

Causal Inference TA Section (3): IV

shengqiaolin@utexas.edu (mailto:shengqiaolin@utexas.edu)

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Today's Goals

- Tutorial: IV Regression
- Replication #1: Hager and Krakowski APSR 2022
 - panel data as main specification
 - IV for causal inference
- Replication #2: López-Cariboni CPS 2022 (panel data with IV)

Tutorial: IV Regression

Acknowledgement: This example is adopted from Dr. Stephen Jessee (<https://laits.utexas.edu/~sjessee/contact.html>)'s lecture on causal inference.

Simulated Data

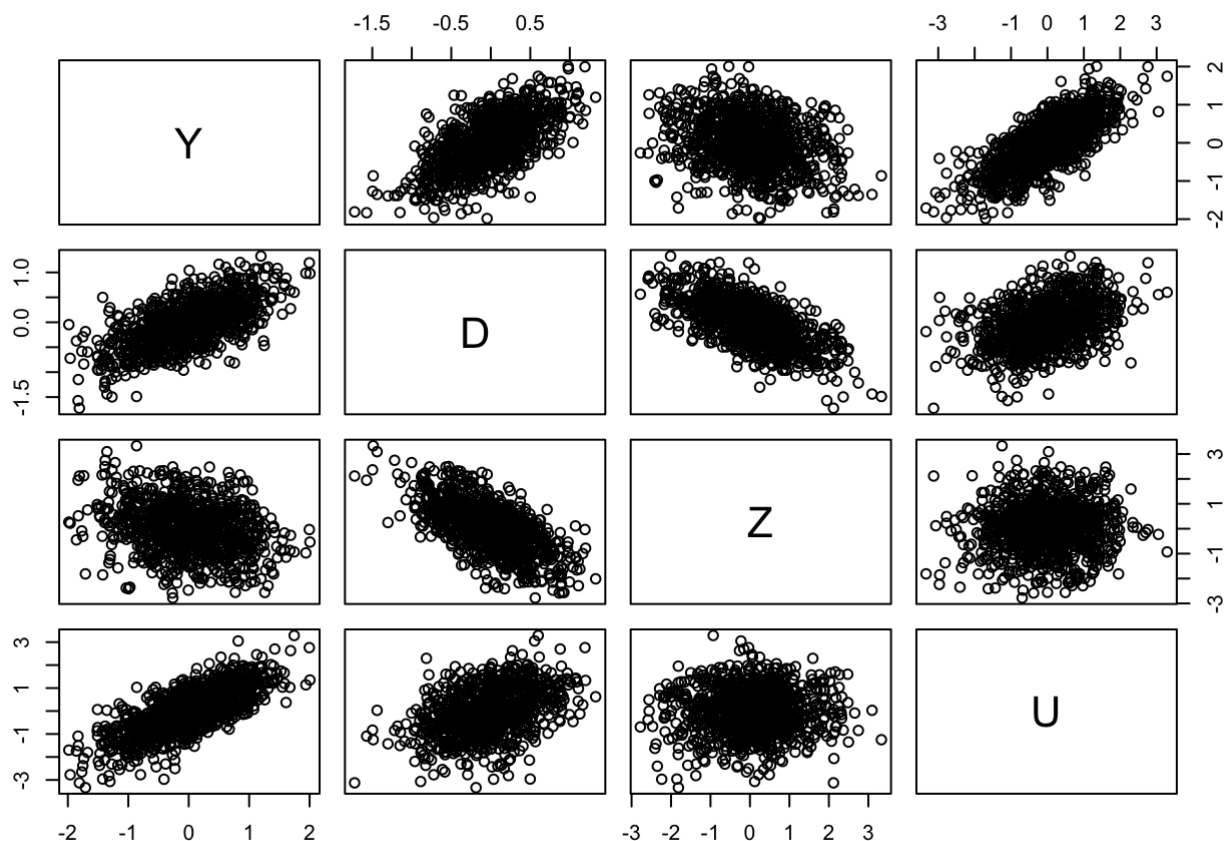
Let's start with generating some simulated data where Y is the outcome of interest, D is the treatment, Z is the instrumental variable, and U is a set of unobserved confounders (sometimes observable).

```
set.seed(12345)
# setting sample size
N <- 1000

Z <- rnorm(N) # instrumental variable
U <- rnorm(N) # unobserved confounders

D <- .2*U - .3*Z + rnorm(N, sd=.3) # (confounded) treatment variable
Y <- .4*U + .6*D + rnorm(N, sd=.4) # dependent variable

pairs(cbind(Y, D, Z, U))
```



Let's do a naive estimation first. The estimated effect is nearly 1 (when real effect is .6).

```
reg.Y.D <- lm(Y ~ D)
summary(reg.Y.D)
```

```
##
## Call:
## lm(formula = Y ~ D)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.94296 -0.35151 -0.01812  0.33887  1.69304
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.004351   0.016809   0.259   0.796
## D            0.962743   0.036676  26.250 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5306 on 998 degrees of freedom
## Multiple R-squared:  0.4084, Adjusted R-squared:  0.4078
## F-statistic: 689 on 1 and 998 DF, p-value: < 2.2e-16
```

