

Foundations of Machine Learning

Classification: decisions and discriminants

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Srinandan (“Sri”) Dasmahapatra [Add WeChat powcoder](#)

2-class classification: determine which class a given input
 $X=x$ belongs to

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2-class classification: determine which class a given input $X=x$ belongs to

decision boundary at equal posterior probability for either class, which are:

$$P(C = +|X = x) = \frac{P(X = x|C = +)P(C = +)}{P(X = x|C = +)P(C = +) + P(X = x|C = -)P(C = -)}$$
$$P(C = -|X = x) = \frac{P(X = x|C = -)P(C = -)}{P(X = x|C = +)P(C = +) + P(X = x|C = -)P(C = -)}$$

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Simplify the notation: introduce $g_{\pm}(x)$

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Simplify the notation: introduce $g_{\pm}(x)$

$$P(C = +|X = x) = \frac{e^{g_+(x)}}{e^{g_+(x)} + e^{g_-(x)}} = \frac{\exp [g_+(x) - g_-(x)]}{1 + \exp [g_+(x) - g_-(x)]}$$
$$P(C = -|X = x) = \frac{e^{g_-(x)}}{e^{g_+(x)} + e^{g_-(x)}} = \frac{1}{1 + \exp [g_+(x) - g_-(x)]}$$

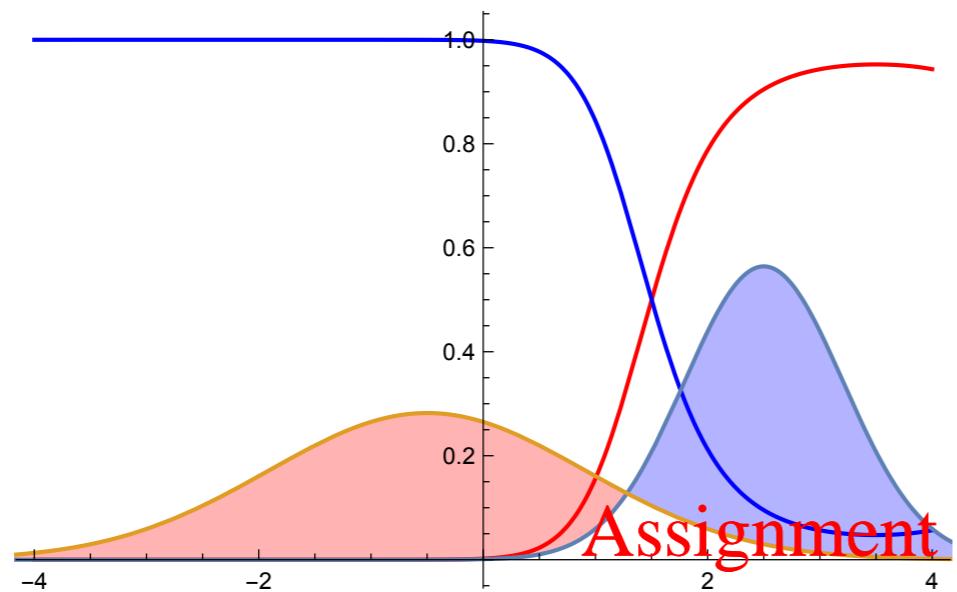
Decision boundary: determine which class a given input belongs to

