Learning from data

Classification

February 3, 2020

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Leads us to the classification setting.

What are we up to?

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- n individuals each is associated with certain quantitative measurements called the predictors or features, say the p measurements of i-th individual are given as (x_{i1}, x_{i2},...,x_{ip})
- Each individual is also categorised using a qualitative variable: dependent variable
- ► Goal: given predictor measurements (x₁, x₂,...,x_p) of certain individual, we wish to model y, the category this individual belongs to Classification problem

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$$y=\mathbf{1}_{\left\{\left(X_{1},\ldots,X_{p}\right)\in B\right\}}$$

or, in general, if y takes J distinct values (a_1,\ldots,a_J) then we can set

$$y = a_1 1_{\{(X_1,...,X_p) \in B_1\}} + \cdots + a_J 1_{\{(X_1,...,X_p) \in B_J\}}$$

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Issues with the proposed solution:

- y could be a qualitative variables, and thus the above method would fail.
- How does one incorporate statistical error (noise)?

Binary classifier: logistic regression

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Writing *X* for the variable "Credit.Amount", the binary classifier is based on modelling:

$$\mathbb{P}(Y=1|X=x)$$

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In other words,

$$p(x) = \frac{e^{\beta_0 + \beta_1 x}}{1 + e^{\beta_0 + \beta_1 x}}$$

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Method: Maximum likelihood estimation

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The likelihood for this training sample is:

$$L(\beta_0, \beta_1) = \prod_{i \in \mathcal{Y}_1} p(x_i) \prod_{i \in \mathcal{Y}_0} [1 - p(x_i)]$$

The maximum likelihood estimates are those $\widehat{\beta_0}$ and $\widehat{\beta_1}$, which maximise $L(\beta_0, \beta_1)$



Finding those $\widehat{\beta_0}$ and $\widehat{\beta_1}$, which maximise

$$L(\beta_0, \beta_1) = \prod_{i \in \mathcal{Y}_1} \frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \prod_{i \in \mathcal{Y}_0} \frac{1}{1 + e^{\beta_0 + \beta_1 x_i}}$$

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Remark: Likelihood estimation is a very general approach, applicable to almost every setting of estimation. Least squares estimation in linear regression is a special case of likelihood estimation.

Interpretation of $\widehat{\beta_1}$: A unit increase in x, changes the odds by a factor of $e^{\widehat{\beta_1}}$.

Prediction

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- Compute

$$\mathbb{P}(Y=1|X=x_0)=p(x_0)=\frac{e^{\widehat{\beta_0}+\widehat{\beta_1}x_i}}{1+e^{\widehat{\beta_0}+\widehat{\beta_1}x_i}}$$

Prediction

- ► Given a new value of x₀ (predictor), we wish to now classify the individual into one of the categories.
- Compute

$$\mathbb{P}(Y = 1 | X = x_0) = p(x_0) = \frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}}$$

▶ If $p(x_0) > 0.5$ (say), then we classify x_0 into the category Y = 1, else into the category Y = 0.

Note: The thresholding at 0.5 is our choice. If we have some prior information then we may want to tweak this a little.



Testing: statistical significance of the model

The hypothesis of interest:

$$H_0$$
 : $\beta_1 = 0$
 H_1 : $\beta_1 \neq 0$

- ► The estimators are tested using the standard Wald's *Z*-test.
- Or, one could also use the likelihood ratio (LR) test.

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Testing: goodness of the model

- The testing phase includes predicting the classification for given set of x_i's, for which the true classification is already know.
- ► We count the number of 1's which are correctly identified as 1, and the number of 1's which are incorrectly identified as 0.
- Similarly, we count the number of 0's which are correctly identified as 0, and the number of 0's which are incorrectly identified as 1.



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For instance, if the task is to determine whether an individual patient suffered a stroke or not, given only the measurement of body temperature. It is important to first find out if there's any statistically significant effect of stroke on body temperature.

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For instance, if the task is to determine whether an individual patient suffered a stroke or not, given only the measurement of body temperature. It is important to first find out if there's any statistically significant effect of stroke on body temperature.

We apply the standard 2-sample *t*-test.

The variables:

```
> names(G.credit)
```

- [1] "Creditability"
- [3] "Duration.of.Credit..month."
- [5] "Purpose"
- [7] "Value.Savings.Stocks"
- [9] "Instalment.per.cent"
- [11] "Guarantors"
- [13] "Most.valuable.available.asset"
- [15] "Concurrent.Credits"
- [17] "No.of.Credits.at.this.Bank"
- [19] "No.of.dependents"
- [21] "Foreign.Worker"

- "Account.Balance"
 "Payment.Status.of.Previous.Credit"
- "Credit.Amount"
- "Length.of.current.employment"
- "Sex...Marital.Status"
- "Duration.in.Current.address"
- "Age..years."
- "Type.of.apartment"
- "Occupation"
- "Telephone"

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> names(G.credit)
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```

We are going to focus on the predictor "Credit.Amount" and the response as "Creditability".

