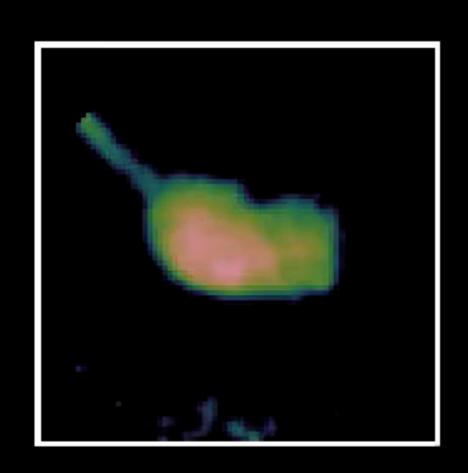
# Machine Learning Methods for Neural Data Analysis

More Hidden Markov Models

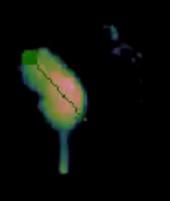
#### Announcements

- Lab 6
  - Add atol=1e-4 to the allclose checks.
- Please submit your 1 page project proposal tonight.

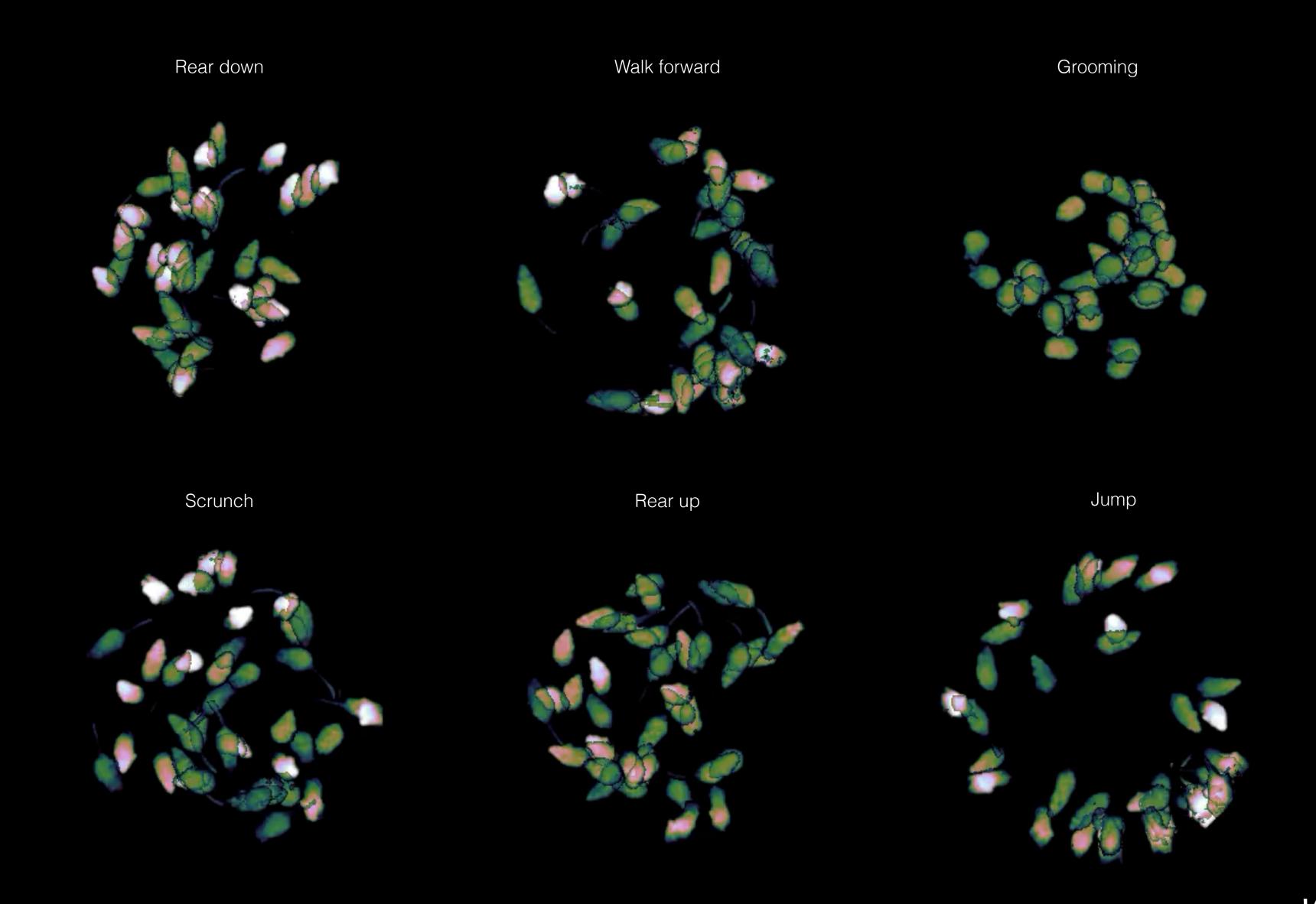
#### Motivating Example: summarizing videos with behavioral states







