

3.

$$(a) P(A \cap B | C) = P(A | B \cap C) P(B | C)$$

$$P(A | B \cap C) = \frac{P(A \cap B \cap C)}{P(B \cap C)}$$

multiply  $\rightarrow P(B | C)$

$$P(A | B \cap C) P(B | C) = \frac{P(A \cap B \cap C)}{P(B \cap C)} \cdot \frac{P(B \cap C)}{P(C)}$$

Hence, True

$$(b) P(A \cap B | C) = P(A | C) P(B | C) \quad (A \text{ \& B are independent})$$

$$P(A \cap B \cap C) = \frac{P(A \cap C) P(B \cap C)}{P(C)} ; P(A \cap B | C) = \frac{P(A \cap B \cap C)}{P(C)}$$

not always true, independent events A & B does not guarantee independence of C,

Hence, False

(c) A is more likely to occur when B is absent i.e.  $B^c$

• for D :  $A | B^c > A | B$

• for  $D^c$  :  $A | B^c > A | B$

this probability is just weighted sum of regional probability

hence,  $P(A | B^c) > P(A | B)$

is True

4)

let  $x$  be random variable such that  
we define  $x \in N$

$$N = \{1, 2, 3, \dots\} \quad (\text{finite})$$

$$P(x=n) = \frac{c}{n^p} \quad [n \geq 1]$$

Now,

$$E(x) = \sum_{n=1}^{\infty} n \cdot \frac{c}{n^p} = c \sum_{n=1}^{\infty} \frac{1}{n^{p-1}}$$

$$E(x^2) = \sum_{n=1}^{\infty} n^2 \frac{c}{n^p} = c \sum_{n=1}^{\infty} \frac{1}{n^{p-2}}$$

let  $p > 2$

therefore,  $\sum \frac{1}{n^{p-1}}$  will converge

but  $\sum \frac{1}{n^{p-2}}$  will diverge.

ex)

say

$p=3$

$$c \sum_{n=1}^{\infty} \frac{1}{n} \rightarrow \infty$$

but  $c \sum_{n=1}^{\infty} \frac{1}{n^2}$  will

be finite.

b)

$$f(x) = \frac{c}{x^p}$$

$x \geq 1$  &  $p > 1$

$$E(x) = \int_1^{\infty} x \cdot \frac{c}{x^p} dx = c \int_1^{\infty} x^{1-p} dx$$

$$E(x^2) = \int_1^{\infty} x^2 \frac{c}{x^p} dx = c \int_1^{\infty} x^{2-p} dx$$

now, let  $p > 2$  say,  $p=2.5$

then,

$$E(x) < \infty$$

$$\text{but } E(x^2) \Rightarrow \infty$$

thus, exist

c)

$$E(e^{-x}) \geq e^{-E(x)}$$

$$> e^{-1} > 0.3679$$

thus, for random variable with

$$E(x) = 1$$

we must have

$$E(e^{-x}) \geq 1/3$$

thus, does not exist



7)

(a) people =  $P_i$   $\therefore i = 0, 1, 2, \dots, n$

starting @ " $P_0$ "

event - 'no one ever tells the rumor back to  $P_0$  in first  $r$  steps'

At any step (i), there are  $(n+1)$  people one of which is  $P_0$ , so, probability of " $P_0$ " is not chosen.

$$P(A) = 1 - \frac{1}{n+1} = \frac{n}{n+1}$$

Also, every step is independent of past, future or any other step, thus according to law of independent events.

