

Q1

N letters are to be put in N separate envelopes.

N=50

$$P(\text{at least one letter is in the correct envelop}) = 1 - P(\text{No letter is in the correct envelop})$$

~~$D_n$~~

$$P(\text{at least one letter is in the correct envelop}) = 1 - \frac{D_n}{n!}$$

$$D_n = \left[ \frac{n!}{e} + \frac{1}{2} \right]$$

[ ]  $\rightarrow$  greatest integer

$$\rightarrow 1 - \frac{D_{50}}{50!} \Rightarrow 1 - \frac{1}{50!} \left[ \frac{50!}{e} + \frac{1}{2} \right] \approx 1 - \frac{1}{e} \approx 0.6321$$

Date \_\_\_\_\_  
Page \_\_\_\_\_

(Q9) Convolution of 2 distribution function is also a distribution function

Let  $F$  &  $G$  be distribution functions of 2 random variables  $X$  &  $Y$   
 $Z = X+Y$

Convolution of  $F$  &  $G$  can be written as

$$H(x) = (F * G)(x) = \int_{-\infty}^{\infty} F(x-y) dG(y)$$

(i) Let  $x_1 < x_2$  for fixed  $y$

$$F(x_1 - y) \leq F(x_2 - y)$$

as  $F$  is ~~monotonic~~ non-decreasing, integrating both sides wrt.  $G$ ,

$$H(x_1) = \int F(x_1 - y) dG(y) \leq \int F(x_2 - y) dG(y) = H(x_2)$$

$\Rightarrow H(x)$  is non-decreasing

(ii) as  $x \rightarrow -\infty$ , for any  $y$ ,  $x-y \rightarrow -\infty \Rightarrow F(x-y) \rightarrow 0$

$$\lim_{x \rightarrow -\infty} H(x) = \int 0 dG(y) = 0$$

$$\text{as } x \rightarrow \infty, F(x-y) \rightarrow 1 \forall y, \lim_{x \rightarrow \infty} H(x) = \int 1 dG(y) = 1$$

(iii) To show:  $H(x) \rightarrow H(x_0)$  as  $x \rightarrow x_0^+$

$\because F$  is right continuous, the integrand  $F(x-y)$  is also right-continuous in  $x$  for each  $y$ .

$\rightarrow$  Thus,  $H(x)$  is right continuous.

$\rightarrow$  As it obeys all properties of a p.d.f, hence a ~~dist~~ distribution.



Q5.

1, 2, 3, ..., N ; n ≤ N tickets are drawn with replacement. Only maximum prize ticket is to be kept.

(M = prize money obtained)

(E(M) = ?)

~~Find the probability of getting the maximum prize ticket~~

$$P(M \leq k) = P(\text{all } n \text{ draws} \leq k) = \left(\frac{k}{N}\right)^n$$

$$P(M = k) = P(M \leq k) - P(M \leq k-1) = \left(\frac{k}{N}\right)^n - \left(\frac{k-1}{N}\right)^n$$

$$\Rightarrow E[M] = \sum_{k=1}^N k \cdot P(M = k) = \sum_{k=1}^N k \left[ \left(\frac{k}{N}\right)^n - \left(\frac{k-1}{N}\right)^n \right]$$

h

Q6.

Let  $X$  &  $Y$  be chosen uniformly on  $[0, d]$

$P(|X-Y| < d/3)$  as  $X$  &  $Y$  are chosen independently & uniformly, their joint distribution is uniform over squared  $[0, d] \times [0, d]$

$$0 \leq X, Y \leq d$$

$$|X-Y| \leq d/3$$

$\Pi$  is a square with area  $d^2 \Rightarrow X-Y = \pm d/3$

The region is a band of width  $2d/3$  along the diagonal  $X=Y$  centered within the square, so desired probability is area of region where  $|X-Y| < d/3$

$$A_{\text{band}} = d^2$$

Area outside band  $|X-Y| \geq d/3 \rightarrow$  this consists of 2 right-angled triangles, one below  $Y=X+d/3$  & one above  $Y=X-d/3$

$$\text{Each of } \Delta \text{ has length} = d - d/3 = 2d/3$$

$$\text{Area of each } \Delta = \frac{1}{2} \times \frac{2d}{3} \cdot \frac{2d}{3} = \frac{2d^2}{9}$$

$$\text{Area outside band} = \frac{4d^2}{9}$$

$$\text{Area inside band} = \frac{5d^2}{9}$$

$$\Rightarrow P(|X-Y| < d/3) = \frac{5d^2/9}{d^2} = \frac{5}{9}$$



Q2.

