2018 Engineering Mathematics II Midterm Solutions

Candidate Avg.	21.53
Max	35
Median	22
Min	3

EM2 TAs: Onyu Kang, Jiho Choi

If you have any question regarding your grade, please contact the following TA.

Total grade points [37 points]

- surname ¬ ~ 人: Onyu Kang,
- surname ~ ㅎ: Jiho Choi

1. [2, 2, 3, 2 points]

- 1-A) [2 points] Onyu Kang
- 1-B) [2 points] Onyu Kang
- 1-C) [3 points] Onyu Kang
- 1-D) [2 points] Onyu Kang

2. [2, 2, 3 points]

- 2-A) [2 points] Onyu Kang
- 2-B) [2 points] Onyu Kang
- 2-C) [3 points] Onyu Kang

3. [2, 3 points]

- 3-A) [2 points] Onyu Kang
- 3-B) [3 points] Onyu Kang

4. [2, 3, 3 points]

- 4-A) [2 points] Jiho Choi
- 4-B) [3 points] Jiho Choi
- 4-C) [3 points] Jiho Choi

5. [2, 2, 4 points]

- 5-A) [2 points] Jiho Choi
- 5-B) [2 points] Jiho Choi
- 5-C) [4 points] Jiho Choi

1. [2, 2, 3, 2 points]

1-A) [2 points] - Onyu Kang

Since the event $A \cap B_i$ (i=1,...,n) are disjoint and cover the entire sample space Ω , it follows that

$$Pr(A) = \sum_{i=1}^{n} Pr(A \cap B_i)$$
 ---[1 pt]

$$Pr(A) = \sum_{i=1}^{n} Pr(A \cap B_i) = (0+0+0+\frac{1}{36}+\frac{1}{36}+\frac{1}{36}) = \frac{1}{12} \quad ---[1 \text{ pt}]$$

Scores were specified for each.

1-B) [2 points] - Onyu Kang

$$\sum_{y} \Pr(Y = y) E[X|Y = y] = \sum_{y} \Pr(Y = y) \sum_{x} x \Pr(X = x|Y = y)$$

$$= \sum_{x} \sum_{y} x \Pr(X = x \mid Y = y) \Pr(Y = y)$$

$$= \sum_{x} \sum_{y} x \Pr(X = x \cap Y = y)$$

$$= \sum_{x} x \Pr(X = x) = E[X] \quad ---[1 \text{ pt}]$$

 $X \sim B(n, p)$, $X_k \sim B(k, p)$,

$$E[X_{k}] = E[X_{k}|X_{1} = 1] \cdot Pr(X_{1} = 1) + E[X_{k}|X_{1} = 0] \cdot Pr(X_{1} = 0)$$

$$= (1 + E[X_{k-1}]) \cdot p + E[X_{k-1}] \cdot (1 - p)$$

$$= p + E[X_{k-1}]$$

$$= 2p + E[X_{k-2}]$$
...
$$= (k-1)p + E[X_{1}]$$

$$= kp ---[1 pt]$$

Scores were specified for each.

1-C) [3 points] - Onyu Kang

C: conditional event = {HT, TH, HH},

$$E=\{HH\} \rightarrow Pr(E|C)=1/3 ---[1 pt]$$

Distinguish Heads into two types

H' (Heads and Big coin) and H (Heads and small coin)

C: Conditional event = {H'T, TH', H'H', H'H, HH'}, $E=\{H'H', H'H, HH', HH\} \rightarrow C \cap E = \{H'H', H'H, HH'\}$

$$Pr(H'T) = 1/100 * 1/2*1/2, Pr(H'H) = Pr(HH') = 1/00*1/2*1/2,$$

Ignore Pr(H'H') and let it be 0.

→
$$Pr(E|C)=1/2$$
 ---[2 pts]

Scores were specified for each.

1-D) [2 points] - Onyu Kang

$$X = \{-1,0,1\}, Y = X^2$$

$$Y=\{1,0,1\}, XY=\{-1,0,1\} = X^3$$

$$\Pr(X = 1 \cap Y = 1) = \frac{2}{3}$$

$$Pr(X=1)Pr(Y=1) = \frac{1}{3} \cdot \frac{2}{3} = \frac{2}{9}$$

These X and Y are dependent variables.

$$E[XY] = \frac{1}{3} \cdot (-1) + \frac{1}{3} \cdot 0 + \frac{1}{3} \cdot 1 = 0$$

$$E[X] = \frac{1}{3} \cdot (-1) + \frac{1}{3} \cdot 0 + \frac{1}{3} \cdot 1 = 0$$

$$E[Y] = \frac{1}{3} \cdot 0 + \frac{2}{3} \cdot 1 = \frac{2}{3}$$

$$E[X]E[Y] = 0 \cdot \frac{2}{3} = 0$$

$$\therefore E[XY] = E[X]E[Y]$$

No partial score. If you have a reasonable answer, you will get full score.

