S. Let
$$W = [W \cdot WA]^T$$

$$= \frac{dr}{k!} \left([W \cdot W] \cdot WA] \left(\frac{1}{J \cdot \Delta k + 1} W_{k-1} \right)^{-1}$$

$$= \frac{dr}{k!} \left([J \cdot \Delta k + 1] W_{k-1} \right)^{-1}$$

$$= \frac{dr}{k!} \left([J \cdot \Delta k + 1] W_{k-1} \right)^{-1}$$

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$$= \frac{dr}{k!} \left([M \cdot M \cdot W_{k-1}] W_{k-1} \right)^{-1}$$

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$$= \frac{dr}{k!} \left([M \cdot M \cdot W_{k-1}] W_{k-1} \right)^{-1}$$

$$= \frac{dr}{k$$

```
7. N (Abels: 4) ..., IN ~ M=0, 02
               N-K
    fraining validation
                                                                                                                     \frac{E\left(\frac{1}{K}\sum_{n=N}^{N}\left(\frac{1}{N^{n-N}}\right)^{2}\right)=\sigma^{2}}{2\frac{1}{K}E\left(\sum_{n=N}^{N}\left(\frac{1}{N^{n}}\right)^{2}=\sigma^{2}}\frac{1}{N^{n}}E\left(\sum_{n=N}^{N}\left(\frac{1}{N^{n}}\right)^{2}=K\sigma^{2}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n}}\frac{1}{N^{n
experted valdation error:
                                                                                                                                        E[(X-E[X])^{L}] = \sigma^{2}
  N-K exemples estimate mean:
                                                                                                                                        To Not Not In
                                                                                                                                               E( { Z (yn-ý)))
Q: espected validation error:
    E( [ [ ] (yn- y) ) = | E [ [ ] (yn= yny + yn )]
                                                                                                      = k \left( E \left[ \sum_{n=N+kn}^{N} y_{n^{2}} \right] - 2 E \left[ \sum_{n=N+kn}^{N} y_{n} \bar{y} \right] + E \left[ \sum_{n=N+kn}^{N} \bar{y}^{2} \right] \right)
= k \left( K \Gamma^{2} - 2 \sum_{n=N+kn}^{N} E \left[ y_{n} \bar{y} \right] + E \left[ \sum_{n=N+kn}^{N} \bar{y}^{2} \right] \right)
                                                                                                                                                                                                                                                                                                                                                         if X, Y: independent
then E[XY] = E(X] E[Y]
                                                                                                           = \( \( \kappa^2 - \rangle \) \( \text{E(yn)} \) \( \frac{\text{E(yn)}}{\text{F(yn)}} \) \( \text{E(yn)} \) \( \frac{\text{E(yn)}}{\text{F(yn)}} \)
                                                                                                                                                                                                                                                                                                                                                                            " values of you range from MK+1 for
                                                                                                                                                                                                                                                                                                                                                                                     and j= wkn=yn depends on yn
                                                                                                                                                                                          (labels generated from distribution with mean = 0)
                                                                                                                                                                                                                                                                                                                                                                                                     that range from 1 to NK
                                                                                                                                                                                                                                                                                                                                                                               1: yn, j: independent
                                                                                                       = K(Koz+ E[Zý]
                                                                                                                                                                                                                                                                                                                                                                         Thus, E[yný] = E[yn] E[x]
                                                                                                                              E[\bar{y}] = 0
E[\bar{y}] = E[Z(\bar{y} - \bar{b}(\bar{y}))^{2}]
= \sum_{n=\mu+c+1}^{\nu} E[(\bar{y} - \bar{b}(\bar{y}))^{2}]
= k. Var(\bar{y})
                                                                                                                                                                                           = K. Vav ( N-1 ( N-1 YN )
                                                                                                                                                                                           = K·(N/c) Var (N/ 4n) ) Bienayme's identity
                                                                                                                                                                                             = K-(NK) = [ NK Var (yn) + Z Z (Go (gi, yi))]
```

ijf [hunnk]

