{sig}|_{m_{B_s^0} \pm 50 \text{ MeV}/c^2}. \quad (4.1)$$

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The efficiency curves as a function of the cut value are shown in Fig. 4.3.

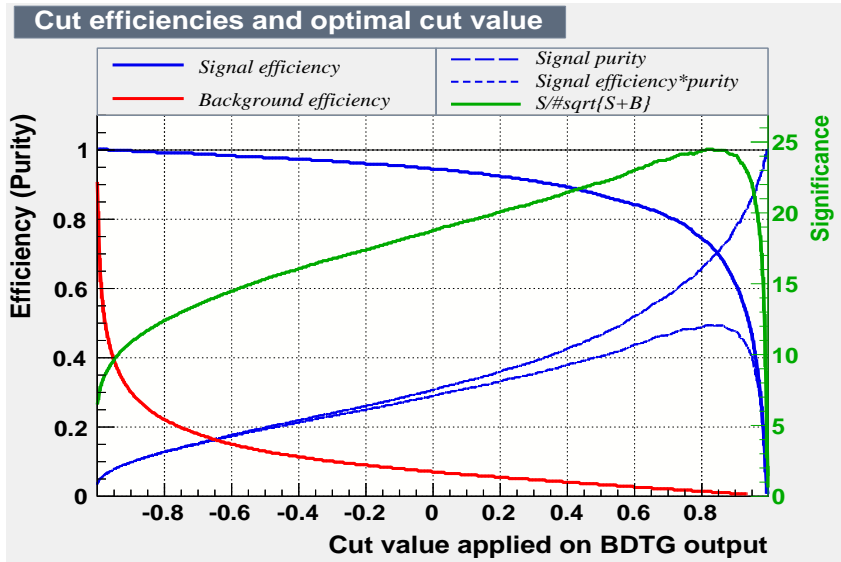


Figure 4.3: Efficiency and purity curves for (blue) signal, (red) background and the (green) FoM curve, as a function of the chosen cut value.

5 Models for signal and background components in invariant mass spectrum

The expected Signal shape, as well as the expected shape for the combinatorial and physical backgrounds have to be known in order to properly describe the invariant mass distribution of $B_s^0 \rightarrow D_s K \pi \pi$ and $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates.

5.1 Signal model

The mass distribution of $B_s^0 \rightarrow D_s K \pi \pi$ signal is modeled using two gaussian functions, which share the same mean μ , but are allowed to have different widths σ_1 and σ_2 . Another double gaussian is used to account for the contribution of $B^0 \rightarrow D_s K \pi \pi$ decays, which are also present in the $m(D_s K \pi \pi)$ spectrum. All parameters of both double gaussians except the core width σ_1 {Why is it fixed and to what? } are allowed to float in the nominal fit. The same approach is used to describe the invariant mass distribution of $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates. A double gaussian is used to model this signal shape, all parameters except the core width σ_1 are allowed to float.

5.2 Background models for $m(D_s \pi \pi \pi)$

Different background sources arise in the invariant mass spectrum of candidates for the normalization mode.

The following backgrounds have to be accounted for:

- combinatorial background: This contribution arises from either a real D_s , which is paired with random tracks to form the B_s^0 candidates, or via real X_d 's, which are combined with three tracks that fake a D_s candidate to form a fake B_s^0 .
- Partially reconstructed $B_s^0 \rightarrow D_s^* \pi \pi \pi$ decays, with $D_s^* \rightarrow D_s \gamma$ or $D_s^* \rightarrow D_s \pi^0$, where the γ/π^0 is not reconstructed in the decay chain.

In both cases of combinatorial background, the distribution in the invariant mass spectrum of B_s^0 candidates is expected to be smooth and decrease with higher masses. Therefore, one exponential function is used to model these contributions.

The shape of the $B_s^0 \rightarrow D_s^* \pi \pi \pi$ contribution is expected to be peaking in the $m(D_s \pi \pi \pi)$ spectrum, with large tails due to the missing momentum, which is carried away by the π^0 or γ . We rely on simulation to estimate the shape of this contribution.

Figure 5.1 shows the fit of the sum of three bifurcated gaussians to the invariant mass distribution of simulated $B_s^0 \rightarrow D_s^* \pi \pi \pi$ event. The pion or photon from $D_s^* \rightarrow D_s(\gamma/\pi^0)$ is excluded from the reconstruction. The obtained shape parameters are used as input values for the nominal $m(D_s \pi \pi \pi)$ mass fit. The yield of this contribution is directly determined in the nominal fit.

