

C.I Discrete Random Variables

PROBLEM 1. The $\text{Im } Z$ is discrete because $\text{Im } X$ and $\text{Im } Y$ are discrete sets.

Let $z \in \mathbb{R}$. If $\{Z = z\} = \emptyset$, then $\{Z = z\}$ is an event because the set \emptyset is always an event. Assume that $\{Z = z\} \neq \emptyset$. We have to consider two cases.

i) $z = 0$. In this case, the only way that $Z(s) = 0$ is if $X(s) = 0$ or $Y(s) = 0$. Therefore,

$$\{Z = 0\} = \{X = 0\} \cup \{Y = 0\}.$$

Since $\{X = 0\}$ and $\{Y = 0\}$ are events, we conclude that $\{Z = 0\}$ are events (recall that, by assumption, X and Y are discrete random variables).

ii) $z \neq 0$. In this case, the functions X and Y can't take the value 0. If $s \in \{Z = z\}$, then $X(s)Y(s) = Z(s) = z$. Therefore $X(s) = z/Y(s)$. Let $y = Y(s)$. Then $X(s) = z/y$ and $Y(s) = y$. In other words, $s \in \{X = z/y\} \cap \{Y = y\}$. On the other hand, if $s \in \{X = z/y\} \cap \{Y = y\}$, then $X(s) = z/y$ and $Y(s) = y$. Therefore, $Z(s) = X(s)Y(s) = (z/y)y = z$ and then $s \in \{Z = z\}$. In summary, we have just proved that

$$\{Z = z\} = \bigcup_{y \in \text{Im } Y, y \neq 0} \left(\{X = z/y\} \cap \{Y = y\} \right).$$

For a given $z \in \mathbb{R}$, $z \neq 0$ and $y \in \text{Im } Y$, the event $\{X = z/y\} \cap \{Y = y\}$ is an event because X and Y are discrete random variable and \mathcal{A} is an event space. Thus, a countable union of these events will remain an event. Hence, $\{Z = z\}$ is an event.

In each case, $\{Z = z\}$ is an event. The map Z satisfies condition (a) and (b) in Definition 1 and therefore Z is a discrete random variable. \triangle

PROBLEM 2. We have $\text{Im}(1_A) = \{0, 1\}$, which is a finite set (therefore discrete).

Let $x \in \mathbb{R}$. We have three cases to consider.

1. $x = 0$. In this case, $\{1_A = 0\} = \overline{A}$. Since A is an event, we know that \overline{A} is also an event. Therefore, $\{1_A = 0\}$ is an event.
2. $x = 1$. In this case, $\{1_A = 1\} = A$ and A is an event. Hence, $\{1_A = 1\}$ is an event.
3. $x \neq 0$ and $x \neq 1$. In this case, $\{1_A = x\} = \emptyset$ because there is no s such that $1_A(s) = x$ (the only possible values are 0 and 1 for 1_A). Since \emptyset is an event, $\{1_A = x\}$ is an event. \triangle

PROBLEM 3.

- a) Let $x \in \mathbb{R}$. If $\{X \leq x\} = \emptyset$, then $\{X \leq x\}$ is an event. Assume that $\{X \leq x\} \neq \emptyset$. Since X is a discrete random variable, the set $\text{Im } X$ is discrete. This means there are only a countable values of $\text{Im } X$ that can be smaller than the number x . List them in decreasing

order, say x_1, x_2, x_3, \dots , with $x_i \geq x_j$, when $i \leq j$ and $x_j \leq x$ for any j . Therefore, we can write

$$\{X \leq x\} = \bigcup_{j=1}^{\infty} \{X = x_j\}.$$

The map X is a discrete random variable. Therefore, each set $\{X = x_j\}$ is an event and this implies that $\bigcup_{j=1}^{\infty} \{X = x_j\}$ is an event. Hence, $\{X \leq x\}$ is an event.