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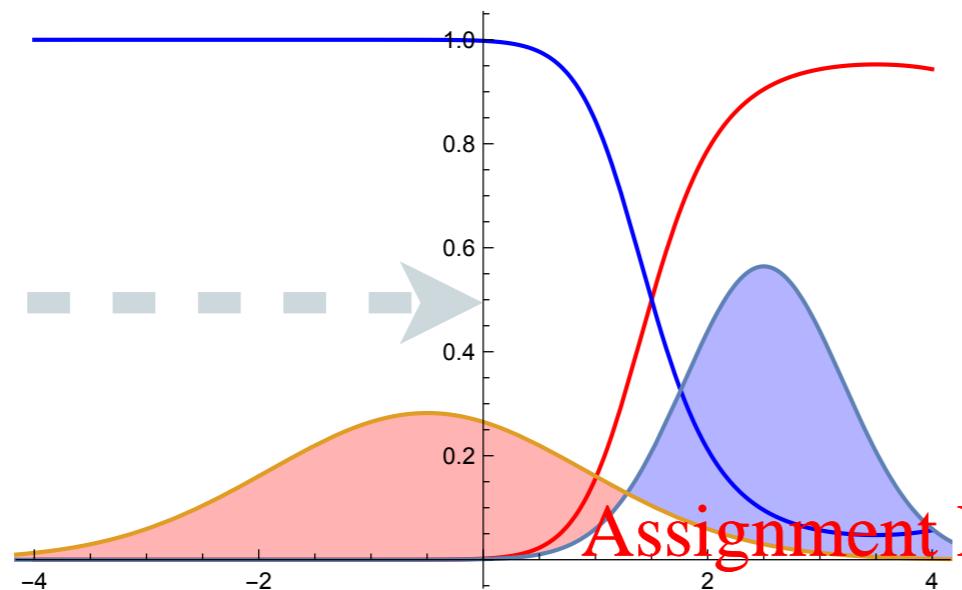
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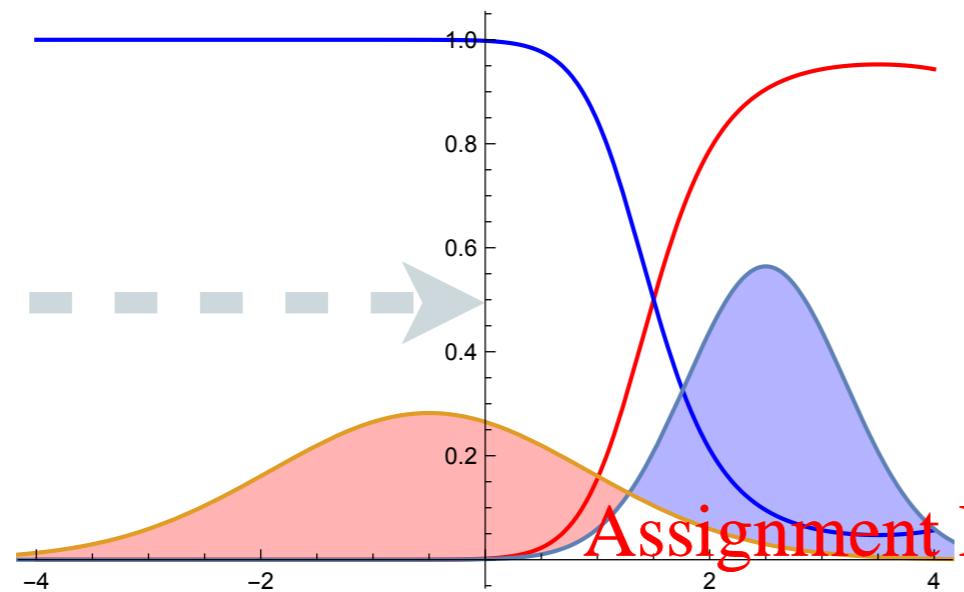
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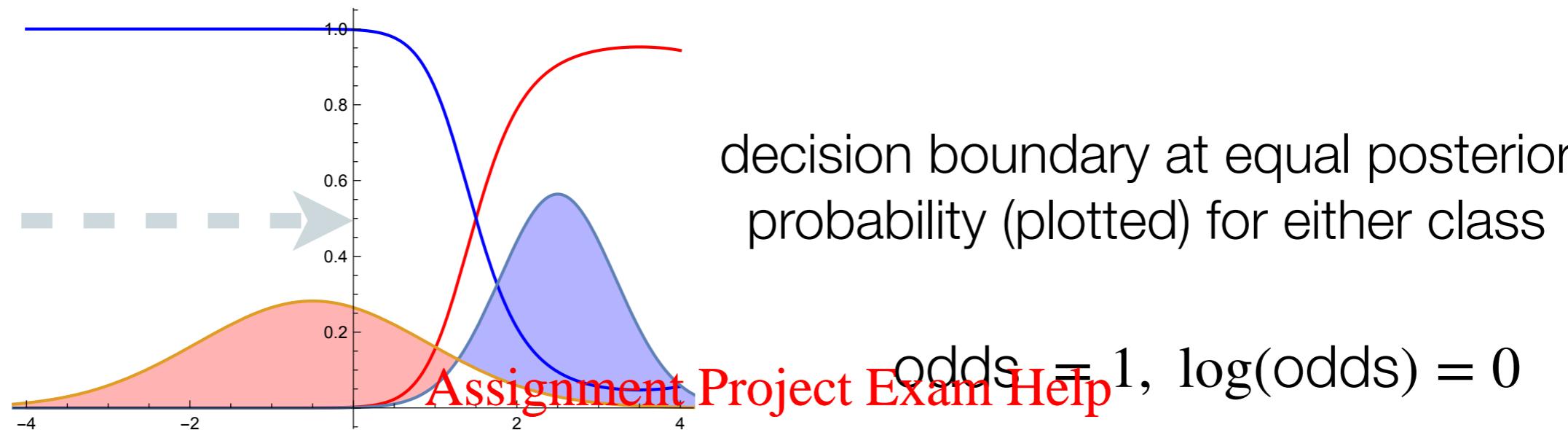
decision boundary at equal posterior probability (plotted) for either class