Ideally, to get the real effect, we want to control U. But it is unrealistic because U is unobserved/unobservable.

```
reg.Y.DU <- lm(Y ~ D + U)
summary(reg.Y.DU)
```

```
##
## Call:
## lm(formula = Y ~ D + U)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.34994 -0.25046 -0.01267  0.25335  1.02902
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.006515   0.012083   0.539    0.59
## D            0.599115   0.028922  20.715 <2e-16 ***
## U            0.401353   0.013128  30.572 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.3814 on 997 degrees of freedom
## Multiple R-squared:  0.6947, Adjusted R-squared:  0.6941
## F-statistic: 1134 on 2 and 997 DF,  p-value: < 2.2e-16
```

Constant Effects Models

Assuming we have an instrumental variable Z which affects Y through D and is independent from U . We start with estimating the effect of Z on Y .

```
reg.Y.Z <- lm(Y ~ Z)
summary(reg.Y.Z)
```

```
##
## Call:
## lm(formula = Y ~ Z)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.0067 -0.4642  0.0171  0.4674  2.0041
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.01427   0.02115  -0.675    0.5
## Z           -0.17199   0.02116  -8.126 1.3e-15 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6681 on 998 degrees of freedom
## Multiple R-squared:  0.06206, Adjusted R-squared:  0.06112
## F-statistic: 66.03 on 1 and 998 DF,  p-value: 1.302e-15
```

The effect of Z on D is -.3 and the effect of D on Y is .6. So the real effect of Z on Y is $-.3 \times .6 = -.18$, which is very closed to the estimated effect.

We are interested in the real effect of D on Y.

```
reg.Y.Z <- lm(Y ~ Z) #Reduced form
reg.D.Z <- lm(D ~ Z) #First stage
stargazer::stargazer(reg.D.Z, reg.Y.Z, type='text', no.space = TRUE)
```

```
##
## =====
##                               Dependent variable:
##                               -----
##                               D           Y
##                               (1)        (2)
## -----
## Z                               -0.296***   -0.172***
##                               (0.011)      (0.021)
## Constant                       -0.014      -0.014
##                               (0.011)      (0.021)
## -----
## Observations                   1,000        1,000
## R2                             0.416        0.062
## Adjusted R2                   0.416        0.061
## Residual Std. Error (df = 998) 0.350        0.668
## F Statistic (df = 1; 998)      711.528***   66.035***
## =====
## Note:                          *p<0.1; **p<0.05; ***p<0.01
```

Then we can calculate the real effect of D on Y:

```
coef(reg.Y.Z)[2] / coef(reg.D.Z)[2]
```

```
##      Z
## 0.5817
```

Two-stage Least Squares (2SLS)

2SLS will be more practical when you have more than 1 IV.

```
reg.D.Z <- lm(D ~ Z)
reg.Y.D_fitted <- lm(Y ~ reg.D.Z$fitted.values)
summary(reg.Y.D_fitted)
```

```
##
## Call:
## lm(formula = Y ~ reg.D.Z$fitted.values)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.0067 -0.4642  0.0171  0.4674  2.0041
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   -0.006164   0.021220  -0.290    0.772
## reg.D.Z$fitted.values  0.581700   0.071584   8.126 1.3e-15 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6681 on 998 degrees of freedom
## Multiple R-squared:  0.06206,    Adjusted R-squared:  0.06112
## F-statistic: 66.03 on 1 and 998 DF,  p-value: 1.302e-15
```

IV Regression

We can also do it in one step:

```
library(AER) # loading Applied Econometrics with R package
reg.iv <- ivreg(Y ~ D | Z)
#reg.iv<- ivreg(Y ~ D + C| C + Z) if you have control variables C
summary(reg.iv)
```

```
##
## Call:
## ivreg(formula = Y ~ D | Z)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.950362 -0.387842 -0.000555  0.382135  1.532142
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.006164   0.017740  -0.347    0.728
## D           0.581700   0.059845   9.720 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5586 on 998 degrees of freedom
## Multiple R-Squared:  0.3445,    Adjusted R-squared:  0.3438
## Wald test: 94.48 on 1 and 998 DF,  p-value: < 2.2e-16
```

Control IV to test Exclusion Restriction? No

Will $Y \sim D + Z$ help to test the exclusion restriction assumption? NO

```
reg.Y.DZ <- lm(Y ~ D + Z)
summary(reg.Y.DZ)
```

```
##
## Call:
## lm(formula = Y ~ D + Z)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.97997 -0.34316 -0.02367  0.34395  1.78548
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.002932   0.016157   0.181    0.856
## D            1.234409   0.046137  26.755 <2e-16 ***
## Z            0.192983   0.021144   9.127 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.51 on 997 degrees of freedom
## Multiple R-squared:  0.4541, Adjusted R-squared:  0.453
## F-statistic: 414.6 on 2 and 997 DF,  p-value: < 2.2e-16
```

