



# Time Varying Covariates in a Cox Model

Steve Simon

# Abstract



Time varying covariates allow you to account for non-proportional hazards and can model settings where patients switch from one therapy to another. You will code data for time-varying covariates, fit time-varying models, and interpret the results.

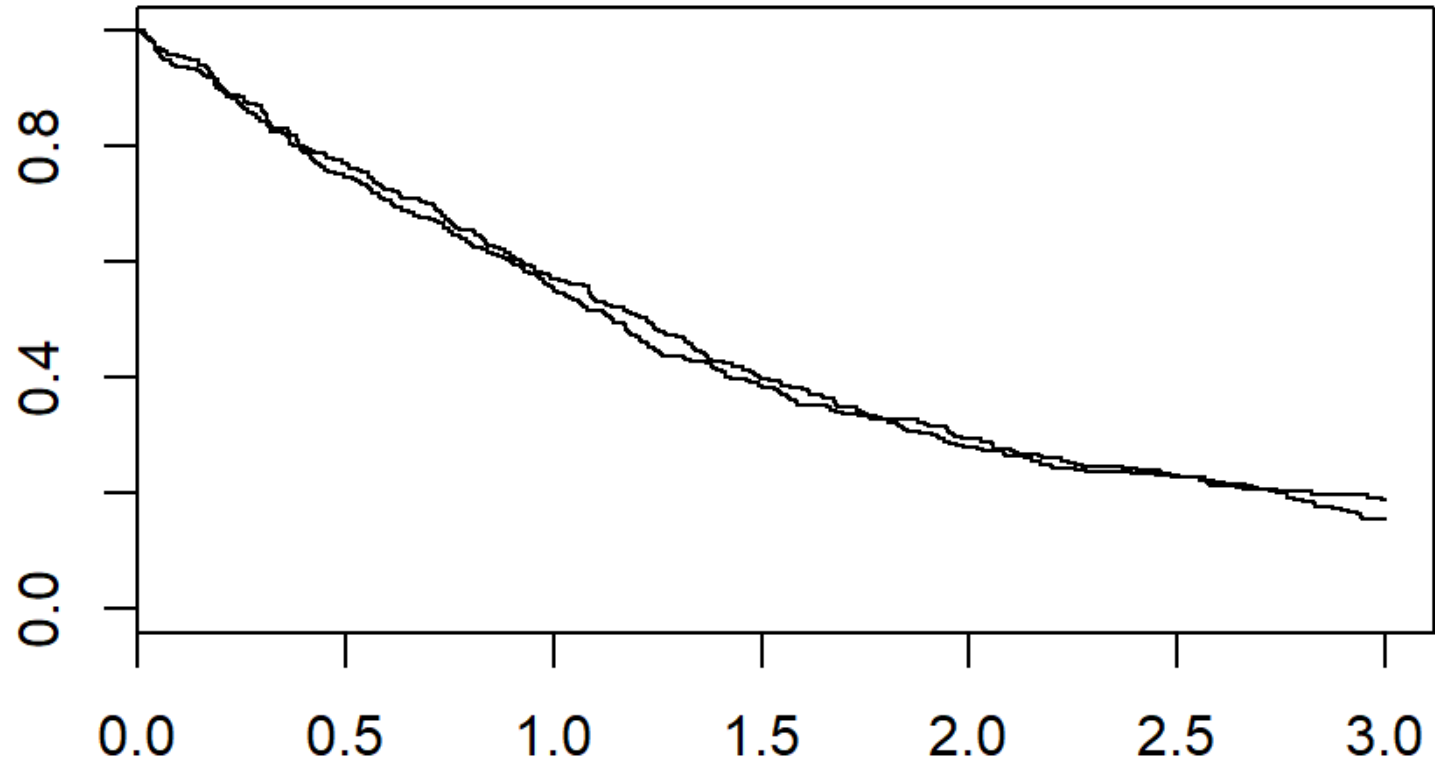


# Testing the Proportional Hazards Assumption

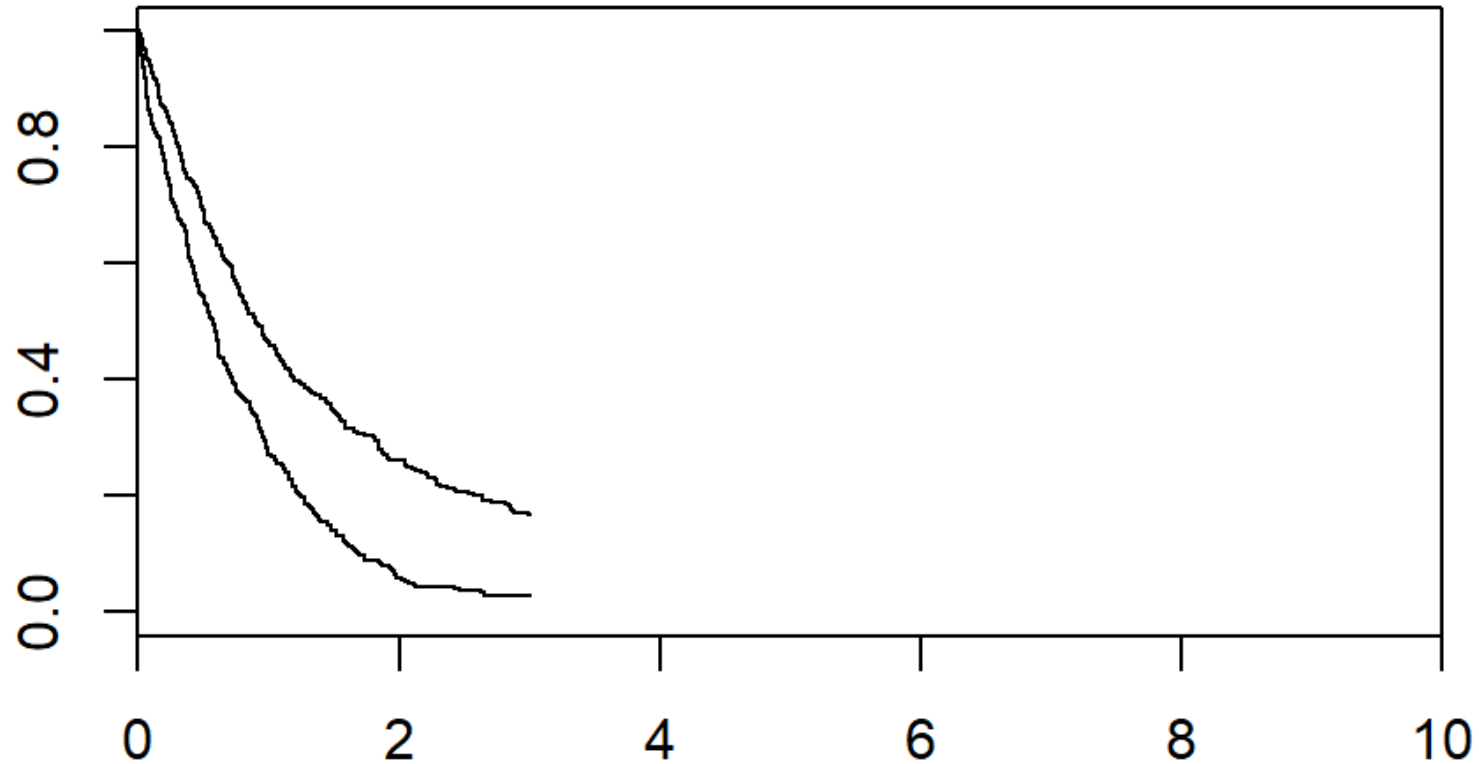
Several approaches

1. Patterns in Kaplan-Meier curves
2. Complementary log-log plot
3. Schoenfeld Residuals
4. Stratified analysis
5. Fit time varying covariates

