

1. Because $p, q \in (0,1)$, $P_{i,j} > 0 \quad \forall i, j \in \{0,1\}$, and hence P is irreducible.

Because an irreducible DTMC on a finite state space is positive recurrent, and

irreducibility + positive recurrence \Rightarrow unique stationary distⁿ exists,

we conclude that there exists a unique $\pi = [\pi_0 \ \pi_1]$ s.t.

$$\pi_0 \geq 0, \ \pi_1 \geq 0, \ \pi_0 + \pi_1 = 1.$$

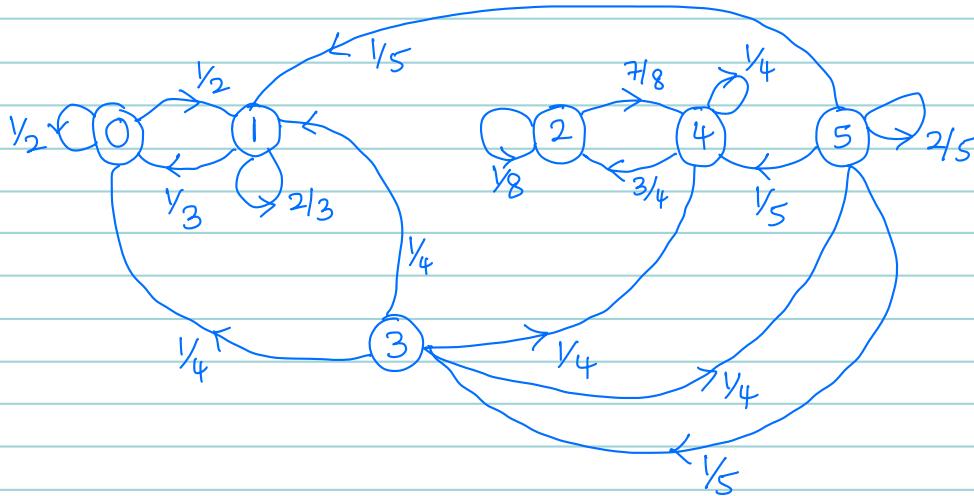
Solving $\pi = \pi P$, we get

$$\pi_0 = (1-p)\pi_0 + q\pi_1 \quad \Rightarrow \quad \pi_1 = \frac{p}{q} \cdot \pi_0.$$

Then, because $\pi_0 + \pi_1 = 1$, we get

$$\pi_0 = \frac{q}{p+q}, \quad \pi_1 = \frac{p}{p+q}.$$

2. The transition diagram is as follows.



a) It is evident from the transition graph that

$$\mathcal{C}_1 = \{0, 1\}, \quad \mathcal{C}_2 = \{2, 4\}, \quad \mathcal{C}_3 = \{3, 5\}$$

are the communicating classes. Furthermore,

$\mathcal{C}_1 \rightarrow$ closed \Rightarrow positive recurrent (because no. of states in \mathcal{C}_1 is finite),

$\mathcal{C}_2 \rightarrow$ closed \Rightarrow positive recurrent,

$\mathcal{C}_3 \rightarrow$ open \Rightarrow transient.

b) If $x \in \mathcal{C}_1$, then $\mathbb{P}(T_{\{0,1\}} < +\infty | X_0 = x) = 1$.

If $x \in \mathcal{C}_2$, then $\mathbb{P}(T_{\{0,1\}} < +\infty | X_0 = x) = 0$.

If $x \in \mathcal{C}_3$, then:

$$\begin{aligned} \mathbb{P}(T_{\{0,1\}} < +\infty | X_0 = 3) &= \mathbb{P}(T_{\{0,1\}} < +\infty, X_1 = 0 | X_0 = 3) \\ &\quad + \mathbb{P}(T_{\{0,1\}} < +\infty, X_1 = 1 | X_0 = 3) \\ &\quad + \mathbb{P}(T_{\{0,1\}} < +\infty, X_1 = 4 | X_0 = 3) \\ &\quad + \mathbb{P}(T_{\{0,1\}} < +\infty, X_1 = 5 | X_0 = 3) \\ &= \gamma_4 \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_1 = 0, X_0 = 3) \\ &\quad + \gamma_4 \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_1 = 1, X_0 = 3) \\ &\quad + \gamma_4 \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_1 = 4, X_0 = 3) \\ &\quad + \gamma_4 \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_1 = 5, X_0 = 3) \end{aligned}$$

