

AI 5090: STOCHASTIC PROCESSES

HOMEWORK 2



CONVERGENCE OF SEQUENCES OF RANDOM VARIABLES

Fix a probability space $(\Omega, \mathcal{F}, \mathbb{P})$.

Assume that all random variables appearing below are defined with respect to this probability space.

1. Let $\Theta \sim \text{Unif}[0, 2\pi]$. For each $n \in \mathbb{N}$, let

$$Y_n = \left| 1 - \frac{\Theta}{\pi} \right|^n.$$

Argue in which of the four senses (a.s., m.s., p., d.) does the sequence $\{Y_n\}_{n=1}^\infty$ converge. Identify the limit random variable in each case, and justify your answers.

2. Let $X_1, X_2, \dots \stackrel{\text{i.i.d.}}{\sim} \text{Ber}(0.5)$.

- (a) Prove that $\{X_n\}_{n=1}^\infty$ converges in distribution. Identify a limit random variable.
(b) Prove that $\{X_n\}_{n=1}^\infty$ cannot converge in probability to any random variable.

Hint for part (b):

Suppose there exists a limit random variable X such that $X_n \xrightarrow{p} X$. Then, from triangle inequality, we have

$$|X_n - X_{n+1}| \leq |X_n - X| + |X_{n+1} - X|.$$

Use the above inequality to prove that for any $\varepsilon > 0$,

$$\mathbb{P}(|X_n - X_{n+1}| > \varepsilon) \leq 2\mathbb{P}\left(|X_n - X| > \frac{\varepsilon}{2}\right).$$

In particular, compute the left-hand side of the above inequality for $\varepsilon = 0.5$, and prove that convergence in probability does not hold.

3. Let $X_1, X_2, \dots \stackrel{\text{i.i.d.}}{\sim} \text{Poisson}(1)$.

For each $n \in \mathbb{N}$, let $S_n = \sum_{i=1}^n X_i$.

- (a) Show that $S_n \sim \text{Poisson}(n)$ for each $n \in \mathbb{N}$.
(b) Without using the central limit theorem, prove from first principles that

$$\frac{S_n - n}{\sqrt{n}} \xrightarrow{d} X, \quad X \sim \mathcal{N}(0, 1).$$

- (c) Prove that

$$\lim_{n \rightarrow \infty} e^{-n} \sum_{k=0}^{\infty} \frac{n^k}{k!} = \frac{1}{2}.$$

Hint for part (c): Use the result in part (b).

4. Suppose that $X_1, X_2, \dots \stackrel{\text{i.i.d.}}{\sim} \text{Ber}(p)$ for some fixed constant $p \in (0, 1)$.

For each $n \in \mathbb{N}$, let $S_n = \sum_{i=1}^n X_i$.

- (a) Derive the moment generating function of S_n .
(b) Show that

$$\mathbb{P}\left(\left|\frac{S_n}{n} - p\right| > \varepsilon\right) = 0$$

for all $\varepsilon > 1$.

(c) Prove that for any $\varepsilon \in (0, 1)$,

$$\mathbb{P}\left(\frac{S_n}{n} - p > \varepsilon\right) \begin{cases} = 0, & p + \varepsilon \geq 1, \\ \leq \exp(-n D((p + \varepsilon) \parallel p)), & 0 < p + \varepsilon < 1, \end{cases}$$

where $D((p + \varepsilon) \parallel p) := (p + \varepsilon) \log \frac{p + \varepsilon}{p} + (1 - p - \varepsilon) \log \frac{1 - p - \varepsilon}{1 - p}$ is the Kullback–Leibler divergence between the probability distributions $\text{Ber}(p + \varepsilon)$ and $\text{Ber}(p)$.

Hint for part (b):

Use Chernoff's bound to obtain an upper bound for $\mathbb{P}(S_n > n(p + \varepsilon))$, and optimise the upper bound.

(d) Prove that for any $\varepsilon > 0$,

$$\mathbb{P}\left(\frac{S_n}{n} - p < -\varepsilon\right) \begin{cases} = 0, & p - \varepsilon \leq 0, \\ \leq \exp(-n D((1 - p + \varepsilon) \parallel (1 - p))), & 0 < p - \varepsilon < 1. \end{cases}$$

Hint for part (d):

Set $Y_n = 1 - X_n$ for each $n \in \mathbb{N}$, and use the result in part (c) on the sum of Y_n 's.

(e) Combining the results in parts (b), (c), and (d) above, prove that

$$\frac{S_n}{n} \xrightarrow{\text{a.s.}} p.$$

5. Suppose that $\{X_n\}_{n=1}^\infty$ is a sequence of random variables satisfying the property

$$\sum_{n=1}^\infty \mathbb{E}[|X_n - 2|^p] < +\infty$$

for some fixed $p > 1$.

Prove that $\{X_n\}_{n=1}^\infty$ converges almost surely. What is the limit random variable?

6. Let $X_1, X_2, \dots \stackrel{\text{i.i.d.}}{\sim} \text{Exp}(1)$.

For each $n \in \mathbb{N}$, let $Y_n = \max\{X_1, \dots, X_n\}$.

