

Assignment - 4

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2025122010

(MA6102) Probability and Random Processes

problem - 1b

Given $f_X(x) = \begin{cases} Pe^{-x} & \text{if } x > 0 \\ (1-P)x^m e^{-x} & \text{if } x \leq 0 \end{cases}$

$$E[X] = \int_0^\infty x Pe^{-x} dx + \int_{-\infty}^0 x (1-P)x^m e^{-x} dx$$

$$= P \cdot \int_0^\infty x^2 e^{-x} dx + (1-P) \int_{-\infty}^0 x^2 \cdot x^m e^{-x} dx$$

$$= P \cdot \frac{1}{2} + (1-P) \int_{-\infty}^0 x^2 \cdot x^m e^{-x} dx$$

$$= P \cdot \frac{1}{2} - (1-P) \int_0^\infty y^2 e^{-y} dy$$

$$= \left[\frac{P}{2} - \frac{(1-P)}{2} \right] = \frac{2P-1}{2}$$

$$\therefore E[X] = \frac{2P-1}{2}$$

$$E[X^2] = \int_0^\infty x^2 Pe^{-x} dx + \int_{-\infty}^0 x^2 (1-P)x^m e^{-x} dx$$

$$= P \int_0^\infty x^2 e^{-x} dx + (1-P) \int_{-\infty}^0 x^2 x^m e^{-x} dx$$

$$= P \cdot \frac{2}{2} + (1-P) \int_{-\infty}^0 y^2 x^m e^{-x} dx$$

$$= P \left(\frac{2}{2} \right) + (1-P) \frac{2}{2} = \frac{2P+2-2P}{2} = \frac{2}{2} = \frac{1}{2}$$

$$\Rightarrow E[X^2] = \frac{2}{\lambda^2}$$

$$\Rightarrow \text{Var}(X) = E[X^2] - E^2[X]$$

$$= \frac{2}{\lambda^2} - \left(\frac{2p-1}{\lambda} \right)^2$$

$$\begin{aligned} &= \frac{2 - (2p-1)^2}{\lambda^2} = \frac{2 - 4p^2 + 4p}{\lambda^2} \\ &\boxed{\text{Var}(X) = \frac{1 + 4p(1-p)}{\lambda^2}} \end{aligned}$$

$$\begin{aligned} &\Rightarrow \int_0^\infty x \lambda e^{-\lambda x} dx \\ &= \lambda \left[x \int e^{-\lambda x} dx - \int 1 \cdot \left(e^{-\lambda x} \right) dx \right]_0^\infty \\ &= \lambda \left[x \cdot \frac{e^{-\lambda x}}{-\lambda} + \int e^{-\lambda x} \frac{dx}{-\lambda} \right] \\ &= \left[-x e^{-\lambda x} + \frac{e^{-\lambda x}}{-\lambda} \right]_0^\infty \\ &= 0 + \left[\frac{e^{-\lambda x}}{-\lambda} \right]_0^\infty = \frac{1}{\lambda} \end{aligned}$$

$$\begin{aligned} &\cancel{\int_0^\infty x^2 \lambda e^{-\lambda x} dx} \\ &= \lambda \left[x^2 \int e^{-\lambda x} dx - \int 2x \left(e^{-\lambda x} \right) dx \right]_0^\infty \\ &= \lambda \left[x^2 \cdot \frac{e^{-\lambda x}}{-\lambda} + 2 \int x \frac{e^{-\lambda x}}{-\lambda} dx \right]_0^\infty \end{aligned}$$

$$\cancel{(x=0)} = 0 + 2 \cdot \frac{1}{\lambda} \cdot \frac{1}{\lambda} = \frac{2}{\lambda^2}$$

$$\boxed{E[X] = \frac{2p-1}{\lambda}}$$

$$\boxed{\text{Var}(X) = \frac{1 + 4p(1-p)}{\lambda^2}}$$

Problem - 3 (for continuous case) 96

i) X and Y be continuous R.v.