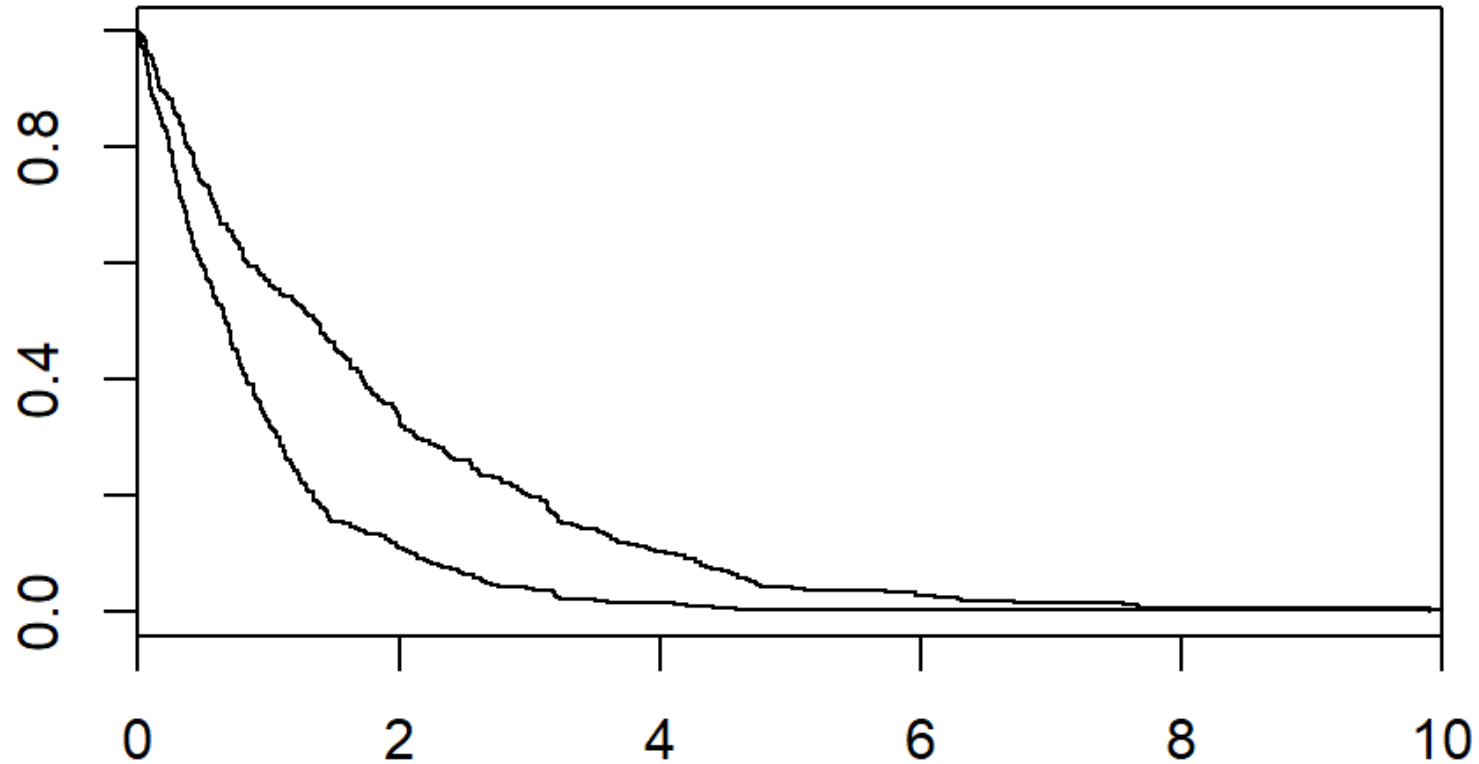
# Patterns in the Kaplan-Meier Curves



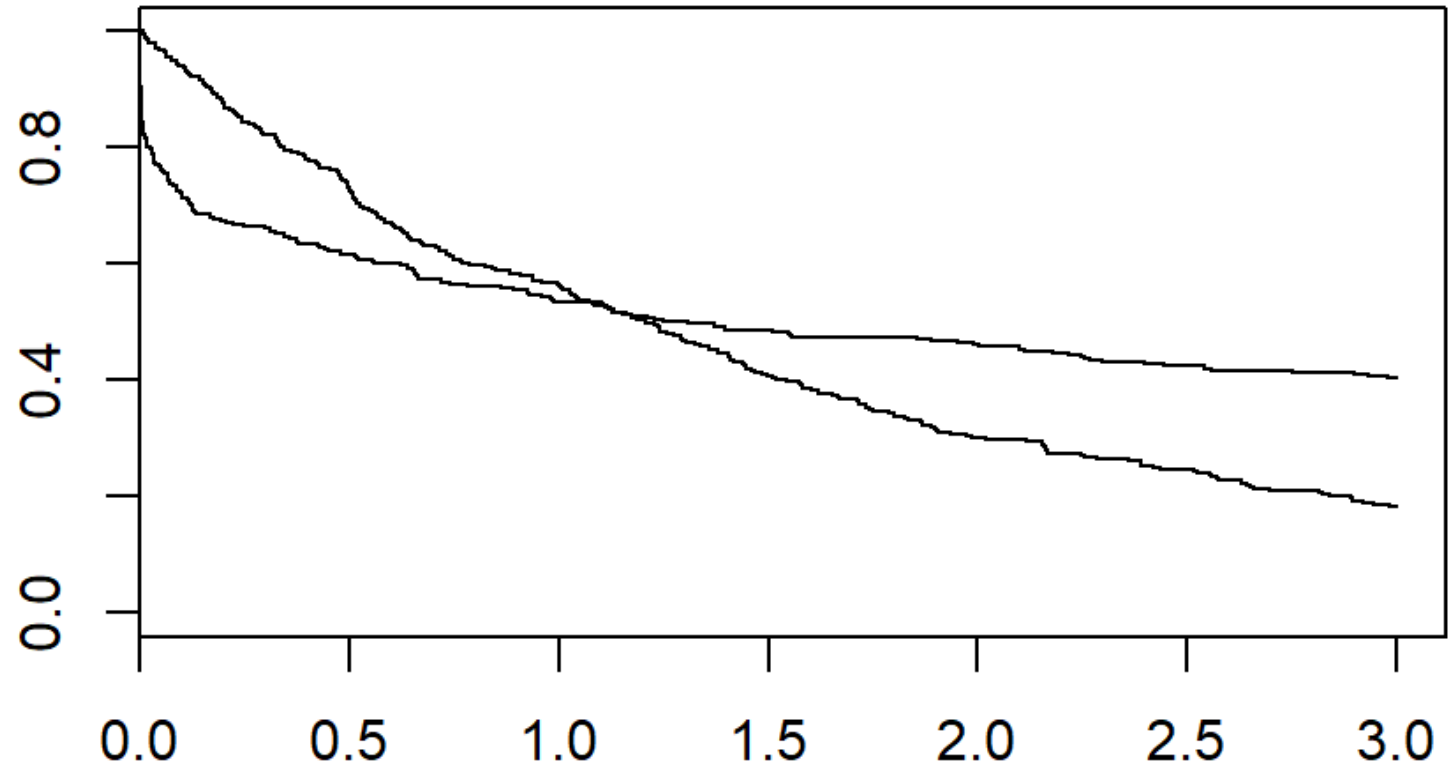
# Patterns in the Kaplan-Meier Curves



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# Patterns in the Kaplan-Meier Curves





# Example: Heroin Clinic Discharge Times

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## Methadone Treatment of Heroin Addicts

Keywords: censored data, Cox regression.

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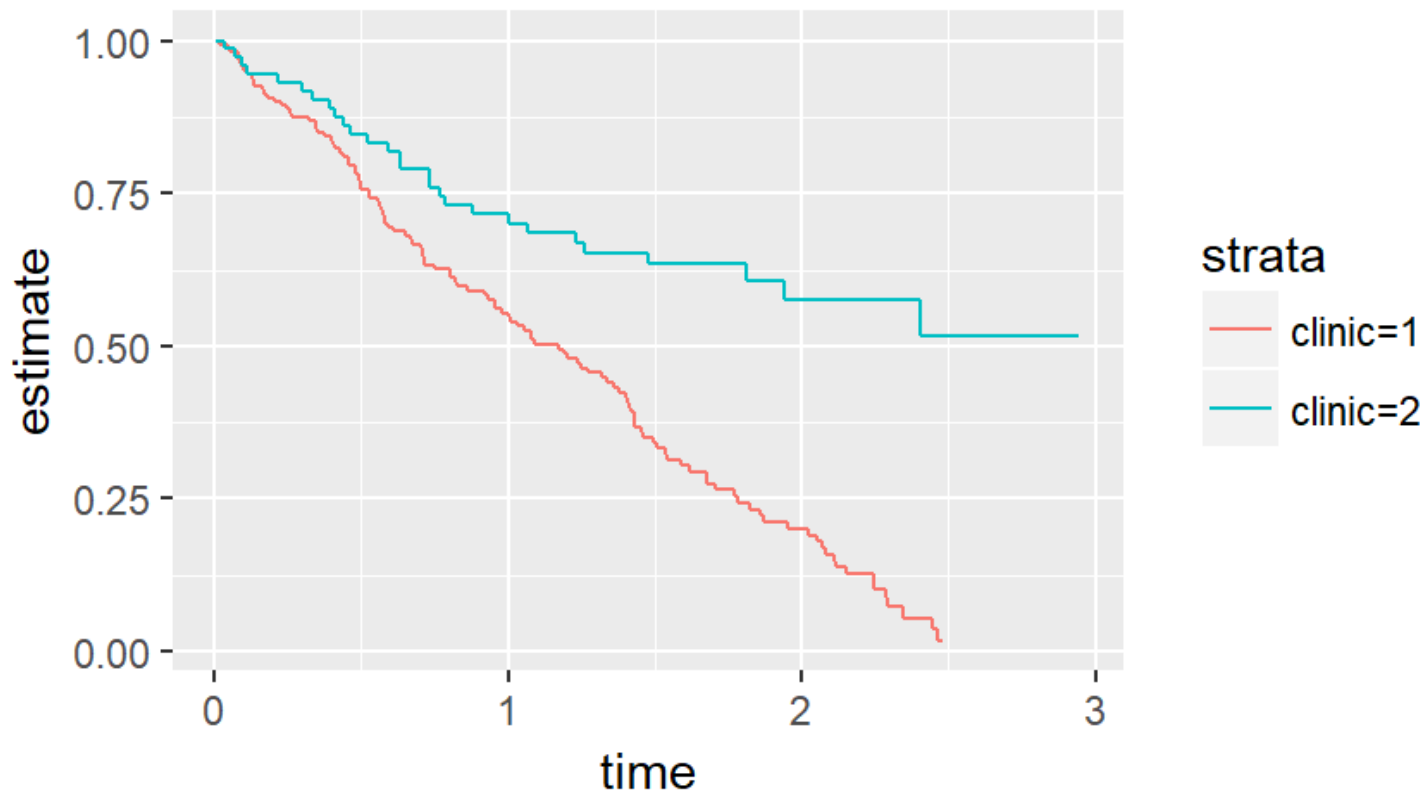
### Description

The data are the times, in days, that heroin addicts spend in a clinic. There are two clinics and the covariates are believed to affect the times spent in the clinic by addicts.