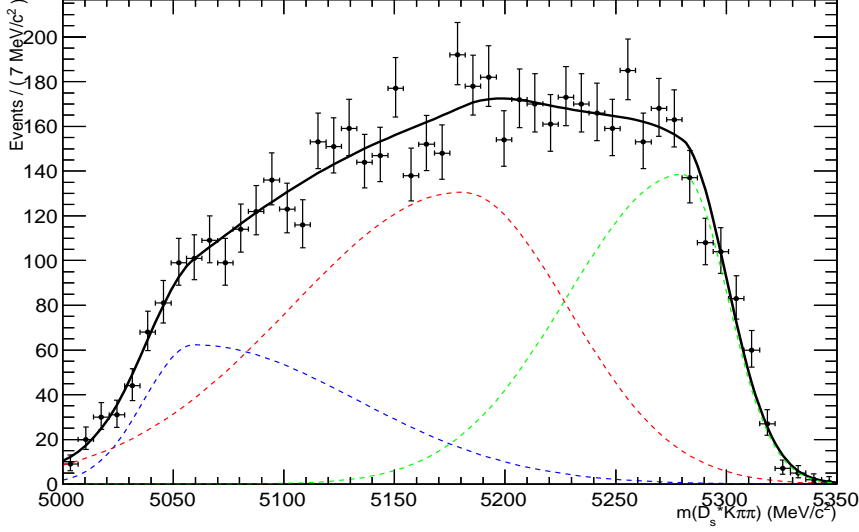


Figure 5.1: Invariant mass distribution of simulated $B_s^0 \rightarrow D_s^* \pi \pi \pi$ events, where the γ/π^0 is excluded from the reconstruction. A fit of the sum of three bifurcated gaussians to this distribution is overlaid.

5.3 Background models for $m(D_s K \pi \pi)$

For the signal channel, the following background sources have to be considered:

- combinatorial background: Same contributions as discussed in Sec. 5.2.
- Partially reconstructed $B_s^0 \rightarrow D_s^* K \pi \pi$ decays, with $D_s^* \rightarrow D_s \gamma$ or $D_s^* \rightarrow D_s \pi^0$, where the γ/π^0 is not reconstructed in the decay chain.
- Partially reconstructed $B^0 \rightarrow D_s^* K \pi \pi$ decays, with $D_s^* \rightarrow D_s \gamma$ or $D_s^* \rightarrow D_s \pi^0$, where the γ/π^0 is not reconstructed in the decay chain.
- miss-identified $B_s^0 \rightarrow D_s \pi \pi \pi$ decays, where one of the pions is wrongly identified as a kaon $\pi \rightarrow K$.
- miss-identified, partially reconstructed $B_s^0 \rightarrow D_s^* \pi \pi \pi$ decays, where one of the pions is wrongly identified as a kaon $\pi \rightarrow K$ and the γ/π^0 from $D_s^* \rightarrow D_s \gamma/\pi^0$ is not reconstructed.

Again the combinatorial background is expected to be flat in the spectrum of the invariant mass of $B_s^0 \rightarrow D_s K \pi \pi$ candidates. An exponential function is used to model this contribution.

The shape of the partially reconstructed $B_s^0/B^0 \rightarrow D_s^* K \pi \pi$ background is taken from simulation. {I think we take them from the normalization mode} A MC sample of

174 $B_s^0 \rightarrow D_s^* K \pi \pi$ events, where the γ/π^0 is excluded from the reconstruction, is generated.
 175 The sum of three bifurcated gaussians is then fitted to the mass distribution of B_s^0
 176 candidates. The distribution and the overlaid fit is shown in Fig. 5.2.