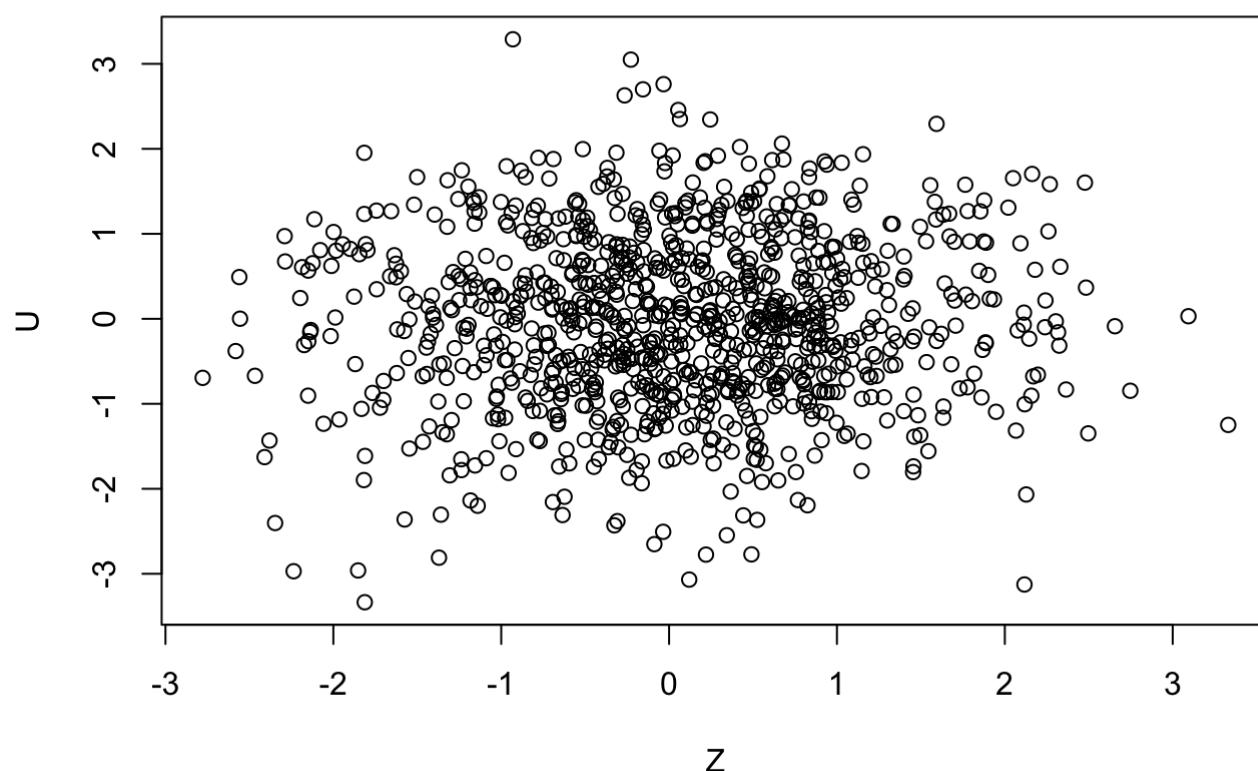
the coefficient on Z is significant even controlling for D! But we created these data such that Z only affects Y through D.

Let's see what happened.

```
reg.U.Z <- lm(U ~ Z)
summary(reg.U.Z)
```

```
##
## Call:
## lm(formula = U ~ Z)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.2315 -0.6975  0.0011  0.7156  3.3577
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.03220   0.03191  -1.009    0.313
## Z           0.03915   0.03193   1.226    0.221
##
## Residual standard error: 1.008 on 998 degrees of freedom
## Multiple R-squared:  0.001504, Adjusted R-squared:  0.000503
## F-statistic: 1.503 on 1 and 998 DF,  p-value: 0.2205
```