Variable	Description
Clinic	1 or 2
Status	0 = still in clinic at end of study (censored) or 1 = departed from clinic
Time	days spent in clinic
Prison	1 = prison record or 0 = no record
Dose	methadone dosage (mg/day)

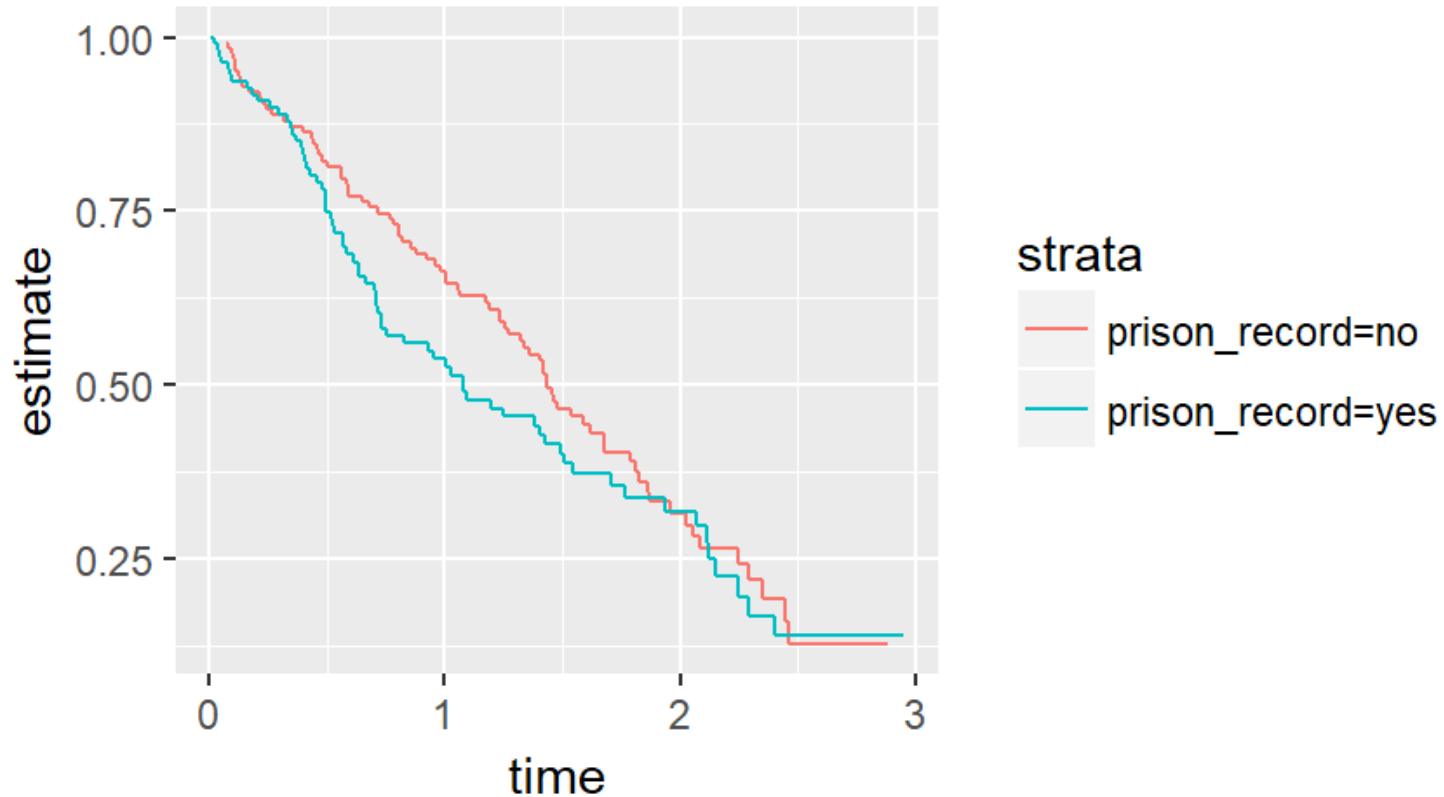


## Example: Heroin Data Set



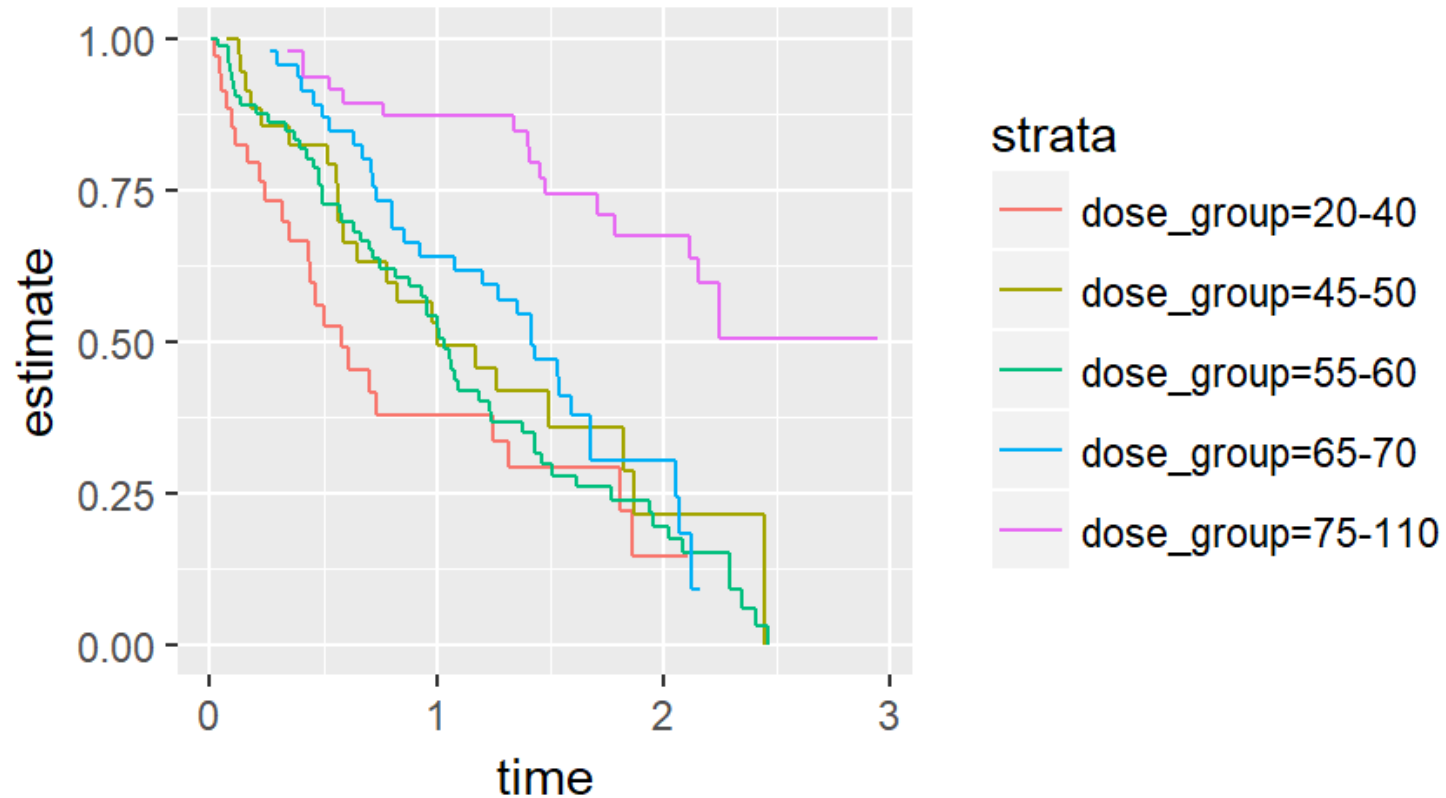


## Example: Heroin Data Set





## Example: Heroin Data Set





## Complementary Log-Log

If the proportional hazards assumption holds then the survival curves for two different groups should be related by