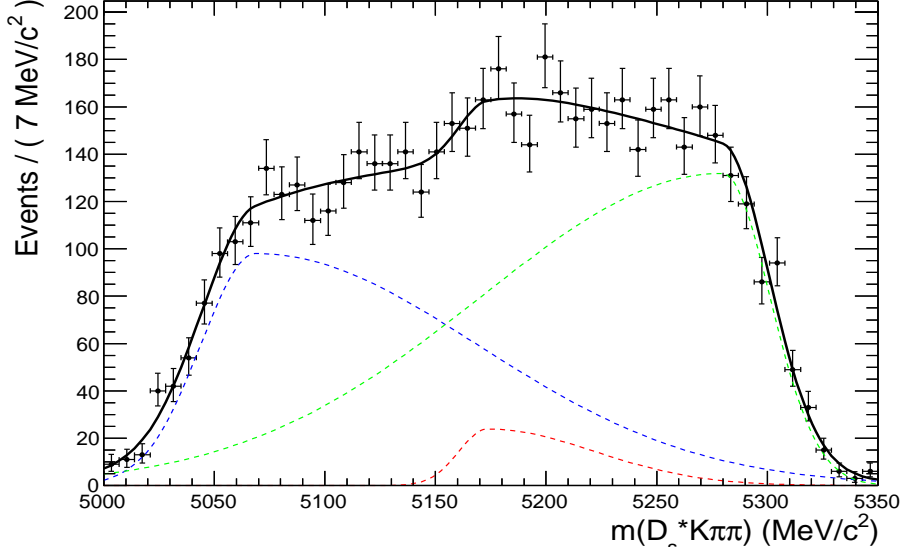


Figure 5.2: Invariant mass distribution of simulated $B_s^0 \rightarrow D_s^* K \pi \pi$ events, where the γ/π^0 is excluded from the reconstruction. A fit of the sum of three bifurcated gaussians to this distribution is overlaid.

177 The obtained shape parameters are used as input values for the nominal $m(D_s K \pi \pi)$
 178 mass fit. For the contribution of the $B^0 \rightarrow D_s^* K \pi \pi$ background, the same shape is used,
 179 but the means μ_i of the bifurcated gaussians are shifted down by $m_{B_s^0} - m_{B^0}$ [5]. The
 180 yield of both contributions are directly determined in the nominal fit.
 181 To determine the shape of miss-identified $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates in the $m(D_s K \pi \pi)$
 182 spectrum, we take a truth matched signal MC sample of our normalization channel. We
 183 then use the PIDCalib package to determine the $\pi \rightarrow K$ fake rate. For every candidate in
 184 our MC sample, a p and η -dependent event weight is computed and assigned. We flip the
 185 particle hypothesis from pion to kaon for the π with the biggest miss-ID weight for each
 186 event and recompute the invariant B_s^0 mass. This distribution is then modelled using two
 187 crystal ball functions. The distribution and fit is shown in Fig. 5.3(left).

188 The expected yield of miss-identified $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates in the $m(D_s K \pi \pi)$
 189 spectrum is computed by multiplying the fake probability of $\propto 3.2\%$, which is derived
 190 from PIDCalib, by the yield of $B_s^0 \rightarrow D_s \pi \pi \pi$ signal candidates, determined in the nominal
 191 mass fit of our normalization channel.

192 In the same way as mentioned above, we can determine the rate of miss-identified, partially
 193 reconstructed $B_s^0 \rightarrow D_s^* \pi \pi \pi$ decays in our sample of $B_s^0 \rightarrow D_s^* K \pi \pi$ decays using PIDCalib
 194 and a MC sample of $B_s^0 \rightarrow D_s^* \pi \pi \pi$ events. The invariant mass distribution we obtain

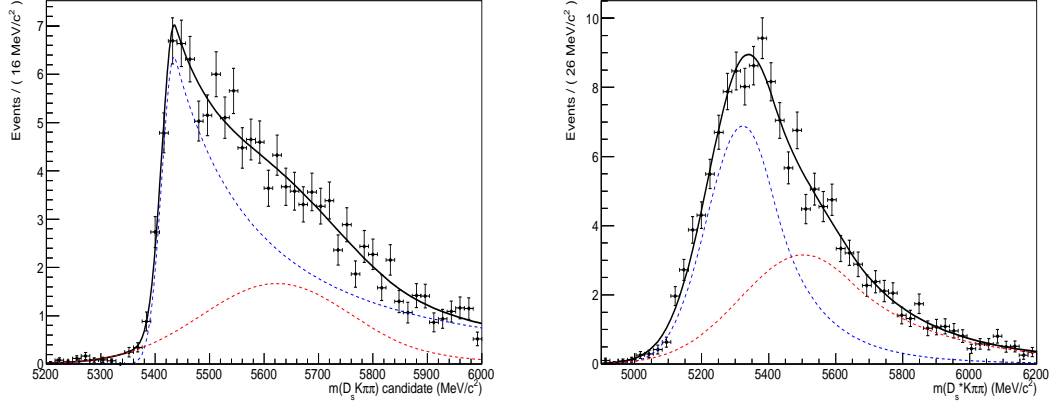


Figure 5.3: Invariant mass distribution of (left) simulated $B_s^0 \rightarrow D_s \pi \pi \pi$ events, where one of the π 's is reconstructed as a K and the miss-ID probability for each event is taken into account. The corresponding distribution for simulated $B_s^0 \rightarrow D_s^* \pi \pi \pi$ events, where the γ/π^0 from the D_s^* is excluded from reconstruction, is shown on the right. A fit of the sum of two crystal ball functions to each of these distributions is overlaid.