Ans:  $P_r(E) = [P(A)]^r$   
 $= \left[ \frac{n}{n+1} \right]^r$

(b) Probability that rumor is told  $r$  times without being reported to any person

$P(E)$  after 1<sup>st</sup> step = 1  $\therefore n/n$

$P(E)$  " 2<sup>nd</sup> " =  $n^{-1}/n$

thus,  $P(E) = \frac{n - (r-1)}{n} = \frac{n-r+1}{n}$

$$P(E) = \left( \frac{n}{n} \right) \left( \frac{n-1}{n} \right) \dots \left( \frac{n-r+1}{n} \right)$$

$$= \frac{n!}{(n-r)! n^r}$$

(a) when each step tells to group of ' $N$ ' random people,

$$P(E) = \frac{{}^n C_N}{{}^{n+1} C_N} = \frac{n(n-1) \dots (n-N+1)}{(n+1)(n) \dots (n-N+1)}$$

(b) No one is told twice.

$$P_r = \prod_{k=1}^r \frac{1}{(n+1-k)(k-1)} \binom{n}{k}$$

$$[(w)q_b(w)x, 0]^{n+1} \binom{n}{k}$$

$$q_b[(w)q_b(w)x, 0]^{n+1} = [x] \int_0^\infty$$

$$q_b(w) [(w)x, 0]^{n+1} = (w)x$$

reduction of both sides

$$[q_b(w) [(w)x, 0]^{n+1}] \int_0^\infty = [(w)x] \int_0^\infty$$

reduction of both sides

$$(w)q_b(w) [(w)x, 0]^{n+1} \int_0^\infty = [x] \int_0^\infty$$

$$(w)q_b(w) [(w)x, 0]^{n+1} \int_0^\infty = [x] \int_0^\infty$$

$$(x < x) q_b = (x)q_b(x) (w)x, 0 \int_0^\infty = 1 - f(x)$$

$$q_b[(w)q_b(w)x, 0]^{n+1} = (x) \int_0^\infty$$

$$q_b[(w)q_b(w)x, 0]^{n+1} = (x) \int_0^\infty$$

$$(w)q_b(w) [(w)x, 0]^{n+1} \int_0^\infty = (x) \int_0^\infty$$

$$q_b[(w)q_b(w)x, 0]^{n+1} = (x) \int_0^\infty$$

radon

(+1)

(+N+1)

8) independent events have independent complements.

$$P\left(\bigcap_{i=1}^n A_i^c\right) = \prod_{i=1}^n P(A_i^c)$$

$$\because P(A_i^c) = 1 - P(A_i)$$

so,

$$\prod_{i=1}^n P(A_i^c) = \prod_{i=1}^n (1 - P(A_i))$$

for any real no.  $1 - x \leq e^{-x}$

this gives,

$$\prod_{i=1}^n (1 - P(A_i)) \leq \prod_{i=1}^n e^{-P(A_i)}$$

$$\Rightarrow \prod_{i=1}^n e^{-P(A_i)} = e^{-\sum_{i=1}^n P(A_i)}$$

thus,

$$P\left(\bigcap_{i=1}^n A_i^c\right) \leq e^{-\sum_{i=1}^n P(A_i)}$$



10)

Show:  $E(x) = \int_0^{\infty} [1 - F(x)] dx$

by showing:

$$\int_0^{\infty} \int_0^{\infty} [0, x(w)] dx dP(w)$$

$$\Rightarrow E(x) = \int_0^{\infty} [1 - F(x)] dx$$

$$x(w) = \int_0^{\infty} [0, x(w)](x) dx$$

taking expectation value of both sides

$$E(x(w)) = E \left[ \int_0^{\infty} [0, x(w)](x) dx \right]$$

applying 'Fubini's theorem'

$$E(x) = \int_0^{\infty} \int_0^{\infty} [0, x(w)](x) dx dP(w)$$

$$= \int_0^{\infty} \left( \int_0^{\infty} [0, x(w)](x) dP(w) \right) dx$$

Now,

$$\int_0^{\infty} [0, x(w)](x) dP(w) = P(x > x)$$

$$= 1 - F(x)$$

$$\# E(x) = \int_0^{\infty} (1 - F(x)) dx$$

$$\# E(x) = \int_0^{\infty} (1 - F(x)) dx$$

$$\# \text{ Therefore, } E(x) = \int_0^{\infty} \int_0^{\infty} [0, x(w)](x) dx dP(w)$$

$$= \int_0^{\infty} (1 - F(x)) dx$$