odds = 1, $\log(\text{odds}) = 0$

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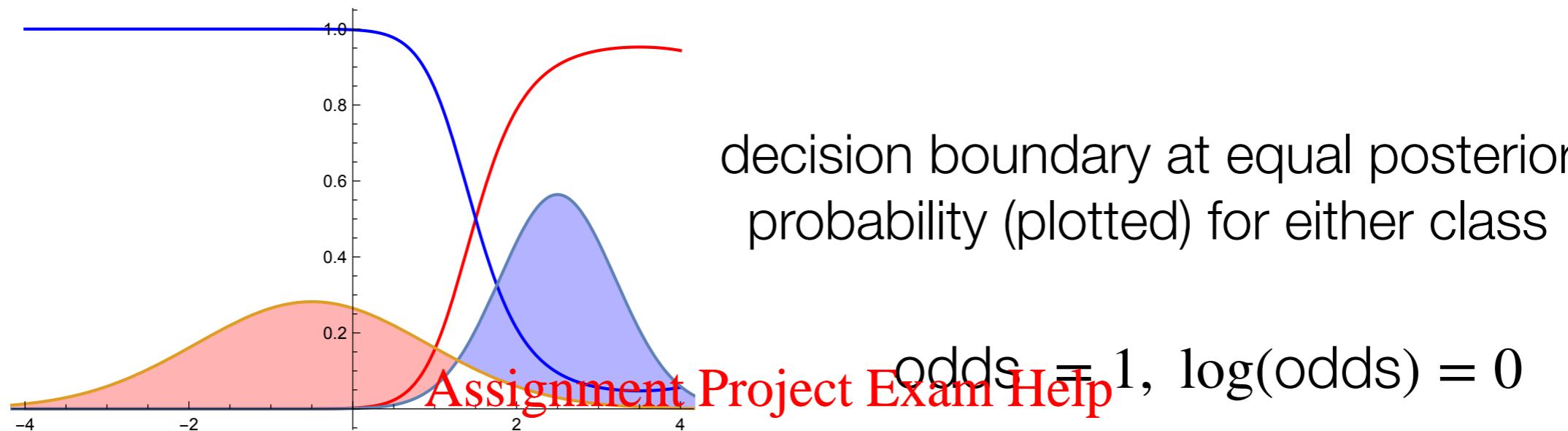
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Decision boundary: determine which class a given input belongs to



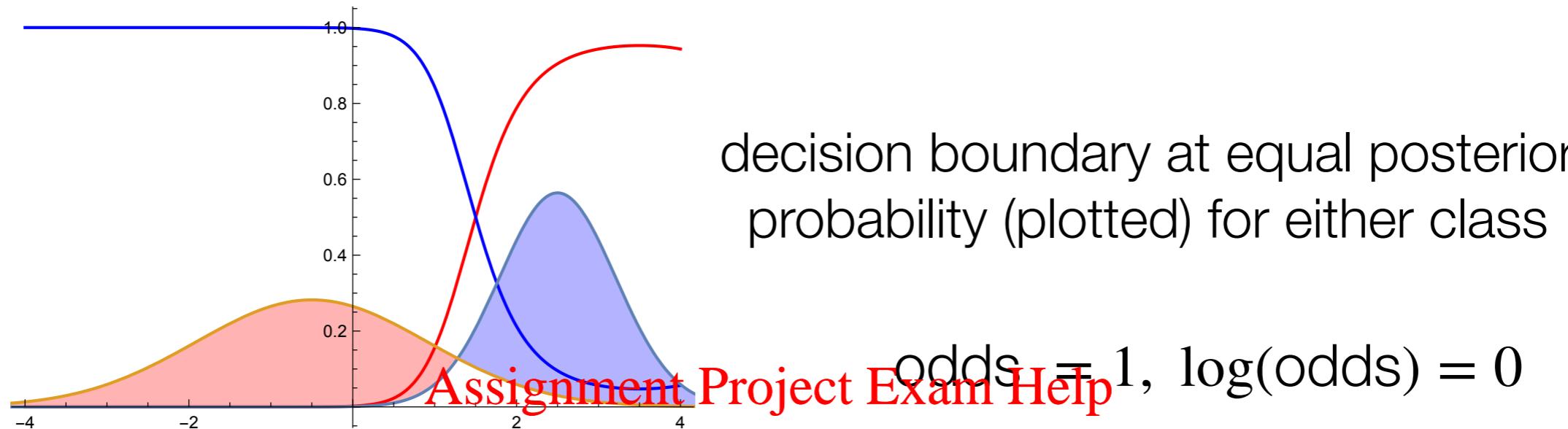
$$\text{odds} = \frac{P(C_+ | X=x)}{P(C_- | X=x)} = \exp [g_+(x) - g_-(x)]$$