b) Let $x \in \mathbb{R}$. We can write

$$\{X < x\} = \{X \leq x\} \cap \overline{\{X = x\}}.$$

In other words, the set $\{X < x\}$ is the set of $s \in S$ that belong to $\{X \leq x\}$ but are not in $\{X = x\}$. From part a), the set $\{X \leq x\}$ is an event and from the fact that X is assumed to be a discrete random variable, $\{X = x\}$ is an event. Therefore, $\{X \leq x\} \cap \overline{\{X = x\}}$ is an event and hence $\{X < x\}$ is an event.

c) We have

$$\{X \geq x\} = \overline{\{X < x\}}.$$

From part b), we know that $\{X < x\}$ is an event, hence $\{X \geq x\}$ is also an event.

d) We have

$$\{X > x\} = \overline{\{X \leq x\}}.$$

From part c), we know that $\{X \leq x\}$ is an event, hence $\{X > x\}$ is also an event.

PROBLEM 4. Assume that X is a discrete random variable. Then $\text{Im } X$ is discrete and $\{X = x\}$ is an event for every $x \in \mathbb{R}$. From Problem 3, part a), the set $\{X \leq x\}$ is an event. Therefore, conditions a) and b) in the statement are satisfied.

Assume that the two conditions in the statement are satisfied. Then, in particular, $\text{Im } X$ is discrete. Also, for an $x \in \mathbb{R}$, the set $\{X > x\}$ is an event because it is the complement of the event $\{X \leq x\}$. Also, for an $x \in \mathbb{R}$, we have

$$\{X < x\} = \bigcup_{j=1}^{\infty} \left\{X \leq x - \frac{1}{j}\right\}.$$

This is a countable unions of the events $\{X \leq x - \frac{1}{j}\}$ and therefore $\{X < x\}$ is an event. But also $\{X \geq x\}$ is also an event because it is the complement of $\{X < x\}$. Let $x \in \mathbb{R}$. We can write

$$\{X = x\} = \{X \leq x\} \cap \{X \geq x\},$$

the intersection of two events! So $\{X = x\}$ is also an event. Hence X is a discrete random variable.

C.II Probability Mass Functions

PROBLEM 5. The probability measure P on S is given by

$$P(\{r\}) = \frac{2}{5}, \quad P(\{b\}) = \frac{2}{5}, \quad P(\{y\}) = \frac{1}{5}.$$

The function $X : S \rightarrow \mathbb{R}$ is given by $X(\{r\}) = -10$, $X(\{b\}) = 10$, and $X(\{y\}) = 20$. Therefore, we have

- $p_X(-10) = P(X = -10) = P(\{r\}) = \frac{2}{5}$.
- $p_X(10) = P(X = 10) = P(\{b\}) = \frac{2}{5}$.
- $p_X(20) = P(X = 20) = P(\{y\}) = \frac{1}{5}$.
- $p_X(x) = 0$, for $x \neq -10, 10, 20$.

PROBLEM 6. A child may or may not identify properly the picture. Let w_1, w_2, w_3 be the words corresponding to the animal in picture p_1, p_2, p_3 . A child will identify correctly a picture if the word w_i put under the picture p_i . Therefore, we can identify an outcome as an ordered list of three symbols from $\{w_1, w_2, w_3\}$. For example, $w_1w_2w_3$ means that the child identified the animal in picture p_1 as w_1 , in picture p_2 as w_2 , and in picture p_3 as w_3 . Therefore, the sample space is

$$S = \{w_1w_2w_3, w_1w_3w_2, w_2w_1w_3, w_3w_2w_1, w_3w_1w_2, w_2w_3w_1\}.$$

If we just keep the numbers

$$S = \{123, 132, 213, 321, 312, 231\}.$$

Each outcome are equally likely to happen, so with $1/6$ chance.