$$S_2(t) = S_1(t)^{hr}$$

Take the logarithm of both sides to get

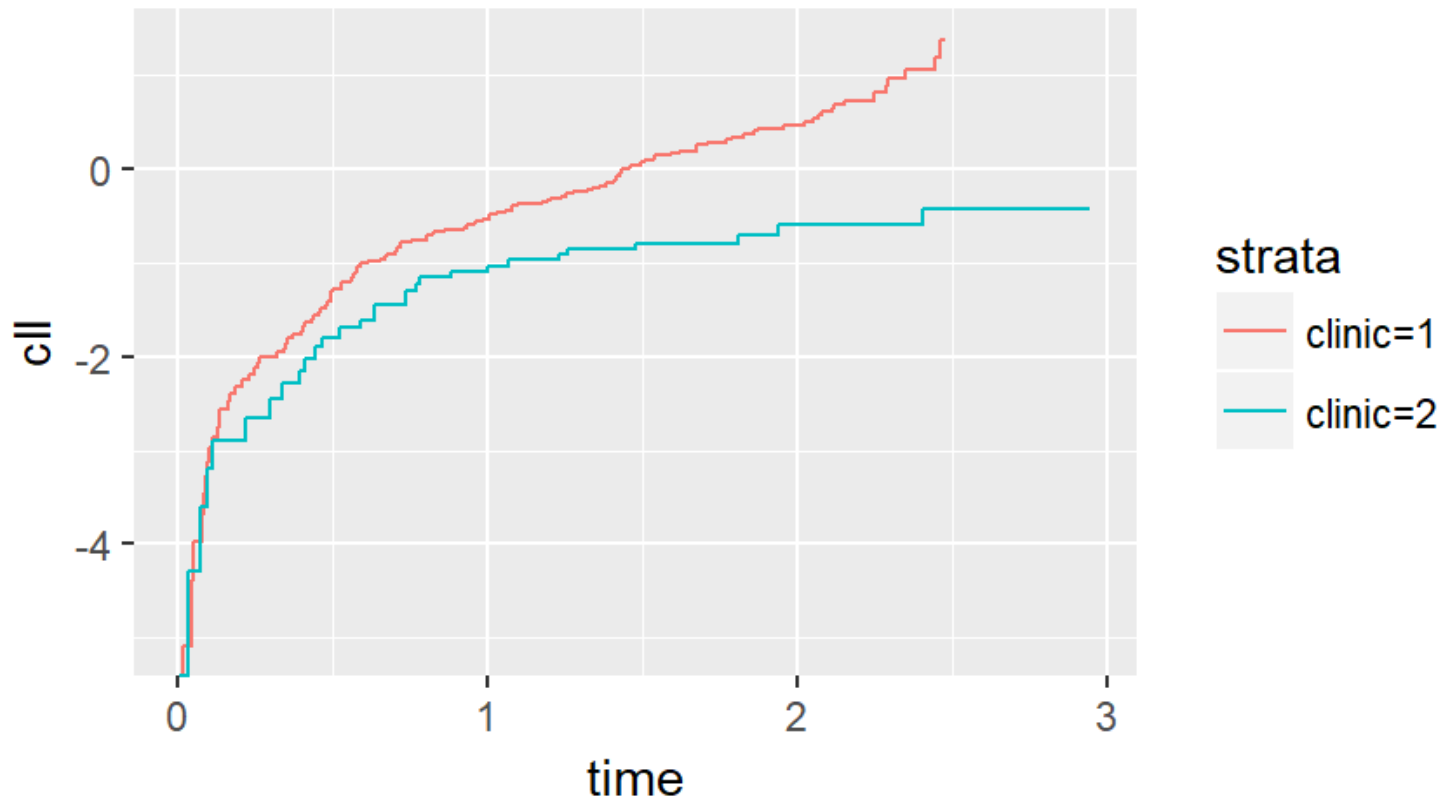
$$\log(S_2(t)) = \log(S_1(t))hr$$

We'd like to take a second log here, but since  $S_2$  and  $S_1$  are always between 0 and 1, their logarithms would be negative. You have to flip this to a positive value and then take a second logarithm.

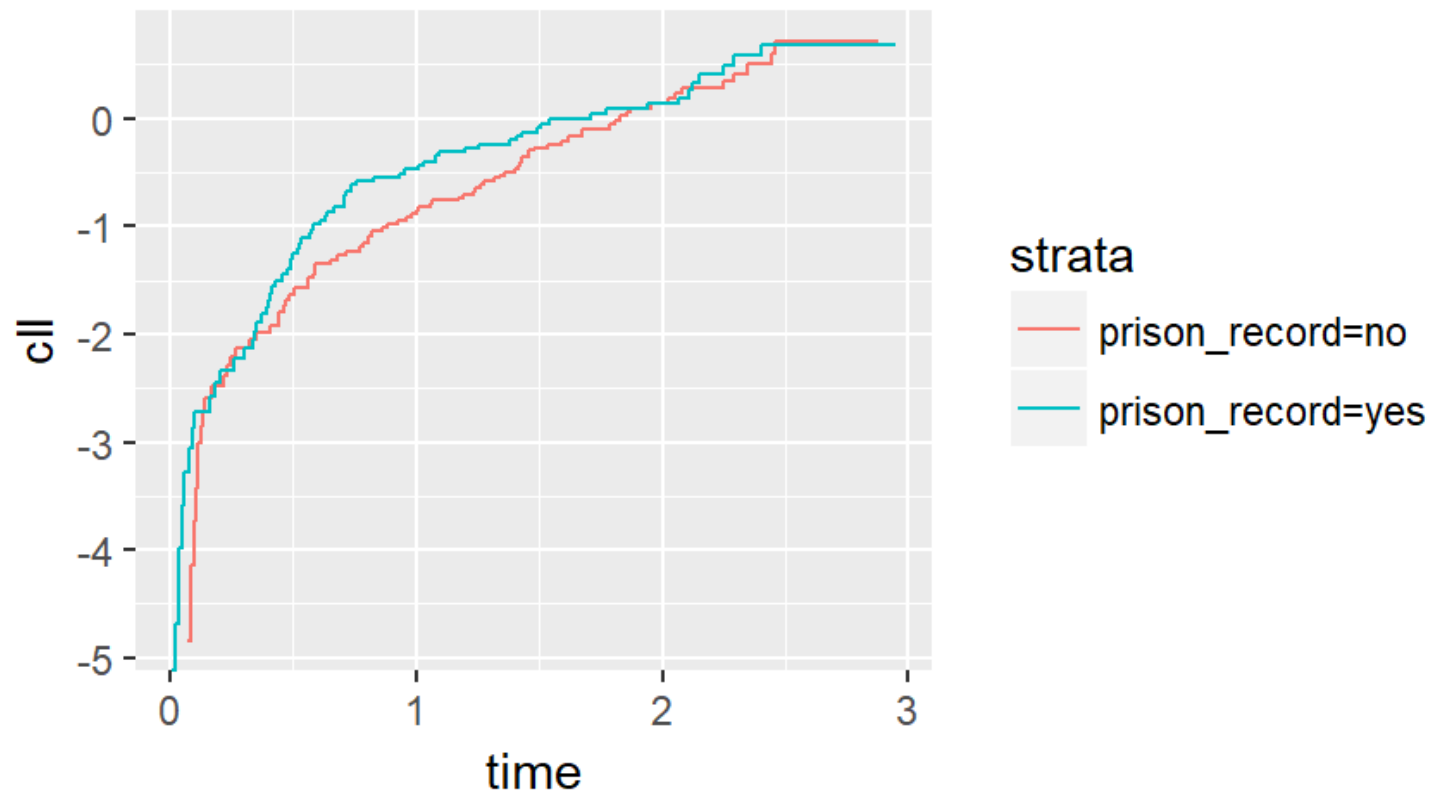
$$\log(-\log(S_2(t))) = \log(-\log(S_1(t))) + \log(hr)$$

If this transformation, the complementary log-log transformation, produces two curves that are separated by a constant for all values of  $t$ , then you have evidence to support the proportional hazards assumption.