Can neglect
 $X_0 = 3$ because
of Markov
property.

$$\begin{aligned} &= \gamma_4 + \gamma_4 + \gamma_4 \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_1 = 5) \\ &= \gamma_2 + \gamma_4 \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_0 = 5) \quad (\text{again, Markov property}) \end{aligned}$$

Similarly, it can be shown that

$$\begin{aligned} \mathbb{P}(T_{\{0,1\}} < +\infty | X_0 = 5) &= \frac{2}{5} \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_1 = 5) \\ &\quad + \gamma_5 \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_1 = 3) \\ &\quad + \gamma_5 \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_1 = 1) \\ &= \gamma_5 + \gamma_5 \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_0 = 3) \\ &\quad + \frac{2}{5} \cdot \mathbb{P}(T_{\{0,1\}} < +\infty | X_0 = 5) \end{aligned}$$

Setting $a = \mathbb{P}(T_{\{0,1\}} < +\infty | X_0 = 3)$

$b = \mathbb{P}(T_{\{0,1\}} < +\infty | X_0 = 5)$, we have

$$a = \frac{y_2}{2} + b/4 \Leftrightarrow 4a = 2 + b$$

$$\frac{3}{5}b = \frac{1}{5} + a/5 \Leftrightarrow 3b = 1 + a$$

Then,

$$3(4a - 2) = 1 + a \Rightarrow 11a = 7 \Rightarrow a = 7/11,$$

$$b = 6/11.$$

3. Let A_n denote the box chosen at time $n \in \mathbb{N}$.

We note that $A_1, A_2, \dots \stackrel{iid}{\sim} \text{Unif}\{1, \dots, N\}$, and $A_n \perp\!\!\!\perp X_{n-1} \forall n$.

Furthermore, we have

$$X_n = X_{n-1} \mathbf{1}_{\{A_n \in \{A_1, \dots, A_{n-1}\}\}} + (x_{n-1} - 1) \mathbf{1}_{\{A_n \notin \{A_1, \dots, A_{n-1}\}\}}.$$

Because $A_n \perp\!\!\!\perp X_{n-1}$ and X_n is a function only of X_{n-1} ,

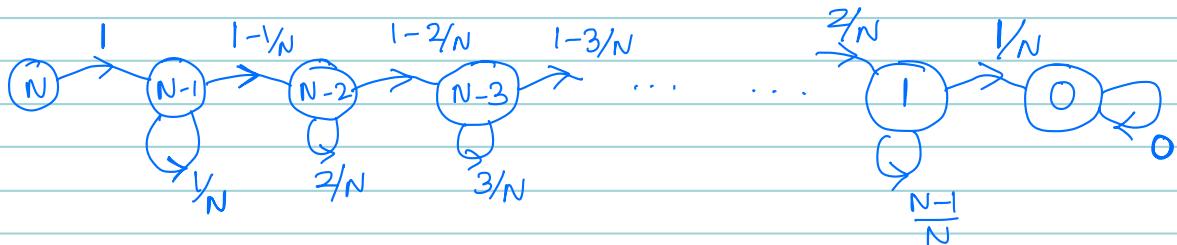
We conclude that $\{X_n\}_{n=0}^{\infty}$ is a DTMC.

- b) The state space $\mathcal{X} = \{N, N-1, \dots, 0\}$.

We note that

$$P(X_n = j \mid X_{n-1} = i) = \begin{cases} \frac{N-i}{N}, & \text{if } j = i, \\ i/N, & \text{if } j = i-1, \\ 0, & \text{otherwise.} \end{cases}$$

The transition graph is as follows.



