# Predicting Student Performance Linear Models (PHP2601), Prof. Ani Eloyan

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# EDA and the Linear Model

#### Introduction

We will be analyzing educational data to understand the predictors of student performance. Specifically, we seek to **understand** whether five predictors — as a subset of an exhaustive list of potential predictors — are significant predictors of student performance.

Testing the significant of a subset of predictors is becoming increasingly important in modern statistical questions, especially with more information becoming available.

We will be using a publicly available dataset from Kaggle that contains information about students and their exam scores.

### Hypothesis to be Tested

#### We are interested in:

- ► Hours Studied
- Attendance
- ► Sleep Hours
- Previous Scores
- ► Tutoring Sessions

We can formalize this question as follows:

$$\qquad \qquad H_A: \{\beta_1 \neq 0\} \cap \dots \cap \{\beta_5 \neq 0\}$$

Observe the 0-indexed variables from p+2 to P.

# Exploratory Data Analysis (EDA)

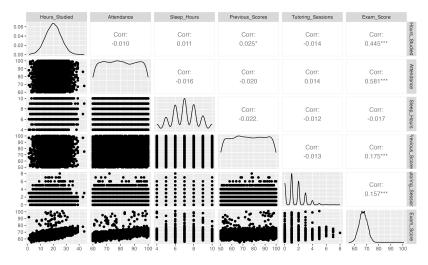


Figure 1: Correlation Matrix

#### Variable Transformations

We will transform the variables to ensure that the assumptions of the linear model are met.

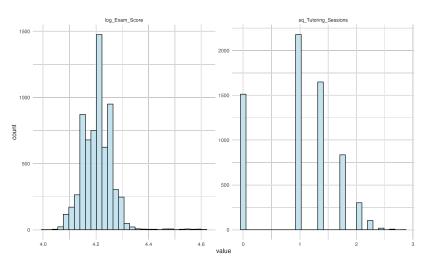


Figure 2: Variable Transformation

#### The Linear Model

Let us begin by discussing the assumptions of linear regression model. In a Gauss-Markov setting, we assume that our linear model is of the form:

$$Y = \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} 1 & X_{12} & X_{13} & \cdots & X_{1(p+1)} \\ 1 & X_{22} & X_{23} & \cdots & X_{2(p+1)} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & X_{n2} & X_{n3} & \cdots & X_{n(p+1)} \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \vdots \\ \beta_p \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{bmatrix}$$

where  $\mathbb{E}[\epsilon]=0$  and  $\mathrm{Var}[\epsilon]=\sigma^2I$  denote the zero-mean and constant variance assumptions. In our case, we begin with p=5, i.e. our design matrix has p+1 columns, accounting for the intercept term.

# Solving for $\hat{\beta}$

We can solve for  $\hat{\beta}$  via the normal equations:

$$\begin{split} \hat{\beta} = & (X^T X)^g X^T Y \\ = & \left[ \begin{bmatrix} 1 & 1 & \cdots & 1 \\ X_{12} & X_{22} & \cdots & X_{n2} \\ \vdots & \vdots & \ddots & \vdots \\ X_{1(p+1)} & X_{2(p+1)} & \cdots & X_{n(p+1)} \end{bmatrix} \begin{bmatrix} 1 & X_{12} & \cdots & X_{1(p+1)} \\ 1 & X_{22} & \cdots & X_{2(p+1)} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & X_{n2} & \cdots & X_{n(p+1)} \end{bmatrix} \right]^g \\ & \left[ \begin{bmatrix} 1 & 1 & \cdots & 1 \\ X_{12} & X_{22} & \cdots & X_{n2} \\ \vdots & \vdots & \ddots & \vdots \\ X_{1(p+1)} & X_{2(p+1)} & \cdots & X_{n(p+1)} \end{bmatrix} \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} \right] \end{split}$$

In our case, all predictors but Sleep Hours are significant predictors of exam scores, even at a 1% level of significance.

### Estimability of the Hypothesis

Question: Can we estimate an object  $K^T\beta$  with our data X?

Formally, we say that if  $\exists~A$  s.t.  $X^TA=K^T$ , i.e.  $K^T$  can be expressed as a linear combination of X and some matrix A, then  $K^T\beta$  is estimable.

In our case, this is straightforward to verify. Can we think of an example when this is not true? (Hint: Dimension "mismatch")

## Distribution of $K^T\beta$

Since  $K^T\beta$  estimable, its best linear unbiased estimator (BLUE) is given by:

$$\begin{split} \mathbf{K_i}^T \hat{\boldsymbol{\beta}} &\sim \mathit{N}(\mathbf{K_i}^T (X^T X)^g X^T X \boldsymbol{\beta}, \sigma^2 \mathbf{K_i}^T (X^T X)^g \mathbf{K_i}) \quad \text{and} \\ \mathbf{K}^T \hat{\boldsymbol{\beta}} &\sim \mathit{N}(\mathbf{K}^T (X^T X)^g X^T X \boldsymbol{\beta}, \sigma^2 \mathbf{K}^T (X^T X)^g \mathbf{K}) \end{split}$$

This object  $K^T\beta$  may seem a bit arbitrary, even useless, at first. However, it is in fact the building block for the test statistic we will construct now!

### Quadratic Form in our Joint Testing Procedure

Suppose  $H := K(X^TX)^gK^T$ , then

$$(K\beta)^T (\sigma^2 H)^{-1} (K\hat{\beta}) \sim \chi^2_{\mathrm{df=rank}(H)}(\lambda)$$

where the non-centrality parameter  $\lambda = \frac{1}{2}(K\beta)^T(\sigma^2H)^{-1}(K\beta)$  is the well-known distributional result of a normal quadratic form.

Finally, our F Statistic:

$$F := \frac{\left((K\beta)^T(\sigma^2H)^{-1}(K\beta)\right)/\mathrm{rank}(H)}{\mathrm{RSS}/(n-p)} \sim \frac{\chi^2(\lambda)}{\chi^2} \sim F_{\mathrm{rank}(H),n-p}(\lambda)$$

We have successfully constructed a statistical test that allows us to test our hypothesis with a simple F-test. In R, we can use the anova() function to perform this test.

#### Results

```
Analysis of Variance Table

Model 1: log_Exam_Score ~ 1

Model 2: log_Exam_Score ~ Hours_Studied + Attendance + Sleep_Hours + Previous_Scores + sq_Tutoring_Sessions

Res.Df RSS Df Sum of Sq F Pr(>F)

1 6606 20.6969
2 6601 7.4771 5 13.22 2334.2 < 2.2e-16 ***
---

Signif. codes: 0 '****' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Figure 3: F-Test Results

The result shows that under the null hypothesis, the probability of getting a more extreme result than our calculate F-test statistics  $\Pr(>F)$  is 2.2e-16.

This evidence would lead us to reject the null hypothesis and conclude that our subset of predictors is indeed a significant predictor of exam scores

# LASSO Regression

# Non-Linear Model