Forecasting with SARIMA Models

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FEP.UP

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Forecasting is estimating how the sequence of observations will continue into the future

Types of data:

- Time series data collected at regular intervals over time
- Cross-sectional data are for a single point in time

Time series models

Time series models use only information on the variable to be forecast $y_{t+1} = f(y_t, y_{t-1}, y_{t-2}, \dots, error)$

where t is time and y_t is the quantity of interest at time t like: sales, electricity demand.

ARIMA models and exponential smoothing

- Useful when predictor variables not known or measured
- Useful if prediction of predictor variables difficult
- Does not lead to understanding of the system

Cross-sectional models

Cross-sectional models assume that variable to be forecast is affected by one or more **predictor variables**

$$y = f(x_1, x_2, \dots, error)$$

where x_1, x_2, \ldots are variables such as current temperature, GDP, population, time of the day, day of the week, etc regression models

Mixed models

 $y_{t+1} = f(y_t, y_{t-1}, y_{t-2}, \dots, x_1, x_2, \dots, error)$ dynamic regression models, panel data models, longitudinal models, transfer function models.

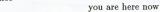
Statistical Forecasting

- Thing to be forecasted: a random variable y
- Forecast distribution: if \mathcal{F} represents all the observations, then $y|\mathcal{F}$ means the random variable y given what we know in \mathcal{F}
- ullet The point forecast is the mean (or median) of $y|\mathcal{F}$
- The forecast variance is $var(y|\mathcal{F})$
- A prediciton interval or interval forecast is a range of values of *y* with high confidence

Prediction and Forecasting

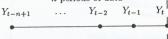
- Forecasting method: algorithm which produces a point forecast
- Statistical model: is a Data Generation Process may be used to
 - construct a probability distribution for y_{n+h} (from where one can obtain a point forecast)
 - construct confidence intervals for the forecasts

a. Point of reference





b. Past data available



c. Future forecasts required^a

$$m$$
 periods ahead $F_{t+1} ext{ } F_{t+2} ext{ } \dots ext{ } F_{t+m}$

d. Fitted values using a model^b

$$F_{t-n+1}$$
 ... F_{t-2} F_{t-1} F_t

- Time

Time

e. Fitting errors

$$(Y_{t-n+1}-F_{t-n+1}),\ldots,(Y_{t-1}-F_{t-1}),(Y_t-F_t)$$

f. Forecasting Errors (when Y_{t+1}, Y_{t+2} , etc., become available)

$$(Y_{t+1}-F_{t+1}), (Y_{t+2}-F_{t+2}), \ldots$$

Strategy



- Split the series into (i) training set (ii) test set
- Choose a method/model
- Start the forecasting procedure /Estimate the model using the training set
- Produce forecasts for the test set to obtain the forecasting accuracy measures
- Optimize parameters or model
- Obtain forecasts

Definitions and notations

- Aim: given a time series with n observations, y_1, y_2, \dots, y_n forecast $y_{n+1}, y_{n+2}, \dots, y_{n+h}$
- n is the origin, n+h is the forecast horizon and h is the number of steps-ahead
- $\hat{y}_n(h)$ denotes de (point) forecast of y_{n+h} , h = 1, 2, ...
- $\hat{y}_n(h)$ is a function of $\underline{y} = (y_1, y_2, \dots, y_n), \ \hat{y}_n(h) = g(\underline{y})$
- Forecast error at \hat{k} steps-ahead $e_n(\hat{k}) = y_{n+1} \hat{y}_n(\hat{k})$
- Criterion to compute $\hat{y}_n(h)$: minimize mean squared error $\sum_{i=1}^h (\hat{y}_n(i) y_{n+i})^2$

Criterion: minimum squared error

$$Ci \left[E(X-e)^2 \right]$$
 minigr $\left[E(X-e)^2 \right]$

- If X is a r.v. with $E(X) = \mu$ and $V(X) = \sigma^2$ then $E(X c)^2$ is minmum for $c = \mu$
- If Y is a r.v. and h(X) is a function of X then $E(Y h(X))^2$ is minimum for h(X) = E(Y|X)E(Y-E(Y|X)) is purposum
- Then

$$\hat{y}_{n}(h) = E(y_{n+h}|y_{1}, y_{2}, \dots, y_{n})$$

$$\hat{y}_{m}(h) = E(y_{m+h}|y_{1}, y_{2}, \dots, y_{n})$$

$$\hat{y}_{m+1} = E(y_{m+1}|y_{m}|y_{m-1}, \dots, y_{n})$$

$$\hat{y}_{m+1} = E(y_{m+1}|y_{m}, y_{m-1}, \dots, y_{n})$$

Criteria to evaluate forecasts

- Scale dependent measures
 - ▶ Mean Error, **ME**, $(1/h) \sum_{i=1}^{h} e_n(i)$