2. [2, 2, 3 points]

2-A) [2 points] - Onyu Kang

D : has DisorderTP : Test PositiveTN : Test Negative

$$Pr(D) = 0.001$$

$$Pr(\overline{D}) = 0.999$$

$$Pr(TP|D) = 0.99$$

$$Pr(TP|\overline{D}) = 0.05$$

$$Pr(D|TP) = \frac{(Pr(TP|D) * Pr(D))}{Pr(TP)}$$
 (Bayes' Theorem)

$$= \frac{(Pr(TP|D) * Pr(D))}{Pr(TP|D) * Pr(D) + Pr(TP|\overline{D}) * Pr(\overline{D})}$$

$$= \frac{0.99 * 0.001}{0.99 * 0.001 + 0.05 * 0.999}$$

$$= 99 / 5094$$

2-B) [2 points] - Onyu Kang

Let r.v. X = # Heads in m tosses, Y=Sequence number of the picked coin

$$E[X \mid Y=i] = m^*(i/n)$$

= 0.0194...

Using
$$\sum_y E[X \mid Y=y] \cdot \Pr(Y=y)$$
, $E[X] = \sum_i m \cdot \left(\frac{i}{n}\right) \cdot \left(\frac{1}{n}\right) \cong m$

No partial score.

2-C) [3 points] - Onyu Kang

Let C1, C2 be events to select coin C1 and C2, respectively. And E be an event of having k Heads among m tosses. We compute Pr(C1 | E).

$$\Pr(\mathsf{E}|\mathsf{C1}) = \binom{m}{k} (\frac{1}{4})^k \cdot (\frac{3}{4})^{m-k}$$

$$\Pr(E|C2) = {m \choose k} (\frac{3}{4})^k \cdot (\frac{1}{4})^{m-k}$$

$$= \frac{\Pr(E|C1)\Pr(C1)}{\Pr(E)}$$

$$= \frac{\Pr(\mathsf{E}|\mathsf{C1})\Pr(\mathsf{C1})}{\Pr(\mathsf{E}|\mathsf{C1})\Pr(\mathsf{C1}) + \Pr(\mathsf{E}|\mathsf{C2})\Pr(\mathsf{C2})}$$

$$= \frac{\binom{m}{k} \left(\frac{1}{4}\right)^k \cdot \left(\frac{3}{4}\right)^{m-k} \cdot \frac{1}{2}}{\binom{m}{k} \left(\frac{1}{4}\right)^k \cdot \left(\frac{3}{4}\right)^{m-k} \cdot \frac{1}{2} + \binom{m}{k} \left(\frac{3}{4}\right)^k \cdot \left(\frac{1}{4}\right)^{m-k} \cdot \frac{1}{2}}$$

$$= \frac{3^{m-k}}{3^{m-k}+3^k}$$

No partial score.

3. [2, 3 points]

3-A) [2 points] - Onyu Kang

 $X \sim Poi(\lambda)$. Consider arrivals of customers as events. Then we can model arrivals as trials with probability of success (female) is p. \Rightarrow Y = # females among n is a Binomial r.v. with parameter n and p.

$$Pr(Y=m \mid X=n) = {n \choose m} \cdot p^m (1-p)^{n-m}$$

No partial score.

3-B) [3 points] - Onyu Kang

$$\begin{aligned} \Pr(\mathsf{Y} = \mathsf{m}) &= \sum_{n=m}^{\infty} \Pr(Y = m \mid X = n) \cdot \Pr(\mathsf{X} = \mathsf{n}) \\ &= \sum_{n=m}^{\infty} \binom{n}{m} \cdot p^m (1-p)^{n-m} \cdot \frac{\lambda^n e^{-\lambda}}{n!} \end{aligned}$$

No partial score.

4. [2, 3, 3 points]

4-A) [2 points] - Jiho Choi

(1 point)

X_i is a Bernoulli [버눌리, 베르누이] random variable.

Since the area of heart is p and the probability that a dart lands on the heart is in proportion to its area, $Pr(X_i = 1) = p$ and $Pr(X_i = 0) = 1 - p$

c.f. Bernoulli distribution (of a random variable) is a special case of the binomial distribution where a single trial is conducted to asks a yes-no question; the question results in a Boolean-valued outcome. **Answer with Binomial R.V. will get 0.5 points**

(1 point)

X is a binomial random variable.

$$X = \sum_{i=1}^{n} X_i$$

$$\Pr(X_i = k) = \binom{n}{k} * p^k * (1 - p)^{n - k}$$

$$X \sim B(n, p)$$

4-B) [3 points] - Jiho Choi

Let the area of the heart equals A.

$$K \sim B(n,A)$$

$$E(k)=n*A, Var(k)=n*A(1-A)$$

(2.5 points)

c.f. Chebyshev's Bound: $\Pr(|X - E[X]| \ge a) \le \frac{Var[X]}{a^2}$

$$\Pr\left(\left|A - \frac{k}{n}\right| \ge \delta\right) = \Pr(\left|A\mathbf{n} - \mathbf{k}\right| \ge \delta n) = \Pr(\left|\mathbf{k} - A\mathbf{n}\right| \ge \delta n) \le \frac{Var(k)}{(\delta n)^2} = \frac{nA(1-A)}{\delta^2 n^2} = \frac{A(1-A)}{\delta^2 n}$$

$$\frac{A(1-A)}{\delta^2 n} \le \frac{1}{4} * \frac{1}{\delta^2 n} \ (\because -A^2 + A = -\left(A - \frac{1}{2}\right)^2 + \frac{1}{4})$$