with joint CDF $F_{XY}(y)$

and marginal CDF $F_X(y)$, $F_Y(y)$

independence \Rightarrow

$$P(X \in A, Y \in B) = P(X \in A) \cdot P(Y \in B)$$

take $A = (-\infty, a]$, $B = (-\infty, y]$

$$\Rightarrow F_{XY}(y) = P(X \leq a, Y \leq y)$$

$$= P(X \in A, Y \in B) = P(X \in A) \cdot P(Y \in B)$$

$$= P(X \leq a) \cdot P(Y \leq y) = F_X(a) \cdot F_Y(y)$$

$$\therefore F_{XY}(y) = F_X(a) \cdot F_Y(y)$$

$$\Leftrightarrow f_{XY}(y) = f_X(a) \cdot f_Y(y)$$

We know $F_{XY}(y) = \int_{-\infty}^y f_{XY}(z) dz$

$$F_X(a) = \int_{-\infty}^a f_X(z) dz ; \quad f_Y(y) = \int_{-\infty}^y f_Y(z) dz$$

$$f_{XY}(y) = \frac{\partial^2}{\partial a \partial y} F_{XY}(y)$$

$$= \frac{\partial^2}{\partial a \partial y} (F_X(a) \cdot F_Y(y))$$

$$= \frac{\partial}{\partial a} (F_X(a) \cdot \frac{\partial (F_Y(y))}{\partial y})$$

$$= \frac{\partial F_X(a)}{\partial a} \cdot \frac{\partial F_Y(y)}{\partial y} = f_X(a) \cdot f_Y(y)$$

$$\therefore f_{XY}(y) = f_X(a) \cdot f_Y(y)$$

Hence

$\therefore X, Y$ independent $\Leftrightarrow f_{XY}(y) = f_X(a) \cdot f_Y(y)$

(q9) For discrete case.

Q) independence $\Rightarrow P(X=a, Y=b) = P_X(a) \cdot P_Y(b)$

$$P(X=a, Y=b) = P_X(a) \cdot P_Y(b)$$

$$\therefore F_{XY}(a, b) = \sum_{a \leq x} \sum_{b \leq y} P_X(a) \cdot P_Y(b)$$

$$= \sum_{a \leq x} P_X(a) \cdot \sum_{b \leq y} P_Y(b) = F_X(a) \cdot F_Y(y)$$

independence \Rightarrow

$$\boxed{F_{XY}(a, b) = F_X(a) \cdot F_Y(b)}$$

(q9) Q.F. $F_{XY}(a, b) = F_X(a) \cdot F_Y(b)$ then

we know

$$P(X=u, Y=v) = F_{XY}(u, v) - F_{XY}(u-1, v) - F_{XY}(u, v-1) + F_{XY}(u-1, v-1)$$

$$\begin{aligned} \therefore P(X=u, Y=v) &= F_X(u) \cdot F_Y(v) - F_X(u-1) \cdot F_Y(v) \\ &\quad - F_X(u) \cdot F_Y(v-1) + F_X(u-1) \cdot F_Y(v-1) \\ &= [F_X(u) - F_X(u-1)] \cdot [F_Y(v) - F_Y(v-1)] \\ &= P_X(u) \cdot P_Y(v) \end{aligned}$$

$\therefore X$ and Y independent

$$\Leftrightarrow F_{XY}(a, b) = F_X(a) \cdot F_Y(b)$$

A.Y

L) For both discrete and continuous.