 - ▶ Mean Absolute Error, **MAE**, $(1/h) \sum_{i=1}^{h} |e_n(i)|$ ▶ Mean Squared Error, **MSE**, $(1/h) \sum_{i=1}^{h} (e_n(i))^2$
- Relative measures: Mean Absolute Percentual Error, MAPE, Relative measures. With |x| = 100 $\frac{1}{h} \sum_{i=1}^{h} |e_n(i)/y_{n+i}| \times 100$ $\frac{1}{h} \sum_{i=1}^{h} \frac{1}{|y_{m+i}-y_{m}(i)|} \times 100$
- Scale independent measures: Mean Absolute Scaled Error, MASE (Hyndman and Koehler, 2006) 5.11) = Ym Define the scaled error as:

$$q_n(i)=rac{e_n(i)}{1/(n-1)\sum_{i=2}^n|y_i-y_{i-1}|}$$
MASE=mean($|q_n(i)|$) preast ever you main fract

Cont.

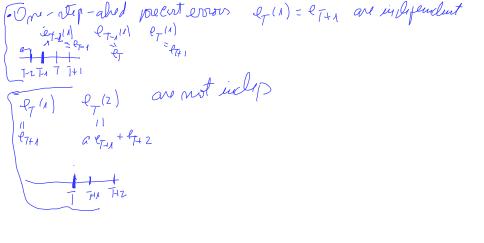
Theil's U-statistic

$$U_1 = \frac{\sqrt{\frac{1}{h} \sum_{i=1}^{h} e_n(i)^2}}{\sqrt{\frac{1}{h} \sum_{i=1}^{h} y_{n+i}^2} + \sqrt{\frac{1}{h} \sum_{i=1}^{h} \hat{y}_n^2(i)}}$$

takes values between 0 and 1 and values near 0 indicate an higher precision in the prediction.

$$\begin{array}{lll} Y_{t} & ARV_{1} & Y_{t} = \alpha \ Y_{t-1} + \theta_{t} & \theta_{t} \ NV(0, \frac{G_{t}}{I+\alpha t}) \\ Y_{1}, & Y_{T} & Y_{T} & Y_{T} & Y_{T}, Y_{T-1}, & Y_{T} & Y_{T} \\ \hline Y_{1} & = \stackrel{?}{Y_{T}} (1) = E \left(Y_{T+1} \mid Y_{T}, Y_{T-1}, & Y_{T} \right) \\ & = E \left(\alpha \ Y_{T} + \theta_{T+1} \mid Y_{T}, Y_{T-1}, & Y_{T} \right) \\ & = \alpha \ E \left(Y_{T} \mid Y_{T}, Y_{T-1}, & Y_{T} \right) + E \left(\theta_{T+1} \mid Y_{T}, Y_{T-1}, & Y_{T} \right) \\ & = \alpha \ Y_{T} + O = \alpha \ Y_{T} \\ \hline \theta_{T} & = (\theta_{T} \cdot Y_{T}) = (\theta_{T+1}) = 0 \\ \hline V\left(\theta_{T} \cdot Y_{T} \right) = E\left(\theta_{T+1} \right) = 0 \\ \hline V\left(\theta_{T} \cdot Y_{T} \right) = V\left(\theta_{T+1} \right) = 0 \\ \hline V\left(\theta_{T} \cdot Y_{T} \right) = V\left(\theta_{T} \cdot Y_{T} \right) = 0 \\ \hline V\left(\theta_{T} \cdot Y_{T} \right) = V\left(\theta_{T} \cdot Y_{T} \right) = 0 \\ \hline V\left(\theta_{T} \cdot Y_{T} \right) = V\left(\theta_{T} \cdot Y_{T} \right) = 0 \\ \hline V\left(\theta_{T} \cdot Y_{T} \right) = V\left(\theta_{T} \cdot Y_{T} \right) = 0 \\ \hline A_{T} & = 0 \\ \hline A$$