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#### The Gaussian HMM

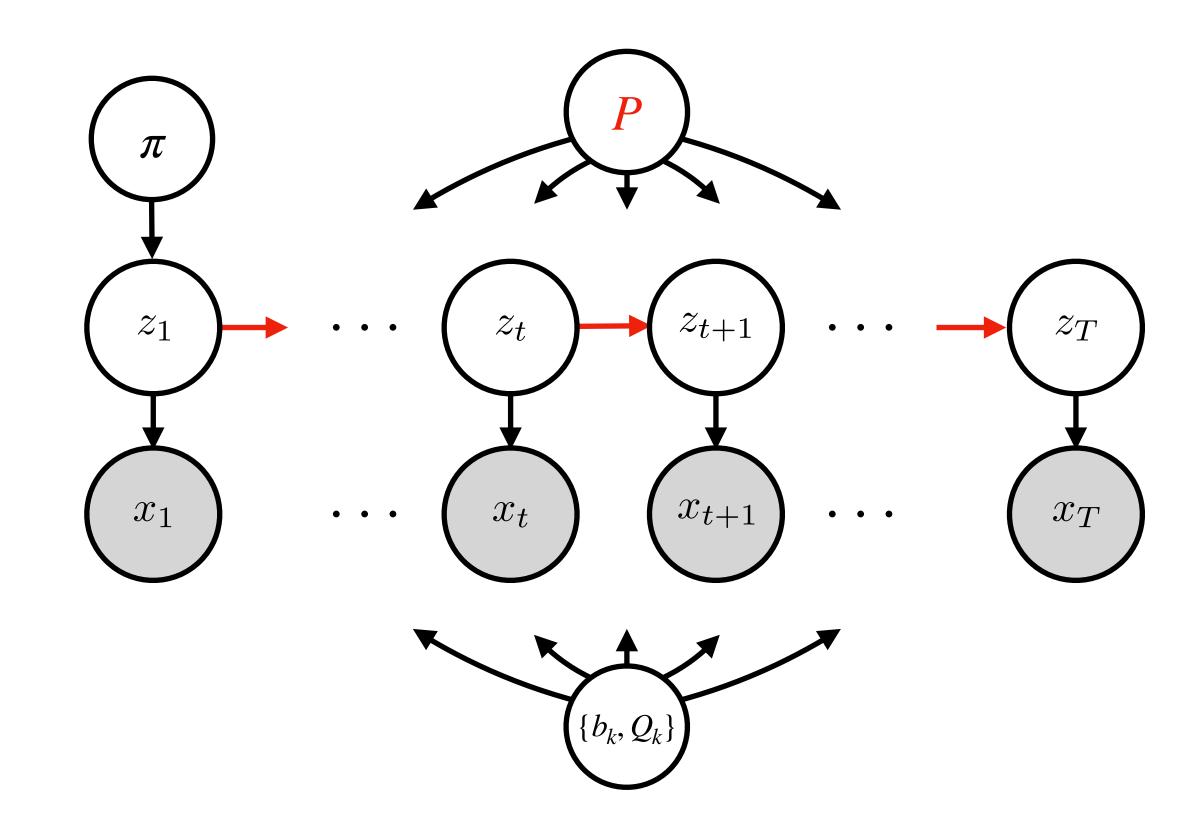
#### **Graphical Model**

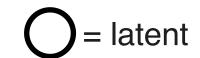
**Transition Probabilities** 

Discrete **Latent States** 

Observations (e.g. PCA loadings of each frame)

> State Means and Covariances









#### The Gaussian HMM

A Gaussian HMM is just a Gaussian mixture model but where cluster assignments are linked across time!

$$z_1 \sim \text{Cat}(\pi),$$

$$z_t \mid z_{t-1} \sim \text{Cat}(P_{z_{t-1}}), \quad \text{for } t = 2, ..., T.$$

$$x_t \mid z_t \sim \mathcal{N}(b_{z_t}, Q_{z_t}) \quad \text{for } t = 1, ..., T$$

Its parameters are  $\Theta = \pi, P, \{b_k, Q_k\}_{k=1}^K$  where  $P \in [0,1]^{K \times K}$  is a row-stochastic transition matrix.

