Stats Final Exam Cheat Sheet

Gamma function

$\Gamma(z) = \int_0^\infty t^{z-1}e^{-t}dt$; for $\mathcal{R}(z) > 0$

$\Gamma(z+1) = z\Gamma(z)$; $\Gamma(n+1) = n!$ for $n \in \mathbb{N}$

Gamma distribution

$f_Y(y) = \frac{\beta^\alpha}{\Gamma(\alpha)}y^{\alpha-1}e^{-\beta y}$; $y, \alpha, \beta > 0$

$E[Y] = \alpha/\beta$; $Var(Y) = \alpha/\beta^2$

$MGF(t) = \left(1 - \frac{t}{\beta}\right)^{-\alpha}$ for $t < \beta$

$\chi^2(k) \sim \text{Gamma}(k/2, 1/2)$

Y_i indip. $\text{Gamma}(\alpha_i, \beta) \Rightarrow \sum_{i=1}^n Y_i \sim \text{Gamma}(\sum_{i=1}^n \alpha_i, \beta)$

$k * \text{Gamma}(\alpha, \beta) = \text{Gamma}(\alpha, \beta/k)$

Inverse Gamma distribution

$X \sim \text{Gamma}(\alpha, \beta)$ $Y = 1/X \sim$ I-Gamma(α, β)

$f_Y(y) = \frac{\beta^\alpha}{\Gamma(\alpha)}y^{-\alpha-1}e^{-\beta/y}$; $y, \alpha, \beta > 0$

$E[Y] = \beta/(\alpha - 1)$; $Var(Y) = \beta^2/[(\alpha-1)^2(\alpha-2)]$

$\text{Mode} = \beta/(\alpha + 1)$

Inverse Gaussian distribution

$Y \sim$ I-G(ϕ, λ)

$f_Y(y) = \frac{\sqrt{\lambda}}{\sqrt{2\pi}}e^{\sqrt{\lambda\phi}}y^{-3/2}\exp\{-\frac{1}{2}\left(\phi y + \frac{\lambda}{y}\right)\}$;

$y, \phi \geq 0, \lambda > 0$

$E[Y] = \left(\frac{\lambda}{\phi}\right)^{1/2}$; $Var(Y) = \sqrt{\frac{\lambda}{\phi^3}}$

Normal Inverse Gamma distribution

$\theta \sim$ N-IGa$(\eta_0, \lambda_0, \alpha_0, \beta_0)$

$\sigma^2 \sim$ IGa(α_0, β_0)

$\mu|\sigma^2 \sim$ N$(\eta_0, \sigma^2/\lambda_0)$

$f(\mu, \sigma^2) \propto \left(\frac{1}{\sigma^2}\right)^{\alpha_0+3/2}\exp\{\lambda_0\eta_0\frac{\mu}{\sigma^2} - \frac{1}{2\sigma^2}(\lambda_0\eta_0^2 + 2\beta_0) - \lambda_0\frac{\mu^2}{2\sigma^2}\}$

Posterior params (from Normal model s.r.s.):

$\alpha = \alpha_0 + \frac{n}{2}$, $\lambda = \lambda_0 + n$

$\eta = \frac{\lambda_0\eta_0 + n\bar{y}}{\lambda_0 + n}$

$\beta = \beta_0 + \frac{n}{2}\upsilon^2 + \frac{n\lambda_0}{2(\lambda_0+n)}(\bar{y} - \eta_0)^2$

$\upsilon = n^{-1}\sum_{i=1}^n(y_i - \bar{y})^2$

Beta distribution

$f_Y(y) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)}y^{\alpha-1}(1-y)^{\beta-1}$; $\alpha, \beta > 0, y \in [0,1]$

$E[Y] = \frac{\alpha}{\alpha+\beta}$; $Var(Y) = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$

Multivariate Normal distribution

$f_Y(y) = \frac{\det(A)^{1/2}e^{-\frac{1}{2}b^TAb}}{(2\pi)^{d/2}}\exp\{-\frac{1}{2}(y^TAy) + b^Ty\}$

