# Cox PH Model Estimation of survival function Cox PH model • Recall the Cox's proportional hazard model $h(t \mid \mathbf{Z}) = h_0(t) \exp(\mathbf{Z}^T \boldsymbol{\beta}).$ • The coefficients $oldsymbol{eta}$ quantifies the effects of covariates on relative risk. • The coefficients $oldsymbol{eta}$ can be estimated by maximizing the partial likelihood function. · How to estimate the baseline? Baseline survival • In practice, we are also interested in estimating the survival probability for a new patient with a given set of covariates $S(t \mid \mathbf{Z})$ . • Note that in the Cox's model $h(t \mid \mathbf{Z}) = h_0(t) \exp(\mathbf{Z}^T \boldsymbol{\beta})$ $\Rightarrow H(t \mid \mathbf{Z}) = H_0(t) \exp(\mathbf{Z}^T \boldsymbol{\beta})$ \Rightarrow \exp[-H(t \mid \mathbf{Z})] = \exp[-H\_0(t) \exp(\mathbf{Z}^T \beta)] $\Rightarrow S(t \mid \mathbf{Z}) = [S_0(t)]^{\exp(\mathbf{Z}^T \boldsymbol{\beta})}$ • After $m{\beta}$ being estimated, we need to estimate the baseline survival $S_0(t)$ .

## Recall Kaplan-Meier estimator

• Recall the relationships between h(t), H(t), and S(t) for discrete r.v.s

$$-h(t_i) = P(T = t_i \mid T \ge t_i)$$

$$-H(t) = \sum_{t_i \le t} h(t_i)$$

$$-S(t) = \prod_{t_i \le t} (1 - h(t_i))$$

# Recall Kaplan-Meier estimator

- In the univariate case, individuals all have the same hazard rate  $h(t_i)$  at  $t_i$ .
   For a subject alive just prior to  $t_i$ ,  $P(\text{Failure at } t_i \mid \text{alive prior to } t_i) = h(t_i)$

$$\frac{d_i}{Y_i} \approx h(t_i) \text{ or } d_i \approx E(d_i \mid R_i) = Y_i h(t_i)$$

$$- \hat{h}(t_i) = \frac{d_i}{Y_i}$$

$$-\widehat{H}(t) = \sum_{t_i \le t} \frac{d_i}{Y_i}$$

$$-\hat{S}(t) = \prod_{t_i \le t} \left(1 - \frac{d_i}{\gamma_i}\right)$$

#### Cox PH model

- Different hazard for different individuals
- Let  $d_i$  be the total number of events at  $t_i$

$$d_i \approx E(d_i \mid R_i) = \sum_{j \in R(t_i)} E(I(\text{subject } j \text{ has event at } t_i) \mid R_i)$$

$$= \sum_{j \in R(t_i)} P(\text{subject } j \text{ has event at } t_i \mid R_i) = \sum_{j \in R(t_i)} h_0(t_i) \exp \left(Z_j^T \beta\right)$$

$$= h_0(t_i) \sum_{j \in R(t_i)} \exp \bigl( Z_j^T \beta \bigr)$$
• Therefore estimate  $h_0(t_i)$  by

$$\hat{h}_0(t_i)$$
 by  $d_i = \sum_{j \in R(t_i)} \expig(Z_j^Tetaig)$ 

#### Baseline estimation

Therefore

$$\widehat{H}_0(t) = \sum_{t_i \leq t} \widehat{h}_0(t_i) = \sum_{t_i \leq t} \frac{d_i}{\sum_{j \in R(t_i)} \exp\left(Z_j^T \beta\right)}$$

The baseline survival function can be estimated by

$$\hat{S}_0(t) = \prod_{t_i \leq t} \left(1 - \hat{h}_0(t_i)\right) = \prod_{t_i \leq t} \left(1 - \frac{d_i}{\sum_{j \in R(t_i)} \exp\left(Z_j^T \beta\right)}\right)$$

• This is referred as Product-Limit method in SAS.

#### Survival estimation

- Therefore the survival function for patients with covariates  $\boldsymbol{Z}$  can be estimated by

$$\hat{S}(t \mid \mathbf{Z}) = \left[\hat{S}_0(t)\right]^{\exp(\mathbf{Z}^T \widehat{\boldsymbol{\beta}})}.$$

- $-\hat{S}_0(t)$  is the estimated baseline survival for individuals with  $\mathbf{Z} = 0$ .
- Sometimes the baseline is meaningless that no individual can have  ${m Z}=0$ .
- For example, if Z=age, all patients are over 30 years old in the study.
- But can construct covariates so that  ${\pmb Z}=0$  is representative by centering the covariates.
- Let  ${f Z}^*={f Z}-{f \overline Z}$ , then  ${f Z}^*=0$  corresponding to the baseline being "average" individuals.
- A good reason to center the covariates in the analysis.

### SAS Example

```
data covin;
input stage age id;
cards;
2 50 1
3 40 2
;
run;
```

- Create a data set to estimate survival functions for two potential patients.
- We will estimate the survival functions

$$S_1(t \mid Stage = 2, age = 50)$$
  
 $S_2(t \mid Stage = 3, age = 40)$ 

## SAS Example

- BASELINE Statement in PROC PHREG
  - OUT=SAS-data-set
  - COVARIATES=SAS-data-set
  - TIMELIST=list (eg. timelist=5,20 to 50 by 10)
  - keyword=name (specifies the statistics to be included in the OUT= data set)
  - METHOD(NELSON, PL)

# SAS Example

proc phreg data=larynx plots(overlay)=survival; class stage; model stime\*censor(0)=stage age/risklimits;/\*computes confidence intervals for hazard ratios\*/baseline out=out covariates=covin survival=\_all\_/method=PL rowid=Id; hazardratio 'H1'age / units=10 cl=both; hazardratio 'H2'stage / cl=both; contrast "age" age 10/estimate=exp; contrast "Stage 1 vs Stage 2" stage 1 -1 0 /estimate=exp; contrast "Stage 2 vs Stage 3" stage 0 1 -1 /estimate=exp; run;