In this exercise, we shall prove formally that almost surely, Y_n grows as $\log n$ for large n .

If the base of the logarithm is not mentioned explicitly, it should be considered to be e .

(a) Show formally that

$$\frac{Y_n}{\log n} \xrightarrow{\text{d.}} 1,$$

and use the reverse implication from class to conclude that $\frac{Y_n}{\log n} \xrightarrow{\text{p.}} 1$.

(b) Based on the conclusion in part (a), show that

$$\frac{Y_n}{\log_2 n} \xrightarrow{\text{p.}} \log 2.$$

(c) Consider the subsequence $\{n_k\}_{k=1}^\infty \subset \mathbb{N}$ given by $n_k = 2^k$ for all $k \in \mathbb{N}$.

Fix $\varepsilon > 0$. For each k , let

$$x_k := e^{-(\varepsilon + \log 2)k}.$$

Prove that

$$(1 - x_k)^{n_k} \geq \exp\left(-\frac{n_k x_k}{1 - x_k}\right) \quad \forall k.$$

Further, deduce that

$$(1 - x_k)^{n_k} \geq \exp(-2n_k x_k)$$

for all sufficiently large values of k .

Hint for part (c):

To prove the first part, use the relation $\log x \geq 1 - \frac{1}{x}$ for any $x > 0$ (this is another way of seeing the well-known inequality $\log x \leq x - 1$ for all $x > 0$).

To deduce the second part, use the fact that x_k converges to 0 as $k \rightarrow \infty$, and therefore $x_k < \frac{1}{2}$ for all sufficiently large values of k .

(d) Using the result in the second half of part (c), prove that for every $\varepsilon > 0$,

$$\sum_{k=1}^{\infty} \mathbb{P}\left(\frac{Y_{n_k}}{\log_2 n_k} > \log 2 + \varepsilon\right) < +\infty.$$

Then, using the Borel–Cantelli lemma, conclude that

$$\mathbb{P}\left(\left\{\omega \in \Omega : \limsup_{k \rightarrow \infty} \frac{Y_{n_k}(\omega)}{\log_2 n_k} \leq \log 2\right\}\right) = 1. \quad (1)$$

(e) Fix $0 < \varepsilon < \log 2$. For each $k \in \mathbb{N}$, let

$$y_k := e^{-(\varepsilon - \log 2)k}.$$

Using the fact that $e^x > x$ for any $x > 0$ (again, another way to see the inequality $\log x \leq x - 1 < x$), prove that

$$(1 - y_k)^{n_k} \leq e^{-\varepsilon k} \quad \forall k.$$

Use this relation to prove that for every $\varepsilon > 0$,

$$\sum_{k=1}^{\infty} \mathbb{P}\left(\frac{Y_{n_k}}{\log_2 n_k} \leq \log 2 - \varepsilon\right) < +\infty,$$

and hence conclude from the Borel–Cantelli lemma that

$$\mathbb{P}\left(\left\{\omega \in \Omega : \liminf_{k \rightarrow \infty} \frac{Y_{n_k}(\omega)}{\log_2 n_k} \geq \log 2\right\}\right) = 1. \quad (2)$$

Epilogue for question 6:

Combining the results in (1) and (2), we see that

$$\mathbb{P}\left(\left\{\omega \in \Omega : \lim_{k \rightarrow \infty} \frac{Y_{n_k}(\omega)}{\log_2 n_k} = \log 2\right\}\right) = 1.$$

That is, $\frac{Y_{n_k}}{\log_2 n_k} \xrightarrow{\text{a.s.}} \log 2$.

This proves that the subsequence $\{Y_{n_k}/(\log_2 n_k)\}_{k=1}^{\infty}$ converges almost surely to the constant random variable taking the value $\log 2$.

We can use this to prove that the entire sequence $\{Y_n/(\log_2 n)\}_{n=1}^{\infty}$ must also converge to the same constant random variable $\log 2$, as follows.

Given any $n \in \mathbb{N}$, find k such that $n_k \leq n < n_{k+1}$ (you can always find at least one such k).

Because $Y_1 \leq Y_2 \leq Y_3 \leq \dots$, it follows that

$$\begin{aligned} Y_{n_k} &\leq Y_n < Y_{n_{k+1}} \\ \Rightarrow \frac{Y_{n_k}}{\log_2 n} &\leq \frac{Y_n}{\log_2 n} < \frac{Y_{n_{k+1}}}{\log_2 n} \\ \Rightarrow \frac{Y_{n_k}}{\log_2(n_k + 1)} &\leq \frac{Y_n}{\log_2 n} < \frac{Y_{n_{k+1}}}{\log_2 n_k} \\ \Rightarrow \frac{\log_2(n_k + 1)}{\log_2 n_k} \cdot \frac{Y_{n_k}}{\log_2 n_k} &\leq \frac{Y_n}{\log_2 n} < \frac{Y_{n_{k+1}}}{\log_2(n_k + 1)} \cdot \frac{\log_2 n_k}{\log_2(n_k + 1)}. \end{aligned}$$

Using (1) and (2), along with the fact that $\lim_{k \rightarrow \infty} \frac{\log_2(n_k + 1)}{\log_2 n_k} = 1$ gives us the desired result.