Decision boundary: determine which class a given input belongs to



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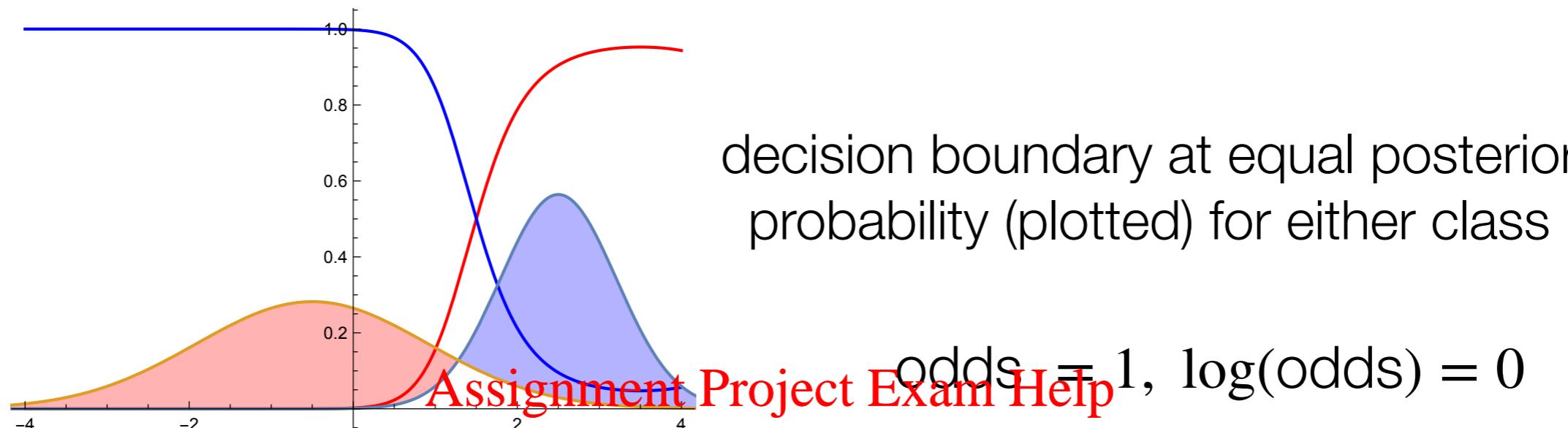
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$$\text{odds} = \frac{P(C=+|X=x)}{P(C=-|X=x)} = \exp [g_+(x) - g_-(x)]$$

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Decision boundary: determine which class a given input belongs to



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$$\text{log-odds} = \ln \left(\frac{P(X=x|C=+)}{P(X=x|C=-)} \right) + \ln \left(\frac{P(C=+)}{P(C=-)} \right)$$

data dependent
(likelihood)

data independent
(prior)

Decision boundaries from probability distributions: zero
log-odds: $0 = g_+(\mathbf{x}) - g_-(\mathbf{x}) = f(\mathbf{x}; \theta) + \theta_0$

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Decision boundaries from probability distributions: zero log-odds: $0 = g_+(\mathbf{x}) - g_-(\mathbf{x}) = f(\mathbf{x}; \theta) + \theta_0$

- If we choose linear functions:

$$g_+(\mathbf{x}) = w_{+0} + \mathbf{w}_+^\top \mathbf{x}, g_-(\mathbf{x}) = w_{-0} + \mathbf{w}_-^\top \mathbf{x}$$

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- Then $f(\mathbf{x}; \theta) = \theta_0 + \theta^\top \mathbf{x}$ is a linear decision boundary

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- where $\theta_0 = w_{+0} - w_{-0}$, and $\theta = \mathbf{w}_+ - \mathbf{w}_-$
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- where $\theta_0 = w_{+0} - w_{-0}$, and $\theta = \mathbf{w}_+ - \mathbf{w}_-$
- This is 2-class softmax (logistic) regression

Reminder: Logistic (2-class)/softmax (K-class) regression

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Reminder: Logistic (2-class)/softmax (K-class) regression

- Choose suitable features Φ to compute scores/activations:
 $s(A_i) := \text{score(class } A_i) = f_i(\Phi(\text{data}))$

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- Recall,

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$$\text{softmax}(s(A_1), \dots, s(A_K)) = \frac{\exp(s(A_1))}{\sum_{k=1}^K \exp(s(A_k))}$$

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$$\text{softmax}(s(A_1), \dots, s(A_K)) = \frac{1}{\sum_{i=1}^K \exp(s(A_i))} [\exp(s(A_1)) + \dots + \exp(s(A_K))]$$