/ (2) = E (/ T+Z | YT) YT-1) -- 14,) = E (a YT+1 + PAZ | YT1 ---, Y1) 5 a E (YT+1 | YT -- Y1) + E (e T+2 | YT1- N1) = a ý_11) + o Ý-(h) = E (YT+h | YT,--, Y,) = E (a YT+h-,+++h | YT,--, Y,) = a yT (h-1) = a yT = 0 as h indeaes the ARU) poent buch to the mean of process (7/2) = /1+2- ×7/1) = a YT+1 + lt2 - a XT(1) = a (YT+1 - YT(1)) + PT+2 = a 27(1) + PTEZ = a PT+1+PT+Z E(fila)=0 = E(x16)=1/2 V Maria Eduarda Silva (FEP.UP) Forecasting with SARIMA Models



Forecasting with AR(1)

- $y_t = ay_{t-1} + e_t$, |a| < 1 $e_t \sim N(0, \sigma^2)$
- Forecast y_{n+h} given y_1, \ldots, y_n
- Note:

$$y_{n+1} = ay_n + e_{n+1}$$

$$y_{n+2} = ay_{n+1} + e_{n+2} = a(ay_n + e_{n+1}) + e_{n+2}$$

$$= a^2y_n + ae_{n+1} + e_{n+2}$$

$$y_{n+3} = ay_{n+2} + e_{n+3} = a(a^2y_n + ae_{n+1} + e_{n+2}) + e_{n+3}$$

$$= a^3y_n + a^2e_{n+1} + ae_{n+2} + e_{n+3}$$

$$\vdots$$

$$\vdots$$

$$y_{n+h} = a^hy_n + a^{h-1}e_{n+1} + \dots + e_{n+h}$$

$$= a^hy_n + \sum_{i=0}^{h-1} a^je_{n+h-j}$$

then *h* steps-ahead forecast is:

$$\hat{y}_{n}(h) = E(y_{n+h}|y_{n})
= E(a^{h}y_{n} + \sum_{j=0}^{h-1} a^{j}e_{n+h-j})|y_{n}
= a^{h}y_{n} + E(\sum_{j=0}^{h-1} a^{j}e_{n+h-j}|y_{n})
= a^{h}y_{n} + \sum_{j=0}^{h-1} a^{j}E(e_{n+h-j}|y_{n})
= a^{h}y_{n}$$

Note that the forecast tends to the mean of the process!

Prediction error

• $\hat{e}_n(h) = (\hat{y}_n(h) - y_{n+h})$ h-steps ahead prediction with origin at n

$$\hat{e}_{n}(h) = y_{n+h} - \hat{y}_{n}(h)
= a^{h}y_{n} - a^{h}y_{n} + \sum_{j=0}^{h-1} a^{j}e_{n+h-j}
= \sum_{j=0}^{h-1} a^{j}e_{n+h-j}$$

and the variance of the prediction error

• Note that $E(\hat{e}_n(h)) = 0$

$$\operatorname{var}(\hat{e}_n(h)) = \operatorname{var}(\hat{y}_n(h) - y_{n+h})$$
$$= \operatorname{var}(\sum_{j=0}^{h-1} a^j e_{n+h-j})$$

but e_t are iid, then

•

•

$$\operatorname{var}(\hat{e}_n(h)) = \sigma_e^2 \sum_{j=0}^{h-1} a^{2j}$$

Interpretation

- For h=1 $\hat{e}_n(1)=e_{n+1}$ that is the reason why e_t are called innovations: e_t represents the news or surprise at each time period, the quantity that is not predictable.
- The forecasts are unbiased: $E(\hat{y}_n(h)) = y_{n+h}$
- The variance of the prediction error depends on: a, σ_e^2 and the horizon h
- The variance of the prediction error increases with *h*.
- The variance of the prediction error tends to the value $\sigma_e^2/(1-a^2)$ which is the same as the variance of the process.
- In practice a and σ_e^2 are replaced by its estimated values \hat{a} and $\hat{\sigma}_e^2$, respectively.
- Assuming $e_t \sim N(0, \sigma_e^2)$ a $\gamma\%$ confidence interval is given by: $\hat{y}_n(h) \pm 2\sqrt{\frac{(1-\hat{a}^{2k})}{(1-\hat{a}^2)}}\hat{\sigma}_e^2$