- c) Noting that 0 is an absorbing state,
we conclude that all states except 0 are transient,
and 0 is positive recurrent (because $\{0\}$ is a closed communicating class).
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4. Let $T_0 = 0$, and for each $k \in \mathbb{N}$, let

$$T_k := \inf \{n > T_{k-1} : X_n = y\}$$

denote the k^{th} return time to y . Note that

$$\mathbb{P}(T_1 < +\infty \mid X_0 = y) = f_{yy} = 1 \quad (\text{because } y \text{ is tve recurrent}),$$

and for any $k \geq 2$, assuming $\mathbb{P}(T_{k-1} < +\infty \mid X_0 = y) = 1$, we have

$$\begin{aligned} \mathbb{P}(T_k < +\infty \mid X_0 = y) &= \mathbb{P}(T_k < +\infty \mid X_{T_{k-1}} = y, T_{k-1} < +\infty, X_0 = y) \\ &= \mathbb{P}(T_k < +\infty \mid X_{T_{k-1}} = y, T_{k-1} < +\infty) \end{aligned}$$

*Strong
Markov
Property*

Hence, $\{T_k\}_{k=0}^{+\infty}$ are a sequence of stopping times, and

$$\mathbb{P}(T_k - T_{k-1} < +\infty \mid X_0 = y) = 1 \quad \forall k \in \mathbb{N}.$$

We then note that

$$\mathbb{P}(V_x^{(k)} = n \mid X_0 = y) = \mathbb{P}(V_x^{(k)} = n, T_{k-1} < +\infty, T_k - T_{k-1} < +\infty \mid X_0 = y)$$

$$\text{law of total probability} \quad \sum_{l=1}^{\infty} \sum_{m=1}^{\infty} \mathbb{P}(V_x^{(k)} = n, T_{k-1} = m, T_k - T_{k-1} = l \mid X_0 = y)$$

$$= \sum_{l=1}^{\infty} \sum_{m=1}^{\infty} \mathbb{P}(V_x^{(k)} = n, T_{k-1} = m, X_m = y, \underbrace{X_{m+1} \neq y, \dots, X_{m+l} \neq y}_{\text{says } T_k - T_{k-1} = l} \mid X_0 = y)$$

$$= \sum_{l=1}^{\infty} \sum_{m=1}^{\infty} \mathbb{P}(T_{k-1} = m, X_m = y \mid X_0 = y)$$

$$\cdot \mathbb{P}(V_x^{(k)} = n, X_{m+1} \neq y, \dots, X_{m+l} \neq y \mid X_0 = y, X_m = y, T_{k-1} = m)$$

$$= \sum_{l=1}^{\infty} \sum_{m=1}^{\infty} \mathbb{P}(T_{k-1} = m, X_m = y \mid X_0 = y)$$

$$\cdot \mathbb{P}(V_x^{(k)} = n, X_{m+1} \neq y, \dots, X_{m+l} \neq y \mid X_m = y)$$

function of block k

$$\text{Markov prop. and} \quad \sum_{l=1}^{\infty} \sum_{m=1}^{\infty} \mathbb{P}(T_{k-1} = m, X_m = y \mid X_0 = y) \cdot \mathbb{P}(V_x^{(1)} = n, X_1 \neq y, \dots, X_l \neq y \mid X_0 = y)$$

$$\text{IID block structure} \quad = \sum_{l=1}^{\infty} \sum_{m=1}^{\infty} \mathbb{P}(T_{k-1} = m \mid X_0 = y) \cdot \mathbb{P}(V_x^{(1)} = n, T_1 = l \mid X_0 = y)$$

$$= \mathbb{P}(V_x^{(1)} = n \mid X_0 = y),$$

thus proving that $\{V_x^{(k)}\}_{k \in \mathbb{N}}$ are identically distributed.