Let  $G_i$  be gift in present  $i$ ,  $i=1,2,3$

Let  $H_2$  be host opens present 2 & shows it is empty.

I chose present 1, initially.

I need  $E[\text{win}/H_2, \text{I switch to present 3}]$

$$= 1000 \cdot P(G_3/H_2) + 0 \cdot P(G_1/H_2) = 1000 \cdot P(G_3/H_2)$$

→ All gifts are equally likely to be chosen

$$\rightarrow P(G_1) = P(G_2) = P(G_3) = \frac{1}{3}$$

Case 1 → If gift in Present 1  $\Rightarrow P(H_2/G_1) = \frac{1}{2}$

Case 2 → If gift in Present 2  $\Rightarrow P(H_2/G_2) = 0$

Case 3 → If gift in Present 3  $\Rightarrow P(H_2/G_3) = 1$

$$\text{Now } P(H_2) = \sum_{i=1}^3 P(H_2/G_i) \cdot P(G_i) = \frac{1}{2} \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + \frac{1}{3} = \frac{1}{2}$$

$$P(G_3/H_2) = \frac{P(H_2/G_3) \cdot P(G_3)}{P(H_2)} = \frac{1 \cdot \frac{1}{3}}{\frac{1}{2}} = \frac{2}{3}$$

Hence

$$E[\text{win}/H_2, \text{switch}] = 1000 \cdot \frac{2}{3} = \text{666.67}$$

diagonal  
no derived probability is  
d/dz / d^2

ists

Date   
Page

Q13  $u \in \mathbb{R}$  .  $\phi(u) = e^{u^2}$   $\forall u \in \mathbb{R}$   
 $X$  is a normal random variable,  $\mu = E[X]$ ,  $\sigma = [E(X-\mu)^2]^{1/2}$

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

Q1  $E[e^{ux}] = e^{u\mu + \frac{1}{2}\sigma^2 u^2}$

LHS  $E[e^{ux}] = E[e^{ux - u\mu + u\mu}] = e^{u\mu} E[e^{u(x-\mu)}]$

$$\begin{aligned} E[e^{ux}] &= e^{u\mu} \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{u(x-\mu)} e^{-\frac{1}{2}\frac{(x-\mu)^2}{\sigma^2}} dx \\ &= e^{u\mu} \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2\sigma^2}[(x-\mu)^2 - 2\sigma^2 u(x-\mu)]} dx \end{aligned}$$

Now

$$\begin{aligned} (x-\mu)^2 - 2\sigma^2 u(x-\mu) &= (x-\mu)^2 - 2\sigma^2 u(x-\mu) + \sigma^4 u^2 - \sigma^4 u^2 = (x-\mu-\sigma^2 u)^2 - \sigma^4 u^2 \\ \Rightarrow (x-\mu-\sigma^2 u)^2 - \sigma^4 u^2 \end{aligned}$$

$$\therefore E(e^{ux}) = e^{u\mu} e^{\frac{\sigma^4 u^2}{2}} \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(x-\mu-\sigma^2 u)^2}{2\sigma^2}} dx$$

$$E(e^{ux}) = e^{u\mu + \frac{1}{2}\sigma^2 u^2} \left( \because \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(x-(\mu+\sigma^2 u))^2}{2\sigma^2}} dx = 1 \right)$$

$\Rightarrow \boxed{E[e^{ux}] = e^{u\mu + \frac{1}{2}\sigma^2 u^2}} \quad \text{f.v.}$



sists of two

(ii)

$$E(\phi(x)) \geq \phi(E(x))$$

Verification

$$E[\phi(x)] \Rightarrow e^{u\mu + \frac{1}{2}\sigma^2 u^2} = e^{u\mu} \cdot e^{\frac{1}{2}\sigma^2 u^2} = \phi(E(x)) \cdot e^{\frac{1}{2}\sigma^2 u^2}$$

Now

$$\therefore e^n \geq 1 \text{ if } n \geq 0$$

$$\& \text{ clearly } \frac{1}{2}\sigma^2 u^2 \geq 0 \Rightarrow$$

$$e^{\frac{1}{2}\sigma^2 u^2} \geq 1$$

$$\therefore E[\phi(x)] \geq \phi(E(x))$$

H.V

$\Rightarrow$  Jensen's inequality holds