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$$P(\text{class } A_i \mid \text{data}) = \frac{\exp(s(A_i))}{\exp \text{softmax}(s(A_1), \dots, s(A_K))} = \frac{\exp(s(A_i))}{\sum_{i=1}^K \exp(s(A_i))}$$

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- $f_i(\text{data } x) = \mathbf{w}_i^\top \Phi(x) = \mathbf{w}_i^\top x$
- Did not seek a decision boundary, just an argmax

Decision boundaries from probability distributions: zero log-odds

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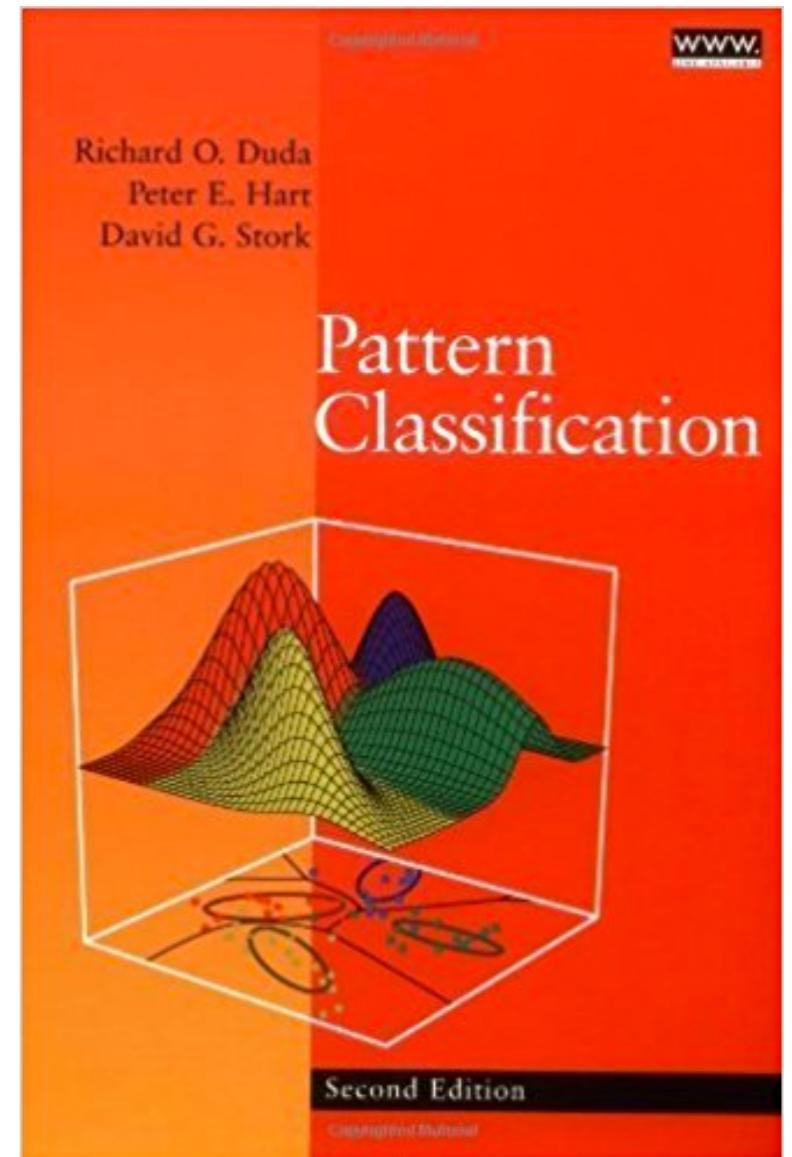
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Decision boundaries from probability distributions: zero log-odds