$$= \frac{\left(\frac{1}{N^{2}}\right)^{2} \frac{N^{2}}{N^{2}} \left(\frac{1}{N^{2}}\right)^{2} \frac{N^{2}}{N^{2}} \left(\frac{1}{N^{2}}\right)^{2} \frac{N^{2}}{N^{2}} \left(\frac{1}{N^{2}}\right)^{2} \frac{N^{2}}{N^{2}} \left(\frac{1}{N^{2}}\right)^{2} \frac{N^{2}}{N^{2}} \left(\frac{1}{N^{2}}\right)^{2} \frac{N^{2}}{N^{2}} \left(\frac{1}{N^{2}}\right)^{2} \left(\frac{$$

$$= \sigma^2 \left(\left(+ \frac{1}{N-K} \right) \right)$$

8. MIGHUS 71,..., MN N MC00

Wo: estimate mean

$$E_{in}[w] = \frac{1}{N} \sum_{n=1}^{N} (w_0 - y_n) \qquad w^* = \frac{1}{N} \sum_{n=1}^{N} y_n$$

[Not : training for N7,2, each time we can partition the Nexamples into Provi.

Let Dr be the new training set, which is;

KIso, let got be the hypothesis learned from Dn

7 Let unor on validation ret [(xn,yn)]: en

$$\frac{1}{2} \frac{1}{N w_0^{*}} = \frac{1}{12} \frac$$

$$= \frac{1}{N(N-1)^2} \sum_{n=1}^{N} \left[N w_0^* - j_n - j_n(N-1) \right]^2$$

$$= \frac{N^2}{N(N-1)} > \frac{N \cdot \text{Ein}[vo^*]}{N}$$

$$= \frac{N^2}{(N1)^2} E N \left[w^{0+1} \right]$$

$$= [(Nwo^{3} - yn) - yn(N1)] = 2(Nwo^{3} - yn)yn(N1) + yn(N1)^{2}$$

$$= (Nwo^{3} - yn)^{2} - 2(Nwo^{3} - yn)yn(N1) + yn(N1)^{2}$$

$$= (Nwo^{3} - yn)^{2} - 2(Nwo^{3} - yn)yn(N1) + yn(N1)^{2}$$

$$= (Nwo^{3} - yn) - 2(Nwo^{3} - yn)yn(N1) + yn(N1)^{2}$$

$$= N^{2} \left[(w^{*})^{2} - 2w^{*}y^{*} + 1 \right]$$

9. Finany dassifier g:
$$P(g(x) = 1 \mid y = +1) = \epsilon_{+}$$
$$P(g(x) = +1 \mid y = -1) = \epsilon_{-}$$

$$diffribution: P(y=-1)=p$$

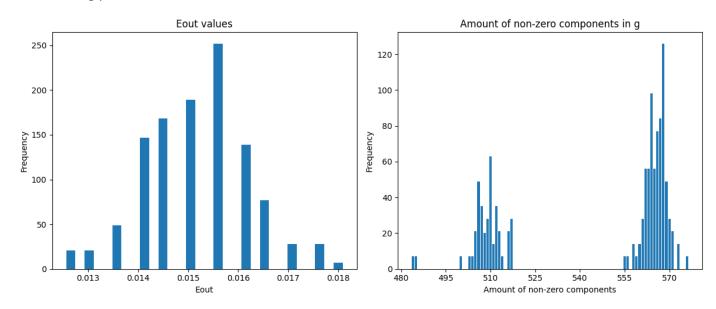
Fork
$$(q) = P(q(x) = -1 | y = +1) P(y = +1) + P(q(x) = +1 | y = -1) P(y = -1)$$

= $(1-p)E_{+} + pE_{-}$

Eout (g) = tent (gi):

HTML homework 5: q10 report

The resulting plot is as follows:



Some additional information that I printed in order to do the comparison between this question and question 11:

```
Statistics for Eout:
Mean: 0.0152
Median: 0.0151
Standard Deviation: 0.0010
Min: 0.0126
Max: 0.0181

Statistics for Non-zero Components:
Mean: 546.5
Median: 563.0
Standard Deviation: 27.2
Min: 484
Max: 576
```

Code¹:

Preprocessing the dataset:

```
1 from liblinear.liblinearutil import *
         2 from itertools import combinations_with_replacement
         3 import numpy as np
         5 import matplotlib.pyplot as plt
        6 from tqdm import tqdm
7 from joblib import Parallel, delayed
Dataset

    mnist.scale for training

    mnist.scale.t for testing

        train_data = 'mnist.scale'
test_data = 'mnist.scale.t'
Empty markdown cell, double-click or press enter to edit.
Read and save data
The training data is saved in the X_train list, with each element being a dictionary of features. example format: {153: 0.0117647, 154: 0.0705882, 155: 0.0705882, ...}
The corresponding label list is saved in y_train.
Similarily, the testing data is saved in X_test and the corresponding label list is saved in y_test.
         1 def read_linear_format(file_path):
               X, y = [], []
with open(file_path, 'r') as f:
                         features = - {
                              for item in parts[1:]:
                                index, value = item.split(":")
features[int(index)] = float(value)
        15  X_train, y_train = read_linear_format(train_data)
16  X_test, y_test = read_linear_format(test_data)
```

Error function:

¹ Most of the screenshots here are from my .ipynb file, since I wrote the explanations here, however, I ran the code in a .py file (which has equivalent contents), and part of the screenshot is from this file.

Explanation of deriving the relationship between C, λ :

Regularized logistic regression Original Formulation in README.md

As the explanation in the liblinear README file, for <u>L1-regularized logistic regression</u> (-s 6), we solve

$$\min_{ec{J}} \sum |w_j| + C \sum \log(1 + \exp(-y_i ec{w}^T ec{x}_i))$$

Components of this equation

In this equation, the <u>L1 regularizer</u> is defined as:

$$\Omega(ec{w}) = \sum |w_j| = ||ec{w}||_1$$

info in Lec14 slide 18

• Note: using L1 regularization encourages sparsity, which means fewer parameters.

And the in-sample error for logistic regression is:

$$E_{in}(ec{w}) = rac{1}{N} \sum \log(1 + \exp(-y_i ec{w}^T ec{x}_i))$$

in textbook p.91

Derivation of C

Since we knew that:

$$E_{aug}(ec{w}) = E_{in}(ec{w}) + rac{\lambda}{N}\Omega(ec{w})$$

We can derive that:

$$E_{aug}(ec{w}) = E_{in}(ec{w}) + rac{\lambda}{N} \sum |w_j|$$

For minimization:

$$\begin{split} \arg\min_{\vec{w}} E_{aug}(\vec{w}) &= \arg\min_{\vec{w}} (E_{in}(\vec{w}) + \frac{\lambda}{N} \sum |w_j|) \\ &= \arg\min_{\vec{w}} \frac{N}{\lambda} (\frac{1}{N} \sum \log(1 + \exp(-y_i \vec{w}^T \vec{x}_i)) + \frac{\lambda}{N} \sum |w_j|) \\ &= \arg\min_{\vec{w}} \frac{1}{\lambda} \sum \log(1 + \exp(-y_i \vec{w}^T \vec{x}_i)) + \sum |w_j| \end{split}$$

Compare to the original equation marked with (\star) , we can see that:

$$C=rac{1}{\lambda}$$

Some additional information:

Select the best λ^*

Selecting optimal λ^* using the following equation:

$$\lambda^* = rg \min_{\log_{10}(\lambda) \in \{-2,-1,0,1,2,3\}} E_{in}(ec{w}_{\lambda})$$

get amount of non-zeros

To get the amount of non-zero components in the model, we can use the function $.get_decfun()$:

[W, b] = model_.get_decfun()

Main code:

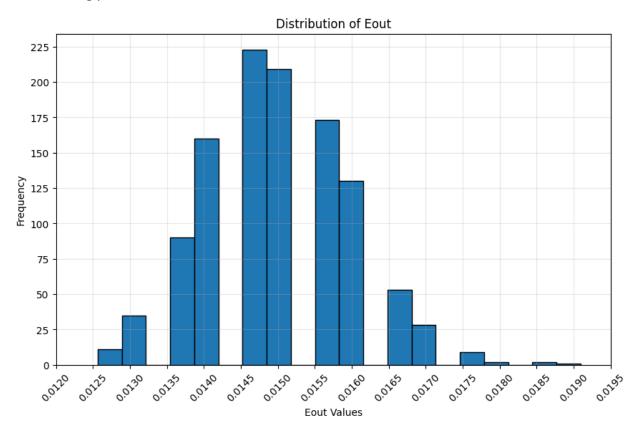
```
def run_single_experiment(experiment):
        np.random.seed(experiment)
        min_Ein = np.inf
        opt_log10_lambda = 0
        for log10_lambda in (-2, -1, 0, 1, 2, 3):
           c = 1 / (10 ** log10_lambda)
            prob = problem(y_train, X_train)
            param = parameter('-s 6 -c ' + str(c))
            model = train(prob, param)
            train_label, _, _ = predict(y_train, X_train, model)
            Ein = ZeroOneError(train_label, y_train)
            if Ein == min_Ein:
                opt_log10_lambda = max(opt_log10_lambda, log10_lambda)
                if opt_log10_lambda == log10_lambda:
                    opt_model = model
            elif Ein < min_Ein:</pre>
                min_Ein = Ein
20
                opt_log10_lambda = log10_lambda
                opt_model = model
        test_label, _, _ = predict(y_test, X_test, opt_model)
        Eout = ZeroOneError(test_label, y_test)
        W = np.array(opt_model.get_decfun()[0])
        non_zero_count = np.count_nonzero(W)
        return Eout, non_zero_count
   experiment_amount = 10
32 results = Parallel(n_jobs=-1)(
                                                                                        # use all cores by setting n_jobs to -1
        delayed(run_single_experiment)(i) for i in tqdm(range(experiment_amount))
36     Eouts, non_zero_count_list = zip(*results)
37 Eouts = list(Eouts)
38 non_zero_count_list = list(non_zero_count_list)
```

Plotting:

```
‡ aim: plotting
‡ subaim: First subplot (Eout values)
ax1.hist(Eouts, bins=30)
ax1.set_title('Eout values')
ax1.set_xlabel('Eout')
ax1.set_ylabel('Frequency')
# subaim: Second subplot (non-zero components)
min_val = int(min(non_zero_count_list))
max_val = int(max(non_zero_count_list))
integer_bins = np.arange(min_val, max_val + 2) - 0.5 # +2 to include max_val, -0.5 for bin edges
ax2.hist(non_zero_count_list,
         bins=integer_bins,
         align='mid',
         rwidth=0.8)
ax2.xaxis.set_major_locator(plt.MaxNLocator(integer=True))
ax2.set_title('Amount of non-zero components in g')
ax2.set_xlabel('Amount of non-zero components')
ax2.set_ylabel('Frequency')
plt.tight_layout()
plt.show()
```

HTML homework 5: q11 report

The resulting picture is as follows:



In order to compare with the previous question, some statistics are:

Statistics for Eout: Mean: 0.0152 Median: 0.0151

Standard Deviation: 0.0010

Min: 0.0126 Max: 0.0181

Statistics for Non-zero Components:

Mean: 546.5 Median: 563.0

Standard Deviation: 27.2

Min: 484 Max: 576

Mean: 0.0150 Median: 0.0151

Min: 0.0126

Max: 0.0191

Statistics for Eout:

Standard Deviation: 0.0010

¹ Left: statistics of q11 result; Right: statistics of q10 result

In this problem, the difference from the previous is that we further divide the training set into a subtrain set (with 8000 exampls), and a validation set (3876 examples).

From the statistics, we can see that most of them are quite similar, I think part of the reason is that, even though 3876 examples are taken out from the original training set, so one may expect that as we train on a smaller subtrain set, the out of sample error should be bigger than in q10 while we have the whole training set used, we got an expected value slightly smaller because we retrain the whole training set after the optimal λ is chosed. We did not use g^- to predict the test set, but g instead.

Apart from this reason, I think this result may also due to the precision, if more digits are taken into consideration, maybe we can observe the difference.

Another cause may be the limited number of possible λ , this constraints the variability of our model, even though there are infinitely many possible models, most of them are alike, while only slightly different by the λ value chosen.