We now prove independence of $V_x^{(k)}$ and $V_x^{(k+1)}$ for all $k \in \mathbb{N}$. We have

$$\mathbb{P}(V_x^{(k)} = m, V_x^{(k+1)} = n \mid X_0 = y)$$

$$= \mathbb{P}(V_x^{(k)} = m, T_{k-1} < +\infty, T_k - T_{k-1} < +\infty, V_x^{(k+1)} = n, T_{k+1} - T_k < +\infty \mid X_0 = y)$$

$$= \sum_{l=1}^{\infty} \sum_{r=1}^{\infty} \sum_{s=1}^{\infty} \mathbb{P}(V_x^{(k)} = m, T_{k-1} = l, T_k - T_{k-1} = r, V_x^{(k+1)} = n, T_{k+1} - T_k = s \mid X_0 = y)$$

Using similar arguments as before (identical distribution part) with IID block structure, it can be shown that

$$\mathbb{P}(V_x^{(k)} = m, V_x^{(k+1)} = n \mid X_0 = y) = \mathbb{P}(V_x^{(1)} = m \mid X_0 = y) \cdot \mathbb{P}(V_x^{(2)} = n \mid X_0 = y)$$

identical distribution of $\{V_x^{(k)}\}_{k \in \mathbb{N}}$

5. We note that every state is accessible from state 0; in particular,

$$\begin{aligned} P_{0,i}^i &= P_{0,1} \cdot P_{1,2} \cdots P_{i-1,i} = \left(1 - \frac{1}{N}\right) \cdots \left(1 - \frac{i-1}{N}\right) \\ &= \frac{N!}{(N-i)! N^i} > 0. \end{aligned}$$

Also, we note that

$$P_{i,0}^i = P_{i,i-1} \times P_{i-1,i-2} \times \cdots \times P_{1,0} = \frac{i!}{N^i} > 0.$$

Thus, the Markov chain is irreducible, and because state space is finite, the Markov chain is positive recurrent.

Let π denote the unique stationary distribution of P . Then,

$$\pi_0 = \frac{1}{N} \pi_1 \Rightarrow \pi_1 = N \pi_0 = N C_1 \cdot \pi_0$$

$$\pi_1 = \left(1 - \frac{1}{N}\right) \pi_0 + \frac{2}{N} \pi_2 \Rightarrow \pi_2 = \frac{N(N-1)}{2} \pi_0 = N C_2 \cdot \pi_0$$

$$\pi_2 = \left(1 - \frac{2}{N}\right)\pi_0 + \frac{3}{N}\pi_3 \Rightarrow \pi_3 = \frac{N(N-1)(N-2)}{6} \pi_0 = N\zeta_3 \cdot \pi_0$$

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In general, we get $\pi_k = N\zeta_k \cdot \pi_0$, $k \in \{0, \dots, N\}$.

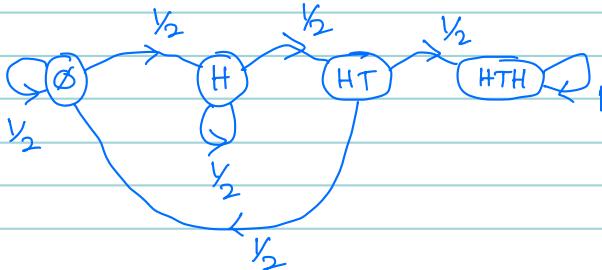
Because $\sum_{k=0}^N \pi_k = 1$, we get

$$\pi_0 = \frac{1}{N\zeta_0 + \dots + N\zeta_N} = \frac{1}{2^N}.$$

$$\text{Thus, } \pi_k = \frac{N\zeta_k}{2^N}, \quad k \in \{0, \dots, N\}.$$

$$\text{Then, it follows that } \mu_{kk} = \frac{1}{\pi_k} = \frac{2^N}{N\zeta_k}, \quad k \in \{0, \dots, N\}.$$

6. Consider a time-homogeneous DTmc with the following transition graph.



$$\Psi_0 := E[\# \text{ tosses required to get "HTH"} | X_0 = \emptyset] = \sum_{n=1}^{\infty} n \cdot P(\tau_{HTH}^{(1)} = n | X_0 = \emptyset)$$

$$\Psi_1 := E[\# \text{ tosses required to get "HTH"} | X_0 = H] = \sum_{n=1}^{\infty} n \cdot P(\tau_{HTH}^{(1)} = n | X_0 = H)$$

$$\Psi_2 := E[\# \text{ tosses required to get "HTH"} | X_0 = HT] = \sum_{n=1}^{\infty} n \cdot P(\tau_{HTH}^{(1)} = n | X_0 = HT)$$

