

# **GSERM - 2018**

## Introduction to Survival Data

June 6, 2018 (afternoon session)

# Survival Analysis

- Models for *time-to-event data*.
- Roots in biostats/epidemiology, plus engineering, sociology, economics.
- Examples...
  - Political careers, confirmation durations, position-taking, bill cosponsorship, campaign contributions, policy innovation/adoption, etc.
  - Cabinet/government durations, length of civil wars, coalition durability, etc.
  - War duration, peace duration, alliance longevity, length of trade agreements, etc.
  - Strike durations, work careers (including promotions, firings, etc.), criminal careers, marriage and child-bearing behavior, etc.

# Characteristics of Time-To-Event Data

- Discrete events (i.e., not continuous),
- Take place over time,
- May not (or *never*) experience the event (i.e., possibility of censoring).

# Survival Data Basics: Terminology

$Y_i$  = the duration until the event occurs,

$Z_i$  = the duration until the observation is “censored”

$T_i$  =  $\min\{Y_i, Z_i\}$ ,

$C_i$  = 0 if observation  $i$  is censored, 1 if it is not.

# Survival Data Basics: The Density

$$f(t) = \Pr(T_i = t)$$

Issues:

- $T_i = t$  iff  $T_i > t - 1, t - 2$ , etc.
- $C_i = 0$  (censoring)

# Survival Data Basics: Survivor Function

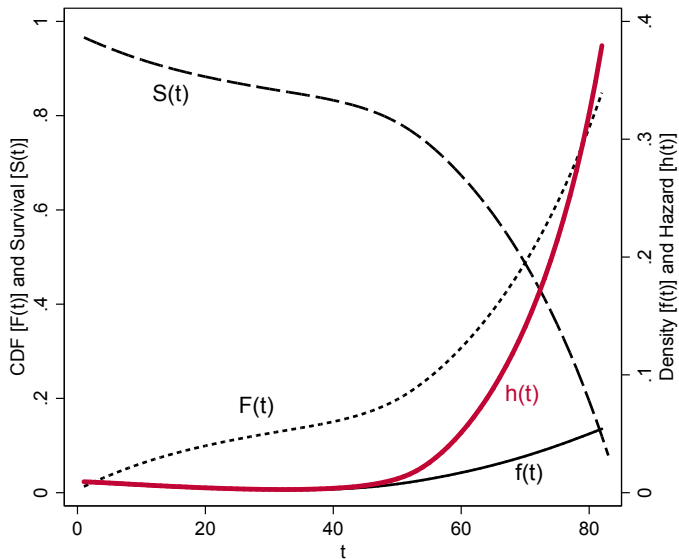
$$\Pr(T_i \leq t) \equiv F(t) = \int_0^t f(t) dt$$

$$\begin{aligned}\Pr(T_i \geq t) \equiv S(t) &= 1 - F(t) \\ &= 1 - \int_0^t f(t) dt\end{aligned}$$

## Survival Data Basics: The Hazard

$$\begin{aligned}\Pr(T_i = t | T_i \geq t) \equiv h(t) &= \frac{f(t)}{S(t)} \\ &= \frac{f(t)}{1 - \int_0^t f(t) dt}\end{aligned}$$

## Example: Human Mortality





# Some Useful Equivalencies

$$f(t) = \frac{-\partial S(t)}{\partial t}$$

Implies

$$\begin{aligned} h(t) &= \frac{\frac{-\partial S(t)}{\partial t}}{S(t)} \\ &= \frac{-\partial \ln S(t)}{\partial t} \end{aligned}$$

# More Useful Things: Integrated Hazard

Define

$$H(t) = \int_0^t h(t) dt.$$

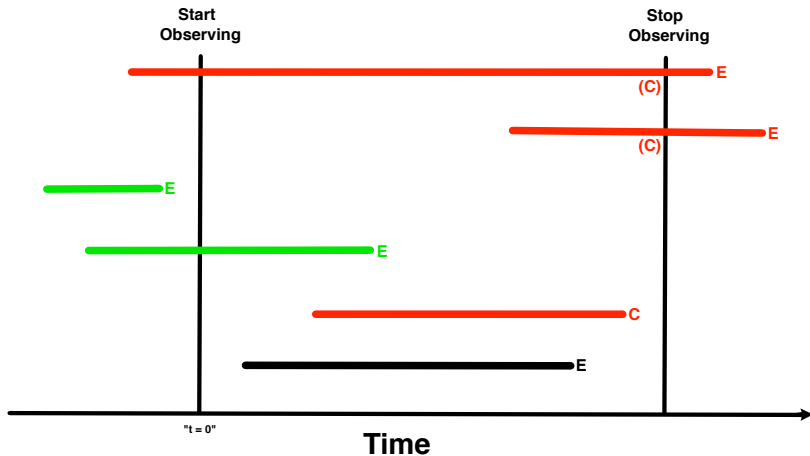
Implies

$$\begin{aligned} H(t) &= \int_0^t \frac{-\partial \ln S(t)}{\partial t} dt \\ &= -\ln[S(t)] \end{aligned}$$

and

$$S(t) = \exp[-H(t)]$$

# Censoring and Truncation



- Defined by the researcher
- Conditionally independent of both  $T_i$  and  $\mathbf{X}_i$
- Doesn't mean that the observation provides no information

# Estimating $S(t)$

Assume  $N$  observations, *absorbing* events, and no ties. Then define

- $n_t$  = number of observations “at risk” for the event at  $t$ , and
- $d_t$  = number of observations which experience the event at time  $t$ .