Forecasting with AR(p): notation

- $y_t = a_1 y_{t-1} + a_2 y_{t-2} + ... + a_p y_{t-p} + e_t$ where
 - $e_t \sim N(0,1)$
 - ▶ $a_1, a_2, ..., a_p$ are such that the roots $z_1, z_2, ldots, z_p$ of the AR polinomial $\phi(z) = 1 a_1 Z a_2 Z^2 ... a_p Z^p$ are $|z_i| > 1$
- Having observed y_1, \ldots, y_n forecast $y_{n+1}, y_{n+2}, \ldots, y_{n+k}$
- Let $\hat{y}_n(k) = \mathrm{E}(y_{n+k}|y_1,\ldots,y_n)$

Forecast with AR(p): 1 step-ahead

$$\hat{y}_n(1) = a_1 y_n + a_2 y_{n-1} + \dots + a_p y_{n+1-p}$$

$$\hat{e}_n(1) = y_{n+1} - \hat{y}_n(1)$$

$$= a_1 y_n + a_2 y_{n-1} + \dots + a_p y_{n+1-p} + e_{n+1}$$

$$-(a_1 y_n + a_2 y_{n-1} + \dots + a_p y_{n+1-p} + e_{n+1})$$

$$= e_{n+1}$$

$$\operatorname{var}(\hat{e}_n(1)) = \sigma_e^2$$

Forecast with AR(p): 2 steps-ahead

$$\hat{y}_{n}(2) = E(a_{1}y_{n+1} + a_{2}y_{n} + \dots + a_{p}y_{n+2-p} + e_{n+2}|y_{1}, \dots, y_{n})$$

$$= a_{1}\hat{y}_{n}(1) + a_{2}y_{n} + \dots + a_{p}y_{n+2-p}$$

$$\hat{e}_{n}(2) = y_{n+2} - \hat{y}_{n}(2)$$

$$= a_{1}y_{n+1} + a_{2}y_{n} + \dots + a_{p}y_{n+2-p} + e_{n+2}$$

$$-(a_{1}\hat{y}_{n}(1) + a_{2}y_{n} + \dots + a_{p}y_{n+2-p})$$

$$= a_{1}\hat{e}_{n}(1) + e_{n+2}$$

$$var(\hat{e}_{n}(2)) = (1 + a_{1}^{2})\sigma_{e}^{2}$$

Then for k-steps ahead:

$$\hat{y}_n(h) = E(a_1 y_{n+h-1} + a_2 y_{n+h-2} + \dots + a_p y_{n+h-p} + e_{n+h} | y_1, \dots, y_n)
= a_1 \hat{y}_n(h-1) + a_2 \hat{y}_n(h-2) + \dots + a_p \hat{y}_n(h-p)$$

where

$$\hat{y}_n(j) = \mathrm{E}(y_{n+j}|y_1,\ldots,y_n) = \left\{ egin{array}{ll} y_{n+j}, & 1 \leq n+j \leq n, \\ 0, & n+j \leq 0 \end{array} \right.$$

and the prediction error and its variance are:

$$\hat{e}_n(h) = y_{n+h} - \hat{y}_n(h)$$

= $a_1 \hat{e}_n(h-1) + a_2 \hat{e}_n(h-2) + \ldots + a_p \hat{e}_n(h-p)$

where

$$\hat{\mathbf{e}}_n(j) = \left\{ egin{array}{ll} \mathbf{e}_{n+j}, & 1 \leq n+j \leq n, \\ 0, & n+j \geq 0 \end{array} \right.$$

NOTE: for fixed n $\hat{e}_n(h_1)$ and $\hat{e}_n(h_2)$ $h_1 \neq h_2$ are correlated!!! To compute $var(\hat{e}_n(h))$, $\hat{e}_n(h)$ is written as a function of e_{n+1}, \ldots, e_{n+h}

Example: MA(1)

$$y_1 - \cdots y_n$$
 y_{m+h} ?
 $\hat{y}_{n}(h) = E(y_{m+h} | y_m - \cdots y_n)$

