Cox Proportional Hazard Model  Estimation	
Introduction  • Cox proportional hazard model $h(t \mathbf{Z}) = h_0(t) \exp(\mathbf{Z}^T \boldsymbol{\beta})$ - The coefficients $\boldsymbol{\beta}$ quantify the effects of covariates on hazard function.  • Therefore $\boldsymbol{\beta}$ is the focus of our inference, while $h_0(t)$ , the baseline hazard, is of less interest. Thus it is treated as "nuisance parameter".  • In the following, we will focus on - Estimate $\boldsymbol{\beta}$ and derive its statistical properties - Hypothesis testing for $\boldsymbol{\beta}$	
<ul> <li>Parameter estimation</li> <li>Cox (1972) constructed a partial likelihood to estimate β.</li> <li>Suppose there is no ties in the event times.</li> <li>Let t₁ &lt; t₂ &lt; ··· &lt; t₂ be the ordered event times.</li> <li>At each event time tᵢ, define</li> <li>Z(ᵢ) denote the covariate vector associated with the individual whose failure time i tᵢ.</li> <li>Rᵢ be the risk set at time tᵢ with Rᵢ = ⟨j: tᵢ ≥ tᵢ⟩.</li> </ul>	

#### Partial likelihood

• Cox (1972) proposed the partial likelihood

$$PL(\boldsymbol{\beta}) = \prod_{i=1}^{D} \frac{\exp(\boldsymbol{z}_{(i)}^{T} \boldsymbol{\beta})}{\sum_{j \in R_{i}} \exp(\boldsymbol{z}_{j}^{T} \boldsymbol{\beta})}.$$

- It only involves  $\pmb{\beta}$ , not the baseline hazard  $h_0(t)$ .
- Does not use information on the actual censoring and event times.
- Only need the rankings to determine the at-risk sets  $\{R_i\}_{i=1}^n$ .

### Partial likelihood

- The partial likelihood is treated as usual likelihood. Thus estimation and inference of  $\beta$  can be carried out by the usual means.
- The maximum partial likelihood estimator (MPLE)

$$\widehat{\pmb{\beta}} = \arg\max_{\beta} PL(\pmb{\beta}) = \arg\max_{\beta} \prod_{i=1}^{D} \frac{\exp\left(\pmb{Z}_{(i)}^{T} \pmb{\beta}\right)}{\sum_{j \in R_{(t_i)}} \exp\left(\pmb{Z}_{j}^{T} \pmb{\beta}\right)}.$$

• The likelihood based testing methods can also be used to make inference about  $\widehat{\beta}$  (more later).

# An example

• Consider a simple data example

ID T  $\delta$  Z

1 9 1 4

2 8 0 5 3 6 1 7

4 10 1 3

• What is the partial likelihood for this data?

### An example

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Contribution to the partial likelihood
 t_i Risk set
        {1, 2, <mark>3</mark>, 4} 7
                                            \varphi_1(\beta) = \frac{1}{\exp(4\beta) + \exp(5\beta) + \exp(7\beta) + \exp(3\beta)}
                                                                                    \exp(4\beta)
        {1, 4}
                                                              \varphi_2(\beta) = \frac{\exp(4\beta)}{\exp(4\beta) + \exp(3\beta)}
                                                                       \varphi_3(\beta) = \frac{\exp(3\beta)}{\exp(3\beta)}
10 {4}
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The partial likelihood is the product of the three terms in the last column.  $PL(\beta) = \varphi_1(\beta)\varphi_2(\beta)\varphi_3(\beta).$ 

### Partial likelihood function

- · How to understand the partial likelihood function?
- We justify the partial likelihood using a simple example.
- Suppose data has only two observations

Individual	Observed time	Corvariates	Hazard
1	$T_1$	$Z_1$	$h_1(t) = h_0(t) \exp(\mathbf{Z}_1^T \boldsymbol{\beta})$
2	$T_2$	$\boldsymbol{Z}_2$	$h_2(t) = h_0(t) \exp(\mathbf{Z}_2^T \boldsymbol{\beta})$

#### Partial likelihood function

Suppose the first event occurred at time t. How likely is it that the subject who had event was subject 1?

 $P(Subject 1 \text{ had event at } t \mid one \text{ event at } t)$  $= \frac{P(\text{Sub 1 failed at t and Sub 2 didn't} | \text{survival to } t)}{\text{survival to } t}$ P(One event at t)

 $P(A_1 \cap A_2^c)$ 

- $=\frac{P(A_1\cap A_2^c)\cup\{A_2\cap A_1^c\})}{P(\{A_1\cap A_2^c\})\cup\{A_2\cap A_1^c\})}$  Let  $A_j$  = the event that subject j had event at  $[t,t+\Delta t)$ , given that he/she still alive at t, for j=1,2.
- $-P(A_1\cap A_2^c)=P(A_1)P(A_2^c)\approx h_1(\mathsf{t})\Delta t(1-h_2(\mathsf{t})\Delta t)\approx h_1(\mathsf{t})\Delta t$
- $-P(\{A_1 \cap A_2^c\} \cup \{A_2 \cap A_1^c\}) = P(A_1 \cap A_2^c) + P(A_2 \cap A_1^c) \approx h_1(t)\Delta t + h_2(t)\Delta t$

#### Partial likelihood

 $P(\text{Subject 1 had event at } t \mid \text{One event at } t)$   $\approx \frac{h_1(t)}{h_1(t) + h_2(t)}$   $= \frac{h_0(t) \exp(Z_1^T \beta)}{h_0(t) \exp(Z_1^T \beta) + h_0(t) \exp(Z_2^T \beta)}$   $= \frac{\exp(Z_1^T \beta)}{\exp(Z_1^T \beta) + \exp(Z_2^T \beta)}$ 

• The baseline  $h_0(t)$  is canceled out.

## Conditional probability

• The partial likelihood

$$\begin{split} PL(\pmb{\beta}) &= \prod_{i=1}^D \varphi_i = \\ \prod_{i=1}^D \frac{\exp \left(\pmb{z}_{(i)}^T \pmb{\beta}\right)}{\sum_{j \in R_i} \exp \left(\pmb{z}_j^T \pmb{\beta}\right)}. \end{split}$$

• The term  $\varphi_i$  characterizes the conditional probability that the event occurred to subject (i), given the information at time  $t_i$ .

# Conditional probability

- The partial likelihood is obtained by taking products of these conditional probabilities over all failure times.
- Note that the arbitrary baseline hazard has been eliminated and resulting likelihood only involves B.
- The exact values for censoring times and events times are not important (nonparametric).
- We only need the risk sets  $R(t_i)$  at each event time.

## Estimation using partial likelihood

• We estimate  $oldsymbol{eta}$  by maximizing the partial likelihood

$$PL(\beta) = \prod_{i=1}^{D} \frac{\exp(z_{(i)}^{T}\beta)}{\sum_{j \in R(t_i)} \exp(z_j^{T}\beta)}$$

 $PL(\beta) = \prod_{i=1}^D \frac{\exp \left( z_{(i)}^T \beta \right)}{\sum_{j \in R(t_i)} \exp \left( z_j^T \beta \right)},$  or equivalently maximizing the log partial likelihood  $\log PL(\beta) = \sum_{i=1}^{D} \{Z_{(i)}^{T}\beta - \log \sum_{j \in R(t_i)} \exp(Z_j^{T}\beta)\}.$ 

- Often numerical procedures such as the Newton-Raphson algorithm can be used to solve for the maximum partial likelihood estimate (MPLE).
- In SAS, PROC PHREG does it for you.