Under this model, the joint probability factors as

$$p(x, z, \Theta) = p(z_1) \prod_{t=1}^{T-1} p(z_{t+1} \mid z_t) \prod_{t=1}^{T} p(x_t \mid z_t)$$

#### Bayesian inference in latent variable models

#### The Expectation-Maximization (EM) algorithm

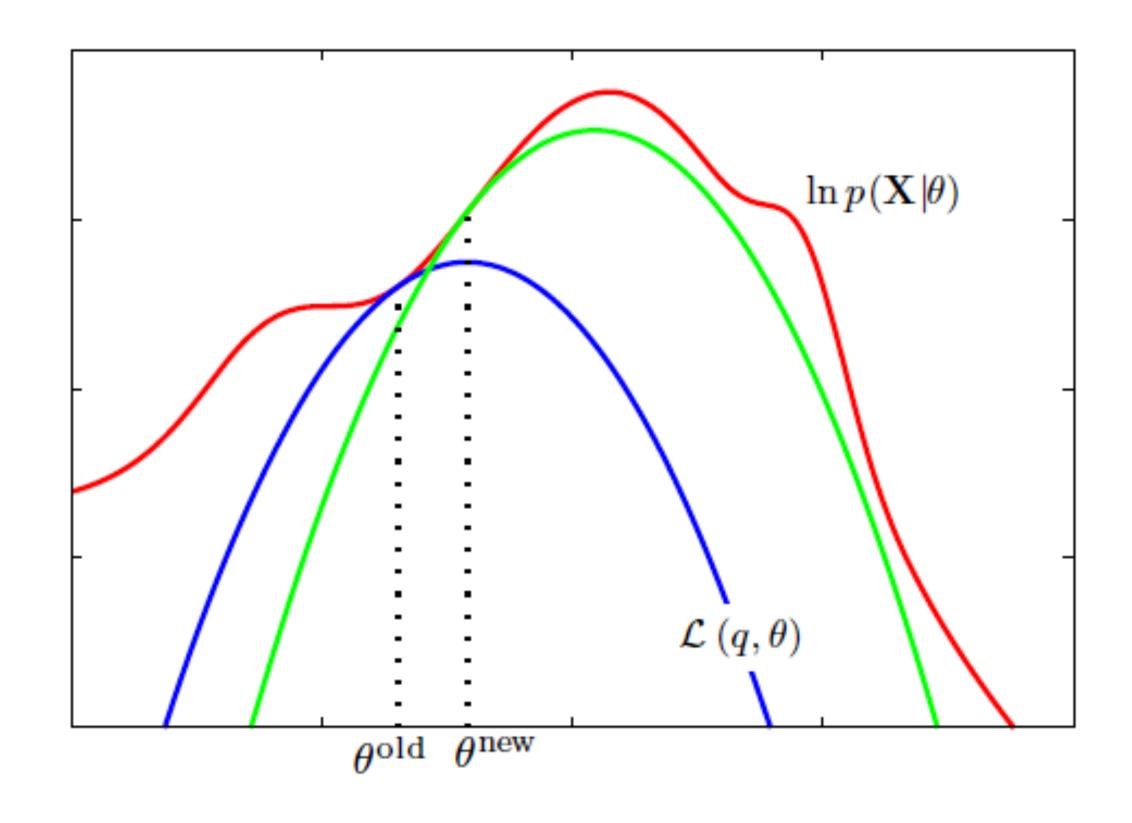
• M-step: Maximize the expected log probability

$$\Theta \leftarrow \arg \max_{\Theta} \mathbb{E}_{q(z)}[\log p(x, z, \Theta)]$$

• **E-step**: Update the posterior over latent variables

$$q \leftarrow p(z \mid x, \Theta)$$

• EM converges to **local maxima** of the log marginal likelihood,  $\log p(x; \Theta)$ 



In the M-step we set,

$$\Theta = \arg \max_{\Theta} \mathbb{E}_{q(z)}[\log p(x, z, \Theta)]$$

$$\mathbb{E}_{q(z)}[\log p(x, z, \Theta)] = \mathbb{E}_{q(z)} \left[ \sum_{t=1}^{T} \log \mathcal{N}(x_t \mid b_{z_t}, Q_{z_t}) + \log \operatorname{Cat}(z_t \mid \pi) \right]$$

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$$= \mathbb{E}_{q(z)} \left[ \sum_{t=1}^{T} \sum_{j=1}^{K} \mathbb{I}[z_t = j] \log \mathcal{N}(x_t \mid b_j, Q_j) \right] + \operatorname{const}$$

In the M-step we set,

$$\Theta = \arg \max_{\Theta} \mathbb{E}_{q(z)}[\log p(x, z, \Theta)]$$

$$\mathbb{E}_{q(z)}[\log p(x, z, \Theta)] = \mathbb{E}_{q(z)} \left[ \sum_{t=1}^{T} \log \mathcal{N}(x_t \mid b_{z_t}, Q_{z_t}) + \log \operatorname{Cat}(z_t \mid \pi) \right]$$

$$= \mathbb{E}_{q(z)} \left[ \sum_{t=1}^{T} \sum_{j=1}^{K} \mathbb{I}[z_t = j] \log \mathcal{N}(x_t \mid b_j, Q_j) \right] + \operatorname{const}$$

$$= \sum_{t=1}^{T} \mathbb{E}_{q(z)}[\mathbb{I}[z_t = k]] \left( -\frac{1}{2} \log |Q_k| - \frac{1}{2} (x_t - b_k)^{\mathsf{T}} Q_k^{-1} (x_t - b_k) \right) + \operatorname{const}$$