Forecast
$$h=2$$
 steps-ahead $X_t=0.4e_{t-1}+e_t,\,\sigma_e^2=1.5$ Note that

$$\sum_{j=0}^{\infty} \pi_j y_{t-j} = e_t$$

$$\frac{1}{2} = \frac{1}{2} + \frac{1$$

Forecasting with SARIMA Models

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$$\begin{array}{l}
\mathcal{L}_{m}(1) = \gamma_{m+1} - \gamma_{m}(1) \\
= \frac{2}{120} \, \pi_{1} \, \gamma_{m+1-1} - \frac{2}{120} \, \pi_{1} \, \gamma_{m+1-1} = \mathcal{L}_{m+1} \\
\gamma_{m}(2) = \frac{1}{120} \left(\gamma_{m+2} \, \left| \, \gamma_{m_{1}} \, \gamma_{m_{2}} \, \gamma$$

```
ARMAtoMA(ar=-0.4,lag.max=50)
```

- [1] -4.000000e-01 1.600000e-01 -6.400000e-02 2.560000e-02 -1.024000e-02
- [9] -2.621440e-04 1.048576e-04 -4.194304e-05 1.677722e-05 -6.710886e-06 [17] -1.717987e-07 6.871948e-08 -2.748779e-08 1.099512e-08 -4.398047e-09
- [25] -1.125900e-10 4.503600e-11 -1.801440e-11 7.205759e-12 -2.882304e-12
- [33] -7.378698e-14 2.951479e-14 -1.180592e-14 4.722366e-15 -1.888947e-15
- [41] -4.835703e-17 1.934281e-17 -7.737125e-18 3.094850e-18 -1.237940e-18
- 1.267651e-20
- [49] -3.169127e-20

```
> ARMAtoMA(ar=-0.3,ma=-0.4,lag.max=50)
```

- [1] -7.000000e-01 2.100000e-01 -6.300000e-02 1.890000e-02 -5.670000e-03
- [9] -4.592700e-05 1.377810e-05 -4.133430e-06 1.240029e-06 -3.720087e-07
- [17] -3.013270e-09 9.039811e-10 -2.711943e-10 8.135830e-11 -2.440749e-11
- [25] -1.977007e-13 5.931020e-14 -1.779306e-14 5.337918e-15 -1.601375e-15
- [33] -1.297114e-17 3.891342e-18 -1.167403e-18 3.502208e-19 -1.050662e-19
- [41] -8.510366e-22 2.553110e-22 -7.659329e-23 2.297799e-23 -6.893396e-24
- [49] -5.583651e-26 1.675095e-26

ARMA(1,1)- forecasts based on the infinite past

- $v_t = av_{t-1} + be_{t-1} + e_t$ where |a| < 1 e |b| < 1
- Having observed y_1, \ldots, y_n forecast $y_{n+1}, y_{n+2}, \ldots, y_{n+h}$
- $\hat{y}_n(1) = aE(y_n|X1,...,y_n) + bE(e_n|X1,...,y_n)$
- Note that
 - $y_t = \sum_{j=0}^{\infty} \psi_j e_{t-j}$ $\sum_{j=0}^{\infty} \pi_j y_{t-j} = e_t$

then the knowledge of y_n, y_{n-1}, \dots is equivalent to the knowledge e_n, e_{n-1}, \ldots

cont

Thus

$$\begin{array}{lll}
\mathring{Y}_{m}(I) &=& E\left(\mathring{Y}_{m}t_{I} \mid \mathring{Y}_{M} \cdots \mathring{Y}_{A}\right) \\
&=& E\left(\mathring{X}_{M} + \mathring{Y}_{M} + \mathscr{C}_{M} + \mathscr{C}_{M} + \mathscr{C}_{M}\right) \\
\mathring{Y}_{n}(1) &=& ay_{n} + be_{n} \\
\mathring{Y}_{n}(h) &=& a\mathring{Y}_{n}(h - \overline{1}), h \geq 2 \\
\mathring{e}_{n}(1) &=& e_{n+1} \\
var(\hat{e}_{n}(1)) &=& \sigma_{e}^{2} \\
\mathring{e}_{n}(k) &=& a\hat{e}_{n}(k - 1) + be_{n+k-1} + e_{n+k}
\end{array} (2)$$

When computing $var(\hat{e}_n(k))$ one must keep in mind that with n fixed the prediction errors are dependent.