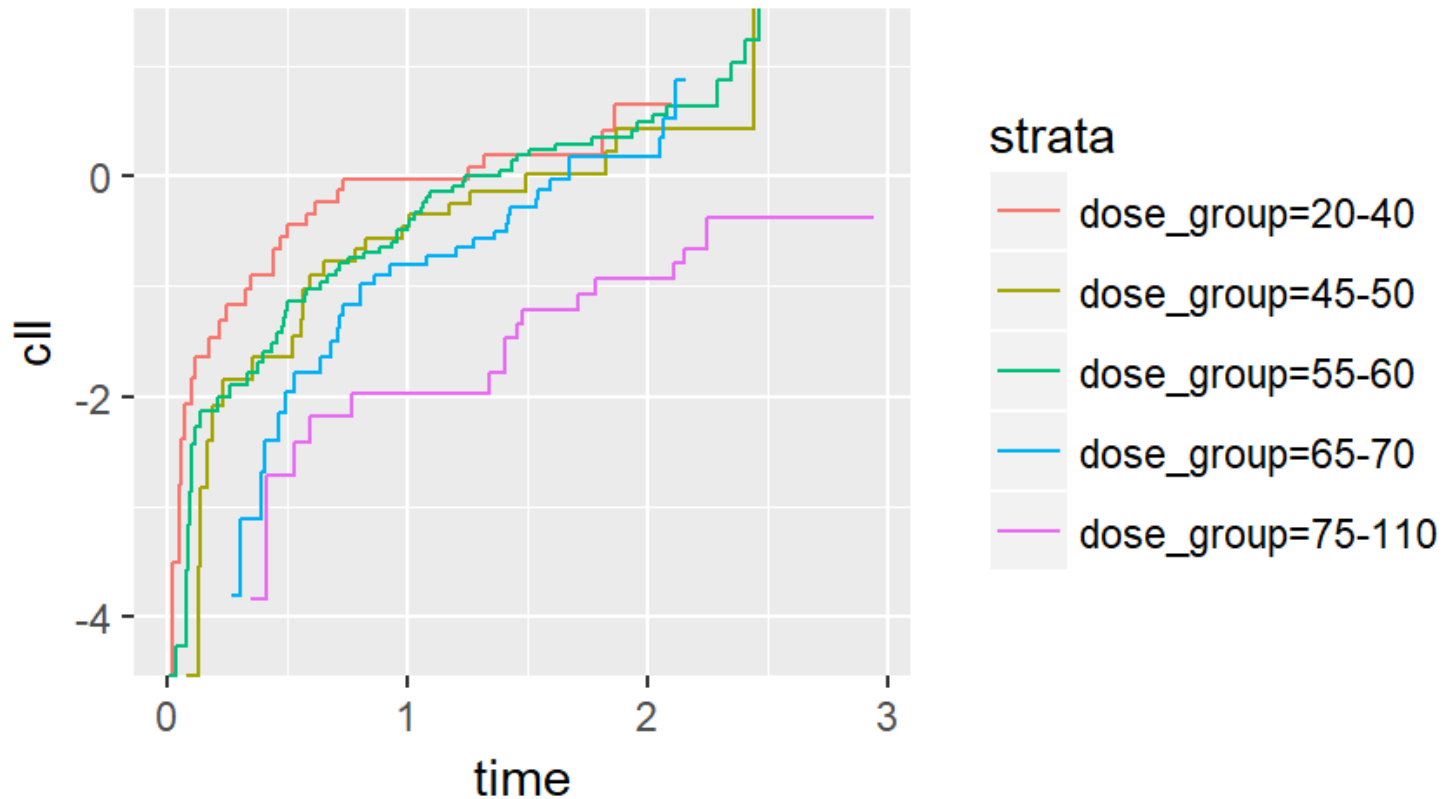
# Complementary Log-Log



# Complementary Log-Log



# Complementary Log-Log





# Review of Likelihood

The partial likelihood is

$$l_p = \prod_i \frac{e^{X_{(i)}\beta}}{\sum_{j \in R_{(i)}} e^{X_j\beta}}$$

The log partial likelihood is

$$L_p = \sum_i (X_{(i)}\beta - \log(\sum_{j \in R_{(i)}} e^{X_j\beta}))$$





## Review of Likelihood

The derivative of the log partial likelihood is

$$\frac{\partial L_p}{\partial \beta} = \sum_{i=1}^m (X_{(i)} - \bar{X}_i(\beta))$$

where  $\bar{X}_i(\beta)$  is a weighted average of all the X's remaining in the risk set and with weights equal to

$$w_{ij} = \frac{e^{X_j \beta}}{\sum_{l \in R_i} e^{X_l \beta}}.$$

A positive derivative implies that we could maximize the log partial likelihood by increasing from the current value of  $\beta$  and a negative derivative implies the opposite.



## Schoenfeld Residuals

The slope of the likelihood function is zero at its maximum, which implies that

$$\sum_{i=1}^m (X_{(i)} - \bar{X}_i(\hat{\beta})) = 0.$$

The individual terms in this sum,

$$X_{(i)} - \hat{X}_{w(i)},$$

are called the Schoenfeld residuals.

# Schoenfeld Residuals



## Interpretation

Normally, the Schoenfeld residuals are standardized.

A time trend is evidence of a violation of the proportional hazards assumption.

Consider this as evidence of an interaction between time and your independent variable.



# Schoenfeld Residuals

## Interpretation

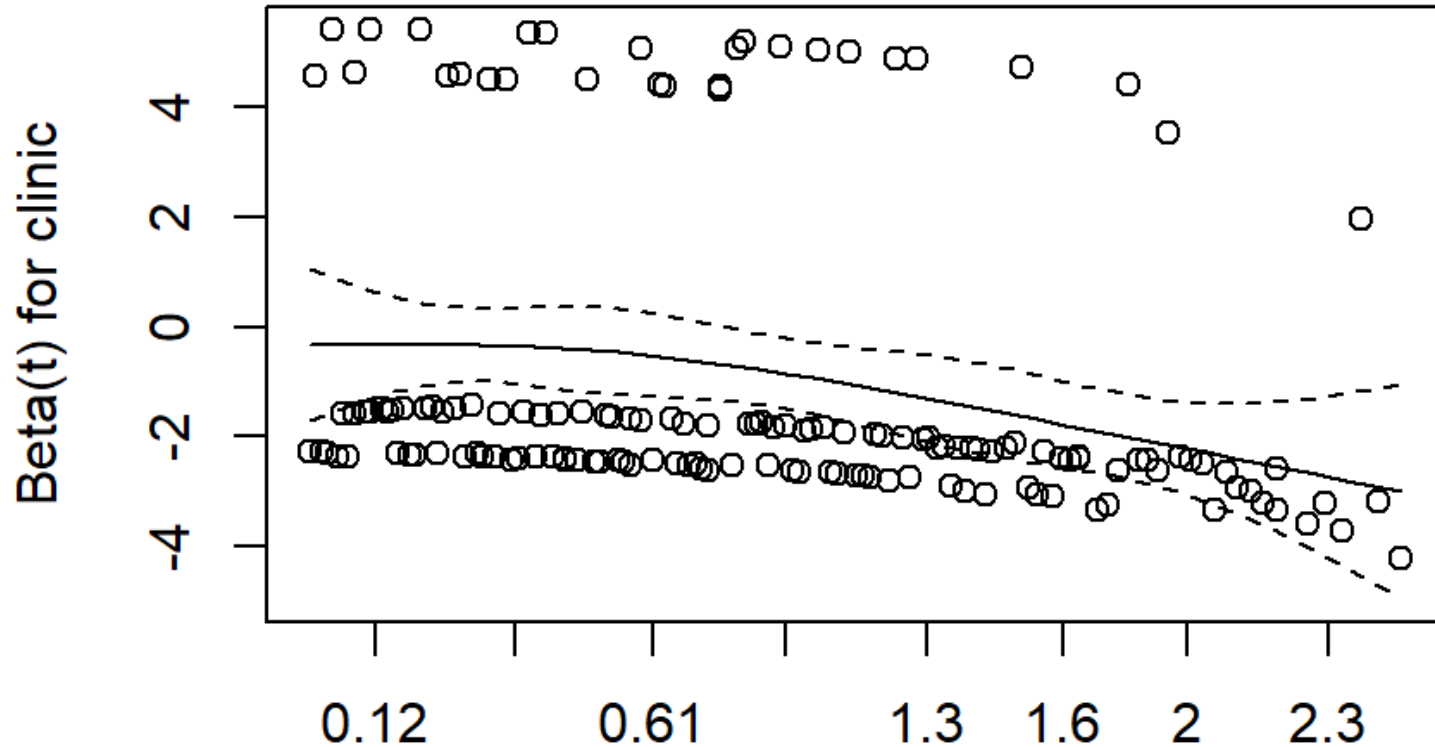
Regions of time where residuals are negative

- the hazard ratio is a bit too large.