In the M-step we set,

$$\Theta = \arg \max_{\Theta} \mathbb{E}_{q(z)}[\log p(x, z, \Theta)]$$

$$\begin{split} \mathbb{E}_{q(z)}[\log p(x, z, \Theta)] &= \mathbb{E}_{q(z)} \left[ \sum_{t=1}^{T} \log \mathcal{N}(x_{t} \mid b_{z_{t}}, Q_{z_{t}}) + \log \operatorname{Cat}(z_{t} \mid \pi) \right] \\ &= \mathbb{E}_{q(z)} \left[ \sum_{t=1}^{T} \sum_{j=1}^{K} \mathbb{I}[z_{t} = j] \log \mathcal{N}(x_{t} \mid b_{j}, Q_{j}) \right] + \operatorname{const} \\ &= \sum_{t=1}^{T} \mathbb{E}_{q(z)}[\mathbb{I}[z_{t} = k]] \left( -\frac{1}{2} \log |Q_{k}| - \frac{1}{2} (x_{t} - b_{k})^{\mathsf{T}} Q_{k}^{-1} (x_{t} - b_{k}) \right) + \operatorname{const} \\ &= \sum_{t=1}^{T} q(z_{t} = k) \left( -\frac{1}{2} \log |Q_{k}| - \frac{1}{2} (x_{t} - b_{k})^{\mathsf{T}} Q_{k}^{-1} (x_{t} - b_{k}) \right) + \operatorname{const} \end{split}$$

#### The M-step

Taking derivatives and setting to zero yields the following updates,

$$T_{k} = \sum_{t=1}^{T} q(z_{t} = k)$$

$$b_{k} = \frac{1}{T_{k}} \sum_{t=1}^{T} q(z_{t} = k) x_{t}$$

$$Q_{k} = \frac{1}{T_{k}} \sum_{t=1}^{T} q(z_{t} = k) (x_{t} - b_{k}) (x_{t} - b_{k})^{T}$$

Note: we only need the posterior marginal probabilities  $q(z_t = k)$ !

#### The posterior is a little trickier than in the Gaussian mixture model

• E-step: Update the posterior over latent variables,

$$q(z) \leftarrow p(z \mid x, \Theta) \propto p(x, z, \Theta) = p(z_1) \prod_{t=1}^{T-1} p(z_{t+1} \mid z_t) \prod_{t=1}^{T} p(x_t \mid z_t)$$

- The normalized posterior no longer has a simple closed form!
- However, we can still efficiently compute the marginal probabilities for the M-step.

#### Computing posterior marginals

• Consider the marginal probability of state *k* at time *t*:

$$q(z_t = k) = \sum_{z_1=1}^K \cdots \sum_{z_{t-1}=1}^K \sum_{z_{t+1}=1}^K \cdots \sum_{z_T=1}^K q(z_1, \dots, z_{t-1}, z_t = k, z_{t+1}, \dots, z_T)$$

#### Computing posterior marginals

• Consider the marginal probability of state k at time t:

$$q(z_{t} = k) = \sum_{z_{1}=1}^{K} \cdots \sum_{z_{t-1}=1}^{K} \sum_{z_{t+1}=1}^{K} \cdots \sum_{z_{T}=1}^{K} q(z_{t}, \dots, z_{t-1}, z_{t} = k, z_{t+1}, \dots, z_{T})$$

$$\propto \left[ \sum_{z_{1}=1}^{K} \cdots \sum_{z_{t-1}=1}^{K} p(z_{1}) \prod_{s=1}^{t-1} p(x_{s} \mid z_{s}) p(z_{s+1} \mid z_{s}) \right] \times \left[ p(x_{t} \mid z_{t}) \right]$$

$$\times \left[ \sum_{z_{t+1}=1}^{K} \cdots \sum_{z_{T}=1}^{K} \prod_{u=t+1}^{T} p(z_{u} \mid z_{u-1}) p(x_{u} \mid z_{u}) \right]$$

#### Computing posterior marginals

• Consider the marginal probability of state k at time t:

$$q(z_{t} = k) = \sum_{z_{1}=1}^{K} \cdots \sum_{z_{t-1}=1}^{K} \sum_{z_{t+1}=1}^{K} \cdots \sum_{z_{T}=1}^{K} q(z_{t}, \dots, z_{t-1}, z_{t} = k, z_{t+1}, \dots, z_{T})$$

$$\propto \left[ \sum_{z_{1}=1}^{K} \cdots \sum_{z_{t-1}=1}^{K} p(z_{1}) \prod_{s=1}^{t-1} p(x_{s} \mid z_{s}) p(z_{s+1} \mid z_{s}) \right] \times \left[ p(x_{t} \mid z_{t}) \right]$$

$$\times \left[ \sum_{z_{t+1}=1}^{K} \cdots \sum_{z_{T}=1}^{K} \prod_{u=t+1}^{T} p(z_{u} \mid z_{u-1}) p(x_{u} \mid z_{u}) \right]$$

$$\triangleq \alpha_{t}(z_{t}) \times p(x_{t} \mid z_{t}) \times \beta_{t}(z_{t})$$