Then

$$\widehat{S(t_k)} = \prod_{t \leq t_k} \frac{n_t - d_t}{n_t}$$

## Variance of $\widehat{S}(t)$

$$\text{Var}[\widehat{S}(t_k)] = \left[\widehat{S}(t_k)\right]^2 \sum_{t \leq t_k} \frac{d_t}{n_t(n_t - d_t)}$$

Note:

- $\text{Var}[\widehat{S}(t_k)]$  is increasing in  $S(t)$ ,
- is also increasing in  $d_t$ , but
- is decreasing in  $n_t$ .

“Nelson-Aalen”:

$$\widehat{H}(t_k) = \sum_{t \leq t_k} \frac{d_t}{n_t}$$

...which gives an alternative estimator for the survival function equal to:

$$\begin{aligned}\widehat{S}(t_k) &= \exp[-\widehat{H}(t_k)] \\ &= \exp \left[ - \sum_{t \leq t_k} \frac{d_t}{n_t} \right]\end{aligned}$$

# Bivariate Hypothesis Testing

	Treatment	Placebo	Total
Event	$d_{1t}$	$d_{0t}$	$d_t$
No Event	$n_{1t} - d_{1t}$	$n_{0t} - d_{0t}$	$n_t - d_t$
Total	$n_{1t}$	$n_{0t}$	$n_t$

Log-Rank Test:

$$Q = \frac{\left[ \sum (d_{1t} - \frac{n_{1t}d_t}{n_t}) \right]^2}{\left[ \frac{n_{1t}n_{0t}d_t(n_t - d_t)}{n_t^2(n_t - 1)} \right]}$$
$$\sim \chi_1^2$$



# A Diversion: Survival Models and Counting Processes

Assume

- Event is *absorbing*,
- $Y_i$  is duration to the event
- $Z_i$  is duration to censoring
- Observe  $T_i = \min(Y_i, Z_i)$ , and
- $C_i$ :
  - $C_i = 0$  if  $T_i = Z_i$ ,
  - $C_i = 1$  if  $T_i = Y_i$ .
- $T_i \neq T_j \forall i \neq j$  (no “ties”)

# Three Key Variables

1. *Counting Process* Indicator:

$$N_i(t) = I(T_i \leq t, C_i = 1)$$

2. *Risk* Indicator:

$$R_i(t) = I(T_i > t)$$

3. *Intensity Process*:

$$\lambda_i(t) dt = R_i(t)h(t)$$

With

$$\Lambda_i(t) = \int_0^t \lambda_i(t) dt$$

we can think of

$$N_i(t) = \Lambda_i(t) + M_i(t)$$

or

$$M_i(t) = N_i(t) - \Lambda_i(t).$$

$$E(X_{t+s} | X_0, X_1, \dots, X_i, \dots, X_t) = X_t \quad \forall s > 0$$

# Data Structure and Organization: Non-Time-Varying

id	durat	censor	timein	timeout	X
1	4	0	30	34	0.12
2	2	1	12	14	0.19
3	5	1	5	10	0.09
...	...	...	...	...	...
N	10	1	21	31	0.22

# Time-Varying Data

id	durat	censor	timein	timeout	X	Z
1	1	0	30	31	0.12	331
1	2	0	31	32	0.12	412
1	3	0	32	33	0.12	405
1	4	0	33	34	0.12	416
2	1	0	12	13	0.19	226
2	2	1	13	14	0.19	296
3	1	0	5	6	0.09	253
3	2	0	6	7	0.09	311
3	3	0	7	8	0.09	327
3	4	0	8	9	0.09	344
3	5	1	9	10	0.09	301
...	...	...	...	...	...	...

# Analyzing Survival Data in R

survival object (non-time-varying):

```
library(survival)
NonTV<-read.csv(NonTVdata.csv)
NonTV.S<-Surv(NonTV$duration, NonTV$censor)
```

survival object (time-varying):

```
TV<-read.csv(TVdata.csv)
TV.S<-Surv(TV$starttime, TV$endtime, TV$censor)
```

# An Example

OECD Cabinet survival [Strom (1985); King et al. (1990)],

$N = 314$  cabinets in 15 countries

Outcome: Duration of cabinet, in months

Covariates (all non-time varying):

- *Fractionalization*
- *Polarization*
- *Formation Attempts*
- **Investiture**
- *Numerical Status*
- *Post-Election*
- *Caretaker*

Also: Indicator for whether the cabinet ended within 12 months of the end of the “constitutional inter-election period” (→ censored)



```
> head(KABL)
  id country durat ciep12 fract polar format invest numst2 eltime2 caret2
1  1      1  0.5     1   656    11     3      1      0      1      0
2  2      1  3.0     1   656    11     2      1      1      0      0
3  3      1  7.0     1   656    11     5      1      1      0      0
4  4      1 20.0     1   656    11     2      1      1      0      0
5  5      1  6.0     1   656    11     3      1      1      0      0
6  6      1  7.0     1   634     6     4      1      1      1      0
```

```
> KABL.S<-Surv(KABL$durat,KABL$ciep12)
```

```
> KABL.S[1:50,]
```

```
[1] 0.5  3.0  7.0 20.0  6.0  7.0  2.0 17.0 27.0 49.0+
[11] 4.0 29.0 49.0+ 6.0 23.0 41.0+ 10.0 12.0  2.0 33.0
[21] 1.0 16.0  2.0  9.0  3.0  5.0  5.0  6.0 45.0+ 23.0
[31] 41.0  7.0 49.0+ 46.0  9.0 51.0+ 10.0 32.0 28.0  3.0
[41] 53.0+ 17.0 59.0+  9.0 52.0+  3.0 23.0 33.0  1.0 30.0
```