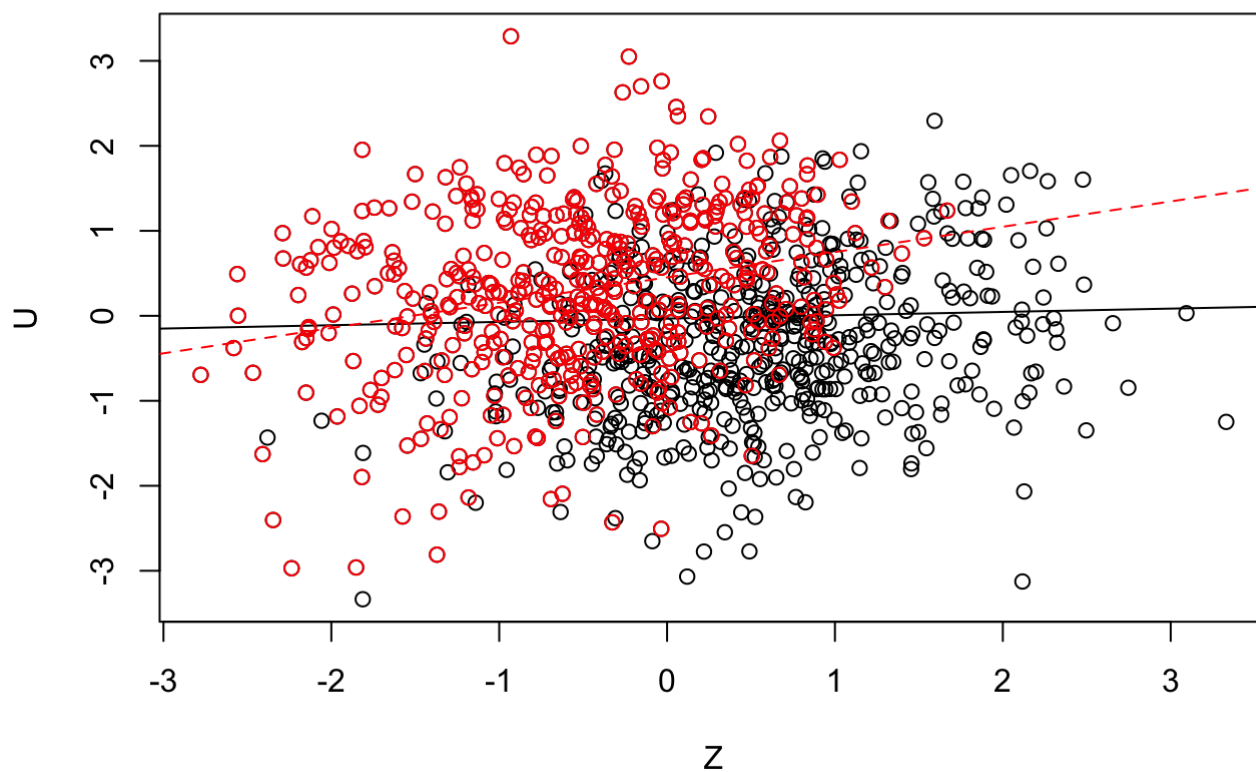
```
plot(Z, U)
```



```
plot(Z, U)
abline(reg.U.Z)
points(Z[D>0], U[D>0], col="red")
reg.U.Z.posD <- lm(U[D>0] ~ Z[D>0])
summary(reg.U.Z.posD)
```

```
##
## Call:
## lm(formula = U[D > 0] ~ Z[D > 0])
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.94792 -0.66253  0.03713  0.69522  3.11393
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   0.45232    0.05069   8.923  < 2e-16 ***
## Z[D > 0]       0.29789    0.05200   5.728  1.8e-08 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9554 on 476 degrees of freedom
## Multiple R-squared:  0.06449,    Adjusted R-squared:  0.06253
## F-statistic: 32.81 on 1 and 476 DF,  p-value: 1.8e-08
```

```
abline(reg.U.Z.posD, col="red", lty=2)
```

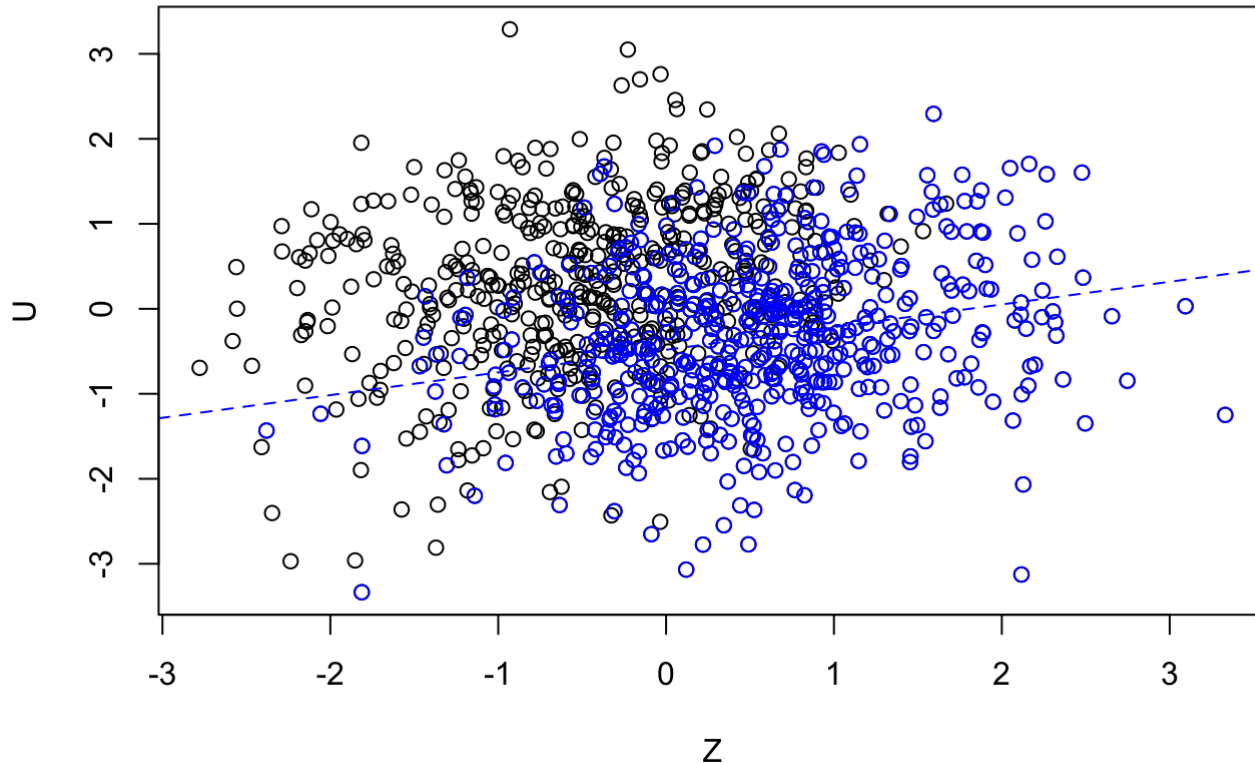


```
plot(Z, U)
points(Z[D<0], U[D<0], col="blue")
reg.U.Z.negD <- lm(U[D<0] ~ Z[D<0])
summary(reg.U.Z.negD)
```



```
##
## Call:
## lm(formula = U[D < 0] ~ Z[D < 0])
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.2089 -0.5604 -0.0299  0.5862  2.3508
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.48192    0.04640  -10.387  < 2e-16 ***
## Z[D < 0]      0.26672    0.04543   5.871 7.73e-09 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.8991 on 520 degrees of freedom
## Multiple R-squared:  0.06216,    Adjusted R-squared:  0.06036
## F-statistic: 34.47 on 1 and 520 DF,  p-value: 7.734e-09
```

```
abline(reg.U.Z.negD, col="blue", lty=2)
```