#### Computing the forward messages $\alpha_t(z_t)$

Consider the forward messages:

$$\alpha_{t}(z_{t}) \triangleq \sum_{z_{1}=1}^{K} \cdots \sum_{z_{t-1}=1}^{K} p(z_{1}) \prod_{s=1}^{t-1} p(x_{s} \mid z_{s}) p(z_{s+1} \mid z_{s})$$

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$$= \sum_{z_{t-1}=1}^{K} \left[ \left( \sum_{z_{1}=1}^{K} \cdots \sum_{z_{t-2}=1}^{K} p(z_{1}) \prod_{s=1}^{t-2} p(x_{s} \mid z_{s}) p(z_{s+1} \mid z_{s}) \right) p(x_{t-1} \mid z_{t-1}) p(z_{t} \mid z_{t-1}) \right]$$

#### Computing the forward messages $\alpha_t(z_t)$

Consider the forward messages:

$$\alpha_{t}(z_{t}) \triangleq \sum_{z_{t-1}=1}^{K} \cdots \sum_{z_{t-1}=1}^{K} p(z_{1}) \prod_{s=1}^{t-1} p(x_{s} \mid z_{s}) p(z_{s+1} \mid z_{s})$$

$$= \sum_{z_{t-1}=1}^{K} \left[ \left( \sum_{z_{1}=1}^{K} \cdots \sum_{z_{t-2}=1}^{K} p(z_{1}) \prod_{s=1}^{t-2} p(x_{s} \mid z_{s}) p(z_{s+1} \mid z_{s}) \right) p(x_{t-1} \mid z_{t-1}) p(z_{t} \mid z_{t-1}) \right]$$

$$= \sum_{z_{t-1}=1}^{K} \alpha_{t-1}(z_{t-1}) p(x_{t-1} \mid z_{t-1}) p(z_{t} \mid z_{t-1})$$

• We can compute these messages recursively!

#### Computing the forward messages $\alpha_t(z_t)$ . Vectorized.

• Let  $\alpha_t = [\alpha_t(z_t = 1), ..., \alpha_t(z_t = K)]^{\mathsf{T}}$  denote the column vector of forward messages. Then,

$$\alpha_t = P^{\mathsf{T}}(\alpha_{t-1} \odot \mathcal{E}_{t-1})$$

where

- $\ell_{t-1} = [p(x_{t-1} \mid z_{t-1} = 1), ..., p(x_{t-1} \mid z_{t-1} = K)]^{\mathsf{T}}$  is the vector of likelihoods,
- O denotes the element-wise product, and
- P is the transition matrix with  $P_{ij} = p(z_t = j \mid z_{t-1} = i)$ .
- For the base case, let  $\alpha_1(z_1) = p(z_1)$ .

#### Computing the backward messages $\beta_t(z_t)$

Now take the backward messages:

$$\beta_{t}(z_{t}) \triangleq \sum_{z_{t+1}=1}^{K} \cdots \sum_{z_{T}=1}^{K} \prod_{u=t+1}^{T} p(z_{u} \mid z_{u-1}) p(x_{u} \mid z_{u})$$

#### Computing the backward messages $\beta_t(z_t)$

Now take the backward messages:

$$\beta_{t}(z_{t}) \triangleq \sum_{z_{t+1}=1}^{K} \cdots \sum_{z_{T}=1}^{K} \prod_{u=t+1}^{T} p(z_{u} \mid z_{u-1}) p(x_{u} \mid z_{u})$$

$$= \sum_{z_{t+1}=1}^{K} p(z_{t+1} \mid z_{t}) p(x_{t+1} \mid z_{t+1}) \sum_{z_{t+2}=1}^{K} \cdots \sum_{z_{T}=1}^{K} \prod_{u=t+2}^{T} p(z_{u} \mid z_{u-1}) p(x_{u} \mid z_{u})$$

#### Computing the backward messages $\beta_t(z_t)$

Now take the backward messages:

$$\beta_{t}(z_{t}) \triangleq \sum_{z_{t+1}=1}^{K} \cdots \sum_{z_{T}=1}^{K} \prod_{u=t+1}^{T} p(z_{u} \mid z_{u-1}) p(x_{u} \mid z_{u})$$

$$= \sum_{z_{t+1}=1}^{K} p(z_{t+1} \mid z_{t}) p(x_{t+1} \mid z_{t+1}) \sum_{z_{t+2}=1}^{K} \cdots \sum_{z_{T}=1}^{K} \prod_{u=t+2}^{T} p(z_{u} \mid z_{u-1}) p(x_{u} \mid z_{u})$$

$$= \sum_{z_{t+1}=1}^{K} p(z_{t+1} \mid z_{t}) p(x_{t+1} \mid z_{t+1}) \beta_{t+1}(z_{t+1})$$

• Again, we can compute the backward messages recursively!