ARMA(1,1)- forecasts based on y_1, \ldots, y_n

In fact we have observed only y_1, \ldots, y_n thus let

$$E(e_k|X1,\ldots,y_n) = \hat{e}_n(-n+k), \ 1 \le k \le n$$

Since $e_t = y_t - ay_{t-1} - be_{t-1}$, we have:

$$\hat{e}_{n}(-n+1) = 0$$
 $\hat{e}_{n}(-n+2) = y_{2} - ay_{1}$
 $\hat{e}_{n}(-n+3) = y_{3} - ay_{2} - b\hat{e}_{n}(-n+2)$
 \vdots
 $\hat{e}_{n}(0) = y_{n} - ay_{n-1} - b\hat{e}_{n}(1)$

cont.

As such

$$\hat{X}_n(1) = ay_n + b\hat{e}_n(0)$$

 $\hat{X}_n(h) = a\hat{X}_n(h-1), h \ge 2$

 $var(\hat{e}_n(h)), h \ge 1$ is computed using (5)

Forecasting with ARMA(p, q)

$$\hat{y}_n(h) = a_1 \hat{y}_n(h-1) + \ldots + a_p \hat{y}_n(h-p) + b_1 E(e_{n+h-1}|y_n,\ldots) + \ldots + b_q E(e_{n+h-q}|y_n,\ldots)$$

where

$$\mathrm{E}(e_{n+k}|y_n,\ldots)=\hat{\mathrm{e}}_n(k)=\left\{\begin{array}{ll}e_{n+k},&1\leq n+k\leq n,\\0,&n+j\geq 0\end{array}\right.$$

and
$$\hat{X}_n(k) = y_{n-k}$$
 if $-(p-1) \le k \le 0$

For h > qthe forecasts are in fact obtained from the AR component of the ARMA model.

Confidence Intervals for the forecasts



Assuming $e_t \sim N(0, \sigma_e^2)$ a 7% confidence interval is given by:

$$\hat{y}_n(h) \pm 2\sqrt{\operatorname{var}(\hat{e}_n(h))}$$

From Hyndman: ARIMA vs Exponential Smoothing

Equivalences

Simple exponential smoothing

- Forecasts equivalent to **ARIMA(0,1,1)**.
- Parameters: $\theta_1 = \alpha 1$.

Holt's method

- Forecasts equivalent to **ARIMA(0,2,2)**.
- Parameters: $\theta_1 = \alpha + \beta 2$ and $\theta_2 = 1 \alpha$.

Damped Holt's method

- Forecasts equivalent to **ARIMA(1,1,2)**.
- Parameters: $\phi_1 = \phi$, $\theta_1 = \alpha + \phi \beta 2$, $\theta_2 = (1 \alpha)\phi$.

Holt-Winters' additive method

- Forecasts equivalent to $ARIMA(0,1,m+1)(0,1,0)_m$.
- Parameter restrictions because ARIMA has m+1 parameters whereas HW uses only three parameters.

Holt-Winters' multiplicative method

No ARIMA equivalence

8. ARIMA models

ARIMA vs ETS

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Forecasting with SARIMA Models

- Expand the SARIMA equation so that y_t is on the left hand side and all other terms are on the right
- Rewrite the equation by replacing t by T + h
- On the right hand side of the equation, replace future observations by their forecasts, future errors by zero and past errors by the corresponding residuals

Begin with h=1 and repeat the above steps are then repeated for $h = 2, 3, \dots$ until all forecasts have been calculated.

SARIMA [1,1,1) x (011,1) (2 m1 = -0.86

 $(1-B^{12})\times(1-B)(1-0.19B)\chi_{t}=(1-0.56B)(1-0.86B^{(2)})$ et $(1-B-B^{12}+B^{13})(1-0.19B)$ $\chi_{\xi} = (1-0.86B^{12}-0.56B+0.56\times0.86B^{13})$ ex

(1-B-B12+B13-0-19B+0-19B2+0-19B13-0-19B14) Xt=J

Xt = 1.19 Xt-1 = 0.19 Xt-2+ Xt-12= 1.19 Xt-13+0.15 Xt-14+ Pt+

(1-1,19B+0.19B2=B12+1,19B13-0.19B19) XE=