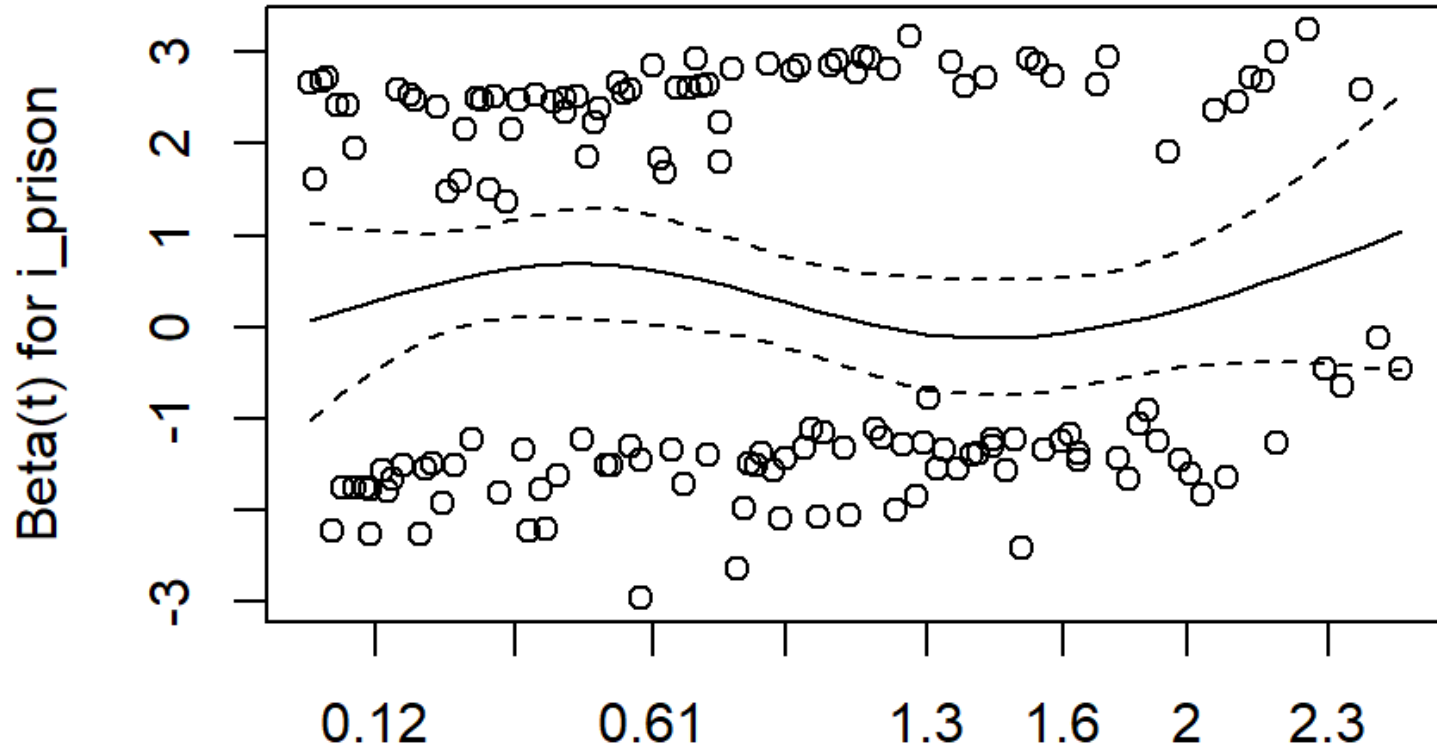
Regions of time where residuals are positive

- the hazard ratio is a bit too small.

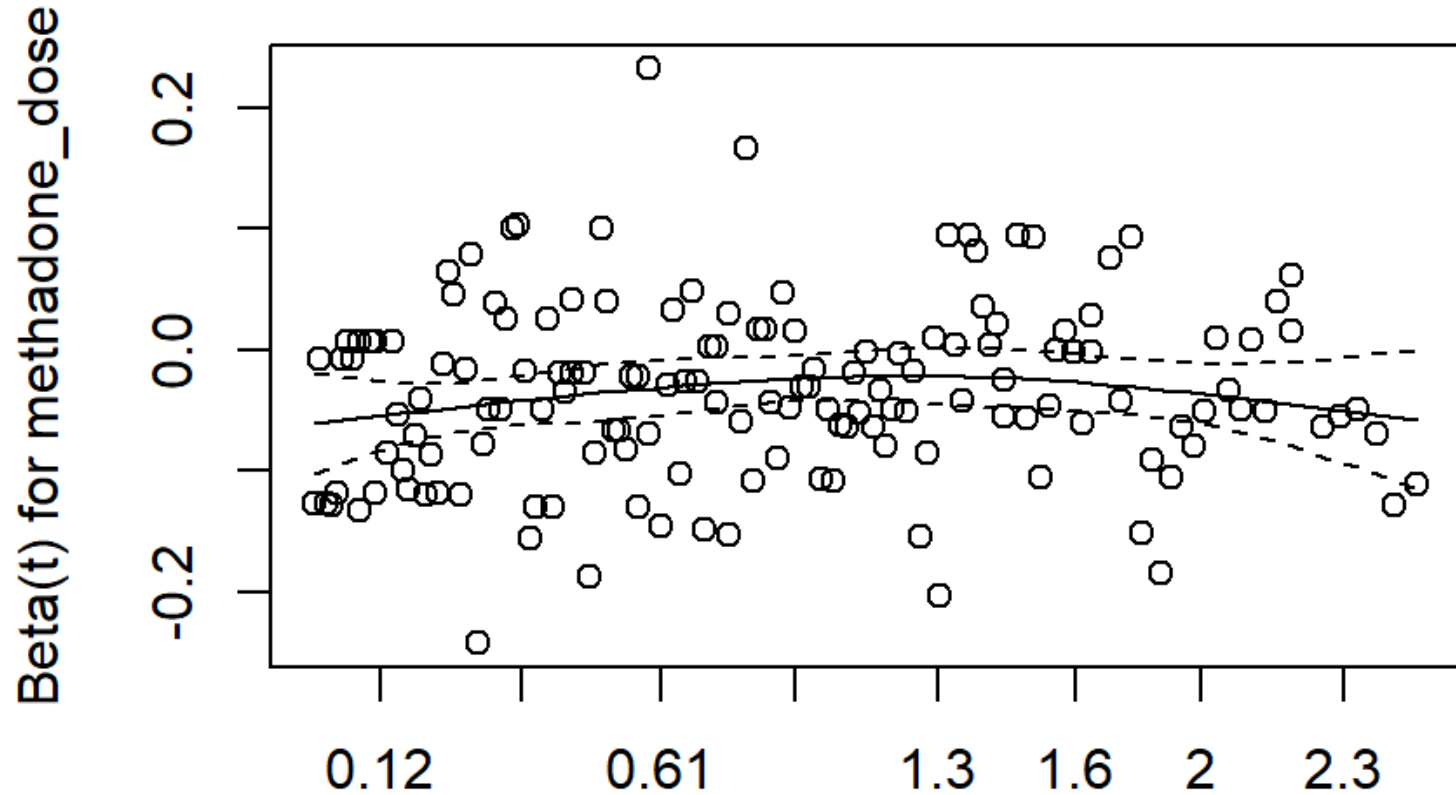
# Schoenfeld Residuals



# Schoenfeld Residuals



## Schoenfeld Residuals



# Stratified Models



The easiest solution for time-varying covariates

Stratification creates a separate baseline hazard for each level of your strata.

Only helpful when the time-varying covariate is a nuisance parameter.



## Stratified Models

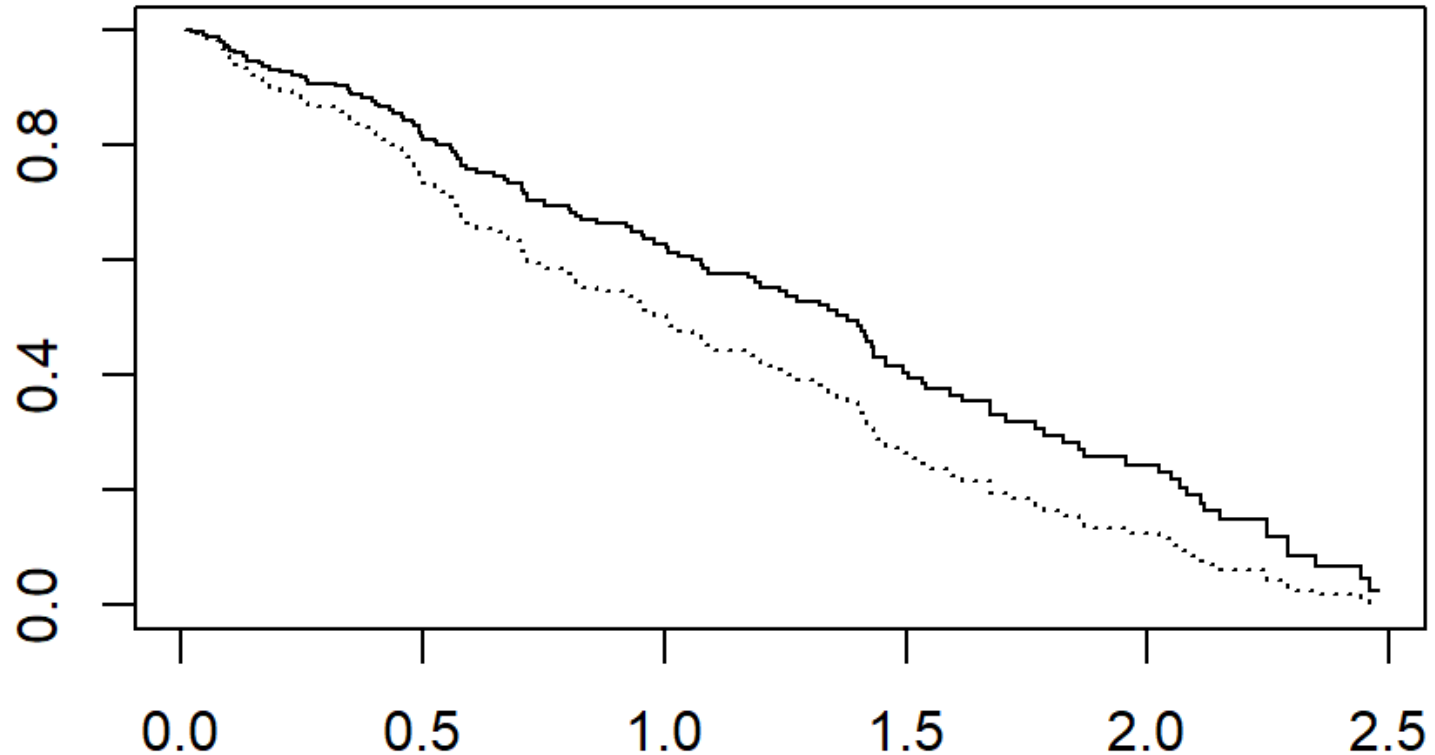


	coef	exp(coef)	se(coef)	z	p
i_prison	0.38960	1.47640	0.16893	2.31	0.021
methadone_dose	-0.03511	0.96549	0.00646	-5.43	5.6e-08