# Example survfit Object

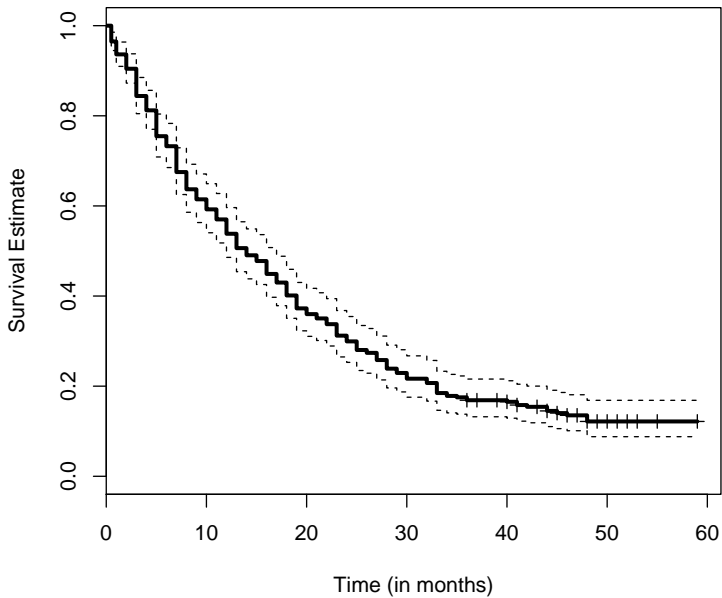
```
> KABL.fit<-survfit(KABL.S~1)
```

```
> str(KABL.fit)
```

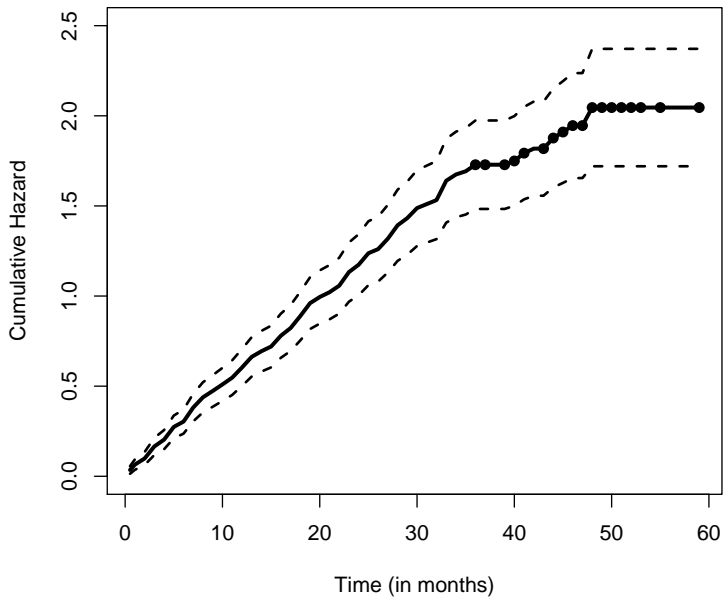
List of 13

```
$ n      : int 314
$ time   : num [1:54] 0.5 1 2 3 4 5 6 7 8 9 ...
$ n.risk  : num [1:54] 314 303 294 284 265 255 237 230 212 200 ...
$ n.event : num [1:54] 11 9 10 19 10 18 7 18 12 7 ...
$ n.censor : num [1:54] 0 0 0 0 0 0 0 0 0 0 ...
$ surv    : num [1:54] 0.965 0.936 0.904 0.844 0.812 ...
$ type    : chr "right"
$ std.err  : num [1:54] 0.0108 0.0147 0.0183 0.0243 0.0271 ...
$ upper    : num [1:54] 0.986 0.964 0.938 0.885 0.856 ...
$ lower    : num [1:54] 0.945 0.91 0.873 0.805 0.77 ...
$ conf.type: chr "log"
$ conf.int : num 0.95
$ call     : language survfit(formula = KABL.S ~ 1)
- attr(*, "class")= chr "survfit"
```

Plotting  $\widehat{S}(t)$



Plotting  $\widehat{H}(t)$



Log-rank test:

```
> survdiff(KABL.S~invest,data=KABL,rho=0)
```

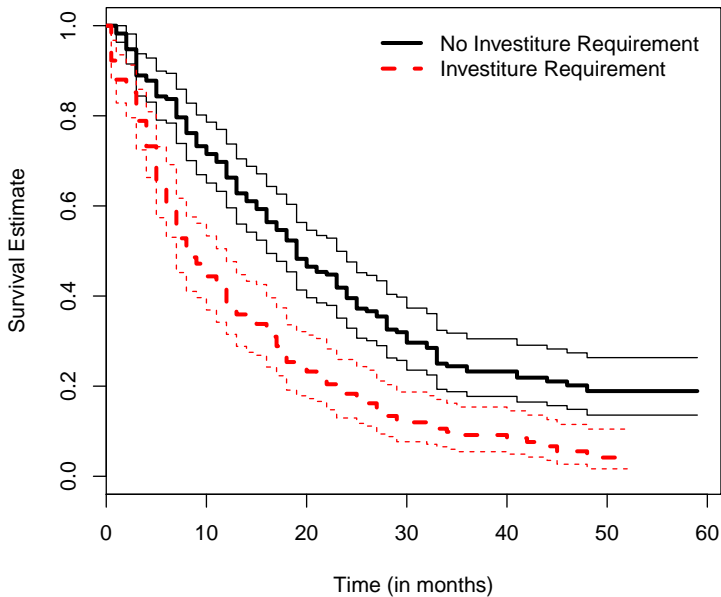
Call:

```
survdiffformula = KABL.S ~ invest, data = KABL, rho = 0)
```