Conditioning on D (in this case high/low values) creates a non-causal association between Z and U it thus also creates a non-causal association b/t Z and Y, which operates through U. So this specification is just misleading.

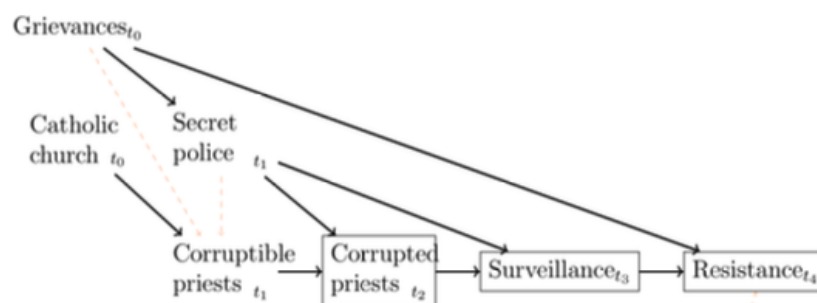
Replication #1: Hager and Krakowski APSR 2022

Hager and Krakowski APSR 2022, Does State Repression Spark Protests? Evidence from Secret Police Surveillance in Communist Poland (10.1017/S0003055421000770)

- Research Question: whether physical surveillance hinder or foster antiregime resistance?
- Data: Communist Poland 1945 - 1989
- Identification: two-way fixed effect panel models + IV
- Result: Secret police -> more protests, less sabotage

Anselm Hager and Krzysztof Krakowski

FIGURE 3. Instrumental Variable Model



Note: Directed acyclic graph of the IV model. Solid black lines refer to causal relationships between variables at the indicated points in time. Dotted orange arrows indicate causal relations assumed to be absent.

- Instrumental variable: corrupted priests
 - Assumptions: causal relationships of these dotted arrows are absent
 - Qualitative and quantitative evidence

Main Finding: Two-ways FE

```
library(readr)
library(plm)
data_t <- read_csv("data_t.csv")
m_2wfe <- plm(strikes ~ commanders, data = data_t, index = c("year", "locality"), model = "within", effect = "twoways")
summary(m_2wfe)
```

```
## Twoways effects Within Model
##
## Call:
## plm(formula = strikes ~ commanders, data = data_t, effect = "twoways",
##      model = "within", index = c("year", "locality"))
##
## Balanced Panel: n = 7, T = 297, N = 2079
##
## Residuals:
##      Min.      1st Qu.      Median      3rd Qu.      Max.
## -5.455487 -0.251754  0.026979  0.139481 17.436081
##
## Coefficients:
##              Estimate Std. Error t-value Pr(>|t|)
## commanders 0.162830    0.011818  13.778 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Total Sum of Squares:    1374.2
## Residual Sum of Squares: 1241.4
## R-Squared:    0.096621
## Adj. R-Squared: -0.057589
## F-statistic: 189.846 on 1 and 1775 DF, p-value: < 2.22e-16
```

IV Without Controls

```
data_agg_late <- readr::read_csv("data_agg_late.csv")
m_iv = ivreg(scale(strikes) ~ commanders | priests_continuous, data = data_agg_late)
summary(m_iv)
```

```
##
## Call:
## ivreg(formula = scale(strikes) ~ commanders | priests_continuous,
##      data = data_agg_late)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -7.0902 -0.2032 -0.2032 -0.2032  5.3956
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.07982    0.04638  -1.721  0.0863 .
## commanders  0.17653    0.02062   8.560 6.33e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.7829 on 295 degrees of freedom
## Multiple R-Squared: 0.3891, Adjusted R-squared: 0.387
## Wald test: 73.27 on 1 and 295 DF, p-value: 6.33e-16
```

Replication #2: López-Cariboni CPS 2022

López-Cariboni 2022 Political Regimes and Informal Social Insurance

(<https://journals.sagepub.com/doi/10.1177/00104140221139378>).

