

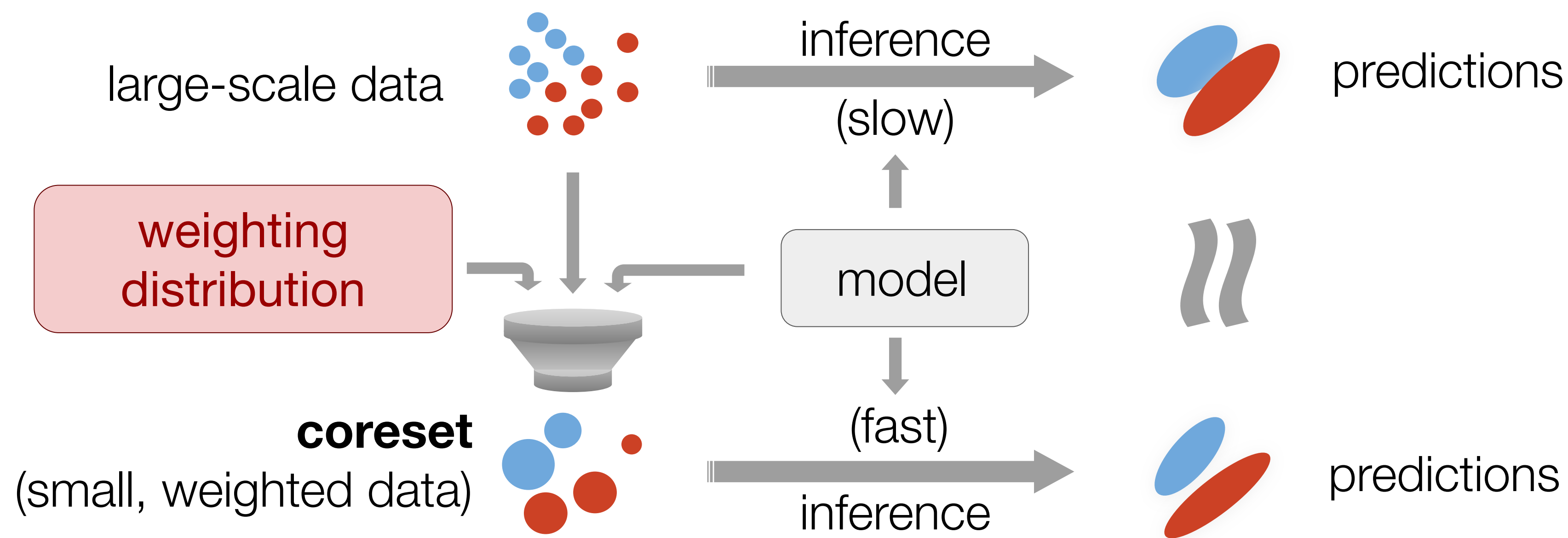


Summary

Automated algorithms in Bayesian statistics have provided practitioners access to reproducible data analysis with complex models. But obtaining scalability, guarantees, and automation together remains a challenge.

Bayesian coresets

Scalable inference with **statistical guarantees** via **data summarization**.



$$\pi_w(\theta) = \frac{1}{Z_w} \pi_0(\theta) \prod_{n=1}^N \exp(f_n(\theta))^{w_n} = \pi_0(\theta) \exp(f(\theta)^T w - \log Z_w)$$

coreset posterior prior log-likelihood data point n weight

$$\begin{aligned} \pi_1 &: \text{exact posterior} & 1 \in \mathbb{R}_{\geq 0}^N, \|1\|_0 = N \\ \pi_w &: \text{sparse coreset posterior} & w \in \mathbb{R}_{\geq 0}^N, \|w\|_0 \leq M \ll N \end{aligned}$$

Contributions

Existing work requires **the choice of a fixed weighting distribution**. Using a novel **information-geometric perspective**, we show this fundamentally limits coreset quality, and develop a **new, tuning-free construction algorithm** with superior accuracy.

