

16-17 学年第一学期答案

二、解：设 A_i 为第 i 个元件正常工作 ($i=1,2,3,4,5$) $A=\{\text{系统正常工作}\}$

$$P(A) = P(A|A_3)P(A_3) + P(A|\bar{A}_3)P(\bar{A}_3)$$

$$\begin{aligned} P(A|A_3) &= P\{(A_1 \cup A_4) \cap (A_2 \cup A_5)\} = P(A_1 \cup A_4)P(A_2 \cup A_5) \\ &= (P(A_1) + P(A_4) - P(A_1A_4))(P(A_2) + P(A_5) - P(A_2A_5)) \\ &= (2p - p^2)^2 = 0.9216 \end{aligned}$$

$$P(A|\bar{A}_3) = P\{A_1A_2 \cup A_4A_5\} = P(A_1A_2) + P(A_4A_5) - P(A_1A_2A_4A_5) = 2p^2 - p^4 = 0.8704$$

$$\begin{aligned} P(A) &= P(A|A_3)P(A_3) + P(A|\bar{A}_3)P(\bar{A}_3) \\ &= (2p - p^2)^2 \times p + (2p^2 - p^4) \times (1 - p) \\ &= 2p^2 + 2p^3 - 5p^4 + 2p^5 = 0.91136 \end{aligned}$$

三、解：1) $1 = \int_{-\infty}^{+\infty} f(x)dx = \int_0^1 \frac{1}{2}dx + A \int_1^{+\infty} \frac{1}{x^2}dx \rightarrow \frac{1}{2} = A \int_1^{+\infty} \frac{1}{x^2}dx \rightarrow -\frac{A}{x} \Big|_1^{+\infty} = \frac{1}{2} \rightarrow A = \frac{1}{2}$

2) 当 $x < 0$ 时, $F(x) = \int_{-\infty}^x f(t)dt = 0$

当 $0 \leq x < 1$ 时

$$F(x) = \int_{-\infty}^x f(t)dt = \int_0^x \frac{1}{2}dt = \frac{x}{2}$$

当 $x \geq 1$ 时

$$F(x) = \int_{-\infty}^x f(t)dt = \int_0^1 f(t)dt + \int_1^x f(t)dt = \frac{1}{2} + \int_1^x \frac{1}{2t^2}dt = \left(1 - \frac{1}{2x}\right)$$

故得分布函数 $F(x) = \begin{cases} 0, & x < 0 \\ \frac{x}{2}, & 0 \leq x < 1 \\ 1 - \frac{1}{2x}, & x \geq 1 \end{cases}$

3) $P\left\{\frac{1}{2} \leq X \leq \frac{5}{2}\right\} = P\left\{\frac{1}{2} < X \leq \frac{5}{2}\right\} = F\left(\frac{5}{2}\right) - F\left(\frac{1}{2}\right) = \frac{4}{5} - \frac{1}{4} = \frac{11}{20} = 0.55$

4) $y = g(x) = (x-1)^2$

当 $y < 0$ 时, $F_Y(y) = 0$

当 $y \geq 0$ 时,

$$F_Y(y) = P\{Y \leq y\} = P\{(X-1)^2 \leq y\} = P\{1-\sqrt{y} \leq X \leq 1+\sqrt{y}\} = F_X(1+\sqrt{y}) - F_X(1-\sqrt{y})$$

则 Y 的 PDF 为

$$f_Y(y) = \frac{1}{2\sqrt{y}}(f_X(1+\sqrt{y}) + f_X(1-\sqrt{y}))$$

$$= \begin{cases} 0 & y < 0 \\ \frac{1}{4\sqrt{y}} \left(\frac{1}{(1+\sqrt{y})^2} + 1 \right) & 0 \leq y < 1 \\ \frac{1}{4\sqrt{y}(1+\sqrt{y})^2} & y > 1 \end{cases}$$

四、解： 1) $f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy = \begin{cases} \int_x^{+\infty} e^{-y} dy = e^{-x}, & x > 0 \\ 0, & \text{else} \end{cases}$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx = \begin{cases} \int_0^y e^{-y} dx = ye^{-y}, & y > 0 \\ 0, & \text{else} \end{cases}$$

$$2) f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{1}{y}, & 0 < x < y \\ 0, & \text{else} \end{cases}$$

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \begin{cases} e^{x-y}, & y > x > 0 \\ 0, & \text{else} \end{cases}$$

3) 因为 $f_X(x) \times f_Y(y) \neq f(x, y)$, 所以 X 与 Y 不独立

$$4) E(X) = \int_{-\infty}^{+\infty} x f_X(x) dx = \int_0^{\infty} x e^{-x} dx = 1$$

$$E(X^2) = \int_{-\infty}^{+\infty} x^2 f_X(x) dx = \int_0^{\infty} x^2 e^{-x} dx = 2, D(X) = 1$$

$$E(Y) = \int_{-\infty}^{\infty} y f_Y(y) dy = \int_0^{+\infty} y^2 e^{-y} dy = 2$$

$$E(Y^2) = \int_{-\infty}^{\infty} y^2 f_Y(y) dy = \int_0^{+\infty} y^3 e^{-y} dy = 6, D(Y) = 2$$

$$E(XY) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f(x, y) dx dy = \int_0^{+\infty} \left(ye^{-y} \int_0^y x dx \right) dy = 3$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 1 \rightarrow \rho_{XY} = \sqrt{2}/2 = 0.717$$

$$\text{方法一: } F_Z(z) = P\{3X + 2Y \leq z\} = \iint_{3x+2y \leq z} f(x, y) dx dy = \int_0^{\frac{z}{5}} \int_x^{\frac{z-3x}{2}} e^{-y} dx dy = \begin{cases} 1 - \frac{5}{3}e^{-\frac{z}{5}} + \frac{2}{3}e^{-\frac{z}{2}} & z > 0 \\ 0, & \text{else} \end{cases}$$