Computing the backward messages  $\beta_t(z_t)$ . Vectorized.

• Let  $\beta_t = [\beta_t(z_t = 1), ..., \beta_t(z_t = K)]^{\mathsf{T}}$  denote the column vector of backward messages. Then,

$$\beta_t = P(\beta_{t+1} \odot \ell_{t+1})$$

• For the base case, let  $\beta_T(z_T) = 1$ .

#### Combining the forward and backward messages

• The posterior marginal probability of state k at time t is,

$$q(z_t = k) \propto \alpha_t(z_t = k) \times p(x_t \mid z_t = k) \times \beta_t(z_t = k)$$
$$= \alpha_{tk} \ell_{tk} \beta_{tk}$$

The probabilities need to sum to one. Normalizing yields,

$$q(z_t = k) = \frac{\alpha_{tk} \ell_{tk} \beta_{tk}}{\sum_{j=1}^{K} \alpha_{tj} \ell_{tj} \beta_{tj}}$$

• Finally, note the marginal is invariant to multiplying  $\alpha_t$  and/or  $\beta_t$  by a constant.

#### Normalizing the messages to prevent underflow

- The messages involve products of probabilities, which quickly underflow.
- We can leverage the scale invariance to renormalize the messages. I.e. replace:

$$\alpha_t = P^\top (\alpha_{t-1} \odot \mathscr{C}_{t-1}) \quad \text{with} \quad \begin{aligned} A_{t-1} &= \sum_k \tilde{\alpha}_{t-1,k} \mathscr{C}_{t-1,k} \\ \tilde{\alpha}_t &= \frac{1}{A_{t-1}} P^\top (\tilde{\alpha}_{t-1} \odot \mathscr{C}_{t-1}) \end{aligned}$$

where  $\tilde{\alpha}_t$  are normalized for numerical stability. As before,  $\tilde{\alpha}_1 = \pi$ .

• This lends a nice interpretation: the forward messages are conditional probabilities  $\tilde{\alpha}_{tk} = p(z_t = k \mid x_{1:t-1})$  and the normalization constants are the marginal likelihoods  $A_t = p(x_t \mid x_{1:t-1})$ .

#### Computing the marginal likelihood

• Finally, we can compute the marginal likelihood alongside the forward messages

$$\log p(x \mid \Theta) = \log \sum_{z_{1}=1}^{K} \cdots \sum_{z_{T}=1}^{K} \left[ p(z_{1}) \prod_{t=1}^{T-1} p(z_{t+1} \mid z_{t}) \prod_{t=1}^{T} p(x_{t} \mid z_{t}) \right]$$

$$= \log \sum_{z_{T}=1}^{K} \alpha_{T}(z_{T}) p(x_{T} \mid z_{T})$$

$$= \log \prod_{t=1}^{T} A_{t} = \sum_{t=1}^{T} \log A_{t}$$

• Again, makes sense since the normalization constants are  $A_t = p(x_t \mid x_{1:t-1})$ .

## The M-step with sufficient statistics

#### **Sufficient statistics**

$$\mathbb{E}_{q(z)}[\log p(x, z, \Theta)] = \sum_{t=1}^{T} q(z_t = k) \left[ -\frac{1}{2} \log |Q_k| - \frac{1}{2} (x_t - b_k)^{\mathsf{T}} Q_k^{-1} (x_t - b_k) \right] + c$$

#### **Sufficient statistics**

$$\mathbb{E}_{q(z)}[\log p(x, z, \Theta)] = \sum_{t=1}^{T} q(z_t = k) \left[ -\frac{1}{2} \log |Q_k| - \frac{1}{2} (x_t - b_k)^{\mathsf{T}} Q_k^{-1} (x_t - b_k) \right] + c$$

$$= \sum_{t=1}^{T} q(z_t = k) \left[ -\frac{1}{2} \log |Q_k| - \frac{1}{2} x_t^{\mathsf{T}} Q_k^{-1} x_t + b_k^{\mathsf{T}} Q_k^{-1} x_t - \frac{1}{2} b_k^{\mathsf{T}} Q_k^{-1} b_k \right] + c$$

#### **Sufficient statistics**

$$\begin{split} \mathbb{E}_{q(z)}[\log p(x,z,\Theta)] &= \sum_{t=1}^{T} q(z_{t}=k) \left[ -\frac{1}{2} \log |Q_{k}| - \frac{1}{2} (x_{t} - b_{k})^{\mathsf{T}} Q_{k}^{-1} (x_{t} - b_{k}) \right] + \mathbf{c} \\ &= \sum_{t=1}^{T} q(z_{t}=k) \left[ -\frac{1}{2} \log |Q_{k}| - \frac{1}{2} x_{t}^{\mathsf{T}} Q_{k}^{-1} x_{t} + b_{k}^{\mathsf{T}} Q_{k}^{-1} x_{t} - \frac{1}{2} b_{k}^{\mathsf{T}} Q_{k}^{-1} b_{k} \right] + \mathbf{c} \\ &= \sum_{t=1}^{T} q(z_{t}=k) \left[ \left\langle -\frac{1}{2} \log |Q_{k}| , 1 \right\rangle + \left\langle -\frac{1}{2} Q_{k}^{-1} , x_{t} x_{t}^{\mathsf{T}} \right\rangle + \left\langle b_{k}^{\mathsf{T}} Q_{k}^{-1} , x_{t} \right\rangle + \left\langle -\frac{1}{2} b_{k}^{\mathsf{T}} Q_{k}^{-1} b_{k} , 1 \right\rangle \right] + \mathbf{c} \end{split}$$