- Research Question: What is the political motivation for deliberate nonenforcement of the law?
- Argument: Democracies * Negative Economic Shock -> Electricity loss
- Data: 110 developing countries 1970-2014
- IV: regional democratic diffusion * Negative Economic Shock

```
library(readr)
dt<- read.csv('dt_replication.csv')
iv_dem <- ivreg(outgap.tdl ~ l.outgap.tdl
  + l.outgap.gdp.hamilton * l.democracy#Estimator
  + as.factor(iso3c)
  + as.factor(year)
  | .
  - l.outgap.gdp.hamilton*l.democracy#Estimator
  + l.outgap.gdp.hamilton*l.wreg.democracy#IV
  ,
  data=dt, na.action=na.omit)
summary(iv_dem)
```

```
##
## Call:
## ivreg(formula = outgap.tdl ~ l.outgap.tdl + l.outgap.gdp.hamilton *
##       l.democracy + as.factor(iso3c) + as.factor(year) | . - l.outgap.gdp.hamilton *
##       l.democracy + l.outgap.gdp.hamilton * l.wreg.democracy, data = dt,
##       na.action = na.omit)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.793263 -0.111132 -0.008912  0.084379  3.961454
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    -0.0963814   0.0755814   -1.275   0.20233
## l.outgap.tdl     0.6423725   0.0142227  45.165 < 2e-16 ***
## l.outgap.gdp.hamilton  0.0022054   0.0068398   0.322   0.74714
## l.democracy     0.0361513   0.1020100   0.354   0.72307
## as.factor(iso3c)ALB  0.1134801   0.1086984   1.044   0.29657
## as.factor(iso3c)ARE  0.0256484   0.0774112   0.331   0.74042
## as.factor(iso3c)ARG  0.0063618   0.1169037   0.054   0.95660
## as.factor(iso3c)ARM -0.0521678   0.0904658  -0.577   0.56421
## as.factor(iso3c)AZE  0.0271074   0.0904714   0.300   0.76448
## as.factor(iso3c)BGD  0.0228123   0.0975748   0.234   0.81516
## as.factor(iso3c)BGR -0.0158363   0.1169311  -0.135   0.89228
## as.factor(iso3c)BHR  0.0758740   0.0803829   0.944   0.34529
## as.factor(iso3c)BIH -0.0649655   0.1390382  -0.467   0.64035
## as.factor(iso3c)BLR -0.0006109   0.0904769  -0.007   0.99461
## as.factor(iso3c)BOL -0.0226212   0.1213126  -0.186   0.85209
## as.factor(iso3c)BRA  0.0100936   0.1127514   0.090   0.92867
## as.factor(iso3c)BRN  0.0260661   0.0802893   0.325   0.74547
## as.factor(iso3c)BWA -0.0284961   0.1321771  -0.216   0.82932
## as.factor(iso3c)CHL  0.0250005   0.1030990   0.242   0.80842
## as.factor(iso3c)CHN  0.0266289   0.0757279   0.352   0.72513
## as.factor(iso3c)CIV  0.0334140   0.0758785   0.440   0.65971
## as.factor(iso3c)CMR  0.0366175   0.0757152   0.484   0.62869
## as.factor(iso3c)COD  0.0645921   0.0758840   0.851   0.39473
## as.factor(iso3c)COG  0.0942302   0.0835898   1.127   0.25971
## as.factor(iso3c)COL -0.0158241   0.1319608  -0.120   0.90456
## as.factor(iso3c)CRI  0.0065985   0.1322744   0.050   0.96022
## as.factor(iso3c)CUB  0.0337645   0.0757317   0.446   0.65574
## as.factor(iso3c)CYP -0.0178146   0.1298952  -0.137   0.89092
## as.factor(iso3c)CZE -0.0641254   0.1335383  -0.480   0.63112
## as.factor(iso3c)DOM -0.0303150   0.1324636  -0.229   0.81900
## as.factor(iso3c)DZA  0.0310692   0.0757261   0.410   0.68163
## as.factor(iso3c)ECU -0.0010899   0.1189519  -0.009   0.99269
## as.factor(iso3c)EGY  0.0270464   0.0757228   0.357   0.72098
## as.factor(iso3c)ERI  0.0318902   0.0975178   0.327   0.74368
## as.factor(iso3c)EST -0.0900163   0.1369359  -0.657   0.51100
## as.factor(iso3c)ETH  0.0585880   0.0810231   0.723   0.46967
## as.factor(iso3c)GAB  0.0305084   0.0757263   0.403   0.68707
## as.factor(iso3c)GEO -0.0180608   0.1032328  -0.175   0.86113
## as.factor(iso3c)GHA  0.0720550   0.0948404   0.760   0.44746
```