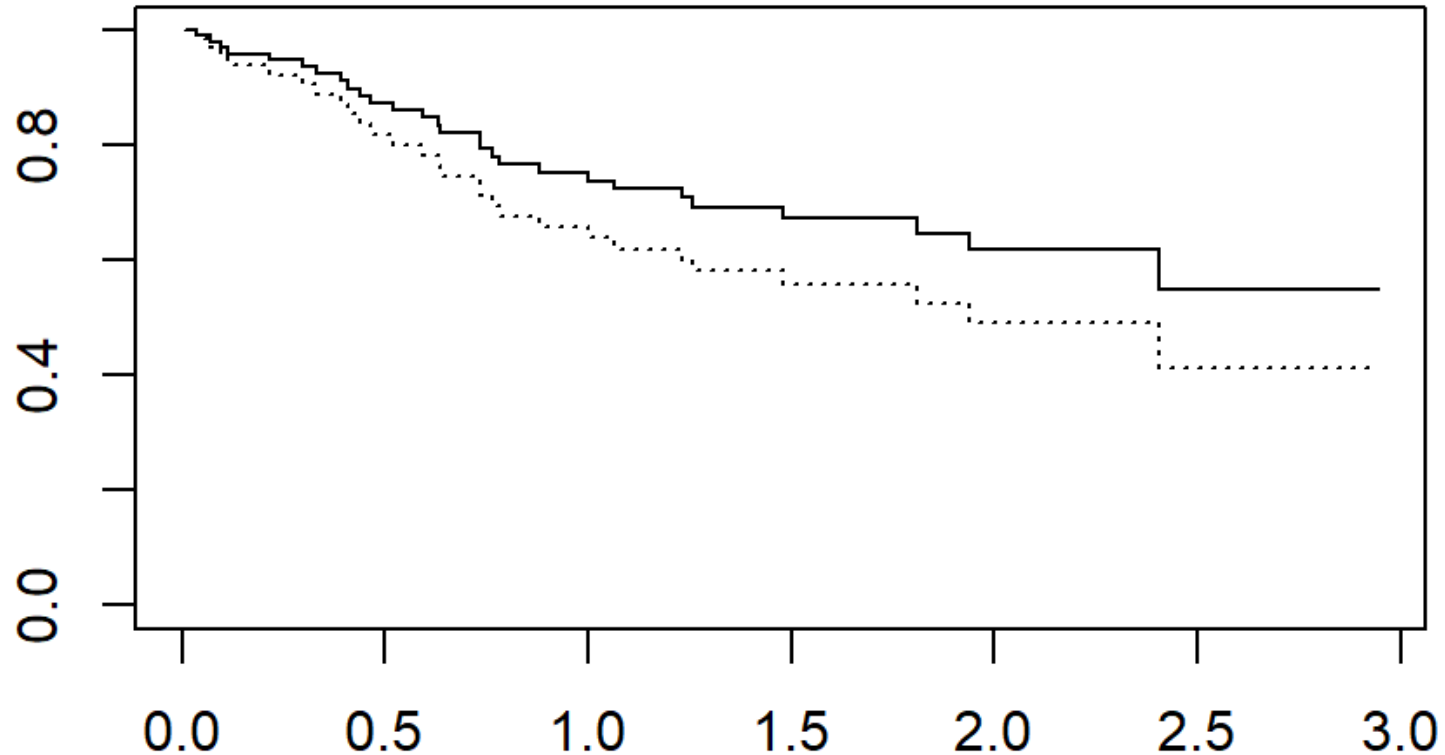
Likelihood ratio test=33.9 on 2 df, p=4.32e-08

n= 238, number of events= 150

## Clinic 1 Survivals: Solid Line is prison\_record=no



## Clinic 2 Survivals: Solid Line is prison\_record=no





# Time-Varying Covariates

Interaction with time

You can address problems with non-proportional hazards by creating an interaction involving time.

- Also consider interactions involving  $\log(\text{time})$ .

# Time-Varying Covariates



Interaction with time, heroin dataset

	coef	exp(coef)	se(coef)	z	p
clinic	0.01940	1.01958	0.34717	0.06	0.9554
tt(clinic)	-1.10331	0.33177	0.34528	-3.20	0.0014
i_prison	0.38997	1.47693	0.16889	2.31	0.0209
methadone_dose	-0.03519	0.96543	0.00644	-5.46	4.7e-08

Likelihood ratio test=76.1 on 4 df, p=1.11e-15

n= 238, number of events= 150

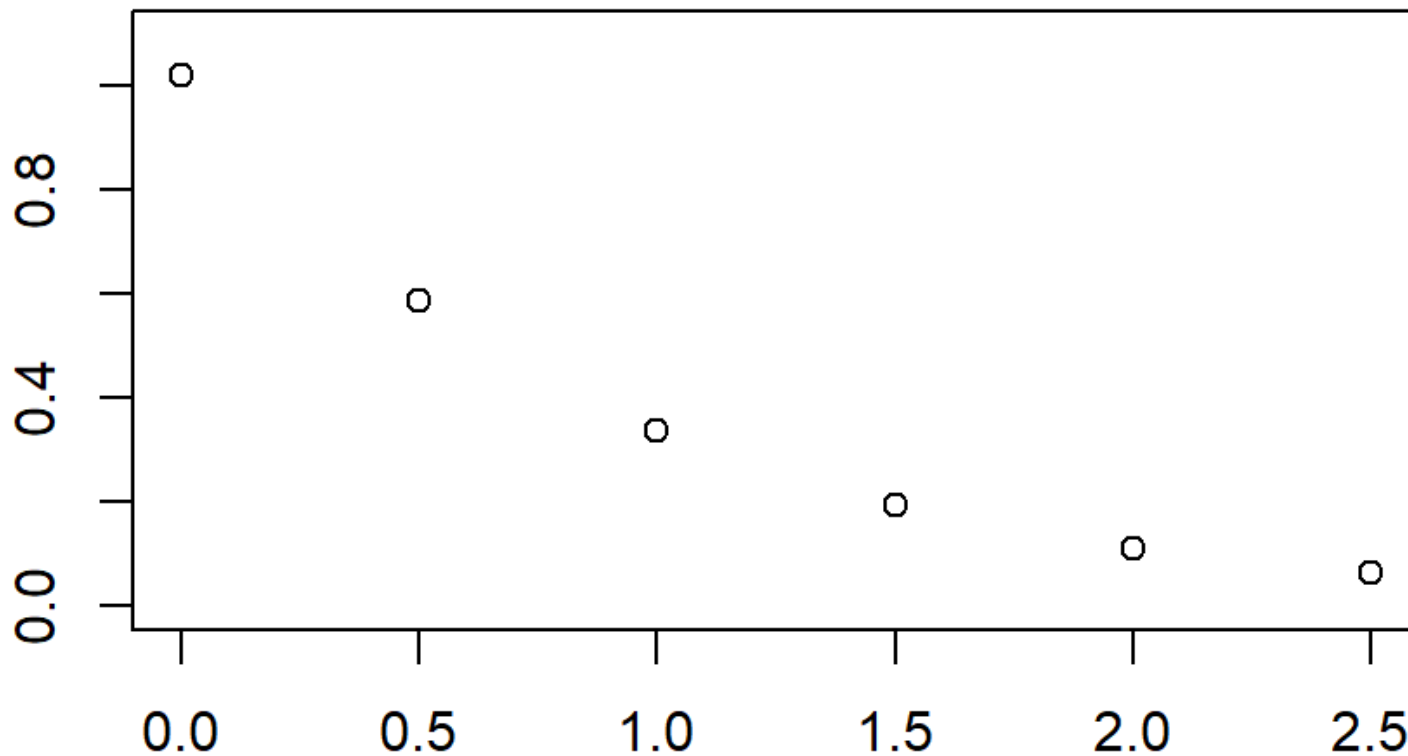
# Time-Varying Covariates



Interaction with time, heroin dataset

	beta_clinic	beta_interaction	t	log_hr	hr
1	0.019	-1.103	0.0	0.019	1.020
2	0.019	-1.103	0.5	-0.532	0.587
3	0.019	-1.103	1.0	-1.084	0.338
4	0.019	-1.103	1.5	-1.636	0.195
5	0.019	-1.103	2.0	-2.187	0.112
6	0.019	-1.103	2.5	-2.739	0.065

# Time-Varying Covariates



# Stanford Transplant Data



[Data](#) pp 45-50 | [Cite as](#)

## Stanford Heart Transplant Data

Authors

[Authors and affiliations](#)

D. F. Andrews, A. M. Herzberg

Chapter

1.6k

Downloads

Part of the [Springer Series in Statistics](#) book series (SSS)

### Abstract

The Stanford Heart Transplantation Program began in October 1967. Patients are admitted to the program after review by a committee, and then they wait for donor hearts to become available. While waiting, some may die or be transferred out of the program, but most receive a transplant. The cut-off date for the data presented in Table 7.1 was in February 1980, and by that time 184 patients had received a transplant.