$$\Psi_3 := E[\# \text{ tosses required to get "HTH"} | X_0 = HTH] = 0.$$

We then have

$$\begin{aligned} P(\tau_{HTH}^{(1)} = n | X_0 = \emptyset) &= P(X_1 \neq HTH, \dots, X_{n-1} \neq HTH, X_n = HTH | X_0 = \emptyset) \\ &= P(X_1 = \emptyset, X_2 \neq HTH, \dots, X_{n-1} \neq HTH, X_n = HTH | X_0 = \emptyset) \\ &\quad + P(X_1 = H, X_2 \neq HTH, \dots, X_{n-1} \neq HTH, X_n = HTH | X_0 = \emptyset) \end{aligned}$$

Markov Property $\hookrightarrow = Y_2 \cdot P(X_2 \neq HTH, \dots, X_{n-1} \neq HTH, X_n = HTH | X_1 = \emptyset)$

can neglect X_0 if given X_1 $+ Y_2 \cdot P(X_2 \neq HTH, \dots, X_{n-1} \neq HTH, X_n = HTH | X_1 = H)$.

We now see that by Markov property,

$$\begin{aligned} & \mathbb{P}(X_2 \neq HTH, \dots, X_{n-1} \neq HTH, X_n = HTH \mid X_0 = \phi) \\ &= \mathbb{P}(X_1 \neq HTH, \dots, X_{n-2} \neq HTH, X_{n-1} = HTH \mid X_0 = \phi) \\ &= \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = \phi). \end{aligned}$$

Similarly,

$$\mathbb{P}(X_2 \neq HTH, \dots, X_{n-1} \neq HTH, X_n = HTH \mid X_0 = H) = \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = H).$$

Thus, we have

$$\mathbb{P}(\tau_{HTH}^{(1)} = n \mid X_0 = \phi) = \frac{1}{2} \cdot \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = \phi) + \frac{1}{2} \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = H),$$

from which we get

$$\begin{aligned} \Psi_0 &= \sum_{n=1}^{\infty} n \left(\frac{1}{2} \cdot \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = \phi) + \frac{1}{2} \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = H) \right) \\ &= \sum_{n=2}^{\infty} n \left(\frac{1}{2} \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = \phi) + \frac{1}{2} \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = H) \right) \\ &\quad \xrightarrow{\text{noting that } \mathbb{P}(\tau_{HTH}^{(1)} = 0 \mid X_0 = \phi) = 0 = \mathbb{P}(\tau_{HTH}^{(1)} = 0 \mid X_0 = H)} \\ &= \frac{1}{2} \left[\sum_{n=2}^{\infty} (n-1) \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = \phi) + \sum_{n=2}^{\infty} \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = \phi) \right] \\ &\quad + \frac{1}{2} \left[\sum_{n=2}^{\infty} (n-1) \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = H) + \sum_{n=2}^{\infty} \mathbb{P}(\tau_{HTH}^{(1)} = n-1 \mid X_0 = H) \right] \\ &= \frac{1}{2} \left[\sum_{m=1}^{\infty} m \mathbb{P}(\tau_{HTH}^{(1)} = m \mid X_0 = \phi) + \sum_{m=1}^{\infty} \mathbb{P}(\tau_{HTH}^{(1)} = m \mid X_0 = \phi) \right] \\ &\quad \xrightarrow{\mathbb{P}(\tau_{HTH}^{(1)} < +\infty \mid X_0 = \phi) = 1 \text{ because HTH is an absorbing state}} \\ &+ \frac{1}{2} \left[\sum_{m=1}^{\infty} m \mathbb{P}(\tau_{HTH}^{(1)} = m \mid X_0 = H) + \sum_{m=1}^{\infty} \mathbb{P}(\tau_{HTH}^{(1)} = m \mid X_0 = H) \right] \\ &\quad \xrightarrow{\mathbb{P}(\tau_{HTH}^{(1)} < +\infty \mid X_0 = H) = 1} \\ &= \frac{1}{2} [\Psi_0 + 1] + \frac{1}{2} [\Psi_1 + 1]. \end{aligned}$$

Along similar lines, we may derive/justify the other recursive relations.