## as.factor(iso3c)GTM	0.0010470	0.1231700	0.009	0.99322
## as.factor(iso3c)HND	0.0043575	0.1174728	0.037	0.97041
## as.factor(iso3c)HRV	-0.0484238	0.1396635	-0.347	0.72883
## as.factor(iso3c)HTI	0.0290805	0.0757351	0.384	0.70102
## as.factor(iso3c)HUN	-0.0397930	0.1346905	-0.295	0.76768
## as.factor(iso3c>IDN	0.0246921	0.0887359	0.278	0.78083
## as.factor(iso3c)IND	0.0083989	0.1299221	0.065	0.94846
## as.factor(iso3c)IRN	0.0311597	0.0759282	0.410	0.68155
## as.factor(iso3c)IRQ	0.0313712	0.0757278	0.414	0.67871
## as.factor(iso3c)ISR	-0.0190116	0.1321169	-0.144	0.88559
## as.factor(iso3c)JAM	-0.0094266	0.1323173	-0.071	0.94321
## as.factor(iso3c)JOR	0.0034274	0.0779095	0.044	0.96491
## as.factor(iso3c)KAZ	-0.0093361	0.0904686	-0.103	0.91781
## as.factor(iso3c)KEN	0.0266775	0.0848121	0.315	0.75313
## as.factor(iso3c)KGZ	0.0192991	0.0904736	0.213	0.83110
## as.factor(iso3c)KHM	0.0137153	0.0998556	0.137	0.89076
## as.factor(iso3c)KOR	-0.0209094	0.1067859	-0.196	0.84477
## as.factor(iso3c)KWT	0.0330063	0.0999465	0.330	0.74124
## as.factor(iso3c)LBN	0.0529039	0.0878224	0.602	0.54695
## as.factor(iso3c)LBY	0.2307820	0.1131075	2.040	0.04140 *
## as.factor(iso3c)LKA	0.0122397	0.0961313	0.127	0.89869
## as.factor(iso3c)LTU	-0.0167540	0.1396229	-0.120	0.90450
## as.factor(iso3c)LVA	-0.0748631	0.1400601	-0.535	0.59303
## as.factor(iso3c)MAR	0.0369274	0.0757190	0.488	0.62580
## as.factor(iso3c)MDA	-0.0246903	0.1394300	-0.177	0.85946
## as.factor(iso3c)MEX	0.0193335	0.0872039	0.222	0.82456
## as.factor(iso3c)MKD	-0.0112741	0.1334112	-0.085	0.93266
## as.factor(iso3c)MLT	-0.0008635	0.1321714	-0.007	0.99479
## as.factor(iso3c)MMR	0.0303363	0.0757495	0.400	0.68883
## as.factor(iso3c)MNE	0.0161723	0.1907842	0.085	0.93245
## as.factor(iso3c)MNG	-0.0047879	0.1305046	-0.037	0.97074
## as.factor(iso3c)MOZ	-0.0359686	0.0876690	-0.410	0.68163
## as.factor(iso3c)MUS	-0.0324916	0.1304515	-0.249	0.80332
## as.factor(iso3c)MYS	0.0227709	0.0757211	0.301	0.76365
## as.factor(iso3c)NAM	0.0367590	0.0920690	0.399	0.68973
## as.factor(iso3c)NER	0.0360422	0.1382634	0.261	0.79436
## as.factor(iso3c)NGA	0.0193410	0.0774048	0.250	0.80271
## as.factor(iso3c)NIC	-0.0240256	0.1162458	-0.207	0.83627
## as.factor(iso3c)NPL	0.0146632	0.0913608	0.160	0.87250
## as.factor(iso3c)OMN	0.0214063	0.0757140	0.283	0.77741
## as.factor(iso3c)PAK	0.0008132	0.0932093	0.009	0.99304
## as.factor(iso3c)PAN	0.0157893	0.1015391	0.155	0.87644
## as.factor(iso3c)PER	0.0127395	0.1014688	0.126	0.90010
## as.factor(iso3c)PHL	0.0019961	0.1112535	0.018	0.98569
## as.factor(iso3c)POL	-0.0589610	0.1334053	-0.442	0.65854
## as.factor(iso3c)PRY	0.0014858	0.0834637	0.018	0.98580
## as.factor(iso3c)QAT	-0.0152580	0.1171755	-0.130	0.89641
## as.factor(iso3c)ROU	-0.0116447	0.1334438	-0.087	0.93047
## as.factor(iso3c)RUS	-0.0038277	0.0925213	-0.041	0.96700
## as.factor(iso3c)SAU	0.0047119	0.0757561	0.062	0.95041
## as.factor(iso3c)SDN	0.0225678	0.0770094	0.293	0.76950
## as.factor(iso3c)SEN	0.0292358	0.0872501	0.335	0.73759