First, let us test if there is indeed any effect of "Creditability" on the "Credit.Amount":

This is Welch's *t*-test which is a slight modification of the standard two sample *t*-test.

Some more evidence of effect of "Creditability" on the "Credit.Amount":

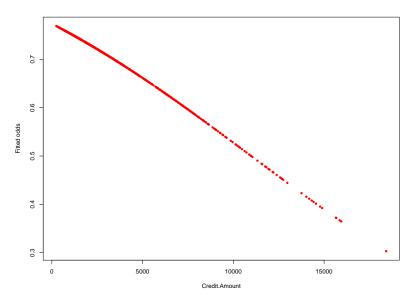
Some more evidence of effect of "Creditability" on the "Credit.Amount":

Conclusion: Indeed creditable group has a different distribution of "Credit.Amount" as compared to that of the non-creditable group. Implying that a binary classification indeed is a good idea.

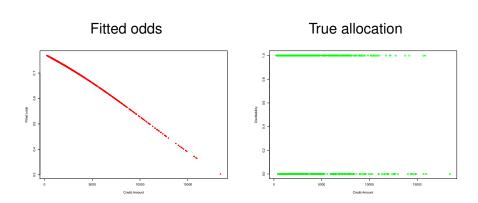
Logistic regression: Creditability ~ Credit.Amount

```
> library(qmodels)
> creditability.on.cramount <- glm(Creditability~Credit.Amount,data=G.credit,
                                  family=binomial)
> summary(creditability.on.cramount)
Call:
glm(formula = Creditability ~ Credit.Amount, family = binomial,
   data = G.credit)
Deviance Residuals:
   Min
             10 Median
                               30
                                       Max
-1.7022 -1.3674 0.7688 0.8269 1.4158
Coefficients:
               Estimate Std. Error z value Pr(>|z|)
              1.229e+00 1.083e-01 11.348 < 2e-16 ***
(Intercept)
Credit.Amount -1.119e-04 2.355e-05 -4.751 2.02e-06 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 1221.7 on 999 degrees of freedom
Residual deviance: 1199.1 on 998 degrees of freedom
ATC: 1203.1
Number of Fisher Scoring iterations: 4
```

Predictions



Compare



Compare

Total Observations in Table: 1000

| <pre> prediction.for.CrAm.c</pre> | | | | |
|------------------------------------|--------|---------|-----------|--|
| creditability.f | | | Row Total | |
| | | | | |
| No I | 20 I | 280 I | 300 I | |
| I | 6.667% | 93.333% | 30.000% | |
| | | | | |
| Yes I | 9 1 | 691 I | 700 | |
| 1 | 1.286% | 98.714% | 70.000% | |
| | | | | |
| Column Total I | 29 | 971 I | 1000 I | |
| | | | | |

Overal, the prediction seems to work fairly well, in that the classifier could correctly predict 71.1% times.

Specifically, the model seems to perform very well when predicting 1s, but it performs poorly in predicting the 0s.



Multiple logistic regression

Consider the same dataset, but now we consider two predictors: "Credit.Amount" and "Duration.of.credit..month.".

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More generally, we could be dealing with many more predictors, but that's not an issue!

Say, we have measurements on *p* predictors or features, and we wish to classify each point into one or the other group (binary classification).

We shall again be interested in modelling:

$$p(x_1,...,x_p) = \mathbb{P}(Y = 1 | X_1 = x_1,...,X_p = x_p)$$

Once again, we propose:

$$\log\left(\frac{p(X_1,\ldots,X_p)}{1-p(X_1,\ldots,X_p)}\right)=\beta_0+\beta_1X_1+\cdots+\beta_pX_p$$

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$$\log\left(\frac{p(X_1,\ldots,X_p)}{1-p(X_1,\ldots,X_p)}\right)=\beta_0+\beta_1X_1+\cdots+\beta_pX_p$$

<u>Estimation</u>: maximum likelihood estimation, i.e., we look for those values of $\beta_0, \beta_1, \dots, \beta_p$ which maximise

$$L(\beta_0, \beta_1, \dots, \beta_p) = \prod_{i \in \mathcal{Y}_1} p(x_{1,i}, \dots, x_{p,i}) \prod_{i \in \mathcal{Y}_0} [1 - p(x_{1,i}, \dots, x_{p,i})]$$

Few remarks:

Interpretation of $\widehat{\beta}_i$ for $i=1,\ldots,p$: Keeping all other variables constant, a unit increase in the *i*-th feature would change the odds by a factor of $e^{\widehat{\beta}_i}$.

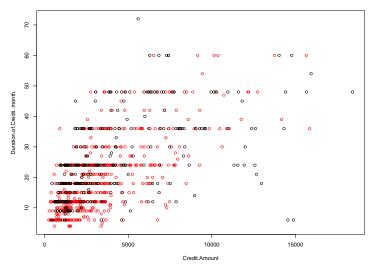
Few remarks:

- Interpretation of $\hat{\beta}_i$ for $i=1,\ldots,p$: Keeping all other variables constant, a unit increase in the *i*-th feature would change the odds by a factor of $e^{\hat{\beta}_i}$.
- ► It is desirable that different features give indeed different information ⇒ multicollinearity.