Let $Y : S \rightarrow \mathbb{R}$. Notice that, if the child successfully matches 2 pictures with their words, then the third picture will be also successfully matched. Therefore, the child may correctly identify 0, 1, or 3 of the pictures presented and $\text{Im } Y = \{0, 1, 3\}$. Then, we have $p_Y(y) = 0$ for any $y \neq 0, 1, 3$. For the other values of y :

- $p_Y(0) = P(Y = 0) = P(\{312, 231\}) = 1/3$.
- $p_Y(1) = P(Y = 1) = P(\{132, 213, 312\}) = \frac{1}{2}$.
- $p_Y(3) = P(Y = 3) = P(\{123\}) = \frac{1}{6}$.

PROBLEM 7. The sample space is all distinct subsets of two numbers from $\{1, 2, 3, 4, 5\}$. There are $\binom{5}{2} = 10$ possible outcomes and all of the outcome are equally likely to occur.

1. We have $\text{Im } X = \{2, 3, 4, 5\}$. The number 1 is missing because in the two balls selected, if ball #1 is selected, then the other ball's number is automatically one of 2, 3, 4, 5. We will present the pmf of X in a table.

x	2	3	4	5
$p_X(x)$	1/10	1/5	3/10	2/5

To compute $p_X(2)$, we first notice that $\{X = 2\} = \{\{1, 2\}\}$ and therefore $P(X = 2) = 1/10$. To compute $p_X(3)$, we first notice that $\{X = 3\} = \{\{1, 3\}, \{2, 3\}\}$ and therefore $P(X = 3) = 2/10 = 1/5$. Similar calculations lead to the values of $p_X(4)$ and $p_X(5)$. Notice that $1/10 + 1/5 + 3/10 + 2/5 = 1$.

2. Removing the parenthesis in the set and considering them as unordered list, the outcomes of S can be explicitly enumerated:

$$S = \{12, 13, 14, 15, 23, 24, 25, 34, 35, 45\}.$$

Therefore, considering all the outcomes and adding the numbers, we see that $\text{Im } X = \{3, 4, 5, 6, 7, 8, 9\}$. We have, more precisely, $X(12) = 3$, $X(13) = 4$, $X(14) = X(23) = 5$, $X(24) = X(15) = 6$, $X(25) = X(34) = 7$, $X(35) = 8$, $X(45) = 9$. Using the same strategy as in a), we find the following values for p_X .

x	3	4	5	6	7	8	9
$p_X(x)$	1/10	1/10	1/5	1/5	1/5	1/10	1/10

PROBLEM 8. If p is a probability mass function, then we know it should satisfy $\sum_{k=1}^{\infty} p(k) = 1$. This gives the following condition:

$$\sum_{k=1}^{\infty} \frac{c}{k(k+1)} = 1.$$

Now, using the trick $\frac{1}{k(k+1)} = \frac{1}{k} - \frac{1}{k+1}$, we see that the series $\sum_{k=1}^{\infty} \frac{1}{k(k+1)}$ is convergent and

$$\sum_{k=1}^{\infty} \frac{1}{k(k+1)} = \lim_{N \rightarrow \infty} \sum_{k=1}^N \left(\frac{1}{k} - \frac{1}{k+1} \right) = \lim_{N \rightarrow \infty} 1 - \frac{1}{N+1} = 1.$$

Therefore, using the properties of series, we see that

$$\sum_{k=1}^{\infty} \frac{c}{k(k+1)} = 1 \iff c \sum_{k=1}^{\infty} \frac{1}{k(k+1)} = 1 \iff c \cdot 1 = 1 \iff c = 1.$$

This means the function p is a pmf if and only if $c = 1$.

C.III Functions of Discrete Random Variables

PROBLEM 9. Setting $X = 1$, $X = 2$, $X = 3$, and $X = 4$ in the expression of Y , we get $Y = 0, 3, 8, 15$. Therefore, $\text{Im } Y = \{0, 3, 8, 15\}$.