Previous State of the Art

Hilbert coresets [CB17,18]: sparse nonnegative least squares

1) discretize log-likelihoods

2) minimize distance to sum

$$\theta_1, \dots, \theta_S \stackrel{\text{i.i.d.}}{\sim} \hat{\pi} \quad \text{weighting distribution} \quad w^* = \arg \min_{w \in \mathbb{R}^N} \left\| \sum_{n=1}^N g_n - \sum_{n=1}^N w_n g_n \right\|_2^2$$

$$g_n = \frac{1}{\sqrt{S}} \begin{bmatrix} f_n(\theta_1) - \bar{f}_n \\ \vdots \\ f_n(\theta_S) - \bar{f}_n \end{bmatrix} \quad \bar{f}_n = \frac{1}{S} \sum_{s=1}^S f_n(\theta_s) \quad \text{s.t. } w \geq 0, \|w\|_0 \leq M$$

How should we pick $\hat{\pi}$, and what is its effect?

Sparse Variational Inference

Key Insight 1: Coreset posteriors form an **exponential family** with natural parameter w and sufficient statistic $f(\theta)$.

Hence, the objective of **sparse variational inference**,

$$w^* = \arg \min_{w \in \mathbb{R}^N} \mathcal{D}_{\text{KL}}(\pi_w \| \pi_1) \quad \text{s.t. } \|w\|_0 \leq M, w \geq 0$$

has tractable gradient:

$$\nabla_w \mathcal{D}_{\text{KL}}(\pi_w \| \pi_1) = \text{Cov}_{\pi_w}[f, f^T(w-1)]$$

Problem: Estimating $\nabla_w \mathcal{D}_{\text{KL}}$ requires sampling from π_w .

Key Insight 2: Sampling from π_w is practical for sparse w .

SparseVI: Iterative Greedy Algorithm

Initialize: weights $w \leftarrow 0 \in \mathbb{R}^N$, active index set $\mathcal{I} \leftarrow \emptyset$

Iterate:

Estimate $\hat{C} := \text{Corr}_{\pi_w}[f, f^T(w-1)] \in \mathbb{R}^N$ via $\theta_1, \dots, \theta_S \stackrel{\text{i.i.d.}}{\sim} \pi_w$

Select data point: add $\arg \max_n \hat{C}_n$ to \mathcal{I}

Update w : SGD for active set \mathcal{I} using $\nabla_w \mathcal{D}_{\text{KL}}$ estimates from

Information-Geometric Perspective

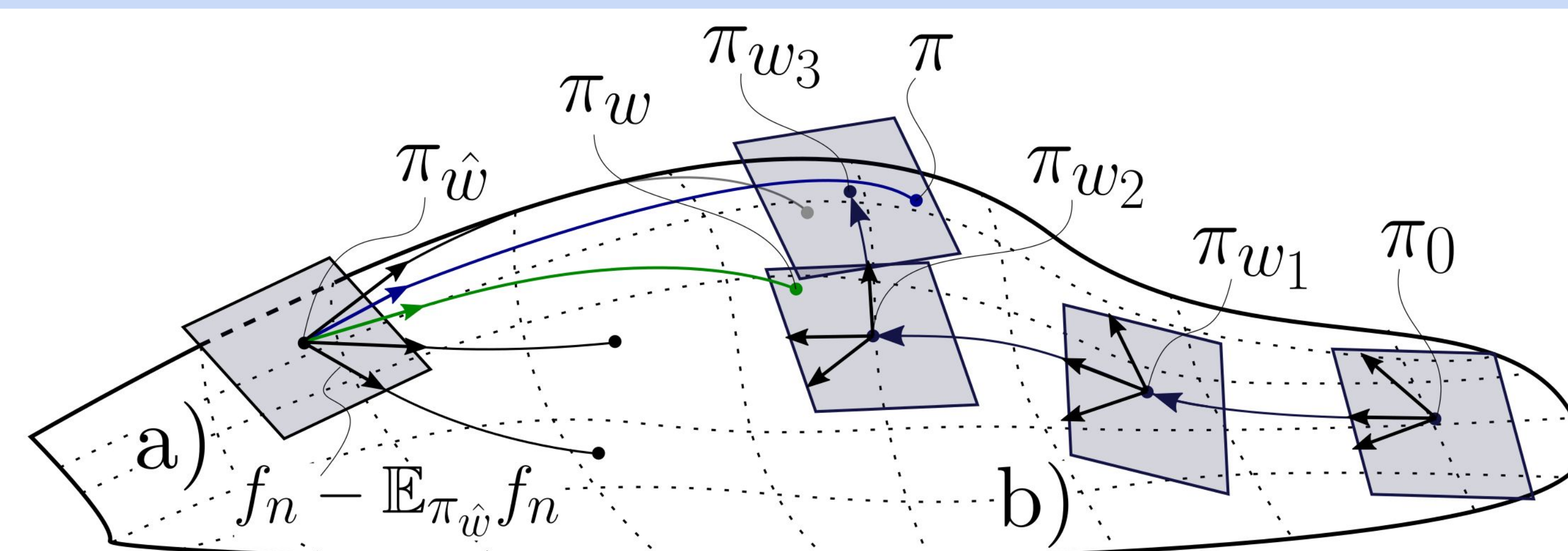
The Fisher information metric on the **coreset posterior manifold** is