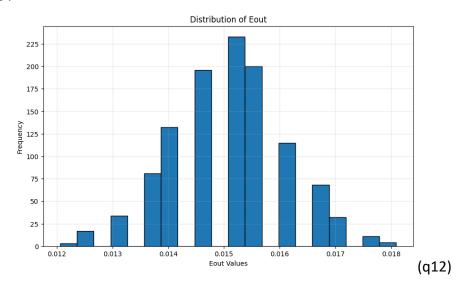
Code:

```
Eouts = []
  for experiment in tqdm(range(1126)):
      np.random.seed(experiment)
      total_size = len(X_train)
      indices = np.random.permutation(total_size)
      subtrain_indices = indices[:8000]
      validation_indices = indices[8000:]
      X_subtrain = [X_train[i] for i in subtrain_indices]
      X_validation = [X_train[i] for i in validation_indices]
      y_subtrain = [y_train[i] for i in subtrain_indices]
      y_validation = [y_train[i] for i in validation_indices]
      subtrain_prob = problem(y_subtrain, X_subtrain)
      min_validation_err = np.inf
      opt_log10_lambda = 0
      for log10_lambda in (-2, -1, 0, 1, 2, 3):
          subtrain_pred_res = []
          c = 1 / (10 ** log10_lambda)
          param = parameter('-s 6 -c ' + str(c))
          model_by_subtrain = train(subtrain_prob, param)
          validation_label, _, _ = predict(y_validation, X_validation, model_by_subtrain)
          validation_err = ZeroOneError(validation_label, y_validation)
          if validation_err == min_validation_err:
              opt_log10_lambda = max(opt_log10_lambda, log10_lambda) # break tie by choosing the larger lambda
          elif validation_err < min_validation_err:</pre>
              min_validation_err = validation_err
              opt_log10_lambda = log10_lambda
      # rerun the model with the best lambda on the whole training set
      whole_train_prob = problem(y_train, X_train)
      param_with_opt_lambda = parameter('-s 6 -c ' + str(1 / (10 ** opt_log10_lambda)))
      whole_train_model = train(whole_train_prob, param_with_opt_lambda)
      test_label, _, _ = predict(y_test, X_test, whole_train_model)
      Eout = ZeroOneError(test_label, y_test)
      Eouts.append(Eout)
61m 42.5s
```

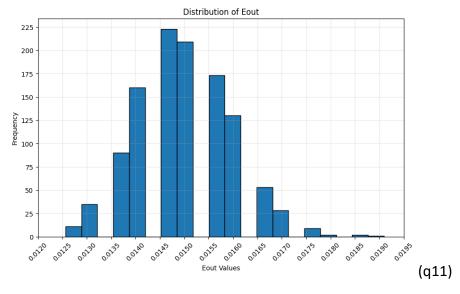
² Other part of codes that are similar to problem 10 (like reading in the data, split the data to get proper form of X and y, the ZeroOneError function...) are omitted in this report, can check for problem 10 for details.

HTML homework 5: q12 report

The resulting picture is as follows:



To find the difference, the plot below is the result of the previous problem:



From the two plots we can see that the subtrain / validation split is more dispersed, and the variability of the 3-fold cross validation is smaller.

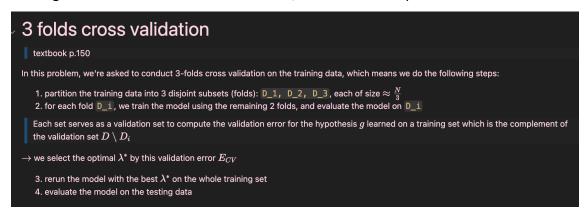
The reason why the subtrain / validation split is more dispersed may due to the fact that the validation set size is not that big compared to the 3-fold version. In the previous question, the validation set size is 3876, and in this question, using 3-fold splits the training set into 3 sets,

with each about 3959 examples, which is slightly higher. Also, E_{CV} is calculated by the mean error over using each of the 3 sets as the validation set, therefore these 2 reasons introduce stability in the 3-fold version.

On the other hand, the subtrain / validation split relies only on one split, if the validation set is not that representative, our selection of λ may be biased, this would magnify the effect of the unstableness in the previous question.