Problem - 3

$X, Y \sim \text{Uniform}[a, b]$, independent

$$(A) E[X] = \frac{a+b}{2}, E[Y] = \frac{a+b}{2}$$

$$f_{XY}(x, y) = f_X(x) \cdot f_Y(y) = \frac{1}{b-a} \cdot \frac{1}{b-a} = \frac{1}{(b-a)^2}$$

$$D = |X - Y| \quad E[D] = E[|X - Y|]$$

$$E[|X - Y|] = \int_a^b \int_a^b |x - y| f_{XY}(x, y) dx dy$$

$$= \int_a^b \int_a^b \frac{1}{(b-a)^2} \cdot |x - y| dx dy$$

$$= \int_a^b \int_a^b \frac{1}{(b-a)^2} (x - y) dx dy$$

$$= \frac{2}{(b-a)^2} \int_a^b \int_a^b (x - y) dx dy$$

$$= \frac{2}{(b-a)^2} \int_a^b \left[xy - \frac{y^2}{2} \right]_a^b dx$$

$$= \frac{2}{(b-a)^2} \int_a^b \left(\frac{x^2}{2} - ax + \frac{a^2}{2} \right) dx$$

$$= \frac{2}{(b-a)^2} \int_a^b \frac{(x-a)^2}{2} dx = \frac{1}{(b-a)^2} \times \frac{(b-a)^3}{3}$$

$$\int_a^b x^2 dx = \frac{b^3 - a^3}{3}$$

$$\therefore E[|X - Y|] = \frac{b-a}{3}$$

↳ expected distance b/w

[881] 3 & 882

$$2 \cdot 8 + 8 = [gA|x] \quad \text{two points}$$

Problem - 4

(a) we know

$$E[X|A] = \sum_{\alpha} \alpha \cdot P(X=\alpha | A)$$

$$\Rightarrow E[X|A] = \sum_{\alpha} \alpha \cdot \frac{P(X=\alpha \cap A)}{P(A)}$$

$$= \sum_{\alpha} \alpha \cdot \frac{P(X=\alpha, A)}{P(A)}$$

$$= \frac{1}{P(A)} \cdot \sum_{\alpha} \alpha \cdot P(X=\alpha, A)$$

$$E[X \cdot 1_A] = \sum_{\alpha} \alpha \cdot P(X=\alpha, A)$$

$$\Rightarrow E[X|A] = E[1_A X]$$

1_A → indicator
R.o.N of
event A

(b)

$$X = D_1 + D_2$$

$$E[X|A^q] = E[D_1 + D_2 | D_1 = q]$$

$$= E[D_1 | D_1 = q] + E[D_2 | D_1 = q]$$

$$= q + E[D_2]$$

$$E[D_2] = 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6} + 4 \cdot \frac{1}{6} + 5 \cdot \frac{1}{6} + 6 \cdot \frac{1}{6}$$

$$= 3.5 \quad [X = X_1 + X_2]$$

$$\Rightarrow E[X|A^q] = q + 3.5$$

$$\therefore E[X|A^q] = q + 3.5$$

for $q \in \{1, 2, 3, 4, 5, 6\}$

Problem 5

Given $F_X(a) = \begin{cases} 1 - \frac{\alpha^3}{24} & \text{if } a \geq 0 \\ 0 & \text{if } a < 0 \end{cases}$

PDF $f_X(a) = \frac{d}{da} F_X(a)$

$$\Rightarrow f_X(a) = \begin{cases} \frac{3\alpha^2}{24} & \text{if } a \geq 0 \\ 0 & \text{if } a < 0 \end{cases}$$

$$E[X] = \int_{-\infty}^{\infty} a \cdot f_X(a) da = \int_0^{\infty} a \cdot \frac{3\alpha^2}{24} da$$

$$= \int_0^{\infty} \frac{3\alpha^3}{24} da = 3\alpha^3 \int_0^{\infty} a^2 da$$

$$= 3\alpha^3 \left[\frac{a^3}{3} \right]_0^{\infty} = \frac{3\alpha^3}{24} = \frac{3\alpha}{2}$$

$$\Rightarrow E[X] = \frac{3\alpha}{2}$$

$$E[X^2] = \int_0^{\infty} a^2 \cdot \frac{3\alpha^2}{24} da = \int_0^{\infty} \frac{3\alpha^3}{24} da$$