#### **Sufficient statistics**

$$\begin{split} \mathbb{E}_{q(z)}[\log p(x,z,\Theta)] &= \sum_{t=1}^{T} q(z_{t}=k) \left[ -\frac{1}{2} \log |Q_{k}| - \frac{1}{2} (x_{t} - b_{k})^{\mathsf{T}} Q_{k}^{-1} (x_{t} - b_{k}) \right] + \mathbf{c} \\ &= \sum_{t=1}^{T} q(z_{t}=k) \left[ -\frac{1}{2} \log |Q_{k}| - \frac{1}{2} b_{k}^{\mathsf{T}} Q_{k}^{-1} b_{k} + b_{k}^{\mathsf{T}} Q_{k}^{-1} x_{t} - \frac{1}{2} x_{t}^{\mathsf{T}} Q_{k}^{-1} x_{t} \right] + \mathbf{c} \\ &= \sum_{t=1}^{T} q(z_{t}=k) \left[ \left\langle -\frac{1}{2} \log |Q_{k}| - \frac{1}{2} b_{k}^{\mathsf{T}} Q_{k}^{-1} b_{k}, 1 \right\rangle + \left\langle b_{k}^{\mathsf{T}} Q_{k}^{-1}, x_{t} \right\rangle + \left\langle -\frac{1}{2} Q_{k}^{-1}, x_{t} x_{t}^{\mathsf{T}} \right\rangle \right] + \mathbf{c} \\ &= \left\langle -\frac{1}{2} \log |Q_{k}| - \frac{1}{2} b_{k}^{\mathsf{T}} Q_{k}^{-1} b_{k}, T_{k} \right\rangle + \left\langle b_{k}^{\mathsf{T}} Q_{k}^{-1}, \mathbf{t}_{k,1} \right\rangle + \left\langle -\frac{1}{2} Q_{k}^{-1}, \mathbf{t}_{k,2} \right\rangle + \mathbf{c} \end{split}$$

where

$$T_k = \sum_{t=1}^{T} q(z_t = k)$$
  $\mathbf{t}_{k,1} = \sum_{t=1}^{T} q(z_t = k) x_t$   $\mathbf{t}_{k,2} = \sum_{t=1}^{T} q(z_t = k) x_t x_t^{\top}$ 

are the weighted sums of sufficient statistics.

#### Solving for the optimal Gaussian parameters

The objective we're trying to maximize is,

$$\mathcal{L}(q,\theta) = \left\langle -\frac{1}{2} \log |Q_k| - \frac{1}{2} b_k^{\mathsf{T}} Q_k^{-1} b_k, T_k \right\rangle + \left\langle b_k^{\mathsf{T}} Q_k^{-1}, \mathbf{t}_{k,1} \right\rangle + \left\langle -\frac{1}{2} Q_k^{-1}, \mathbf{t}_{k,2} \right\rangle + c$$

Taking the partial derivative wrt  $b_k$  and setting equal to zero,

$$\frac{\partial}{\partial b_k} \mathcal{L}(q, \theta) = Q_k^{-1} \mathbf{t}_{k,1} - Q_k^{-1} b_k T_k = 0$$

$$\Longrightarrow b_k^{\star} = \frac{\mathbf{t}_{k,1}}{T_k} = \frac{1}{T_k} \sum_{t=1}^T q(z_t = k) x_t$$

#### Solving for the optimal Gaussian parameters

Plug in the optimum

$$\mathcal{L}(q,\theta) = \left\langle -\frac{1}{2} \log |Q_k| - \frac{1}{2} \frac{\mathbf{t}_{k,1}^{\top}}{T_k} Q_k^{-1} \frac{\mathbf{t}_{k,1}}{T_k}, T_k \right\rangle + \left\langle -\frac{1}{2} Q_k^{-1}, \mathbf{t}_{k,2} \right\rangle + \left\langle \frac{\mathbf{t}_{k,1}^{\top}}{T_k} Q_k^{-1}, \mathbf{t}_{k,1} \right\rangle + c$$

$$= \left\langle -\frac{1}{2} \log |Q_k|, T_k \right\rangle + \left\langle -\frac{1}{2} Q_k^{-1}, \mathbf{t}_{k,2} - \frac{\mathbf{t}_{k,1} \mathbf{t}_{k,1}^{\top}}{T_k} \right\rangle + c$$

