CSC311 Assignment3

Peiqing Yu 2020/11/21 1003769004

Q1(a).

Computing the log-sum-exp of a=[-100000 -100000 -100000]

Unstable: -inf

Stable: -99998.90138771133

Computing the log-sum-exp of b=[100000 100000 100000]

Unstable: inf

Stable: 100001.09861228867

Q1(b).

Proof: Let
$$S = \max_{j=0}^{k} \{a_j\}$$
, Let $Y = \log(\sum_{j=0}^{k} \exp(a_i))$
Then LHS = $\log(\sum_{j=0}^{k} \exp(a_i)) = y$

$$\sum_{j=0}^{k} \exp(a_i) = e^{y}$$

$$e^{-s} \sum_{j=0}^{k} \exp(a_i) = e^{y} e^{-s}$$

$$\sum_{j=0}^{k} \exp(a_{i-s}) = e^{y-s}$$

$$Y - S = \log(\sum_{j=0}^{k} \exp(a_{i-s}))$$

$$Y = \log(\sum_{j=0}^{k} \exp(a_{i-s})) + S$$

$$= RHS$$

This numerical stable version is more robust to underflow or overflow, because we shift the values in the exponent by the maximum value of log(p(x, i)). Therefore, if we have a[i] very large, we can use the maximum value to adjust it and have the largest value be exp(0) which prevents overflow. For each value of a is small, we can also use the maximum value to adjust it and prevent underflow.

Q2(a).

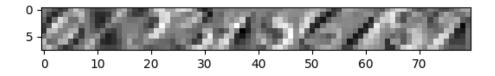
The average conditional log-likelihood on train set is -0.12462443666863023. The average conditional log-likelihood on test set is -0.1966732032552555.

Q2(b).

The accuracy on train set is 0.9814285714285714. The accuracy on test set is 0.97275.

The code to compute the accuracy is:

Q2(c). The leading eigenvectors for each covariance matrix plotted side by side are:



Q3(a).

$$P(\theta|D) = \frac{P(D|\theta) P(\theta)}{P(D)} \propto P(D|\theta) P(\theta)$$

$$\propto \left(\prod_{i=1}^{k} \prod_{k=1}^{k} \theta_{k} X_{k}^{(i)} \right) \left(\theta_{i}^{a_{i}-1} \theta_{2}^{a_{2}-1} \cdots \theta_{k}^{a_{k}-1} \right)$$

$$\propto \left(\theta_{i}^{x_{i}^{(i)}} X_{i}^{(k)} \cdots \theta_{i}^{x_{i}^{(k)}} \theta_{i}^{a_{i}-1} \right) - \cdots \left(\theta_{k}^{x_{k}^{(i)}} X_{k}^{(k)} \cdots \theta_{k}^{x_{k}^{(k)}} \theta_{k}^{a_{k}-1} \right)$$

$$\propto \theta_{i}^{\sum_{i=1}^{k} X_{i}^{(i)} + a_{i}-1} \cdots \theta_{i}^{\sum_{k=1}^{k} X_{k}^{(i)} + a_{k}-1}$$

$$\propto \theta_{i}^{\sum_{i=1}^{k} X_{i}^{(i)} + a_{i}-1} \cdots \theta_{i}^{\sum_{k=1}^{k} X_{k}^{(i)} + a_{k}-1}$$

By examing this, we can identify the posterior distribution $P(\theta|D)$ follows dirichlet distribution with parameters b_1, b_2, \cdots, b_k where $b_k = \int_{-1}^{k} x_k^{(i)} + a_k = N_k + a_k$

Q3(b).

$$\frac{\partial}{\partial MAP} = \underset{\theta}{\operatorname{argmax}} p(\theta|\theta)$$

$$= \underset{\theta}{\operatorname{argmax}} p(\theta)p(D|\theta)$$

$$= \underset{\theta}{\operatorname{argmax}} \log (p(\theta)p(D|\theta))$$

$$= \underset{\theta}{\operatorname{argmax}} \log (p(\theta)p(D|\theta))$$

$$= \underset{\theta}{\operatorname{argmax}} \log (p(\theta)p(D|\theta))$$

$$= \underset{\theta}{\operatorname{constant}} + (N_1 + a_1 - 1) \log (\theta_1) + (N_2 + a_2 - 1) \log (\theta_2)$$

$$= \underset{\theta}{\operatorname{since}} \frac{\sum_{i=1}^{k} \theta_i}{\sum_{i=1}^{k-1} (N_1 + a_2 - 1) \log (\theta_i)} + (N_2 + a_2 - 1) \log (1 - \sum_{i=1}^{k-1} \theta_i)$$

$$= \underset{\theta}{\operatorname{dil}(\theta)} = \underset{\theta}{\operatorname{Ni+ai-1}} - \underset{\theta}{\operatorname{Ni+ai-1}} \log (1 - \sum_{i=1}^{k-1} \theta_i)$$

$$= \underset{\theta}{\operatorname{dil}(\theta)} = \underset{\theta}{\operatorname{Ni+ai-1}} - \underset{\theta}{\operatorname{Ni+ai-1}} = 0$$

$$\frac{\partial i}{\partial k} = \frac{Ni + \alpha i - 1}{Nk + \alpha k - 1}$$
We assure that $\frac{\partial i}{\partial k}$ is maximized by checking l''

Since we have $\frac{\partial i}{\partial k} + \frac{\partial i}{\partial k} + \cdots + \frac{\partial k}{\partial k} = \frac{\sum_{i=1}^{k} \hat{\theta} i}{\hat{\theta} k} = \frac{1}{\hat{\theta} k}$

$$\frac{Ni + \alpha_i - 1}{Nk + \alpha_k - 1} + \frac{Nz + \alpha_z - 1}{Nk + \alpha_k - 1} + \cdots + \frac{N_k + \alpha_k - 1}{Nk + \alpha_k - 1} = \frac{1}{\hat{\theta} k}$$

$$\frac{\sum_{i=1}^{k} Ni + \alpha_i - k}{Nk + \alpha_k - 1} = \frac{1}{\hat{\theta} k}$$

$$\hat{\theta} = \frac{Nk + \alpha_k - 1}{\left(\sum_{i=1}^{k} Ni - \alpha_i\right) - k}$$

Thus the jth entry of Omap is Nj+aj-1 , where j=1.2,..., k

Q3(c).

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We know $P(X^{(N+1)}|D) = \int P(X^{(N+1)}|\theta) P(\theta|D) d\theta$; suppose $X^{(N+1)}k = 1$.

Then
$$\int P(x^{(N+1)}|\theta) P(\theta|D) d\theta$$

= $\int \theta_k^{(N+1)} P(\theta|D) d\theta$
= $\int \theta_k P(\theta|D) d\theta$

which follows distribution of 03(4)