Code:

The markdown screenshot below describes the modification of this problem, other parts like reading in the data and the error function..., are the same as problem 10:



```
def run_single_experiment(experiment):
    np.random.seed(experiment)
    total_size = len(X_train)
    indices = np.random.permutation(total_size)
    fold_size = total_size // 3
    fold1_indices = indices[:fold_size]
    fold2_indices = indices[fold_size:2*fold_size]
    fold3_indices = indices[2*fold_size:]
    X_fold1 = [X_train[i] for i in fold1_indices]
    X_fold2 = [X_train[i] for i in fold2_indices]
    X_fold3 = [X_train[i] for i in fold3_indices]
    y_fold1 = [y_train[i] for i in fold1_indices]
    y_fold2 = [y_train[i] for i in fold2_indices]
    y_fold3 = [y_train[i] for i in fold3_indices]
    min_Ecv = np.inf
    opt_log10_lambda = 0
    for log10_lambda in (-2, -1, 0, 1, 2, 3):
        each_fold_as_valid_err = []
        c = 1 / (10 ** log10_lambda)
        for i in range(1,4):
            if i == 1:
               X_{subtrain} = X_{fold2} + X_{fold3}
                y_subtrain = y_fold2 + y_fold3
                X_{validation} = X_{fold1}
                y_validation = y_fold1
            elif i == 2:
                X_{subtrain} = X_{fold1} + X_{fold3}
                y_subtrain = y_fold1 + y_fold3
                X_{validation} = X_{fold2}
               y_validation = y_fold2
                X_subtrain = X_fold1 + X_fold2
                y_subtrain = y_fold1 + y_fold2
                X_{validation} = X_{fold3}
                y_validation = y_fold3
            subtrain_prob = problem(y_subtrain, X_subtrain)
            param = parameter('-s 6 -c ' + str(c))
            model_by_subtrain = train(subtrain_prob, param)
            validation_label, _, _ = predict(y_validation, X_validation, model_by_subtrain)
            validation_err = ZeroOneError(validation_label, y_validation)
            each_fold_as_valid_err.append(validation_err)
        Ecv_each_lambda = np.mean(each_fold_as_valid_err)
```

```
if Ecv_each_lambda == min_Ecv:
    opt_log10_lambda = max(opt_log10_lambda) # break tie by choosing larger lambda
elif Ecv_each_lambda < min_Ecv:
    min_Ecv = Ecv_each_lambda
    opt_log10_lambda = log10_lambda

whole_train_prob = problem(y_train, X_train)
param_with_opt_lambda = parameter('-s 6 -c ' + str(1 / (10 ** opt_log10_lambda)))
whole_train_model = train(whole_train_prob, param_with_opt_lambda)

test_label, _, _ = predict(y_test, X_test, whole_train_model)
return ZeroOneError(test_label, y_test)

n_jobs = -1
Eouts = Parallel(n_jobs=n_jobs)(
    delayed(run_single_experiment)(experiment)
    for experiment in tqdm(range(1126))

166m 7.7s</pre>
```

13, elathe not sives:

i'. We couldingto descent (instead of gradient descent):

$$W_{i}^{(trl)} = \underset{w_{i} \in \mathbb{R}}{\operatorname{avg}} \underset{w_{i} \in \mathbb{R}}{\operatorname{Av}} \left(y_{n} - \sum_{j \neq i} w_{j}^{(t)} x_{n,j} - w_{i} x_{n,i} \right)^{2} + \underset{N}{\lambda_{i}} \left(\sum_{j \neq i} \left[w_{j}^{(t)} \right] + \left[w_{i} \right] \right) + \underset{N}{\lambda_{i}} \left(\sum_{j \neq i} \left[w_{j}^{(t)} \right]^{2} + w_{i}^{2} \right)$$

expand the update equation:

$$\frac{|w|^2}{N} \sum_{n=2}^{N} (\gamma_{n,i})^2 + \frac{\lambda_1}{N} \sum_{t \in [w]} |w_j^{(t)}| + \frac{\lambda_1}{N} |w_i| + \frac{\lambda_2}{N} \sum_{t \in [w]} |w_j^{(t)}|^2 + \frac{\lambda_2}{N} |w_i|^2$$