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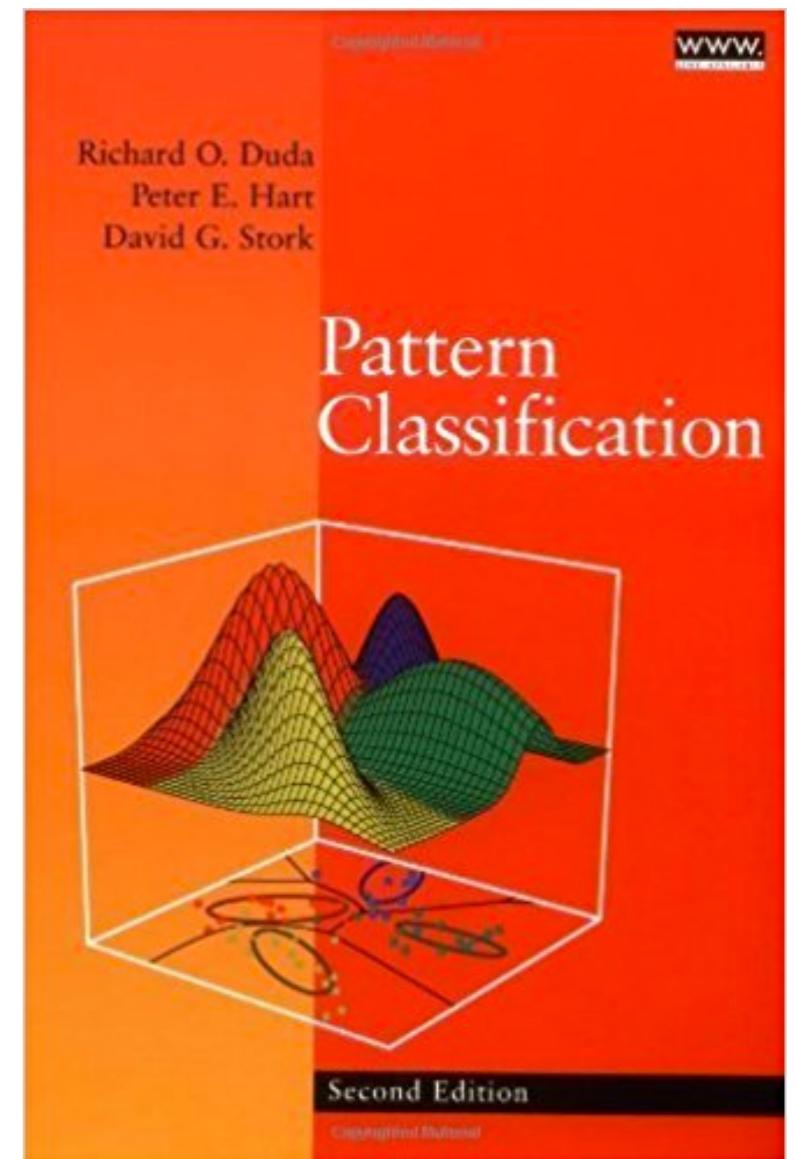