	N	Observed	Expected	(O-E) <sup>2</sup> /E	(O-E) <sup>2</sup> /V
invest=0	172	137	178.7	9.72	30.5
invest=1	142	134	92.3	18.81	30.5

Chisq= 30.5 on 1 degrees of freedom, p= 3.26e-08

# Comparing $\widehat{S}(t)$ s



# A General Parametric Model

$$f(t) = \lim_{\Delta t \rightarrow 0} \frac{\Pr(t \leq T < t + \Delta t)}{\Delta t}$$

$$\begin{aligned} S(t) &= \Pr(T \geq t) \\ &= 1 - \int_0^t f(t) dt \\ &= 1 - F(t) \end{aligned}$$

$$\begin{aligned} h(t) &= \frac{f(t)}{S(t)} \\ &= \lim_{\Delta t \rightarrow 0} \frac{\Pr(t \leq T < t + \Delta t | T \geq t)}{\Delta t} \end{aligned}$$

$$L = \prod_{i=1}^N [f(T_i)]^{C_i} [S(T_i)]^{1-C_i}$$

$$\ln L = \sum_{i=1}^N \{C_i \ln [f(T_i)] + (1 - C_i) \ln [S(T_i)]\}$$

$$\ln L | \mathbf{X}, \boldsymbol{\beta} = \sum_{i=1}^N \{C_i \ln [f(T_i | \mathbf{X}, \boldsymbol{\beta})] + (1 - C_i) \ln [S(T_i | \mathbf{X}, \boldsymbol{\beta})]\}$$



# The Exponential Model

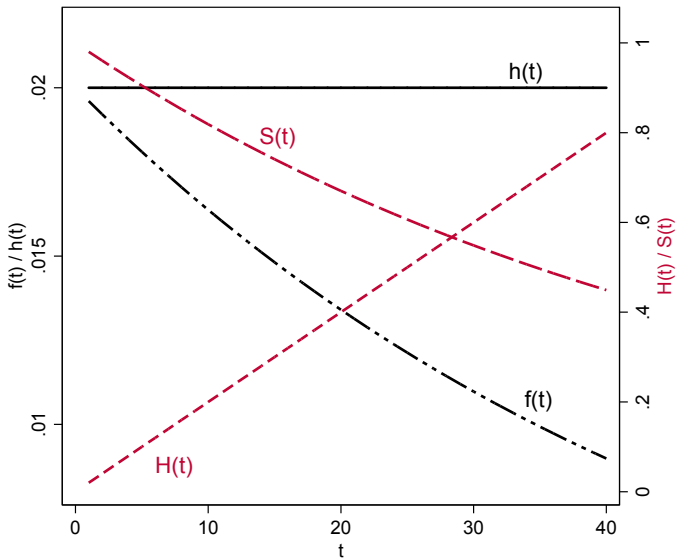
$$h(t) = \lambda$$

$$\begin{aligned} H(t) &= \int_0^t h(t) dt \\ &= \lambda t \end{aligned}$$

$$\begin{aligned} S(t) &= \exp[-H(t)] \\ &= \exp(-\lambda t) \end{aligned}$$

$$\begin{aligned} f(t) &= h(t)S(t) \\ &= \lambda \exp(-\lambda t) \end{aligned}$$

# The Exponential Model, Illustrated



$$\lambda_i = \exp(\mathbf{X}_i\beta).$$

$$S_i(t) = \exp(-e^{\mathbf{X}_i\beta} t).$$

# Exponential (log-)Likelihood

$$\begin{aligned}\ln L &= \sum_{i=1}^N \left\{ C_i \ln \left[ \exp(\mathbf{X}_i \beta) \exp(-e^{\mathbf{X}_i \beta} t) \right] + \right. \\ &\quad \left. (1 - C_i) \ln \left[ \exp(-e^{\mathbf{X}_i \beta} t) \right] \right\} \\ &= \sum_{i=1}^N \left\{ C_i \left[ (\mathbf{X}_i \beta) (-e^{\mathbf{X}_i \beta} t) \right] + (1 - C_i) (-e^{\mathbf{X}_i \beta} t) \right\}\end{aligned}$$

## Exponential: “AFT”

$$\ln T_i = \mathbf{X}_i \gamma + \epsilon_i$$

$$T_i = \exp(\mathbf{X}_i \gamma) \times u_i$$

$$\epsilon_i = \ln T_i - \mathbf{X}_i \gamma$$

## Interpretation: Hazard Ratios

$$\text{HR}_k = \frac{\widehat{h(t)|X_k = 1}}{\widehat{h(t)|X_k = 0}}$$

$$h_i(t) = \exp(\beta_0)\exp(\mathbf{X}_i\beta)$$

$$\begin{aligned}\text{HR}_k &= \frac{\widehat{h(t)|X_k = 1}}{\widehat{h(t)|X_k = 0}} \\&= \frac{\exp(\hat{\beta}_0 + X_1\hat{\beta}_1 + \dots + \hat{\beta}_k(1) + \dots)}{\exp(\hat{\beta}_0 + X_1\hat{\beta}_1 + \dots + \hat{\beta}_k(0) + \dots)} \\&= \frac{\exp(\hat{\beta}_k \times 1)}{\exp(\hat{\beta}_k \times 0)} \\&= \exp(\hat{\beta}_k)\end{aligned}$$

$$\begin{aligned}\text{HR}_k &= \frac{\hat{h}(t)|X_k + \delta}{\hat{h}(t)|X_k} \\ &= \exp(\delta \hat{\beta}_k)\end{aligned}$$

$$\text{HR}_{\frac{i}{j}} = \frac{\exp(\mathbf{X}_i \hat{\beta})}{\exp(\mathbf{X}_j \hat{\beta})}$$