$$f_Z(z) = F'_Z(z) = \begin{cases} \frac{1}{3} \left(e^{-\frac{z}{5}} - e^{-\frac{z}{2}} \right), & z > 0 \\ 0, & \text{else} \end{cases}$$

$$\text{方法二: } f_Z(z) = \int_{-\infty}^{\infty} f\left(\frac{z-2y}{3}, y\right) \left(\frac{z-2y}{3}\right)' dy = \frac{1}{3} \int_{\frac{z}{5}}^{\frac{z}{2}} e^{-y} dy = \begin{cases} \frac{1}{3} \left(e^{-\frac{z}{5}} - e^{-\frac{z}{2}} \right), & z > 0 \\ 0, & \text{else} \end{cases}$$

$$0 < \frac{z-2y}{3} < y \rightarrow 2y < z < 5y \rightarrow z/5 < y < z/2$$

五、解：1) $\mu = E(X) = \frac{\theta}{2}$, 令 $\mu = \frac{\theta}{2} = \bar{X}$ 得 θ 的矩估计量 $\hat{\theta}_M = 2\bar{X}$.

$$\hat{\theta}_{MLE} = \max(X_1, X_2, \dots, X_n) = X_{(n)}$$

$$2) E(\hat{\theta}_M) = E(2\bar{X}) = 2E(\bar{X}) = 2E(X) = 2 \cdot \frac{\theta}{2} = \theta \quad \text{无偏}$$

$$D(\hat{\theta}_M) = D(2\bar{X}) = 4D(\bar{X}) = 4 \frac{D(X)}{n} = 4 \cdot \frac{\theta^2}{12n} \xrightarrow{n \rightarrow \infty} 0$$

$$E(\hat{\theta}_M) = \theta, \lim_{n \rightarrow +\infty} D(\hat{\theta}_M) = 0 \quad \text{故 } \hat{\theta}_M \text{ 为 } \theta \text{ 的相合估计。}$$

或者方法二：按照 $\bar{X} \xrightarrow{P} \frac{\theta}{2}$, 故 $2\bar{X} \xrightarrow{P} \theta$, 所以 $\hat{\theta}_M$ 为 θ 的相合估计

$$3) \hat{\theta}_{MLE} \text{ 的 PDF 为 } f_{X_{(n)}}(x) = n[F(x)]^{n-1}f(x) = \begin{cases} \frac{nx^{n-1}}{\theta^n} & 0 \leq x \leq \theta \\ 0 & \text{else} \end{cases}$$

$$E(\hat{\theta}_{MLE}) = \int_0^\theta x f_{X_{(n)}}(x) dx = \frac{n}{\theta^n} \int_0^\theta x^n dx = \frac{n\theta}{n+1} \neq \theta \quad \text{有偏} \quad \frac{n\theta}{n+1} \xrightarrow{n \rightarrow \infty} \theta$$

$$D(\hat{\theta}_{MLE}) = \int_0^\theta x^2 f_{X_{(n)}}(x) dx - \left(\frac{n\theta}{n+1}\right)^2 = \frac{n\theta^2}{n+2} - \left(\frac{n\theta}{n+1}\right)^2 = \frac{n\theta^2}{(n+2)(n+1)^2} \xrightarrow{n \rightarrow \infty} 0$$

$$\lim_{n \rightarrow +\infty} E(\hat{\theta}_{MLE}) = \theta, \lim_{n \rightarrow +\infty} D(\hat{\theta}_{MLE}) = 0 \quad \text{故 } \hat{\theta}_{MLE} \text{ 为 } \theta \text{ 的相合估计。}$$

$$4) e^\theta \text{ 的矩估计是 } e^{\hat{\theta}_M} = e^{2\bar{X}}$$

因为 e^θ 具有单值反函数, 所以其最大似然估计为 $e^{\hat{\theta}_{MLE}} = e^{X_{(n)}}$

六、解： $n_1 = n_2 = 12, \bar{X} = 5.25, \bar{Y} = 1.5, S_1^2 = 0.93182, S_2^2 = 1, S_1 = 0.9653, S_2 = 1$

1) 检验统计量 $Z = S_1^2 / S_2^2 \sim F(n_1 - 1, n_2 - 1)$

如果 H_0 成立，则 Z 的取值应该在 1 附近，拒绝域为 $(0, F_{1-\alpha/2}) \cup (F_{\alpha/2}, +\infty)$

或 $S_1^2 / S_2^2 \leq F_{1-\alpha/2}(n_1 - 1, n_2 - 1)$ or $S_1^2 / S_2^2 \geq F_{\alpha/2}(n_1 - 1, n_2 - 1)$

拒绝域为 $(0, 0.2879) \cup (3.4737, +\infty)$

带入样本观察值，检验统计量的观察值为 $Z = 0.9318$ 在接受域内，故接受 H_0 ，认为两者的方差相等。

2) 求解均值差 $\mu_1 - \mu_2$ 的置信区间所需的枢轴量为

$$W = \frac{(\bar{X} - \bar{Y}) - (\mu_1 - \mu_2)}{S_w \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t(n_1 + n_2 - 2)$$

$$S_w^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2} = 0.9659 \quad S_w = 0.9828$$

$$\text{置信区间为} \quad \left((\bar{X} - \bar{Y}) \mp t_{\alpha/2}(n_1 + n_2 - 2) S_w \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} \right)$$

$(2.9179, 4.5821)$ 。 $\mu_1 - \mu_2$ 置信区间下限大于 0，我们认为 μ_1 比 μ_2 大