Few remarks:

- Interpretation of $\hat{\beta}_i$ for $i=1,\ldots,p$: Keeping all other variables constant, a unit increase in the *i*-th feature would change the odds by a factor of $e^{\hat{\beta}_i}$.
- It is desirable that different features give indeed different information ⇒ multicollinearity. This is again tested by the same VIF

The plot for the three variables looks like:



Multiple Logistic Regression: output

```
> summary(creditability.on.cramdur)
Call:
alm(formula = Creditability ~ Credit.Amount + Duration.of.Credit..month..
   family = binomial, data = G.credit)
Deviance Residuals:
   Min
             10 Median
                              30
                                     Max
-1.8249 -1.2734 0.7164 0.8533 1.5020
Coefficients:
                           Estimate Std. Error z value Pr(>|z|)
                        1.670e+00 1.466e-01 11.390 < 2e-16 ***
(Intercept)
Credit.Amount
                       -2.300e-05 3.059e-05 -0.752 0.452
Duration.of.Credit..month. -3.412e-02 7.282e-03 -4.685 2.8e-06 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 1221.7 on 999 degrees of freedom
Residual deviance: 1176.6 on 997 degrees of freedom
ATC: 1182.6
Number of Fisher Scoring iterations: 4
```



Predictions

Total Observations in Table: 1000

| prediction.for.CrAmDur.c | | | | |
|--------------------------|---------|-----------|-------------|--|
| creditability.f | | | Row Total I | |
| | | | | |
| No | I 37 I | 263 I | 300 | |
| | 12.333% | 87.667% I | 30.000% | |
| | - | | | |
| Yes | I 29 I | 671 I | 700 I | |
| | 4.143% | 95.857% I | 70.000% | |
| | - | | | |
| Column Total | | 934 I | 1000 I | |
| | | | | |

Overall efficiency of the model has actually reduced a bit as compared to the logistic model with single predictor "Credit.Amount".

Although the model still performs poorly in predicting the 0s, but it is better than the single predictor model.

However, the model's efficiency is classifying correctly identifying the 1s has considerably reduced.



Comparing the two binary classifiers

Predicting creditability using "Credit.Amount"

Total Observations in Table: 1000

| prediction.for.CrAm.c | | | | |
|-----------------------|----------|---------|-------------|--|
| creditability.f | l No l | Yes | Row Total I | |
| | - | | | |
| No | l 20 l | 280 | 300 I | |
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| | - | | | |
| Column Total | l 29 l | 971 | 1000 | |
| | - | | | |

Predicting creditability using "Credit.Amount" and "Duration"

Total Observations in Table: 1000

| 1 | prediction.fo | or.CrAmDur. | c |
|----------------|---------------|-------------|------|
| | No I | | |
| No I | 37 | 263 | 300 |
| I | 12.333% | | |
| Yes I | | 671 | 700 |
| | | 95.857% | |
| Column Total I | 66 I | 934 | 1000 |
| | | | |

Which one would we choose?

Comparing the two binary classifiers

Predicting creditability using "Credit.Amount"

Predicting creditability using "Credit.Amount" and "Duration"

| Total | Observations | in | Table: | 1000 |
|-------|--------------|----|--------|------|
|-------|--------------|----|--------|------|

| prediction.for.CrAm.c | | | | |
|-----------------------|--------|---------|-----------|--|
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| No I | 37 | l 263 l | 300 I | |
| 1 | 12.333% | 87.667% | 30.000% | |
| | | | | |
| Yes I | 29 | 671 | 700 | |
| 1 | 4.143% | 95.857% | 70.000% I | |
| | | | | |
| Column Total I | 66 | 934 | 1000 | |
| | | | | |

Which one would we choose?

Sensitivity: The ability to correctly identify the actual "positives"

Specificity: The ability to correctly identify the actual "negatives"



ROC curve and AUC

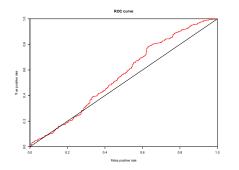
Receiver operating characteristic (ROC) curve, historically, originated from communications theory, as a simple tool to plot true positive rate (sensitivity) and false positive rate (specificity) for various values of the threshold.

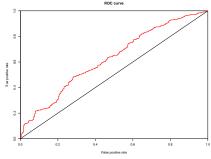
ROC curve and AUC

Receiver operating characteristic (ROC) curve, historically, originated from communications theory, as a simple tool to plot true positive rate (sensitivity) and false positive rate (specificity) for various values of the threshold.

ROC for predicting creditability using "Credit.Amount"

ROC for predicting creditability using "Credit.Amount" and "Duration"





ROC curve and AUC

ROC curve provides compelling visual depiction of the goodness of models.

However, it is often preferable to be able to quote a single, which is the **area under the curve (AUC)**.

This is the area under the ROC curve. A good model should have a high value of AUC (close to 1).

For the two models we considered, the AUC were:

simple logistic: 0.5548 multiple logistic: 0.6257

Another approach: KNN classifier