



Time Series Analytics

111-1 Homework #03

Due at 23h59, October 2, 2022; files uploaded to NTU-COOL

1. (10%) y_t is a stationary process with the autocovariance function γ_k . Define $\bar{y} = \frac{1}{n} \sum_{t=1}^n y_t$. Show that

$$V[\bar{y}] = \frac{\gamma_0}{n} + \frac{2}{n} \sum_{k=1}^{n-1} \left(1 - \frac{k}{n}\right) \gamma_k = \frac{1}{n} \sum_{k=-n+1}^{n-1} \left(1 - \frac{|k|}{n}\right) \gamma_k.$$

2. (10%) Assume x_t is a stationary process and define $y_t = \begin{cases} x_t & \text{for odd } t \\ x_t + 3 & \text{for even } t \end{cases}$
 (a) Show that $\text{COV}[y_t, y_{t-k}]$ is independent of t for all lags k .
 (b) Is y_t stationary?
3. (10%) Let y_t be a stationary process with an autocovariance function γ_k .
 (a) Show that $z_t = \nabla y_t = y_t - y_{t-1}$ is stationary by finding the mean and autocovariance function for z_t .
 (b) Show that $w_t = \nabla^2 y_t = z_t - z_{t-1} = y_t - 2y_{t-1} + y_{t-2}$ is stationary.
4. (15%) Let x_t be stationary with $E[x_t] = 0, V[x_t] = 1$, autocorrelation function ρ_k . Define that μ_t is a nonconstant function and σ_t is a positively nonconstant **function** (that is to say: μ_t and σ_t are both deterministic and in function of t). Now we observe a time series formulated as

$$y_t = \mu_t + \sigma_t x_t.$$

 (a) Find the mean and autocovariance function of y_t .
 (b) Show that the autocorrelation of y_t depends only on the lag k . Is y_t stationary?
 (c) Let $\mu_t = \mu_0$ be a constant value. Justify that y_t is still nonstationary.
5. (5%) Let x_t be the series of the “expected” measurements during the production process. Because the measuring tool itself won’t be perfect, we observe $y_t = x_t + e_t$, assuming x_t and e_t are independent. In general, we call x_t the **signal** and e_t the **measurement (white) noise**.

If x_t is stationary with the autocorrelation function ρ_k , show that y_t is also a stationary process with

$$\text{corr}(y_t, y_{t-k}) = \frac{\rho_k}{1 + \frac{\sigma_e^2}{\sigma_x^2}}, \text{ for } k \geq 1.$$

$\frac{\sigma_x^2}{\sigma_e^2}$ is usually referred to as the **signal-to-noise ratio**, or **SNR** for short.

The larger the SNR, the closer the autocorrelation function of the observed series y_t is to the autocorrelation function of the desired signal x_t .

6. (10%) Suppose $y_t = \alpha_0 + \sum_{i=1}^q [\alpha_i \cos(2\pi f_i t) + \beta_i \sin(2\pi f_i t)]$, where $\alpha_0, f_1, f_2, \dots, f_q$ are constants and $\alpha_1, \alpha_2, \dots, \alpha_q, \beta_1, \beta_2, \dots, \beta_q$ are independent random variables with zero means and variances $V[\alpha_i] = V[\beta_i] = \sigma_i^2$. Show that y_t is stationary and find its autocovariance function.

(Hint: show $\text{COV}[y_t, y_s]$ depends only on $t - s$.)