when we exlude the γ/π^0 , flip the the particle hypothesis $\pi \rightarrow K$ and apply the event weights given by the fake rate, is shown in Fig. 5.3 (right). The fit of two crystal ball functions to this distribution is overlaid. The yield of this contribution is determined from the yield of $B_s^0 \rightarrow D_s^* \pi \pi \pi$ candidates in the nominal mass fit of our normalization channel, multiplied by the miss-ID probability of $\propto 3.6\%$.

6 Massfits for signal and normalization channel

This section describes the nominal fits to the invariant mass distribution of $B_s^0 \rightarrow D_s K \pi \pi$ and $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates after all selection steps, described in the previous Sections, are applied. The obtained yields are summarized in Tab. 6.1.

6.1 Fit to $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates

An unbinned maximum likelihood fit is performed simultaneously to the invariant mass distribution of $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates, for 7 and 8 TeV data. As discussed in Sec. 5.1, the fit is given as the sum of the double gaussian signal model, the sum of three bifurcated gaussians to model the partially reconstructed $B_s^0 \rightarrow D_s^* \pi \pi \pi$ background, as well as an exponential to account for combinatorial background. The invariant mass distribution and the fit to it is shown in Fig. 6.1.

The determined number of $B_s^0 \rightarrow D_s \pi \pi \pi$ decays is 8693 ± 102 for 2011 data and 19881 ± 159 for 2012 data. The determined yield for the partially reconstructed $B_s^0 \rightarrow D_s^* \pi \pi \pi$ background is (2011) 16904 ± 299 and (2012) 38437 ± 589 , while the yield for the

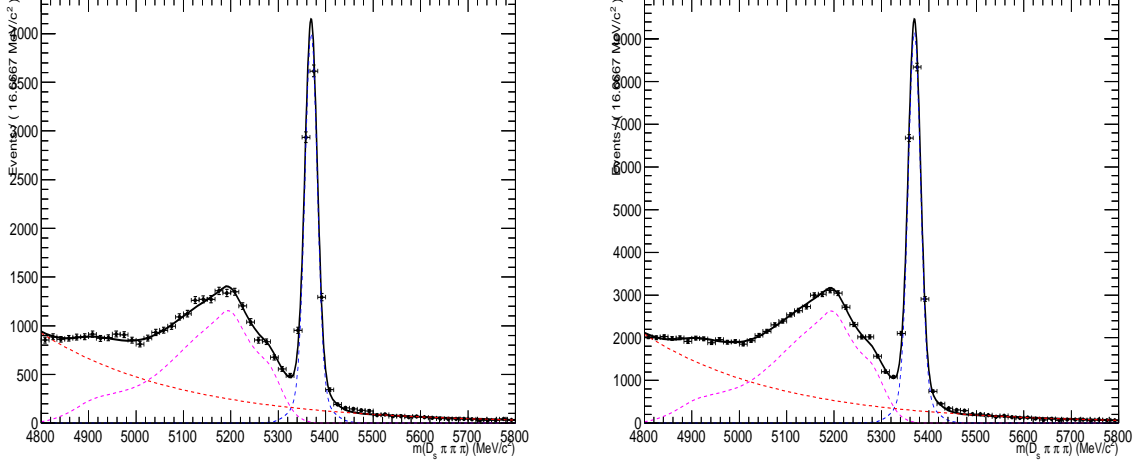


Figure 6.1: Invariant mass distribution of $B_s^0 \rightarrow D_s \pi \pi \pi$ candidates for (left) 2011 and (right) 2012 data. A fit described in the text is overlaid. The dashed lines show the (green) partially reconstructed and (red) combinatorial background, as well as the (blue) signal component.

combinatorial background is (2011) 16066 ± 304 and (2012) 35285 ± 596 .