$$= 3\alpha^3 \cdot \int_0^{\infty} \frac{1}{24} da = 3\alpha^3 \left[-\frac{1}{24} a \right]_0^{\infty}$$

$$= \frac{3\alpha^3}{24} = \frac{3\alpha^2}{8}$$

$$\Rightarrow E[X^2] = 3\alpha^2$$

$$\therefore \text{Var}(X) = E[X^2] - E[X]^2 = 3\alpha^2 - \left(\frac{3\alpha}{2}\right)^2$$

$$= 3\alpha^2 - \frac{9\alpha^2}{4} = \frac{3\alpha^2}{4}$$

$$\therefore \text{Var}(X) = \frac{3\alpha^2}{4}; E(X) = \frac{3\alpha}{2}$$

Problem - 6

Given

$$f_{X,Y}(x,y) = C(y^2 - x^2)e^{-y} \quad \text{for } 0 < y < \infty$$

$$\int_0^{\infty} \int_0^y C(y^2 - x^2)e^{-y} dx dy = 1$$

$$\Rightarrow C \cdot \int_0^{\infty} e^{-y} \left[\int_{-y}^y (y^2 - x^2) dx \right] dy = 1 \quad \text{Eqn 1}$$

$$\int_{-y}^y (y^2 - x^2) dx = y^2(y+y) - \left[\frac{x^3}{3} \right]_y^{-y}$$

$$= 2y^3 - \left[\frac{y^3}{3} + \frac{y^3}{3} \right] = 2y^3 - \frac{2y^3}{3}$$

$$\frac{\partial E}{\partial y} = \frac{\partial \text{Eqn 1}}{\partial y} = \frac{4y^3}{3}$$

Now

$$\int_0^{\infty} e^{-y} \cdot \frac{4y^3}{3} dy = \frac{4}{3} \int_0^{\infty} y^3 \cdot e^{-y} dy$$

$$= \frac{4}{3} \times 6 = 8$$

$$\Rightarrow C \cdot 8 = 1 \Rightarrow C = \frac{1}{8}$$

$$\int_0^{\infty} y^3 \cdot e^{-y} dy$$

$$= \left[-y^3 \cdot e^{-y} \right]_0^{\infty} + \int_0^{\infty} 3y^2 e^{-y} dy$$

$$= 3 \int_0^{\infty} y^2 e^{-y} dy = 3 \left[-y^2 e^{-y} \right]_0^{\infty} + \int_0^{\infty} 2y e^{-y} dy$$

$$= 6 \int_0^{\infty} y e^{-y} dy$$

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$$\Rightarrow C = \frac{1}{8}$$

f(x,y) = marginal PDF

(ii) marginal PDF $f_X(y)$

$$f_Y(y) = \int_{-\infty}^{\infty} f_{XY}(x,y) dx = \int_{-\infty}^y \frac{1}{8} (y^2 - x^2) e^{-y} dx$$

$$= \frac{e^{-y}}{8} \int_{-\infty}^y (y^2 - x^2) dx = \frac{e^{-y}}{8} \cdot \frac{4y^3}{3}$$

marginal $f_X(y)$

$$\Rightarrow f_X(y) = \frac{1}{6} y^3 e^{-y}, y > 0$$

(iii) marginal PDF $f_X(\alpha)$

$$-\alpha \leq x \leq \alpha \Rightarrow y \geq |\alpha|$$

$$f_X(\alpha) = \int_{|\alpha|}^{\infty} f_{XY}(x,y) dy$$

$$y = |\alpha|$$

$$= \int_{|\alpha|}^{\infty} \frac{1}{8} (y^2 - \alpha^2) e^{-y} dy$$

$$= \frac{1}{8} \left[\int_{|\alpha|}^{\infty} y^2 e^{-y} dy - \int_{|\alpha|}^{\infty} \alpha^2 e^{-y} dy \right]$$