$E[Y] = A^{-1}b$; $MGF(t) = \exp\{\mu t + \sigma^2 t^2/2\}$

$Y_1|Y_2 = y_2 \sim N(\mu_1 + \Sigma_{12}\Sigma_{22}^{-1}(y_2 - \mu_2); \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21})$

$X|\mu \sim N(\mu, \sigma^2)$ and $\mu \sim N(\gamma, \tau^2)$

$X \sim N(\gamma, \sigma^2 + \tau^2)$

t-Student distribution

$f_Y(y) = \frac{\Gamma(\frac{\nu+1}{2})}{\sqrt{\nu\pi}\Gamma(\frac{\nu}{2})}\left(1 + \frac{y^2}{\nu}\right)^{-(\nu+1)/2}$

$\nu > 0$

Uniform distribution

$Y \sim Unif(a, b)$

$f_Y(y) = \frac{1}{b-a}$ for $y \in [a, b]$ otherwise 0

$F_Y(y) = \frac{y-a}{b-a}$ for $y \in [a, b]$

$E[Y] = \frac{1}{2}(a + b)$; $Var(Y) = \frac{1}{12}(b - a)^2$

Hypergeometric distribution

$f_Y(y) = \frac{\binom{K}{y}\binom{N-K}{n-y}}{\binom{N}{n}}$

$E[Y] = n\frac{K}{N}$; $Var(Y) = n\frac{N-K}{N}\frac{K}{N}\frac{N-n}{N-1}$

Negative Binomial distribution

$Y = $ "n. of failures till r-th success"

$f_Y(k) = \binom{k+r-1}{k}(1-p)^kp^r$; $\frac{\Gamma(k+r)}{k!\Gamma(r)}$

$E[Y] = \frac{r(1-p)}{p}$; $Var(Y) = \frac{r(1-p)}{p^2}$

Weibull distribution

$Y \sim Wei(\gamma, \beta)$ $y, \beta, \gamma > 0$

$f(y) = \gamma\beta^{-\gamma}y^{\gamma-1}e^{-(y/\beta)^\gamma}$

$E[Y] = \beta\Gamma(1 + 1/\gamma)$

Series

$\sum_{i=0}^n r^i = \frac{1-r^{n+1}}{1-r}$ for $|r| < 1$

$e^x = \sum_{n=0}^\infty x^n/n! = \lim_{n\to\infty}\left(1 + \frac{x}{n}\right)^n$

$\sum_{i=0}^n \frac{1}{x^p}$ for $p > 1$

Log base change

$\log_a(x) = \frac{\log_b(x)}{\log_b(a)}$

Derivatives

$\frac{d}{dx}\sin(x) = \cos(x)$; $\frac{d}{dx}\cos(x) = -\sin(x)$

$\frac{d}{dx}\arcsin(x) = \frac{1}{\sqrt{1-x^2}}$; $\frac{d}{dx}\arccos(x) = -\frac{1}{\sqrt{1-x^2}}$

$\frac{d}{dx}\tan(x) = \frac{1}{\cos^2(x)}$; $\frac{d}{dx}\arctan(x) = \frac{1}{1+x^2}$

Integrals

$\int x^k dx = \frac{x^{k+1}}{k+1} + C$ $k \neq 1$; $\int \frac{1}{x}dx = \log(|x|) + C$

$\int \tan(x)dx = \log(\frac{1}{\cos(x)}) + C$

$\int \frac{1}{a^2+x^2}dx = \frac{1}{a}\arctan(\frac{x}{a}) + C$

$\int \frac{1}{\sqrt{a^2+x^2}}dx = \arcsin(\frac{x}{a}) + C$

$\int_0^\infty y^\alpha e^{-ky}dy < \infty$ if $\alpha > -1$ $(k > 0)$

Taylor expansions

$e^x = \sum_{i=0}^k \frac{x^i}{i!} + o(x^k)$ for $x \to 0$

$\log(1 + x) = \sum_{i=1}^k(-1)^{i-1}\frac{x^i}{i} + o(x^k)$ for $x \to 0$

$\sin(x) = \sum_{i=0}^n \frac{(-1)^i x^{2i+1}}{(2i+1)!} + o(x^{2n})$

$\cos(x) = \sum_{i=0}^n \frac{(-1)^i x^{2i}}{(2i)!} + o(x^{2n})$

Matrix

$$X = \begin{pmatrix} A & B \\ C & D \end{pmatrix}$$

If X is a 2×2 matrix:

$X^{-1} = \frac{1}{\det X}\begin{pmatrix} D & -B \\ -C & A \end{pmatrix}$

X is positive definite if $A > 0$ and $\det X > 0$

If X is a block Matrix X^{-1} has elements:

$(A - BD^{-1}C)^{-1}$

$-A^{-1}B(D - CA^{-1}B)^{-1}$

$-(D - CA^{-1}B)^{-1}CA^{-1}$

$(D - CA^{-1}B)^{-1}$

$i^{\psi\psi} = (i_{\psi\psi} - i_{\psi\lambda}(i_{\lambda\lambda})^{-1}i_{\lambda\psi})^{-1}$

$i^{\psi\lambda} = -i^{\psi\psi}i_{\psi\lambda}i_{\lambda\lambda}^{-1}$

Order Statistics

i.i.d continuous case.

$f_{Y_{(1)},...,Y_{(n)}}(y_1, ., y_n) = n!f(y_1)f(y_2)...f(y_n)$

$f_{Y_{(r)}}(y_r) = \frac{n!}{(n-r)!(r-1)!}[F(y_r)]^{r-1}f(y_r)[1 - F(y_r)]^{n-r}$

$f_{Y_{(r)}Y_{(s)}}(y, z) = \frac{n!}{(n-s)!(s-r-1)!(r-1)!}[F(y)]^{r-1}f(y)[F(z) - F(s)]^{s-r-1}f(z)[1 - F(z)]^{n-r}$

Profile pivot

Likelihood profile pivot $\theta = (\psi, \lambda)$

$W_{P_e}(\psi) = (\hat{\psi} - \psi)^T j_P(\hat{\psi})(\hat{\psi} - \psi)$

$W_{P_u}(\psi) = U_P(\psi)^T i_P^{\psi\psi}(\psi, \hat{\lambda_\psi})U_P(\psi)$

we can use $j^{\psi\psi}(\theta) = (j_{\psi\psi} - j_{\psi\lambda}(j_{\lambda\lambda})^{-1}j_{\lambda\phi})^{-1}$

Confidence intervals and Test

Binomial: for $n\theta, n(1 - \theta) > 5$

$q_2(\theta, \hat{\theta}) = \frac{\hat{\theta} - \theta}{\sqrt{\frac{\theta(1-\theta)}{n}}}$

$\frac{\hat{\theta} + (z_{1-\alpha/2})^2/2n}{1 + (z_{1-\alpha/2})^2/n} \pm \frac{z_{1-\alpha/2}/\sqrt{n}}{1 + (z_{1-\alpha/2})^2/n}\sqrt{\hat{\theta}(1 - \hat{\theta}) + \frac{(z_{1-\alpha/2})^2}{4n}}$

Clopper Pearson:

$\{\theta \in (0, 1) : \mathbb{P}_\theta(Y \leq y) \geq \alpha/2$ and $\mathbb{P}_\theta(Y \geq y) \geq \alpha/2\}$

Exact distribution test: $n\hat{\theta} \sim Bi(n, \theta)$

Mid-pvalue: $\alpha_{mid}^{oss} = \sup_{\theta \in \Theta_0}\{\mathbb{P}_\theta(t(Y) > t(y)) + \frac{1}{2}\mathbb{P}_\theta(t(Y) = t(y))\}$

Netwon - Raphson

$x_{n+1} = x_n - \frac{f(x_n)}{f'(x_n)}$

Total Variance

$Var(Y) = E[Var(Y|X)] + Var(E[Y|X])$

Delta Method

$\sqrt{n}(Y_n - \theta) \xrightarrow{D} U$

$\sqrt{n}(g(Y_n) - g(\theta)) \xrightarrow{D} g'(\theta)U$

Posterior Predicitive Density

if $(y^*|y, \theta) = (y^*|\theta)$
$p(y^*|y) = \int p(y^*|\theta)p(\theta|y)d\theta$

Another topic

some stuff