6.2 Fit to $B_s^0 \rightarrow D_s K \pi \pi$ candidates

Fig. 6.2 shows the invariant mass distribution of $B_s^0 \rightarrow D_s K \pi \pi$ candidates. A simultaneous unbinned maximum likelihood fit is overlaid, which consists of two double gaussian models for the B^0 and B_s^0 signal, two sums of three bifurcated gaussians for the $B_s^0/B^0 \rightarrow D_s^* K \pi \pi$ partially reconstructed background contributions and two sums of double crystal ball functions for the single miss-ID $B_s^0 \rightarrow D_s \pi \pi \pi$ and the partially reconstructed, miss-identified $B_s^0 \rightarrow D_s^* \pi \pi \pi$ decays.

The extracted signal yields are (2011) 367 ± 26 and (2012) 825 ± 40 .

Decay	yield 2011	yield 2012
$B_s^0 \rightarrow D_s K \pi \pi$	367 ± 26	825 ± 40
$B_s^0 \rightarrow D_s \pi \pi \pi$	8693 ± 102	19881 ± 159

Table 6.1: Summary of signal yields from the fits to 2011 and 2012 data.

7 Efficiency corrections

Several relative efficiency corrections are needed to measure the branching fraction of $B_s^0 \rightarrow D_s K \pi \pi$ with respect to $B_s^0 \rightarrow D_s \pi \pi \pi$. Precise knowledge of the efficiency related

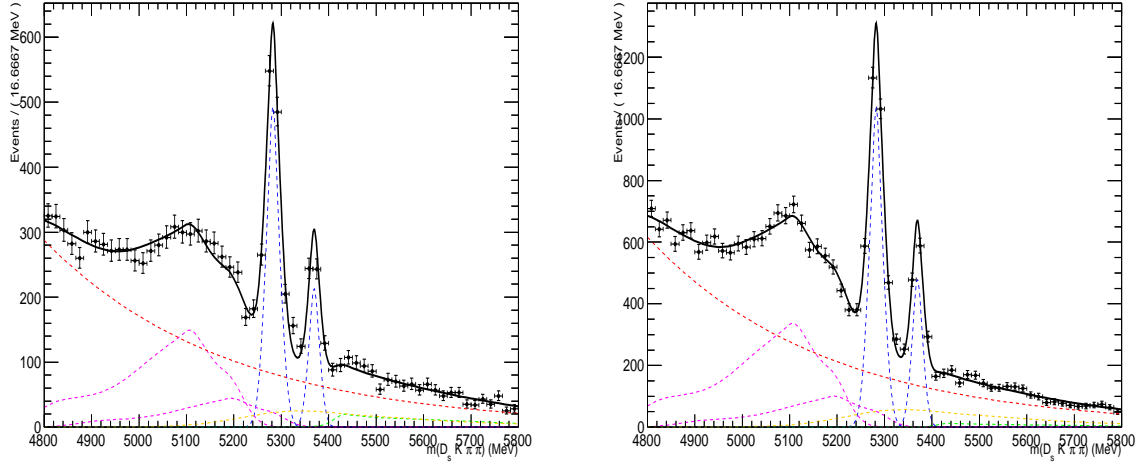


Figure 6.2: Invariant mass distribution of $B_s^0 \rightarrow D_s K \pi \pi$ candidates for (left) 2011 and (right) 2012 data. A fit described in the text is overlaid. The dashed lines show the (green) partially reconstructed and (red) combinatorial background, as well as the (blue) signal component. Additional, the dashed magenta line depicts the miss-ID background and the dashed yellow line shows the miss-identified, partially reconstructed background component.

to the detector acceptance, PID requirements, used trigger lines and offline selections are crucial for both, the determination of γ and the branching ratio measurement.

7.1 Relative efficiency for BR measurement

For the branching ratio measurement, the relative efficiency is given by

$$\epsilon_{rel} = \epsilon_{rel}^{acc} \cdot \epsilon_{rel}^{sel} \cdot \epsilon_{rel}^{pid}, \quad (7.1)$$

where $\epsilon = \frac{\epsilon_{Norm}}{\epsilon_{Sig}}$ is the ratio of the efficiency for the signal and normalization mode. To evaluate these efficiencies, we rely on simulation. The three efficiencies given in Eq. 7.1 are:

- ϵ_{rel}^{acc} : This is the relative efficiency due to the geometrical acceptance of the LHCb detector. All tracks are required to have a polar angle between 10 and 400 mrad and a minimal momentum of $|p| > 1.6$ GeV/c in order to be recorded for further analysis. Since the particle species of one track differs between the signal and normalization mode, the efficiencies caused by the geometrical acceptance are expected to be different for the two channels.
- ϵ_{rel}^{sel} : The relative selection efficiency due to trigger and offline requirements.
- ϵ_{rel}^{pid} : The relative PID efficiency due to the identification likelihood requirements for tracks from both modes. This is evaluated using efficiencies from $D^{*+} \rightarrow$

242 $D^0(K^-\pi^+)\pi^+$ calibration data, which is weighted by the expected momentum (p)
 243 distribution taken from simulation.

244 Using the definition given in Eq. 7.1, the branching ratio can be expressed as

$$\frac{\mathcal{B}(B_s^0 \rightarrow D_s K \pi \pi)}{\mathcal{B}(B_s^0 \rightarrow D_s \pi \pi \pi)} = \frac{\mathcal{Y}(B_s^0 \rightarrow D_s K \pi \pi)}{\mathcal{Y}(B_s^0 \rightarrow D_s \pi \pi \pi)}, \cdot \epsilon_{rel} \quad (7.2)$$

245 where $\mathcal{Y}(x)$ represents the yield of the respective channel.

246 The single efficiencies, as well as the total selection efficiency, for the signal and normaliza-
 247 tion channel, is given in Table 7.1.

Efficiency (%)	$B_s^0 \rightarrow D_s K \pi \pi$	$B_s^0 \rightarrow D_s \pi \pi \pi$
2011 ϵ^{acc}	15.84 ± 0.04	9.85 ± 0.04
2012 ϵ^{acc}	16.11 ± 0.04	yyy
2011 ϵ^{sel}	0.658 ± 0.011	0.894 ± 0.013
2012 ϵ^{sel}	0.574 ± 0.008	yyy
2011 ϵ^{pid}	74.88 ± 0.85	92.64 ± 0.47
2012 ϵ^{pid}	74.30 ± 0.85	-
2011 total ϵ	0.078 ± 0.002	0.082 ± 0.001
2012 total ϵ	0.069 ± 0.002	zz

Table 7.1: Efficiencies due to the detector acceptance, selection requirements and PID cuts for the signal and normalization mode. All values are obtained using simulated events.

248 8 Systematic errors

249 Several systematic errors contribute to the overall uncertainty on the brachning fractions.
 250 We consider the most significant ones:

- 251 • Particle identification
- 252 • Signal and background models
- 253 • Determination of the selection efficiency with MC
- 254 • MC statistics
- 255 • BDTG efficiency

256 Text with short description of every source and how we determine systematic uncer-
 257 tainty.

Source	Uncertainty on $\frac{\mathcal{B}(B_s^0 \rightarrow D_s K \pi \pi)}{\mathcal{B}(B_s^0 \rightarrow D_s \pi \pi \pi)}$ [%]
PID	
Fit model	
MC efficiency determination	
MC statistics	
BDTG efficiency	
Total	

Table 8.1: Summary of considered systematic uncertainties on the branching ratio determination.

9 Results and summary

Using the definition of the branching ratio given in Eq. 7.2, we compute from the measured yields and efficiencies:

$$\frac{\mathcal{B}(B_s^0 \rightarrow D_s K \pi \pi)}{\mathcal{B}(B_s^0 \rightarrow D_s \pi \pi \pi)} = 0.051 \pm 0.002 \pm 0.xxx, \quad (9.1)$$

where the uncertainties are statistical and systematical, respectively. Further discussion...

263 **A Appendix**

264 **A.1 Re-weghted MC observables**

265 The following figures show the distributions of observables used during the multivariate
266 selection stage, for data, monte carlo and re-weighted monte carlo.

267 **A.2 Toys for normalization fit**

268 To validate the fit model used to describe the $m(D_s\pi\pi\pi)$ spectrum, we produce 1000
269 pseudo samples from our fit pdf and fit them with the same nominal pdf model.
270 A pull of a certain fit parameter is defined as

$$p = \frac{x - x_0}{\Delta x}, \tag{1.1}$$

271 where x is the fitted value, x_0 the generated value and Δx the uncertainty on x . Given
272 the fit is correctly implemented and unbiased, one expects the distribution of the pulls for
273 every fit parameter to be centered around 0, with a gaussian width of 1.

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