$$G(w) = \mathbb{E}_{\pi_w}[(\nabla_w \log \pi_w)(\nabla_w \log \pi_w)^T] = \nabla_w^2 \log Z_w = \text{Cov}_{\pi_w}[f]$$

We show that **past constructions operate on a fixed tangent space**, whereas **SparseVI** is a Riemannian optimization algorithm that adapts *iteratively* to the manifold geometry.

Theorem: Both Hilbert coreset construction and *each iteration* of SparseVI are equivalent to local alignment of geodesic initial tangents:

$$w^* = \arg \min_{w \in \mathbb{R}^N} \|\xi_{\hat{w} \rightarrow 1} - \xi_{\hat{w} \rightarrow w}\|_{G(\hat{w})} \quad \text{s.t. } \|w\|_0 \leq M, w \geq 0$$

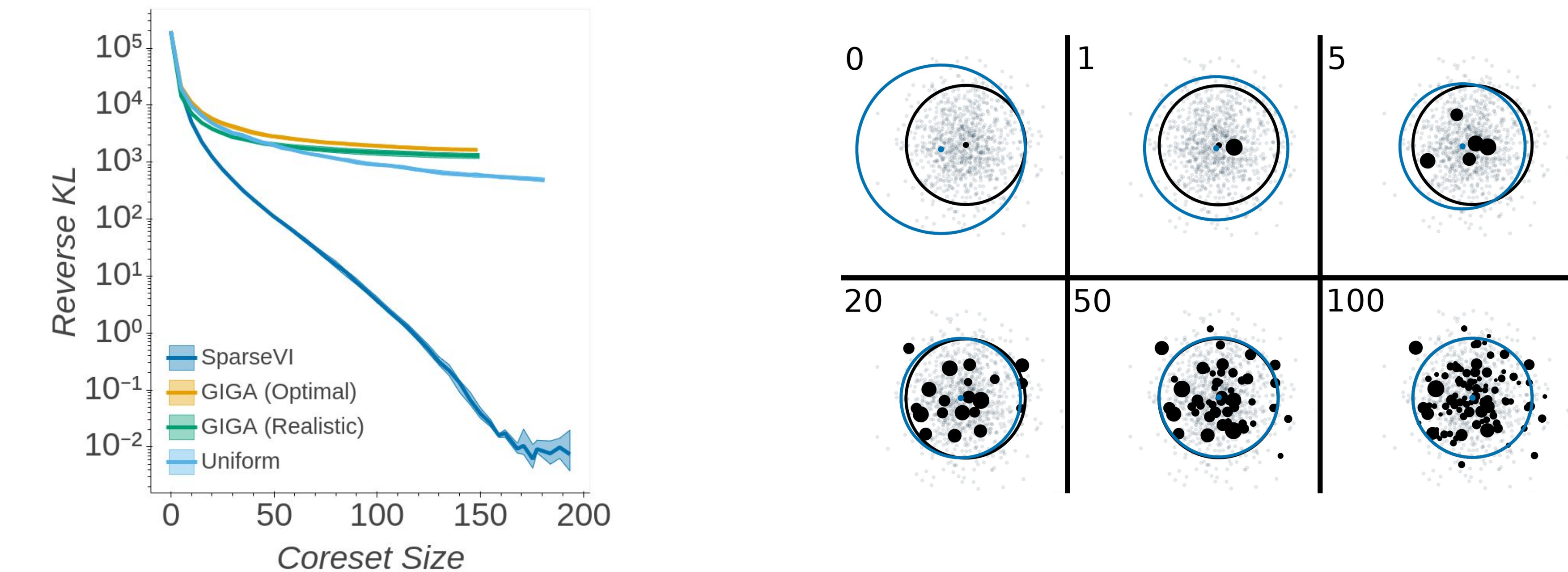


Experimental Results

Coreset post. convergence of **SparseVI** vs. **Uniform** subsampling and GIGA [CB18] with weighting distributions: **Exact (Optimal)** / **Noisy (Realistic)**

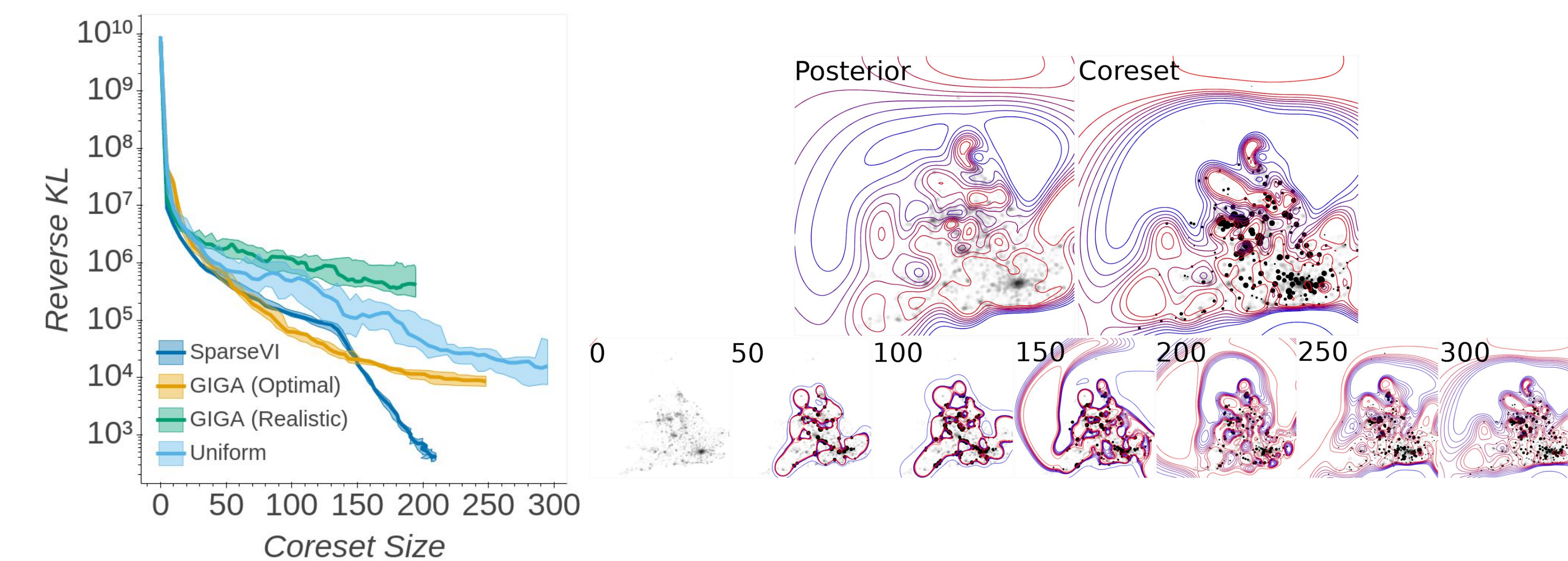
Synthetic Gaussian

200 dimensions, 1K samples



Basis Function Regression

301 dimensions, 10K records (2018 UK Land Registry Dataset)



Logistic & Poisson Regression

6 datasets, 2-15 dimensions, 500 data points