Using Theorem 3, with $g(x) = x^2 - 1$, we have

$$p_Y(y) = \sum_{x \in g^{-1}(y)} P(X = x).$$

For

- $y = 0$, we have $g^{-1}(0) = \{x \in \text{Im } X : g(x) = 0\} = \{1\}$;
- $y = 3$, we have $g^{-1}(3) = \{x \in \text{Im } X : g(x) = 3\} = \{2\}$;
- $y = 8$, we have $g^{-1}(8) = \{x \in \text{Im } X : g(x) = 8\} = \{3\}$;
- $y = 15$, we have $g^{-1}(15) = \{x \in \text{Im } X : g(x) = 15\} = \{4\}$.

Therefore,

- $p_Y(0) = P(X = 1) = 0.4$.
- $p_Y(3) = P(X = 2) = 0.3$.
- $p_Y(8) = P(X = 3) = 0.2$.
- $p_Y(15) = P(X = 4) = 0.1$.
- $p_Y(y) = 0$ for any other values y different from 0, 3, 8, 15. \triangle

PROBLEM 10. Setting $X = 1, 2, 3, 4$ in the expression of Y , we get $Y = 1, 0, -1, 0$ respectively. Therefore, $\text{Im } Y = \{-1, 0, 1\}$.

Using Theorem 3 again, but with $g(x) = \sin(\frac{\pi}{2}x)$, we have

$$p_Y(y) = \sum_{x \in g^{-1}(y)} P(X = x).$$

For

- $y = -1, g^{-1}(-1) = \{3\};$
- $y = 0, g^{-1}(0) = \{2, 4\};$
- $y = 1, g^{-1}(1) = \{1\}.$

Therefore,

- $p_Y(-1) = P(X = 3) = 0.2$.
- $p_Y(0) = P(X = 2) + P(X = 4) = 0.3 + 0.1 = 0.4$.
- $p_Y(1) = P(X = 1) = 0.4$. \triangle

C.IV Expectation and Variance

PROBLEM 11. Let t_1 be the label “the dimensions of the trailer are $8 \times 10 \times 30$.” and let t_2 be the label “the dimensions of the trailer are $8 \times 10 \times 40$.” Given a trailer, the possible outcome is the trailer is of type t_1 or of type t_2 . Therefore, $S = \{t_1, t_2\}$ with $P(\{t_1\}) = 0.3$ and $P(\{t_2\}) = 0.7$.

Let X be the map given the volume of a trailer. We have

$$X(t_1) = 8 \cdot 10 \cdot 30 = 2400 \quad \text{and} \quad X(t_2) = 8 \cdot 10 \cdot 40 = 3200.$$

Therefore, we get

$$\text{Exp}(X) = X(t_1)P(X = 2400) + X(t_2)P(X = 3200) = (2400)(0.3) + (3200)(0.7) = 2960.$$

The average volume shipped per trailer load is 2960ft³. \triangle

PROBLEM 12. A firm can be assign one or the two contracts. Therefore, we can generate the set of outcomes as couple of letters taken from $\{a, b, c\}$. For example AA means A was assigned to the two contracts, but AB or BA means that A and B was assigned to one of the contracts. The sample space S is

$$S = \{aa, ab, ba, ac, ca, bb, bc, cb, cc\}.$$

Since the firms are assigned a contract at random, each outcome are equally likely to occur, so with $1/9$.

a) In the first scenario, assume that X is the possible profit made by firm A after the contracts were assigned. Therefore, this means

- $X(aa) = 180,000$.
- $X(ab) = X(ba) = X(ac) = X(ca) = 90,000$.
- $X(bb) = X(bc) = X(cb) = X(cc) = 0$.