1. (10%) y_t is a stationary process with the autocovariance function γ_k . Define $\bar{y} = \frac{1}{n} \sum_{t=1}^n y_t$. Show that

$$V[\bar{y}] = \frac{\gamma_0}{n} + \frac{2}{n} \sum_{k=1}^{n-1} \left(1 - \frac{k}{n}\right) \gamma_k = \frac{1}{n} \sum_{k=-n+1}^{n-1} \left(1 - \frac{|k|}{n}\right) \gamma_k.$$

Given. y_t is a stationary process with autocovariance function γ_k

$$\bar{y} = \frac{1}{n} \sum_{t=1}^n y_t$$

Proof of \bar{y} : $V[\bar{y}] = \frac{\gamma_0}{n} \left[1 + 2 \sum_{k=1}^{n-1} \left(1 - \frac{k}{n}\right) \gamma_k \right] = \text{Var} \left[\frac{1}{n} \sum_{t=1}^n y_t \right]$

$$= \frac{1}{n^2} \text{Var} \left[\sum_{t=1}^n y_t \right]$$

$$= \frac{1}{n^2} \text{Cov} \left[\sum_{t=1}^n y_t, \sum_{s=1}^n y_s \right]$$

$$= \frac{1}{n^2} \sum_{t=1}^n \sum_{s=1}^n \gamma_{t-s}$$

Theorem: Variance of sum

$$\text{Var} \left(\sum_{i=1}^n X_i \right) = \sum_{i=1}^n \text{Var}(X_i) + 2 \sum_{i < j} \text{Cov}(X_i, X_j)$$

Let $k = t-s$

$$V[\bar{y}] = \frac{1}{n^2} \sum_{t=1}^n \sum_{s=t-k+1}^n \gamma_k$$

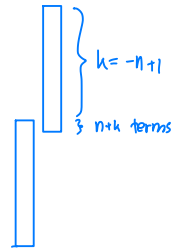
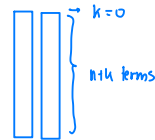
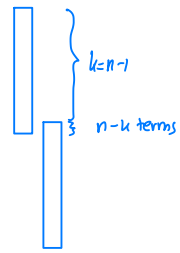
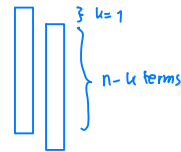
$$= \frac{1}{n} \sum_{t=1}^n \sum_{k=t+1}^{n-t} \gamma_k$$

$$= \frac{1}{n^2} \left(\sum_{k=1}^{n-1} (n-k) \gamma_k + \sum_{k=-n+1}^0 (n+k) \gamma_k \right)$$

$$= \frac{1}{n^2} \sum_{k=-n+1}^{n-1} ((n-k) \gamma_k + (n+k) \gamma_k)$$

$$= \frac{1}{n^2} \sum_{k=-n+1}^{n-1} (n-|k|) \gamma_k$$

$$= \frac{1}{n} \sum_{k=-n+1}^{n-1} \left(1 - \frac{|k|}{n}\right) \gamma_k$$



2. (10%) Assume x_t is a stationary process and define $y_t = \begin{cases} x_t & \text{for odd } t \\ x_t + 3 & \text{for even } t \end{cases}$
- (a) Show that $\text{COV}[y_t, y_{t-k}]$ is independent of t for all lags k .
- (b) Is y_t stationary?

Given: x_t is stationary

$$y_t = \begin{cases} x_t & \text{for odd } t \\ x_t + 3 & \text{for even } t \end{cases}$$

- a) Show that $\text{COV}[y_t, y_{t-k}]$ is independent of t for all lags k

$$\begin{aligned} \text{COV}[y_t, y_{t-k}] &= E[(y_t - \mu)(y_{t-k} - \mu)] \\ &= E[y_t y_{t-k} - \mu y_t - \mu y_{t-k} + \mu^2] \\ &= E[y_t]E[y_{t-k}] - \cancel{(E[y_t])^2} - E[y_t]E[y_{t-k}] + \cancel{(E[y_t])^2} \\ &= 0 \end{aligned}$$

Since $\text{COV}[y_t, y_{t-k}] = 0$, and x_t is stationary, the $\text{COV}[y_t, y_{t-k}]$ is independent of t for all lags k .