## as.factor(iso3c)SGP	0.0142157	0.0757210	0.188	0.85109
## as.factor(iso3c)SLV	0.0074315	0.1149239	0.065	0.94845
## as.factor(iso3c)SUR	0.0575049	0.1515375	0.379	0.70436
## as.factor(iso3c)SVK	-0.0298885	0.1352356	-0.221	0.82510
## as.factor(iso3c)SVN	-0.0235229	0.1334988	-0.176	0.86015
## as.factor(iso3c)TGO	0.0418883	0.0784486	0.534	0.59341
## as.factor(iso3c)THA	0.0095160	0.1049851	0.091	0.92778
## as.factor(iso3c)TJK	0.0349217	0.0905161	0.386	0.69967
## as.factor(iso3c)TKM	0.0158482	0.0905251	0.175	0.86104
## as.factor(iso3c)TTO	0.0351470	0.1314149	0.267	0.78914
## as.factor(iso3c)TUN	0.0310974	0.0757267	0.411	0.68135
## as.factor(iso3c)TUR	0.0055834	0.1250410	0.045	0.96439
## as.factor(iso3c)TZA	-0.0051167	0.0877885	-0.058	0.95353
## as.factor(iso3c)UKR	-0.0234375	0.1340214	-0.175	0.86119
## as.factor(iso3c)URY	0.0064197	0.1126347	0.057	0.95455
## as.factor(iso3c)UZB	0.0181445	0.0905778	0.200	0.84124
## as.factor(iso3c)VEN	-0.0169432	0.1137131	-0.149	0.88156
## as.factor(iso3c)VNM	-0.0034668	0.0835089	-0.042	0.96689
## as.factor(iso3c)YEM	0.0101083	0.0908929	0.111	0.91146
## as.factor(iso3c)ZAF	0.0163869	0.0962576	0.170	0.86483
## as.factor(iso3c)ZMB	0.0272836	0.0791966	0.345	0.73049
## as.factor(iso3c)ZWE	0.0385686	0.0757188	0.509	0.61053
## as.factor(year)1978	0.0477726	0.0564564	0.846	0.39751
## as.factor(year)1979	0.0384090	0.0565107	0.680	0.49676
## as.factor(year)1980	0.0262766	0.0564452	0.466	0.64159
## as.factor(year)1981	0.1114896	0.0561780	1.985	0.04728 *
## as.factor(year)1982	0.0521906	0.0556874	0.937	0.34873
## as.factor(year)1983	0.1124997	0.0558479	2.014	0.04405 *
## as.factor(year)1984	-0.0073257	0.0561587	-0.130	0.89622
## as.factor(year)1985	0.0390559	0.0566671	0.689	0.49074
## as.factor(year)1986	-0.0021277	0.0561284	-0.038	0.96976
## as.factor(year)1987	0.0333909	0.0569295	0.587	0.55756
## as.factor(year)1988	0.1060890	0.0566294	1.873	0.06111 .
## as.factor(year)1989	0.1647638	0.0574957	2.866	0.00419 **
## as.factor(year)1990	0.0882197	0.0569067	1.550	0.12118
## as.factor(year)1991	0.1050157	0.0572741	1.834	0.06681 .
## as.factor(year)1992	0.0332995	0.0590216	0.564	0.57267
## as.factor(year)1993	0.0073011	0.0602568	0.121	0.90357
## as.factor(year)1994	0.1031109	0.0601549	1.714	0.08661 .
## as.factor(year)1995	0.0390873	0.0613603	0.637	0.52416
## as.factor(year)1996	0.0934199	0.0605137	1.544	0.12274
## as.factor(year)1997	0.0689713	0.0588758	1.171	0.24150
## as.factor(year)1998	0.0538141	0.0597448	0.901	0.36780
## as.factor(year)1999	0.0628991	0.0596550	1.054	0.29179
## as.factor(year)2000	0.1288616	0.0592191	2.176	0.02963 *
## as.factor(year)2001	0.1081492	0.0595244	1.817	0.06933 .
## as.factor(year)2002	0.0666299	0.0602498	1.106	0.26886
## as.factor(year)2003	0.0608043	0.0601372	1.011	0.31205
## as.factor(year)2004	0.0467735	0.0611736	0.765	0.44457
## as.factor(year)2005	0.0730868	0.0610259	1.198	0.23115
## as.factor(year)2006	0.0690148	0.0605076	1.141	0.25413
## as.factor(year)2007	0.1190551	0.0601148	1.980	0.04774 *

```
## as.factor(year)2008      0.1324141  0.0597912  2.215  0.02686 *
## as.factor(year)2009      0.0777981  0.0614147  1.267  0.20533
## as.factor(year)2010      0.0723541  0.0603809  1.198  0.23089
## as.factor(year)2011      0.0608862  0.0604628  1.007  0.31401
## as.factor(year)2012      0.0250867  0.0618298  0.406  0.68496
## as.factor(year)2013      0.0596776  0.0617932  0.966  0.33424
## as.factor(year)2014      0.0927240  0.0617555  1.501  0.13334
## as.factor(year)2015     -0.0312008  0.3234857 -0.096  0.92317
## as.factor(year)2016     -0.0218120  0.3235011 -0.067  0.94625
## l.outgap.gdp.hamilton:l.democracy -0.0607667  0.0256632 -2.368  0.01795 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.3057 on 3096 degrees of freedom
## Multiple R-Squared:  0.4455, Adjusted R-squared:  0.4183
## Wald test: 16.41 on 152 and 3096 DF, p-value: < 2.2e-16
```

Another (easier) way to run a 2WFE with IV

```
library(fixest)#Fastest Fixed Effects
iv_dem<- feols(outgap.tdl ~ l.outgap.tdl|iso3c+year|l.outgap.gdp.hamilton * l.democracy~
l.outgap.gdp.hamilton * l.wreg.democracy,se='iid',data=dt)
etable(iv_dem)
```

```
##                               iv_dem
## Dependent Var.:                outgap.tdl
##
## l.outgap.gdp.hamilton          0.0022 (0.0068)
## l.democracy                    0.0362 (0.1020)
## l.outgap.gdp.hamilton:l.democracy -0.0608* (0.0257)
## l.outgap.tdl                   0.6424*** (0.0142)
## Fixed-Effects:                -----
## iso3c                          Yes
## year                           Yes
##
## S.E. type                      IID
## Observations                   3,249
## R2                             0.44551
## Within R2                      0.42776
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```