# Example: King et al. (1990) Data

```
> summary(KABL)
```

id	country	durat	ciep12
Min. : 1.00	Min. : 1.000	Min. : 0.50	Min. :0.0000
1st Qu.: 79.25	1st Qu.: 4.000	1st Qu.: 6.00	1st Qu.:1.0000
Median :157.50	Median : 7.000	Median :14.00	Median :1.0000
Mean :157.50	Mean : 7.182	Mean :18.44	Mean :0.8631
3rd Qu.:235.75	3rd Qu.:10.000	3rd Qu.:28.00	3rd Qu.:1.0000
Max. :314.00	Max. :15.000	Max. :59.00	Max. :1.0000

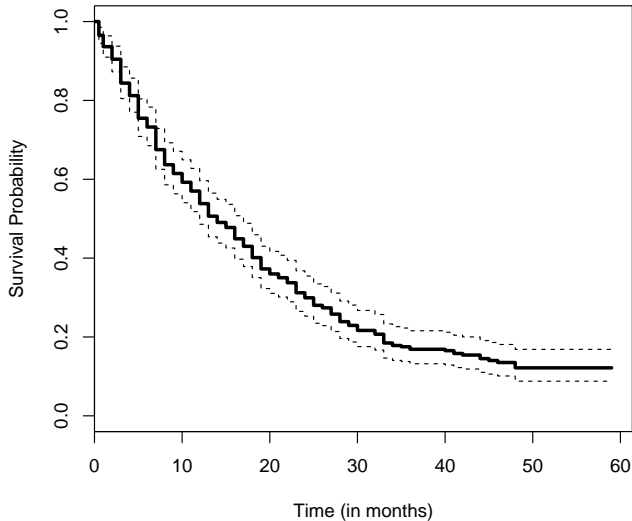
fract	polar	format	invest
Min. :349.0	Min. : 0.00	Min. :1.000	Min. :0.0000
1st Qu.:677.0	1st Qu.: 3.00	1st Qu.:1.000	1st Qu.:0.0000
Median :719.0	Median :14.50	Median :1.000	Median :0.0000
Mean :718.8	Mean :15.29	Mean :1.904	Mean :0.4522
3rd Qu.:788.0	3rd Qu.:25.00	3rd Qu.:2.000	3rd Qu.:1.0000
Max. :868.0	Max. :43.00	Max. :8.000	Max. :1.0000

numst2	eltime2	caret2
Min. :0.0000	Min. :0.0000	Min. :0.00000
1st Qu.:0.0000	1st Qu.:0.0000	1st Qu.:0.00000
Median :1.0000	Median :0.0000	Median :0.00000
Mean :0.6306	Mean :0.4873	Mean :0.05414
3rd Qu.:1.0000	3rd Qu.:1.0000	3rd Qu.:0.00000
Max. :1.0000	Max. :1.0000	Max. :1.00000



# Cabinet Durations: Kaplan-Meier



# Exponential Model (AFT form)

```
> KABL.S<-Surv(KABL$durat,KABL$ciep12)
> xvars<-c("fract","polar","format","invest","numst2","eltime2","caretk2")
> MODEL<-as.formula(paste(paste("KABL.S ~ ", paste(xvars,collapse="+"))))
> KABL.exp.AFT<-survreg(MODEL,data=KABL,dist="exponential")
> summary(KABL.exp.AFT)
```

Call:

```
survreg(formula = MODEL, data = KABL, dist = "exponential")
```

	Value	Std. Error	z	p
(Intercept)	3.72460	0.630834	5.90	3.54e-09
fract	-0.00116	0.000905	-1.29	1.98e-01
polar	-0.01610	0.006097	-2.64	8.28e-03
format	-0.09097	0.045544	-2.00	4.58e-02
invest	-0.36937	0.139398	-2.65	8.06e-03
numst2	0.51464	0.129233	3.98	6.83e-05
eltime2	0.72316	0.134999	5.36	8.47e-08
caretk2	-1.30035	0.259566	-5.01	5.45e-07

Scale fixed at 1

Exponential distribution

Loglik(model)= -1025.6    Loglik(intercept only)= -1100.7

Chisq= 150.21 on 7 degrees of freedom, p= 0

Number of Newton-Raphson Iterations: 4

n= 314

# Exponential Model (hazard form)

```
> KABL.exp.PH<-(-KABL.exp.AFT$coefficients)
```

```
> KABL.exp.PH
```

(Intercept)	fract	polar	format	invest
-3.724598700	0.001163784	0.016098468	0.090965318	0.369367997
numst2	eltime2	caretk2		
-0.514643548	-0.723161401	1.300349770		

# Exponential: Hazard Ratios

```
> KABL.exp.HRs<-exp(-KABL.exp.AFT$coefficients)
```

```
> KABL.exp.HRs
```

(Intercept)	fract	polar	format	invest	numst2
0.02412278	1.00116446	1.01622875	1.09523102	1.44681993	0.59771361
eltime2	caretk2				
0.48521587	3.67058030				

# Hazard Ratios: Interpretation

- On average, an investiture requirement *increases* the *hazard* of cabinet failure by  $100 \times (1.447 - 1) = 44.7$  percent.
- On average, an investiture requirement *decreases* the predicted *survival* time by
$$100 \times [1 - \exp(-0.369)] = 100 \times (1 - 0.691)$$
$$= 30.1 \text{ percent.}$$