- b) Is y_t stationary

Consider t and k are both odd, meaning $t-k$ is even

$$\hookrightarrow \text{COV}[y_t, y_{t-k}] = \text{COV}[x_t, x_{t+3}] = E[(x_t - E[x_t])(x_{t+3} - E[x_{t+3}])]$$

Consider t and k are both even, meaning $t-k$ is even

$$\hookrightarrow \text{COV}[y_t, y_{t-k}] = \text{COV}[x_t + 3, x_{t+3} + 3] = E[(x_{t+3} - E[x_{t+3}])(x_{t+3} - E[x_{t+3}])]$$

Consider t and k are odd & even, meaning $t-k$ is odd

$$\hookrightarrow \text{COV}[y_t, y_{t-k}] = \text{COV}[x_t, x_t] = E[(x_t - E[x_t])(x_t - E[x_t])]$$

Consider t and k are even & odd, meaning $t-k$ is odd

$$\hookrightarrow \text{COV}[y_t, y_{t-k}] = \text{COV}[x_t + 3, x_t] = E[(x_{t+3} - E[x_{t+3}])(x_t - E[x_t])]$$

For all cases, y_t depends on t

Since y_t is $x_t + 3$ when t is even, and x_t when t is odd, y_t depends on t and thus, y_t is not a stationary time series.

3. (10%) Let y_t be a stationary process with an autocovariance function γ_k .

(a) Show that $z_t = \nabla y_t = y_t - y_{t-1}$ is stationary by finding the mean and autocovariance function for z_t .

(b) Show that $w_t = \nabla^2 y_t = z_t - z_{t-1} = y_t - 2y_{t-1} + y_{t-2}$ is stationary.

a) $z_t = \nabla y_t = y_t - y_{t-1}$

Since y_t is a stationary process, it must have a constant mean μ over time and autocovariance function γ_k

Mean (Expectation)

$$E[z_t] = E[y_t - y_{t-1}]$$

$$E[z_t] = E[y_t] - E[y_{t-1}]$$

$$= \mu - \mu$$

$$= 0$$

Note: the mean function is a constant and is independent of time

Autocovariance

The autocovariance function of $\{z_t\}$ is $\text{Cov}(z_t, z_{t-k})$

$$\text{Cov}(z_t, z_{t-k}) = \text{Cov}(y_t - y_{t-1}, y_{t-k} - y_{t-k-1}) \quad \leftarrow \text{Apply the distributive property of covariance}$$

$$= \text{Cov}(y_t, y_{t-k}) + \text{Cov}(y_t, -y_{t-k-1}) + \text{Cov}(-y_{t-1}, y_{t-k}) + \text{Cov}(-y_{t-1}, -y_{t-k-1})$$

$$= \text{Cov}(y_t, y_{t-k}) - \text{Cov}(y_t, y_{t-k-1}) - \text{Cov}(y_{t-1}, y_{t-k}) + \text{Cov}(y_{t-1}, y_{t-k-1})$$

$$= \gamma_k - \gamma_{k+1} - \gamma_{k-1} + \gamma_k$$

Note: the autocovariance function is independent of time

Since both the mean and autocovariance function are independent of time, z_t is stationary //

b) $w_t = \nabla z_t = z_t - z_{t-1} = y_t - 2y_{t-1} + y_{t-2}$

Since it is proven that the first-order difference of y_t , " z_t ", is stationary, then the first-order difference of z_t , " w_t " must also be stationary. //

4. (15%) Let x_t be stationary with $E[x_t] = 0, V[x_t] = 1$, autocorrelation function ρ_k . Define that μ_t is a nonconstant function and σ_t is a positively nonconstant **function** (that is to say: μ_t and σ_t are both deterministic and in function of t). Now we observe a time series formulated as

$$y_t = \mu_t + \sigma_t x_t.$$

- (a) Find the mean and autocovariance function of y_t .
 (b) Show that the autocorrelation of y_t depends only on the lag k . Is y_t stationary?
 (c) Let $\mu_t = \mu_0$ be a constant value. Justify that y_t is still nonstationary.