The expectation is then calculated as followed:

$$\begin{aligned}
 \text{Exp}(X) &= 180,000P(X = 180,000) + 90,000P(X = 90,000) + 0P(X = 0) \\
 &= 180,000P(\{aa\}) + 90,000(P(\{ab, ba, ac, ca\}) + 0 \\
 &= \frac{180,000}{9} + \frac{90,000 \cdot 4}{9} \\
 &= 20,000 + 40,000 \\
 &= 60,000
 \end{aligned}$$

b) Let Y be the possible possible made by firms A and B after the contrasts were assigned. Therefore, this means

- $X(aa) = X(bb) = X(ab) = X(ba) = 180,000$.
- $X(ac) = X(ca) = X(bc) = X(cb) = 90,000$.
- $X(cc) = 0$.

The expectation is then calculated as followed:

$$\begin{aligned}
 \text{Exp}(X) &= 180,000P(X = 180,000) + 90,000P(X = 90,000) + 0P(X = 0) \\
 &= 180,000P(\{aa, bb, ab, va\}) + 90,000P(\{ac, ca, bc, cb\}) \\
 &= \frac{(180,000)(4)}{9} + \frac{(90,000)(4)}{9} \\
 &= 80,000 + 40,000 \\
 &= 120,000.
 \end{aligned}$$

△

PROBLEM 13. We have $\text{Im } X = \{1, 2, 3, 4, 5, 6\}$ and each value of X has a chance of $1/6$ to occur. Therefore,

$$\text{Exp}(X) = \frac{1}{6} + \frac{2}{6} + \frac{3}{6} + \frac{4}{6} + \frac{5}{6} + \frac{6}{6} = \frac{21}{6} = 3\frac{1}{2}.$$

The variance is calculated using formula in the Theorem 6. We first have $\text{Im } X^2 = \{1, 4, 9, 16, 25, 36\}$ and

$$\text{Exp}(X) = \frac{1}{6} + \frac{4}{6} + \frac{9}{6} + \frac{16}{6} + \frac{25}{6} + \frac{36}{6} = \frac{91}{6} = 30\frac{1}{6}.$$

Therefore,

$$\text{Var}(X) = \text{Exp}(X^2) - (\text{Exp}(X))^2 = \frac{91}{6} - \frac{21}{6} = \frac{70}{6} = 11\frac{2}{3}.$$

Thus,

$$\sigma = \sqrt{\text{Var}(X)} = \sqrt{70/6} \approx 3.4157.$$

△

PROBLEM 14. By the formula in Theorem 6, we have

$$\text{Var}(aX + b) = \text{Exp}((aX + b)^2) - (\text{Exp}(aX + b))^2.$$

We have $(aX + b)^2 = a^2X^2 + 2abX + b^2$ and $\text{Exp}(aX + b) = a\text{Exp}(X) + b$. Therefore,

$$\begin{aligned}\text{Var}(aX + b) &= a^2\text{Exp}(X^2) + 2ab\text{Exp}(X) + b^2 - a^2(\text{Exp}(X))^2 - 2ab\text{Exp}(X) - b^2 \\ &= a^2\text{Exp}(X^2) - a^2\text{Exp}(X)^2 \\ &= a^2\text{Var}(X).\end{aligned}$$

△

C.V Conditional Expectation and the Partition Theorem

PROBLEM 15. Let $B = \{X = x\}$. Then, we have

$$E(g(X)|B) = \sum_{y \in \text{Im } g(X)} yP(g(X) = y|B).$$

However, if B has occurred, then $X = x$ and the sum over $\text{Im } g(X)$ is restricted to the value $y = g(x)$. Hence,

$$E(g(X)|B) = g(x)P(g(X) = g(x)) = g(x)P(X = x).$$

△

C.VI Examples of Discrete Random Variables

PROBLEM 16. The map X has a discrete range, that is $\text{Im } X = \{0, 1, 2, 3, \dots, 30\}$. However it does not have a binomial distribution because the probability that there is rain on a given day varies from day to day. Therefore, the parameter p is not fixed. △

PROBLEM 17.