Decision boundaries from probability distributions: zero log-odds

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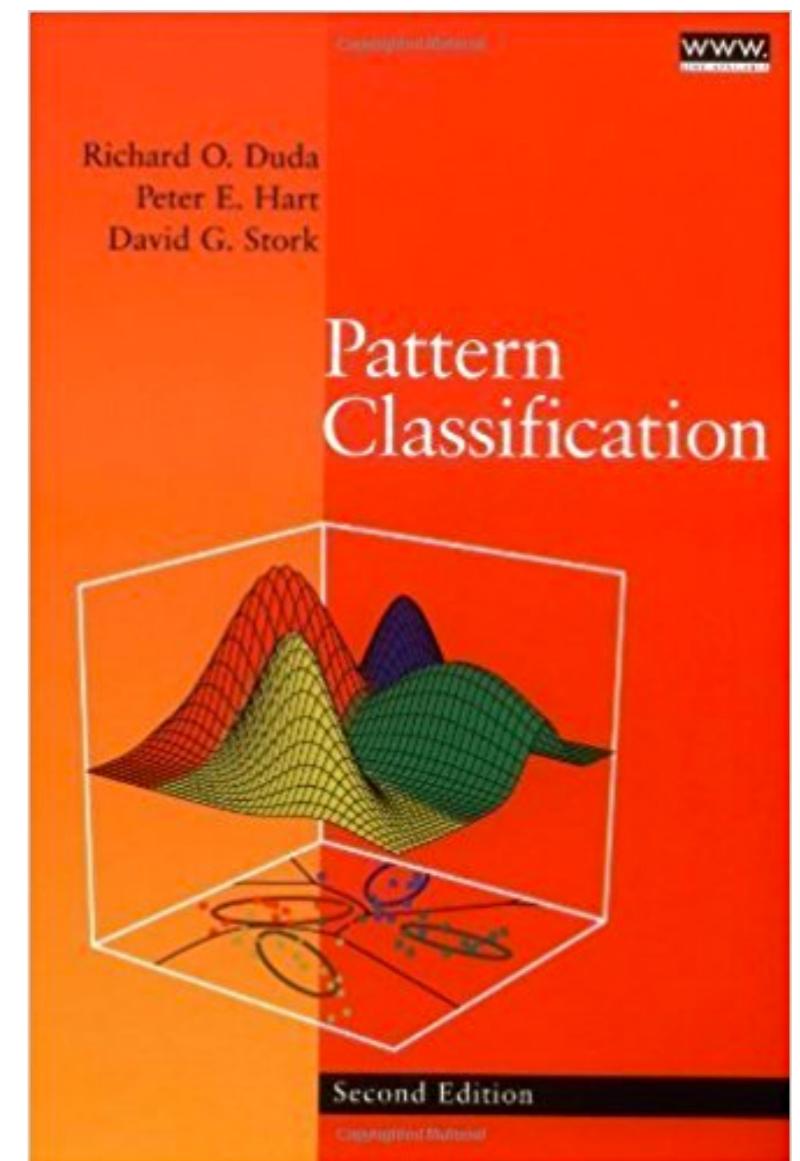
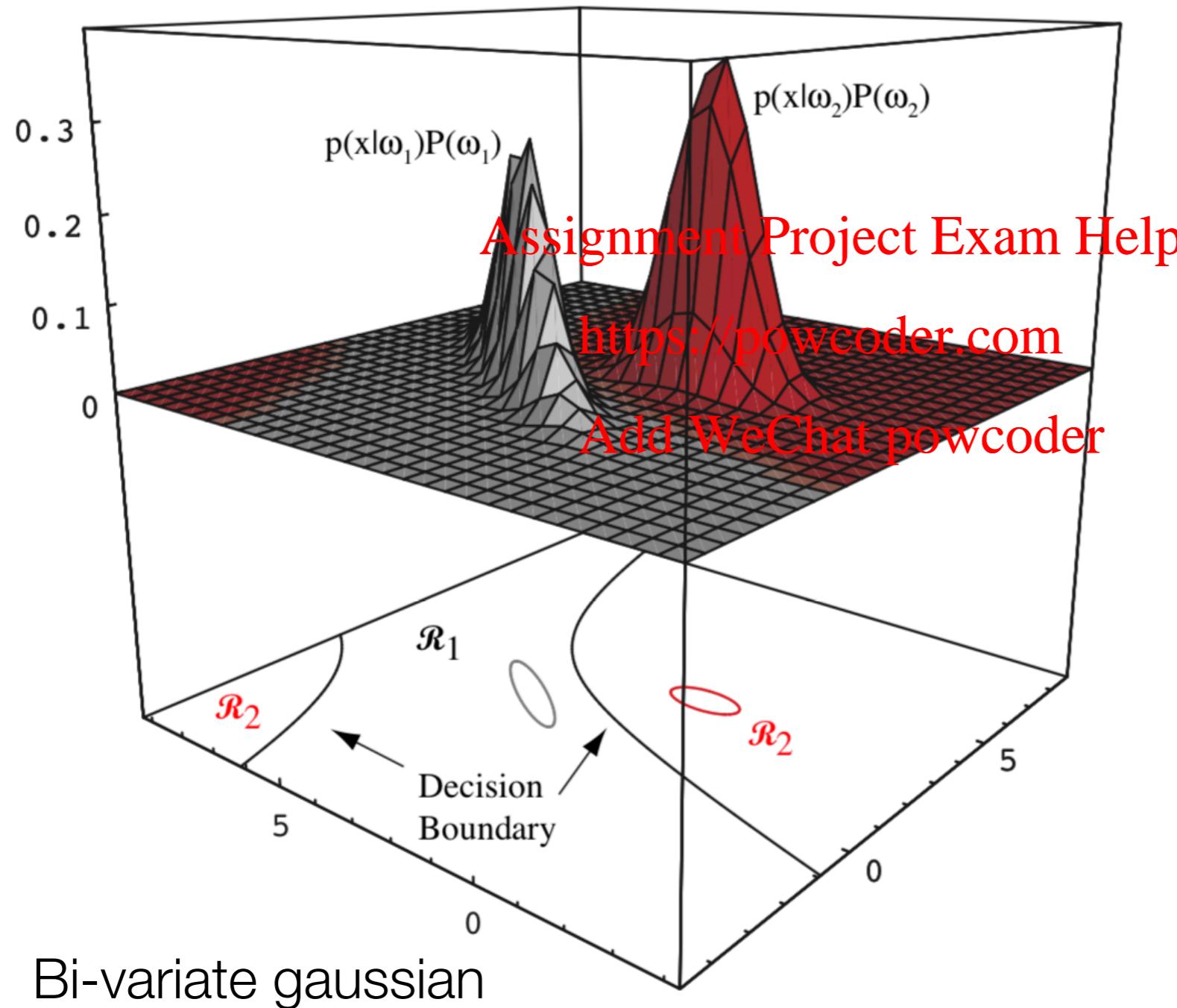
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Notice the contours and
line segments at the
bottom of the cover graphic

Decision boundaries from probability distributions: zero log-odds



Notice the contours and line segments at the bottom of the cover graphic

Log of ratio of two multi-dimensional Gaussians

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Log of ratio of two multi-dimensional Gaussians

$$p(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^p |\boldsymbol{\Sigma}|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)$$