a) Mean function of y_t

$$\begin{aligned} E(y_t) &= E(\mu_t + \sigma_t x_t) \\ &= \mu_t + \sigma_t E[x_t] \end{aligned}$$

$$E(y_t) = \mu_t //$$

Autocovariance function of y_t

$$\begin{aligned} \text{Cov}(y_t, y_{t-k}) &= \text{Cov}(\mu_t + \sigma_t x_t, \mu_{t-k} + \sigma_{t-k} x_{t-k}) \\ &= \text{Cov}(\sigma_t x_t, \sigma_{t-k} x_{t-k}) \\ &= \sigma_t \sigma_{t-k} \text{Cov}(x_t, x_{t-k}) \\ &= \sigma_t \sigma_{t-k} [\text{Corr}(x_t, x_{t-k}) \sqrt{\text{Var}(x_t) \text{Var}(x_{t-k})}] \\ &= \sigma_t \sigma_{t-k} (\rho_k \sqrt{1 \cdot 1}) \end{aligned}$$

$$\text{Cov}(y_t, y_{t-k}) = \sigma_t \sigma_{t-k} \rho_k //$$

b) Consider the variance

$$\begin{aligned} \text{Var}(y_t) &= \text{Var}(\mu_t + \sigma_t x_t) \\ &= \text{Var}(\sigma_t x_t) \\ &= \sigma_t^2 \text{Var}(x_t) \\ \text{Var}(y_t) &= \sigma_t^2 \end{aligned}$$

Consider the autocorrelation function

$$\begin{aligned} \text{Corr}(y_t, y_{t-k}) &= \frac{\text{Cov}(y_t, y_{t-k})}{\sqrt{\text{Var}(y_t) \text{Var}(y_{t-k})}} \\ &= \frac{\sigma_t \sigma_{t-k} \rho_k}{\sqrt{\sigma_t^2 \sigma_{t-k}^2}} \\ \text{Corr}(y_t, y_{t-k}) &= \rho_k \end{aligned}$$

Since the autocorrelation of y_t is ρ_k

for all cases, autocorrelation of y_t does not depend on t since x_t is stationary.

Thus, autocorrelation of y_t only depends on the lag k //

However, since $y_t = \mu_t + \sigma_t x_t$, and μ_t and σ_t are deterministic and are functions of t , y_t is NOT stationary

- c) By having a mean function that is constant over time t , the mean function becomes stationary.

$$\hookrightarrow y_t = \mu + \sigma_t x_t$$

The autocovariance is still $\text{Cov}(y_t, y_{t-k}) = \sigma_t \sigma_{t-k} \rho_k$.

While the mean function is constant, the autocovariance function is not.

Thus, y_t is still nonstationary //

5. (5%) Let x_t be the series of the "expected" measurements during the production process. Because the measuring tool itself won't be perfect, we observe $y_t = x_t + e_t$, assuming x_t and e_t are independent. In general, we call x_t the **signal** and e_t the **measurement (white) noise**.

If x_t is stationary with the autocorrelation function ρ_k , show that y_t is also a stationary process with

$$\text{corr}(y_t, y_{t-k}) = \frac{\rho_k}{1 + \frac{\sigma_e^2}{\sigma_x^2}}, \text{ for } k \geq 1.$$

$\frac{\sigma_x^2}{\sigma_e^2}$ is usually referred to as the **signal-to-noise ratio**, or **SNR** for short.

The larger the SNR, the closer the autocorrelation function of the observed series y_t is to the autocorrelation function of the desired signal x_t .

a) Mean function of y_t

$$E[y_t] = E[x_t + e_t]$$

$$= E[x_t] + E[e_t] = \mu_x$$

- Note that x_t is stationary, meaning the mean function $E[x_t]$ is a constant
- Since "white noise" is random, due to it being the measurement of the environment, the mean function $E[e_t]$ doesn't depend on time t
- Thus $E[y_t]$ is stationary.