# Comparing Predicted Survival

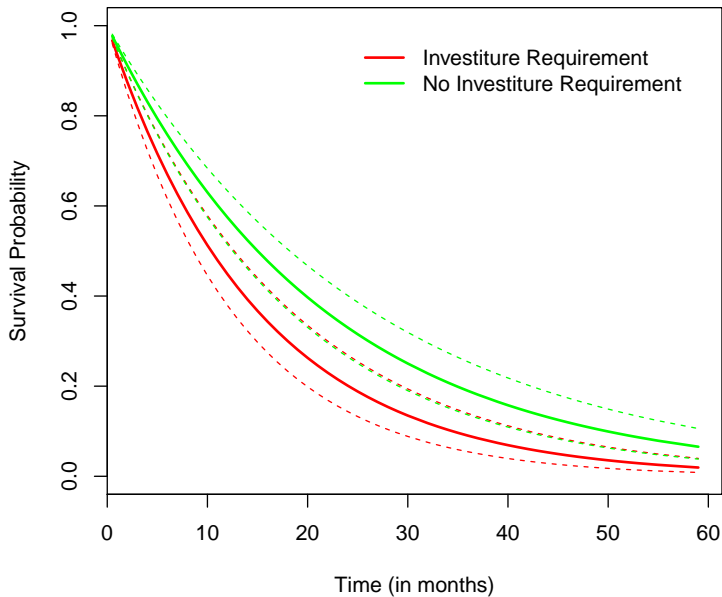
Can use predict, or...

```
KABL.exp<-flexsurvreg(MODEL,data=KABL,dist="exp")

FakeInvest<-t(c(mean(KABL$fract),mean(KABL$polar),mean(KABL$format),1,
               mean(KABL$numst2),mean(KABL$eltime2),mean(KABL$caretk2)))
FakeNoInvest<-t(c(mean(KABL$fract),mean(KABL$polar),mean(KABL$format),0,
                  mean(KABL$numst2),mean(KABL$eltime2),mean(KABL$caretk2)))

plot(KABL.exp,FakeInvest,mark.time=FALSE,col.obs="black",
     lty.obs=c(0,0,0),xlab="Time (in months)",ylab="Survival Probability")
lines(KABL.exp,FakeNoInvest,mark.time=FALSE,col.obs="black",
      lty.obs=c(0,0,0),col=c(rep("green",times=3)))
```

# Comparing Predicted Survival



# The Weibull Model, I

$$h(t) = \lambda p(\lambda t)^{p-1}$$

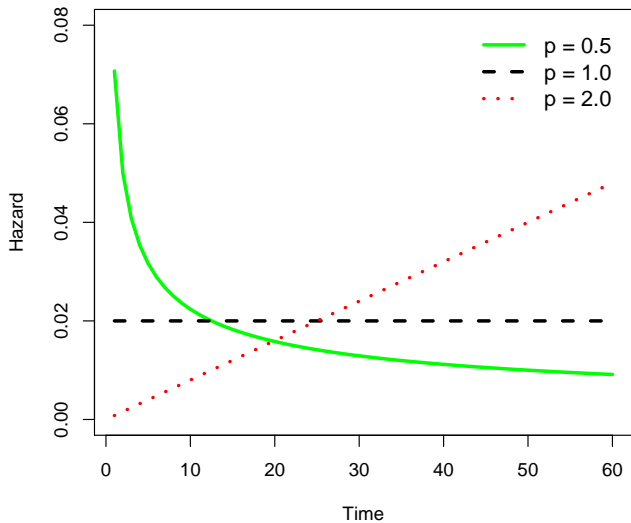
$$\begin{aligned} S(t) &= \exp \left[ - \int_0^t \lambda p(\lambda t)^{p-1} dt \right] \\ &= \exp(-\lambda t)^p \end{aligned}$$

$$f(t) = \lambda p(\lambda t)^{p-1} \times \exp(-\lambda t)^p$$

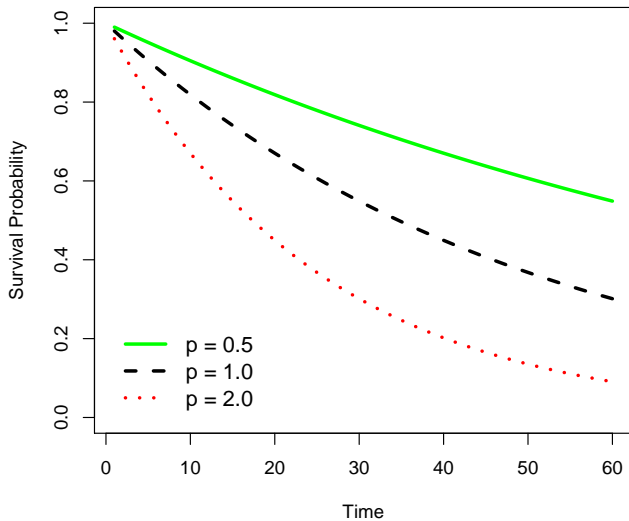


- $p = 1 \rightarrow$  exponential model
- $p > 1 \rightarrow$  rising hazards
- $0 < p < 1 \rightarrow$  declining hazards