#### Solving for the optimal Gaussian parameters

Let 
$$\Lambda_k = Q_k^{-1}$$
,

$$\mathcal{L}(q,\theta) = \left\langle \frac{1}{2} \log |\Lambda_k|, T_k \right\rangle + \left\langle -\frac{1}{2} \Lambda_k, \mathbf{t}_{k,2} - \frac{\mathbf{t}_{k,1} \mathbf{t}_{k,1}^{\top}}{T_k} \right\rangle + c$$

Taking the partial derivative wrt  $\Lambda_k$  and setting equal to zero,

$$\frac{\partial}{\partial \Lambda_k} \mathcal{L}(q, \theta) = \frac{T_k}{2} \Lambda_k^{-1} - \frac{1}{2} \left( \mathbf{t}_{k,2} - \frac{\mathbf{t}_{k,2} \mathbf{t}_{k,1}^{\mathsf{T}}}{T_k} \right) = 0$$

$$\implies (\Lambda_k^{-1})^* = Q_k^* = \frac{1}{T_k} \left( \mathbf{t}_{k,2} - \frac{\mathbf{t}_{k,1} \mathbf{t}_{k,1}^\mathsf{T}}{T_k} \right)$$

#### In summary...

• **E-step**: Compute the posterior probabilities:

$$q(z_t = k) \leftarrow p(z_t = k \mid x_t, \Theta)$$
 via the forward-backward algorithm.

Compute weighted sums of sufficient statistics:

$$T_k = \sum_{t=1}^{T} q(z_t = k)$$
  $\mathbf{t}_{k,1} = \sum_{t=1}^{T} q(z_t = k) x_t$   $\mathbf{t}_{k,2} = \sum_{t=1}^{T} q(z_t = k) x_t x_t^{\top}$ 

M-step: Update the parameters.

$$b_k \leftarrow \frac{\mathbf{t}_{k,1}}{T_k} \qquad Q_k \leftarrow \frac{1}{T_k} \left( \mathbf{t}_{k,2} - \frac{\mathbf{t}_{k,1} \mathbf{t}_{k,1}^{\mathsf{T}}}{T_k} \right)$$

ullet Note: The updates are equivalent if we use normalized sufficient statistics, each divided by T.

#### Stochastic EM for the Gaussian mixture model

- On iteration i, grab a sub-sequence (aka **mini-batch**) of length M.
- **E-step**: Compute the posterior probabilities for each data point in the mini-batch:

$$q(z_{m} = k) \leftarrow p(z_{m} = k \mid x_{m}, \Theta) \propto \frac{\pi_{k} \mathcal{N}(x_{m} \mid b_{k}, Q_{k})}{\sum_{j=1}^{K} \pi_{j} \mathcal{N}(x_{m} \mid b_{j}, Q_{j})}$$

Compute **normalized sufficient statistics** for the mini-batch:

$$\bar{T}_{k}^{(i)} = \frac{1}{M} \sum_{m=1}^{M} q(z_{m} = k) \qquad \bar{\mathbf{t}}_{k,1}^{(i)} = \frac{1}{M} \sum_{m=1}^{M} q(z_{m} = k) x_{m} \qquad \bar{\mathbf{t}}_{k,2}^{(i)} = \frac{1}{M} \sum_{m=1}^{M} q(z_{m} = k) x_{m} x_{m}^{\top}$$

Fold the normalized stats from this mini-batch into the running average via a convex combination with step size  $\alpha \in [0,1]$ :

$$\bar{T}_k \leftarrow (1 - \alpha)\bar{T}_k + \alpha\bar{T}_k^{(i)} \qquad \bar{\mathbf{t}}_{k,1} \leftarrow (1 - \alpha)\bar{\mathbf{t}}_{k,1} + \alpha\bar{\mathbf{t}}_{k,1}^{(i)} \qquad \bar{\mathbf{t}}_{k,2} \leftarrow (1 - \alpha)\bar{\mathbf{t}}_{k,2} + \alpha\bar{\mathbf{t}}_{k,2}^{(i)}$$

• M-step: Update the parameters.

$$b_k \leftarrow \frac{\overline{\mathbf{t}}_{k,1}}{\overline{T}_k} \qquad Q_k \leftarrow \frac{1}{\overline{T}_k} \left( \overline{\mathbf{t}}_{k,2} - \frac{\overline{\mathbf{t}}_{k,1} \overline{\mathbf{t}}_{k,1}^{\mathsf{T}}}{\overline{T}_k} \right)$$

#### Conclusion

- Hidden Markov models (HMMs) are just mixture models with dependencies across time.
- The EM algorithm is nearly the same as for mixture models, but we use the forward-backward algorithm to compute posterior marginal probabilities.
- With exponential family likelihoods, the M-step only needs weighted sums of sufficient statistics.
- Stochastic EM generalizes the EM algorithm to work with mini-batches of data and rolling averages of the sufficient statistics. It can be seen as SGD with natural gradients.