= arg mir
$$\frac{\gamma wi}{N} \sum_{n=1}^{N} \left(\sum_{j \neq i} w_{j}^{(R)} \chi_{n,j} \chi_{n,i} - \chi_{n,i} \right) + \left(\frac{N}{N} \left(\chi_{n,i} \right)^{2} + \frac{\lambda_{2}}{N} \right) w_{i}^{2} + \frac{\lambda_{1}}{N} |w_{i}|$$
with

avgran
$$\frac{\lambda_{i}}{\lambda_{i}} = \frac{\lambda_{i}}{\lambda_{i}} = \frac{\lambda_$$

$$\frac{\partial}{\partial x} = -\frac{1}{N} \left[\sum_{n=1}^{N} \left(\sum_{j\neq i} w_{j}^{(e)} \chi_{n,j} \chi_{n,i} - \chi_{n,i} \right) \right] - \frac{\lambda_{1}}{N}$$

$$\frac{2}{N} \left(\frac{\sum_{j\neq i} \chi_{n,i}^{(e)} \chi_{n,j} \chi_{n,i} - \chi_{n,i}}{N} + \frac{\lambda_{2}}{N} \right)$$

else if vi < o , then we get:

argum
$$\frac{2W_i}{N} \stackrel{N}{\approx} \left(\stackrel{Z}{\downarrow} W_j^{(t)} \chi_{n,j} \chi_{n,i} - \chi_{n,i} \right) + \left(\stackrel{N}{\approx} \frac{\chi_{n,i}}{N} + \frac{\lambda_z}{N} \right) w_i^2 - \frac{\lambda_1}{N} w_i$$
with

$$\frac{1}{\sqrt{N}} \left(\sum_{n=1}^{N} \left(\sum_{j\neq i} w_{j}^{(e)} \chi_{n,j} \chi_{n,i} - \chi_{n,i} \right) \right) + \frac{\lambda_{1}}{\sqrt{N}}$$

$$\frac{1}{\sqrt{N}} \left(\sum_{n=1}^{N} \left(\sum_{j\neq i} w_{j}^{(e)} \chi_{n,j} \chi_{n,i} - \chi_{n,i} \right) \right) + \frac{\lambda_{1}}{\sqrt{N}}$$

in For can 1:

For can 2:

$$-\frac{2}{N}\left[\sum_{i=1}^{N}\left(\sum_{j\neq i}^{N}V_{j}^{(t)}\chi_{n,j}\chi_{n,i}-\gamma_{n,i}\right)\right]+\frac{\lambda_{1}}{N}<0$$

$$\frac{2}{N}\left[\sum_{i\neq i}^{N}\left(\sum_{j\neq i}^{N}W_{j}^{(t)}\chi_{n,j}\chi_{n,i}-\gamma_{n,i}\right)\right]-\frac{\lambda_{1}}{N}<0$$

$$\frac{2}{N}\left[\sum_{i\neq i}^{N}W_{j}^{(t)}\chi_{n,j}\chi_{n,i}-\gamma_{n,i}\right]-\gamma_{n,i}$$

$$\frac{2}{N}\left(\sum_{j\neq i}^{N}W_{j}^{(t)}\chi_{n,j}\chi_{n,i}-\gamma_{n,i}\right)-\gamma_{n,i}$$

then we should update wi as

$$\frac{1}{\sqrt[3]{\sum_{k=1}^{N} \sum_{j \neq i} (X_{i}, i)} \left(\frac{1}{\sqrt[3]{\sum_{k=1}^{N} \sum_{j \neq i} (X_{i}, i)} \frac{1}{\sqrt[3]{\sum_{k=1}$$

Thurshie, we have
$$d = sign(\sum_{n=1}^{N} (\sum_{j \neq i} w_j^{(e)} \chi_{n,j} \chi_{n,i} - \chi_{n,c}))$$

$$B = \left| -\sum_{n=1}^{N} \sum_{j \neq i} (\sum_{j \neq i} w_j^{(t)} \chi_{n,j} \chi_{n,i} - \chi_{n,c}) \right| - \frac{\lambda_1}{N}$$

$$= \left(\frac{\sum_{i \neq j} (\chi_{n,i})^2}{N} + \frac{\lambda_2}{N} \right)$$