# Weibull Hazards Illustrated



# Weibull Survival



$$\lambda_i = \exp(\mathbf{X}_i\beta)$$

$$T_i = \exp(\mathbf{X}_i\gamma) \times \sigma u_i$$

Means:

$$\rho = 1/\sigma$$

$$\beta = -\gamma/\sigma$$

# Weibull Example (AFT)

```
> KABL.weib.AFT<-survreg(MODEL,data=KABL,dist="weibull")  
> summary(KABL.weib.AFT)
```

Call:

```
survreg(formula = MODEL, data = KABL, dist = "weibull")
```

	Value	Std. Error	z	p
(Intercept)	3.69641	0.491590	7.52	5.51e-14
fract	-0.00106	0.000705	-1.50	1.33e-01
polar	-0.01508	0.004677	-3.22	1.26e-03
format	-0.08675	0.035133	-2.47	1.35e-02
invest	-0.33019	0.106991	-3.09	2.03e-03
numst2	0.46352	0.100367	4.62	3.87e-06
eltime2	0.66381	0.104265	6.37	1.93e-10
caretk2	-1.31758	0.201065	-6.55	5.64e-11
Log(scale)	-0.26079	0.049971	-5.22	1.80e-07

Scale= 0.77

Weibull distribution

Loglik(model)= -1013.5    Loglik(intercept only)= -1100.6

Chisq= 174.23 on 7 degrees of freedom, p= 0

Number of Newton-Raphson Iterations: 5

n= 314

# Weibull Example (hazard)

```
> KABL.weib.PH<-(-KABL.weib.AFT$coefficients)/(KABL.weib.AFT$scale)
```

```
> KABL.weib.PH
```

(Intercept)	fract	polar	format	invest
-4.797770943	0.001374065	0.019573990	0.112598478	0.428574214
numst2	eltime2	caretk2		
-0.601628072	-0.861597589	1.710156135		

# Weibull Hazard Ratios

```
> KABL.weib.HRs<-exp(KABL.weib.PH)
```

```
> KABL.weib.HRs
```

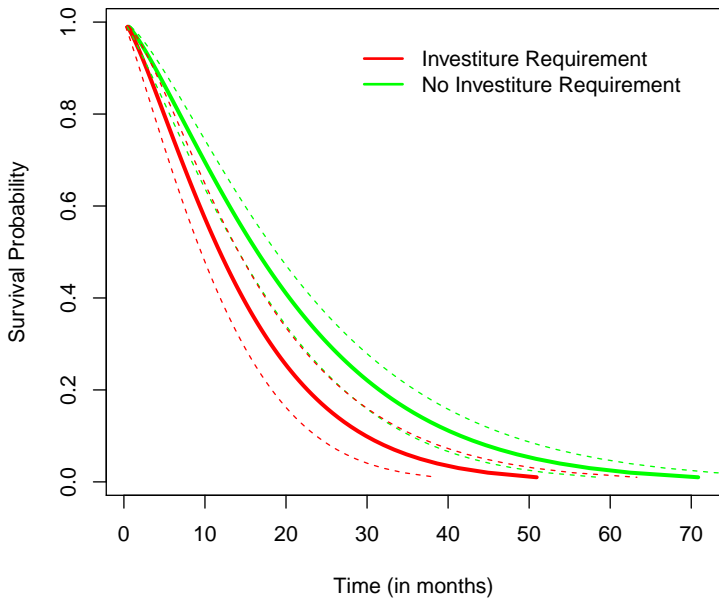
(Intercept)	fract	polar	format	invest	numst2
0.008248112	1.001375009	1.019766817	1.119182466	1.535067285	0.547918858
eltime2	caretk2				
0.422486583	5.529824807				

Interpretation:

- On average, an investiture requirement *increases* the *hazard* of cabinte failure by  $100 \times (1.535 - 1) = 53.5$  percent.



# Comparing Predicted Survival Curves



# The Gompertz Model (hazard)

$$h(t) = \exp(\lambda) \exp(\gamma t)$$

$$S(t) = \exp \left[ -\frac{e^\lambda}{\gamma} (e^{\gamma t} - 1) \right]$$

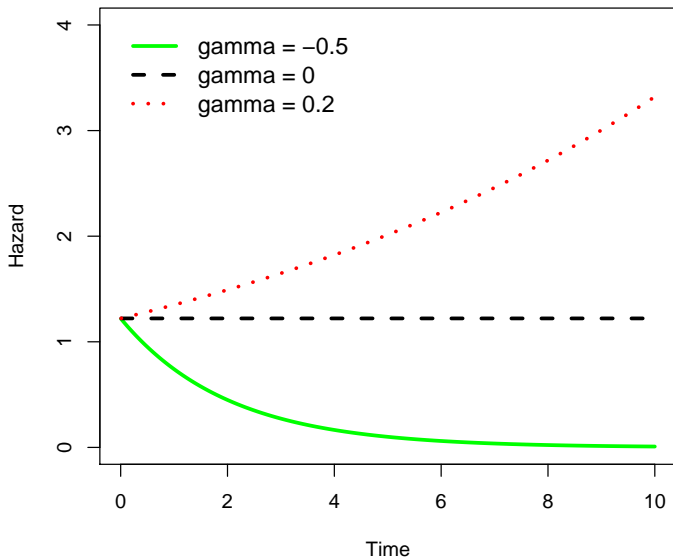
with

$$\lambda_i = \exp(\mathbf{X}_i \beta)$$

$\gamma$  is for “Gompertz”

- $\gamma = 0 \rightarrow$  constant hazard
- $\gamma > 0 \rightarrow$  rising hazard
- $\gamma < 0 \rightarrow$  declining hazard

# Gompertz Hazards



# Gompertz Estimates

```
> library(flexsurv)
> KABL.Gomp<-flexsurvreg(MODEL,data=KABL,dist="gompertz")
> KABL.Gomp
```

```
Call:
flexsurvreg(formula = MODEL, data = KABL, dist = "gompertz")
```