# Stanford Transplant Data



jaso: original data

birth.dt: birth date

accept.dt: acceptance into program

tx.date: transplant date

fu.date: end of followup

fustat: dead or alive

surgery: prior bypass surgery

age: age (in years)

futime: followup time

wait.time: time before transplant

transplant: transplant indicator

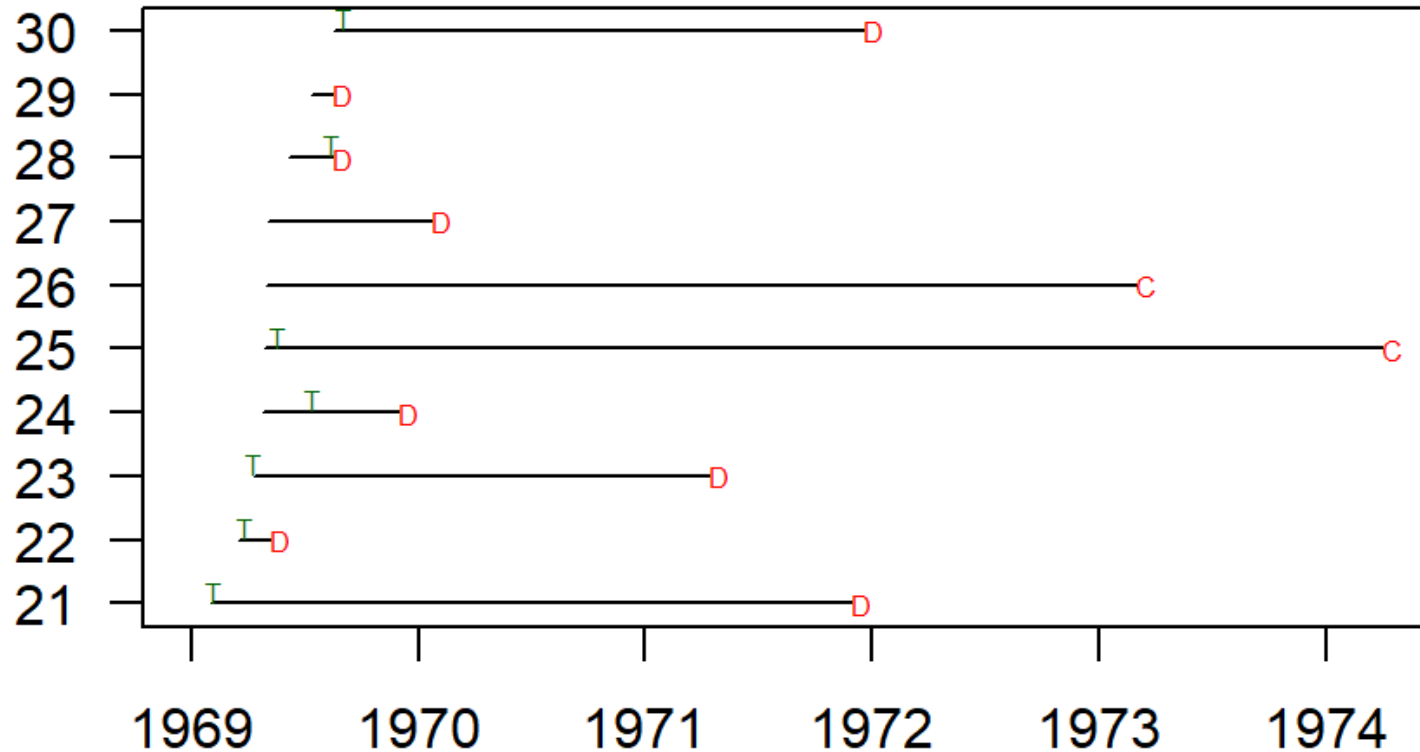
mismatch: mismatch score

hla.a2: particular type of mismatch

mscore: another mismatch score

reject: rejection occurred

## Stanford Transplant Data, Patients 21-30





## Stanford Transplant Data, Naive Analysis

	coef	exp(coef)	se(coef)	z	p
transplant	-1.7171	0.1796	0.2785	-6.16	7.1e-10
age	0.0589	1.0607	0.0150	3.91	9.1e-05
surgery	-0.4190	0.6577	0.3712	-1.13	0.26

Likelihood ratio test=45.9 on 3 df, p=6.11e-10

n= 103, number of events= 75

# Stanford Transplant Data



## Patient #21

id	accept.dt	tx.date	fu.date	fustat	transplant
21	1969-02-01	1969-02-08	1971-11-29	1	1

id	time1	time2	event	transplant
21	7 days	1031 days	D	1

id	start	stop	event	transplant
21	0	7	0	0
21	7	1031	1	1

# Stanford Transplant Data



## Patient #22

id	accept.dt	tx.date	fu.date	fustat	transplant
22	1969-03-18	1969-03-29	1969-05-07	1	1

id	time1	time2	event	transplant
22	11 days	50 days	D	1

id	start	stop	event	transplant
22	0	11	0	0
22	11	50	1	1

# Stanford Transplant Data



## Patient #23

id	accept.dt	tx.date	fu.date	fustat	transplant
23	1969-04-11	1969-04-13	1971-04-13	1	1

id	time1	time2	event	transplant
23	2 days	732 days	D	1

id	start	stop	event	transplant
23	0	2	0	0
23	2	732	1	1

# Stanford Transplant Data



## Patient #24

id	accept.dt	tx.date	fu.date	fustat	transplant
24	1969-04-25	1969-07-16	1969-11-29	1	1

id	time1	time2	event	transplant
24	82 days	218 days	D	1

id	start	stop	event	transplant
24	0	82	0	0
24	82	218	1	1

# Stanford Transplant Data



## Patient #25

id	accept.dt	tx.date	fu.date	fustat	transplant
25	1969-04-28	1969-05-22	1974-04-01	0	1

id	time1	time2	event	transplant
25	24 days	1799 days	C	1

id	start	stop	event	transplant
25	0	24	0	0
25	24	1799	0	1



# Stanford Transplant Data



## Patient #26

id	accept.dt	tx.date	fu.date	fustat	transplant
26	1969-05-01	<NA>	1973-03-01	0	0

id	time1	time2	event	transplant
26	NA days	1400 days	C	0

id	start	stop	event	transplant
26	0	1400	0	0

# Stanford Transplant Data



## Patient #27

id	accept.dt	tx.date	fu.date	fustat	transplant
27	1969-05-04	<NA>	1970-01-21	1	0

id	time1	time2	event	transplant
27	NA days	262 days	D	0

id	start	stop	event	transplant
27	0	262	1	0

# Stanford Transplant Data



## Patient #28

id	accept.dt	tx.date	fu.date	fustat	transplant
28	1969-06-07	1969-08-16	1969-08-17	1	1

id	time1	time2	event	transplant
28	70 days	71 days	D	1

id	start	stop	event	transplant
28	0	70	0	0
28	70	71	1	1

# Stanford Transplant Data



## Patient #29

id	accept.dt	tx.date	fu.date	fustat	transplant
29	1969-07-14	<NA>	1969-08-17	1	0

id	time1	time2	event	transplant
29	NA days	34 days	D	0

id	start	stop	event	transplant
29	0	34	1	0

# Stanford Transplant Data



## Patient #30

id	accept.dt	tx.date	fu.date	fustat	transplant
30	1969-08-19	1969-09-03	1971-12-18	1	1

id	time1	time2	event	transplant
30	15 days	851 days	D	1

id	start	stop	event	transplant
30	0	15	0	0
30	15	851	1	1



## Stanford Transplant Data, Time-varying Model

	coef	exp(coef)	se(coef)	z	p
transplant	0.0141	1.0142	0.3082	0.05	0.964
age	0.0306	1.0310	0.0139	2.20	0.028
surgery	-0.7733	0.4615	0.3597	-2.15	0.032

Likelihood ratio test=10.7 on 3 df, p=0.0134  
n= 170, number of events= 75

# Conclusion



What have you learned today?

1. There are several ways to check for/control for time varying covariates: patterns in Kaplan-Meier curves, the complementary log-log plot, Schoenfeld Residuals, stratified analysis, and fitting time varying covariates.
2. You can fit time-varying covariates using start/stop coding or a time transfer function.