$$= \frac{1}{8} \left[\left[y^2 \cdot e^{-y} \right]_{|\alpha|}^{\infty} + \int_{|\alpha|}^{\infty} e^{-y} \cdot 2y dy + \alpha^2 \left[e^{-y} \right]_{|\alpha|}^{\infty} \right]$$

$$= \frac{1}{8} \left[e^{-|\alpha|} \left(|\alpha|^2 + 2|\alpha| - \alpha^2 \right) + \alpha^2 e^{|\alpha|} \right]$$

$$= \frac{1}{8} \left[e^{-|\alpha|} (|\alpha|^2 + 2|\alpha| + 2 - \alpha^2) \right]$$

$$\text{marginal } f_X(\alpha) = \frac{1}{8} \left[e^{-|\alpha|} (2|\alpha| + 2) \right] \approx \frac{|\alpha| + 1}{4} \cdot e^{-|\alpha|}$$

$$\Rightarrow f_X(\alpha) = \frac{|\alpha| + 1}{4} \cdot e^{-|\alpha|}, -\infty < \alpha < \infty$$

Problem - 7

Given x_1, x_2, x_3 be independent Continuous
with common PDF $f(x)$

$$\begin{aligned}
 P(x_1 < x_2 < x_3) &= \iiint_{x_1 < x_2 < x_3} f_{x_1, x_2, x_3}(x_1, x_2, x_3) dx_1 dx_2 dx_3 \\
 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x_1) \cdot f(x_2) \cdot f(x_3) dx_3 \cdot f(x_1) \cdot f(x_2) dx_2 dx_1 \\
 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x_1) \cdot \underbrace{\int_{-\infty}^{\infty} f(x_2) (1 - F(x_2)) dx_2}_{1 - F(x_2)} dx_1 \\
 &= \int_{-\infty}^{\infty} f(x_1) \cdot \left[\int_{-\infty}^{\infty} f(x_2) (1 - F(x_2)) dx_2 \right] dx_1
 \end{aligned}$$

$$\Rightarrow \int_{-\infty}^{\infty} f(x_1) (1 - F(x_1)) dx_1$$

$$\text{let } t = F(x_1)$$

$$dt = f(x_1) dx_1$$

$$\Rightarrow \int_{t=0}^{1} (1-t) dt = \left[t - \frac{t^2}{2} \right]_{0}^{1} F(x_1)$$

$$\left(1 - \left(\frac{1}{2} - \left(F(x_1) - F^2(x_1) \right) \right) \right) = \frac{1}{2} - \left(F(x_1) - \frac{F^2(x_1)}{2} \right)$$

$$= \frac{(1 - F(x_1))^2}{2}$$

$$\boxed{P(x_1 < x_2 < x_3) = \frac{(1 - F(x_1))^2}{2}}$$

$$\Rightarrow P(X_1 < x_2 < x_3) = \int_{-\infty}^{\infty} f_{X_1}(x_1) \cdot (1 - F_{X_1}(x_1))^2 dx_1$$

$$P(X_1 < x_2 < x_3) = \frac{1}{2} \int_{-\infty}^{\infty} (1 - F_{X_1}(x_1))^2 f_{X_1}(x_1) dx_1$$

$$t = F_{X_1}(x_1) \Rightarrow dt = f_{X_1}(x_1) dx_1$$

$$\frac{1}{2} \int_0^1 (1-t)^2 dt = \frac{1}{2} \left[\frac{(1-t)^3}{3} \right]_0^1 = \frac{1}{6}$$

$[0, 1] = a, [1, \infty) = b$

$$\therefore P(X_1 < x_2 < x_3) = \frac{1}{6}$$

$$E[X_1 X_2] = E[X_1] + E[X_2]$$

$$E[X_1 X_2] = E[X_1] + E[X_2] \Rightarrow$$

Prob $E[X_1 X_2] \neq E[X_1] E[X_2]$ wird gew

Prob $E[X_1 X_2] \neq E[X_1] E[X_2]$ ab $X_1 \neq X_2$

$$E[X_1 X_2] = E[X_1] E[X_2]$$