Variance of y_t

$$\text{Var}[y_t] = \text{Var}[x_t + e_t]$$

$$= \text{Var}[x_t] + \text{Var}[e_t]$$

$$= \sigma_x^2 + \sigma_e^2$$

Autocovariance function of y_t

$$\text{Cov}[y_t, y_{t-k}] = \text{Cov}[x_t + e_t, x_{t-k} + e_{t-k}]$$

$$= \text{Cov}[x_t, x_{t-k}]$$

$$= \rho_k$$

Note: the autocovariance function is independent of time

Since both the mean and autocovariance function are independent of time, y_t is stationary //

Autocorrelation function of y_t

$$\text{Corr}[y_t, y_{t-k}] = \frac{\text{Cov}[y_t, y_{t-k}]}{\sqrt{\text{Var}(y_t) \text{Var}(y_{t-k})}}$$

$$= \frac{\rho_k}{\sqrt{(\sigma_x^2 + \sigma_e^2)(\sigma_x^2 + \sigma_e^2)}}$$

$$= \frac{\rho_k}{\sigma_x^2 + \sigma_e^2}$$

$$\text{Corr}[y_t, y_{t-k}] = \frac{\rho_k}{1 + \frac{\sigma_e^2}{\sigma_x^2}} //$$

6. (10%) Suppose $y_t = \alpha_0 + \sum_{i=1}^q [\alpha_i \cos(2\pi f_i t) + \beta_i \sin(2\pi f_i t)]$, where $\alpha_0, f_1, f_2, \dots, f_q$ are constants and $\alpha_1, \alpha_2, \dots, \alpha_q, \beta_1, \beta_2, \dots, \beta_q$ are independent random variables with zero means and variances $V[\alpha_i] = V[\beta_i] = \sigma_i^2$. Show that y_t is stationary and find its autocovariance function.

(Hint: show $\text{COV}[y_t, y_s]$ depends only on $t - s$.)

Mean function of y_t

$$\begin{aligned} E[y_t] &= E \left[\alpha_0 + \sum_{i=1}^q (\alpha_i \cos(2\pi f_i t) + \beta_i \sin(2\pi f_i t)) \right] \\ &= \alpha_0 + \sum_{i=1}^q E[\alpha_i] \cos(2\pi f_i t) + E[\beta_i] \sin(2\pi f_i t) \\ &= \alpha_0 \end{aligned}$$

Note: the mean function is a constant and is independent of time

Autocovariance function of y_t

$$\begin{aligned} \text{Cov}[y_t, y_s] &= \text{Cov} \left[\sum_{i=1}^k \alpha_i \cos(2\pi f_i t) + \beta_i \sin(2\pi f_i t), \sum_{i=1}^k \alpha_i \cos(2\pi f_i s) + \beta_i \sin(2\pi f_i s) \right] \\ &= \sum_{i=1}^k \text{Cov}(\alpha_i \cos(2\pi f_i t) + \beta_i \sin(2\pi f_i t), \alpha_i \cos(2\pi f_i s) + \beta_i \sin(2\pi f_i s)) \quad \leftarrow \text{Theorem: } \text{Var}[X] = \text{Cov}[X, X] \\ &= \sum_{i=1}^k \text{Var}[\alpha_i] (\cos(2\pi f_i t) + \sin(2\pi f_i s)) + \sum_{i=1}^k \text{Var}[\beta_i] (\cos(2\pi f_i t) + \sin(2\pi f_i s)) \\ &= \frac{\sigma^2}{2} \sum_{i=1}^k [\cos(2\pi f_i (t-s)) + \sin(2\pi f_i (t+s))] + \frac{\sigma^2}{2} \sum_{i=1}^k [\cos(2\pi f_i (t-s)) + \sin(2\pi f_i (t+s))] \end{aligned}$$

$$\text{Cov}[y_t, y_s] = \sigma^2 \sum_{i=1}^k \cos(2\pi f_i (t-s))$$

Note: the autocovariance function is only dependent on $t-s$

Since both the mean and autocovariance function are independent of time, y_t is stationary //