Estimates:

	data	mean	est	L95%	U95%	exp(est)	L95%	U95%
shape		NA	0.02320	0.01150	0.03490	NA	NA	NA
rate		NA	0.01520	0.00407	0.05680	NA	NA	NA
fract	719.00000		0.00140	-0.00039	0.00319	1.00000	1.00000	1.00000
polar	15.30000		0.01890	0.00666	0.03120	1.02000	1.01000	1.03000
format	1.90000		0.10700	0.01590	0.19800	1.11000	1.02000	1.22000
invest	0.45200		0.41200	0.13700	0.68600	1.51000	1.15000	1.99000
numst2	0.63100		-0.60800	-0.86800	-0.34900	0.54400	0.42000	0.70500
eltime2	0.48700		-0.87300	-1.15000	-0.59400	0.41800	0.31600	0.55200
caretk2	0.05410		1.46000	0.94500	1.98000	4.32000	2.57000	7.24000

N = 314, Events: 271, Censored: 43

Total time at risk: 5789.5

Log-likelihood = -1018.317, df = 9

AIC = 2054.634

# The Log-Logistic Model

$$\ln(T_i) = \mathbf{X}_i\beta + \sigma\epsilon_i$$

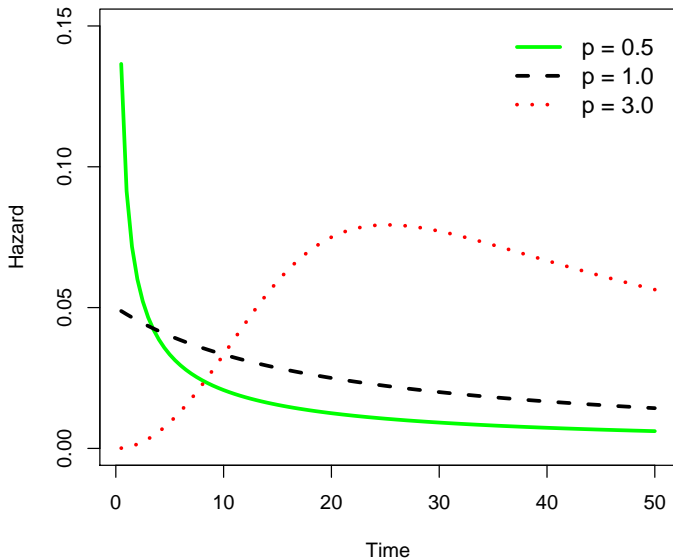
$$S(t) = \frac{1}{1 + (\lambda t)^p}$$

$$h(t) = \frac{\lambda p(\lambda t)^{p-1}}{1 + (\lambda t)^p}$$

$$\begin{aligned} f(t) &= \frac{\lambda p(\lambda t)^{p-1}}{1 + (\lambda t)^p} \times \frac{1}{1 + (\lambda t)^p} \\ &= \frac{\lambda p(\lambda t)^{p-1}}{[1 + (\lambda t)^p]^2} \end{aligned}$$

$$\lambda_i = \exp(\mathbf{X}_i\beta)$$

# Log-Logistics Illustrated



## Example: Log-Logistic

```
> KABL.loglog<-survreg(MODEL,data=KABL,dist="loglogistic")
> summary(KABL.loglog)
```

Call:

```
survreg(formula = MODEL, data = KABL, dist = "loglogistic")
```

	Value	Std. Error	z	p
(Intercept)	3.333841	0.54735	6.09	1.12e-09
fract	-0.000913	0.00079	-1.15	2.48e-01
polar	-0.019092	0.00588	-3.24	1.18e-03
format	-0.096975	0.04315	-2.25	2.46e-02
invest	-0.357403	0.12876	-2.78	5.51e-03
numst2	0.479507	0.12104	3.96	7.45e-05
eltime2	0.627837	0.12405	5.06	4.16e-07
caretk2	-1.252349	0.23151	-5.41	6.32e-08
Log(scale)	-0.568276	0.05116	-11.11	1.14e-28

Scale= 0.567

Log logistic distribution

Loglik(model)= -1024    Loglik(intercept only)= -1099

Chisq= 150.05 on 7 degrees of freedom, p= 0

Number of Newton-Raphson Iterations: 4

n= 314



## Other Parametric Survival Models

- Log-Normal
- Rayleigh (Weibull w/ $p = 2$ )
- Logistic
- $t$
- Generalized Gamma

## R:

- `survreg` (in `survival`)
- `rms` package
- `flexsurv` package
- `eha` package
- `SurvRegCensCov` package (Weibull models)

## Notes on parametric models with time-varying covariate data:

- Stata handles time-varying data with `aplomb`.
- R does not.
  - `survreg` (in the `survival` package) will not estimate models with time-varying data (it will not take a survival object of the form `Surv(start,stop,censor)`).
  - `psm` (in the `rms` package) will also not accept time-varying data.
  - `aftreg` and `phreg` (part of the `eha` package) will accept time-varying data. `phreg` accepts survival objects of the form `Surv(start,stop,censor)`. `aftreg` does as well, and notes in its documentation that “(I)f there are [sic] more than one spell per individual, it is essential to keep spells together by the `id` argument. This allows for time-varying covariates.” In practice, this functions somewhat inconsistently.
- Recommendations: If you want to use R to fit parametric survival models with time-varying covariate data, stick with proportional hazards formulations, and use `phreg`. Also, Weibull models tend to be easier to fit than exponentials in this framework.