(0.5 points)

$$\frac{1}{4} * \frac{1}{\delta^2 n} \le 0.05$$

$$\frac{1}{4} * \frac{100}{5} * \frac{1}{\delta^2} \le n$$

$$\therefore n \ge \frac{5}{\delta^2}$$

c.f. $\frac{20A(1-A)}{\delta^2} \le n$, Answer without the bound of $A(1-A) \le \frac{1}{4}$ also gets 2.5 points

4-C) [3 points] - Jiho Choi

c.f. Chernoff's Bound: $\Pr(X \ge a) = \Pr(e^{tX} \ge e^{ta}) \le \frac{E[e^{tX}]}{e^{ta}}$

$$\Pr(X \ge a) \le \min_{t>0} \frac{E[e^{tX}]}{e^{ta}}$$

→
$$\Pr(|X - \mu| \ge \delta \cdot \mu) \le 2e^{-\frac{\mu \cdot \delta^2}{3}}$$
 Refer to Theorem 4.5 & Corollary 4.6

1)

Yes, with the MGF of k, we can apply Chernoff's inequality. (Apply Theorem 4.5)

$$\Pr\left(\left|A - \frac{k}{n}\right| \ge \delta\right) = \Pr(|An - k| \ge \delta n) = \Pr(|X - \mu| \ge \delta \frac{\mu}{A}), \ \mu = A * n$$

$$\Pr\left(|X - \mu| \ge \delta \frac{\mu}{A}\right) \le 2e^{-\frac{\mu_A^{\delta^2}}{3}}$$
 (Theorem 4.5 & Corollary 4.6)

2)

$$\begin{split} &\Pr(|\mathsf{A} \mathsf{n} - \mathsf{k}| \geq \delta n) \\ &= \Pr(\mathsf{X} - \mathsf{n} \mathsf{A} \geq \delta n) + \Pr(\mathsf{X} - \mathsf{n} \mathsf{A} \leq -\delta n) \\ &= \Pr(\mathsf{X} \geq \delta n + nA) + \Pr(\mathsf{X} \leq -\delta n + nA) \\ &= \Pr\left(e^{tX} \geq e^{t(\delta n + nA)}\right) + \Pr\left(e^{sX} \geq e^{s(-\delta n + nA)}\right) \leq \frac{E\left[e^{tX}\right]}{e^{t(\delta n + nA)}} + \frac{E\left[e^{sX}\right]}{e^{s(-\delta n + nA)}} \ (t > 0, \ s < 0) \end{split}$$

C.f. Answer "No" with reasonable approach will get some partial credit.

5. [2, 2, 4 points] Jiho Choi

5-A) [2 points] - Jiho Choi

Poisson Approximation

Poisson random variable Yi, Throw m=2n balls into n bins $\mu=\frac{2n}{n}=2$

$$Y_i \sim Poi(\frac{2n}{n})$$

$$Pr(Y_1 = 2, Y_2 = 2, Y_3 = 2, ..., Y_n = 2) = \prod_{i=1}^{n} Pr(Y_i = 2) = \left(\frac{e^{-2}2^2}{2!}\right)^n = 2^n e^{-2n}$$

$$\Pr(X_1 = 2, X_2 = 2, X_3 = 2, ..., X_n = 2) \le e\sqrt{2n} * \prod_{i=1}^n \Pr(Y_i = 2) = e\sqrt{2n} * 2^n e^{-2n}$$

5-B) [2 points] - Jiho Choi

$$\begin{split} &\Pr(\mathbf{X}_1 = 2, \mathbf{X}_2 = 2, \mathbf{X}_3 = 2, ..., \mathbf{X}_n = 2) = \binom{2n}{2} \binom{2n-2}{2} \binom{2n-4}{2} * ... * \binom{2}{2} * \frac{1}{n^{2n}} \ \text{(1.5 points)} \\ &= \frac{(2n)!}{(2!)^n} \frac{1}{n^{2n}} = \frac{(2n)!}{(2n)^n} = \frac{(2n)!}{(2n)^n} \ \text{(0.5 points)} \end{split}$$

5-C) [4 points] - Jiho Choi

E_n: A group consists of bin n~k are all non-empty

$$\Pr(\mathsf{Max}.\mathsf{Cluster} \geq \mathsf{k}) = \Pr(\mathsf{E}_1 \cup \mathsf{E}_2 \cup ... \ \cup \mathsf{E}_{\mathsf{n-k+1}}) \leq \Pr(\mathsf{E}_1) + \Pr(\mathsf{E}_2) + \cdots + \Pr(\mathsf{E}_{\mathsf{n-k+1}})$$

=
$$\Pr(E_1) + \Pr(E_2) + \dots + \Pr(E_{n-k+1}) \le 2 *$$
Probability of the Poisson Case

$$= 2 * (n - k + 1) \cdot (1 - \frac{1}{e})^{k}$$

c.f. Answer with reasonable approach will get some partial credit.