- In this case, if it was explicitly mentioned “The number of students in a sample of X students who took the SAT”, then we could model the distribution of X on the binomial distribution with $n = 100$ and $q = 0.45$. Unfortunately, it is not mentioned and therefore we can’t model the distribution with a binomial distribution.
- It can’t be model by a binomial distribution because there is not enough information to find the parameter q . We will see later that the distribution of the scores of the 100 students can be model by a normal distribution.
- If X_j is the random variable “The student labeled j scored above average on the SAT”, then X : “the number of students in the sample who scored above average on the SAT”, which is equal to $X_1 + X_2 + \dots + X_{100}$, has a binomial distribution. In this case, $n = 100$ and the value of q is not possible to find. We would need more information to compute an approximate value for q . For example, with the additional assumption that the distribution of the student’s scores is a Normal distribution, then we can assume that $q = 0.5$, because $P(X > \mu) = 0.5$ for any normal distribution.
- It can’t be model by a binomial distribution because there is not enough information to find the parameter q . We would need additional information on the average time of a student to complete the test. We will see later that the distribution of the random variable will be modeled by a Normal distribution. △

PROBLEM 18. By Definition 3, we have

$$\text{Exp}(X) = \sum_{k=0}^n kP(X = k) = \sum_{k=0}^n k \frac{n!}{k!(n-k)!} q^k (1-q)^{n-k}.$$

The term with $k = 0$ disappears and we enter into the following chain of equalities:

$$\begin{aligned} \text{Exp}(X) &= \sum_{k=1}^n \frac{kn!}{k!(n-k)!} q^k (1-q)^{n-k} \\ &= \sum_{k=0}^{n-1} \frac{(k+1)n!}{(k+1)!(n-k-1)!} q^{k+1} (1-q)^{n-k-1} \\ &= nq \sum_{k=0}^{n-1} \frac{(n-1)!}{k!(n-1-k)!} q^k (1-q)^{n-k-1} \\ &= nq(q + (1-q))^{n-1} \\ &= nq. \end{aligned}$$

To compute the $\text{Var}(X)$, we use the formula Theorem 6. We have $\text{Exp}(X) = nq$ from the previous calculations. We need to compute $\text{Exp}(X^2)$:

$$\text{Exp}(X^2) = \sum_{k=0}^n k^2 P(X^2 = k^2) = \sum_{k=0}^n k^2 P(X = k),$$

where $\{X^2 = k^2\} = \{X = k\}$ because X assumes only non-negative integer values. Therefore,

$$\begin{aligned} \text{Exp}(X^2) &= \sum_{k=0}^n \frac{k^2 n!}{k!(n-k)!} q^k (1-q)^{n-k} \\ &= \sum_{k=1}^n \frac{k^2 n!}{k!(n-k)!} q^k (1-q)^{n-k} \\ &= nq \sum_{k=0}^{n-1} \frac{(k+1)(n-1)!}{k!(n-1-k)!} q^k (1-q)^{n-1-k} \\ &= nq \left(\sum_{k=0}^{n-1} \frac{k(n-1)!}{k!(n-1-k)!} q^k (1-q)^{n-1-k} + \sum_{k=0}^{n-1} \frac{(n-1)!}{k!(n-1-k)!} q^k (1-q)^{n-1-k} \right) \\ &= nq \left(\sum_{k=1}^{n-1} \frac{k(n-1)!}{k!(n-1-k)!} q^k (1-q)^{n-1-k} + 1 \right) \\ &= nq \left((n-1)q \sum_{k=0}^{n-2} \frac{(n-2)!}{k!(n-2-k)!} q^k (1-q)^{n-2-k} + 1 \right) \\ &= nq((n-1)q + 1) \\ &= nq(nq - q + 1) \\ &= n^2 q^2 + nq(1-q). \end{aligned}$$

Hence,

$$\text{Var}(X) = n^2 q^2 + nq(1-q) - n^2 q^2 = nq(1-q).$$

△