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Precision matrix $\Lambda = \boldsymbol{\Sigma}^{-1}$
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For 2D features

$$p(x_1, x_2 | C = a) = \frac{\sqrt{|\Lambda_a|}}{2\pi} \exp \left\{ \frac{1}{2} [x_1 - \mu_{a1} \ x_2 - \mu_{a2}] \begin{bmatrix} \Lambda_{a11} & \Lambda_{a12} \\ \Lambda_{a21} & \Lambda_{a22} \end{bmatrix} \begin{bmatrix} x_1 - \mu_{a1} \\ x_2 - \mu_{a2} \end{bmatrix} \right\}$$

$$p(x_1, x_2 | C = b) = \frac{\sqrt{|\Lambda_b|}}{2\pi} \exp \left\{ -\frac{1}{2} [x_1 - \mu_{b1} \ x_2 - \mu_{b2}] \begin{bmatrix} \Lambda_{b11} & \Lambda_{b12} \\ \Lambda_{b21} & \Lambda_{b22} \end{bmatrix} \begin{bmatrix} x_1 - \mu_{b1} \\ x_2 - \mu_{b2} \end{bmatrix} \right\}$$

Precision matrix $\Lambda = \boldsymbol{\Sigma}^{-1}$
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Log of ratio of two multi-dimensional Gaussians

$$p(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^p |\boldsymbol{\Sigma}|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)$$

Precision matrix $\boldsymbol{\Lambda} = \boldsymbol{\Sigma}^{-1}$

For 2D features

$$p(x_1, x_2 | C = a) = \frac{\sqrt{|\boldsymbol{\Lambda}_a|}}{2\pi} \exp \left\{ \frac{1}{2} [x_1 - \mu_{a1} \ x_2 - \mu_{a2}] \begin{bmatrix} \Lambda_{a11} & \Lambda_{a12} \\ \Lambda_{a21} & \Lambda_{a22} \end{bmatrix} \begin{bmatrix} x_1 - \mu_{a1} \\ x_2 - \mu_{a2} \end{bmatrix} \right\}$$

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Log-odds (no prior) for 2-class classification

$$\ln \left(\frac{p(\mathbf{x}|C_+)}{p(\mathbf{x}|C_-)} \right) = \ln \left(\frac{|\boldsymbol{\Lambda}_+|}{|\boldsymbol{\Lambda}_-|} \right) + \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_-)^T \boldsymbol{\Lambda}_- (\mathbf{x} - \boldsymbol{\mu}_-) - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_+)^T \boldsymbol{\Lambda}_+ (\mathbf{x} - \boldsymbol{\mu}_+)$$

Log of ratio of posteriors of two Gaussians: for equal and unequal covariances

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Log of ratio of posteriors of two Gaussians: for equal and unequal covariances

$$\begin{aligned}g_-(\mathbf{x}) &= \ln P(C_-) + \frac{1}{2} \ln |\Lambda_-| - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_-)^T \Lambda_- (\mathbf{x} - \boldsymbol{\mu}_-) \\&= \ln P(C_-) + \frac{1}{2} \ln |\Lambda_-| - \frac{1}{2} (\mathbf{x}^T \Lambda_- \mathbf{x} - 2\boldsymbol{\mu}_-^T \Lambda_- \mathbf{x} + \boldsymbol{\mu}_-^T \Lambda_- \boldsymbol{\mu}_-) \\g_+(\mathbf{x}) &= \ln P(C_+) + \frac{1}{2} \ln |\Lambda_+| - \frac{1}{2} (\mathbf{x}^T \Lambda_+ \mathbf{x} - 2\boldsymbol{\mu}_+^T \Lambda_+ \mathbf{x} + \boldsymbol{\mu}_+^T \Lambda_+ \boldsymbol{\mu}_+)\end{aligned}$$

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For unequal covariances, $\Lambda_+ \neq \Lambda_-$ decision boundary is quadratic
Your turn: what is the quadratic term in $g_+(\mathbf{x}) - g_-(\mathbf{x})$

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Log of ratio of posteriors of two Gaussians: for equal and unequal covariances

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Decision boundary for equal covariances $\Lambda_+ = \Lambda_- = \Lambda$ is linear:

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$$g_+(\mathbf{x}) - g_-(\mathbf{x}) = \ln \frac{P(C_+)}{P(C_-)} + (\boldsymbol{\mu}_+ - \boldsymbol{\mu}_-)^T \Lambda \mathbf{x} + \mathbf{x}\text{-independent term}$$

Log of ratio of posteriors of two Gaussians: for equal and unequal covariances

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Your turn: derive the linear equation

Discriminants (decision boundaries) for 2-class gaussians:
linear if covariances equal, quadratic if not

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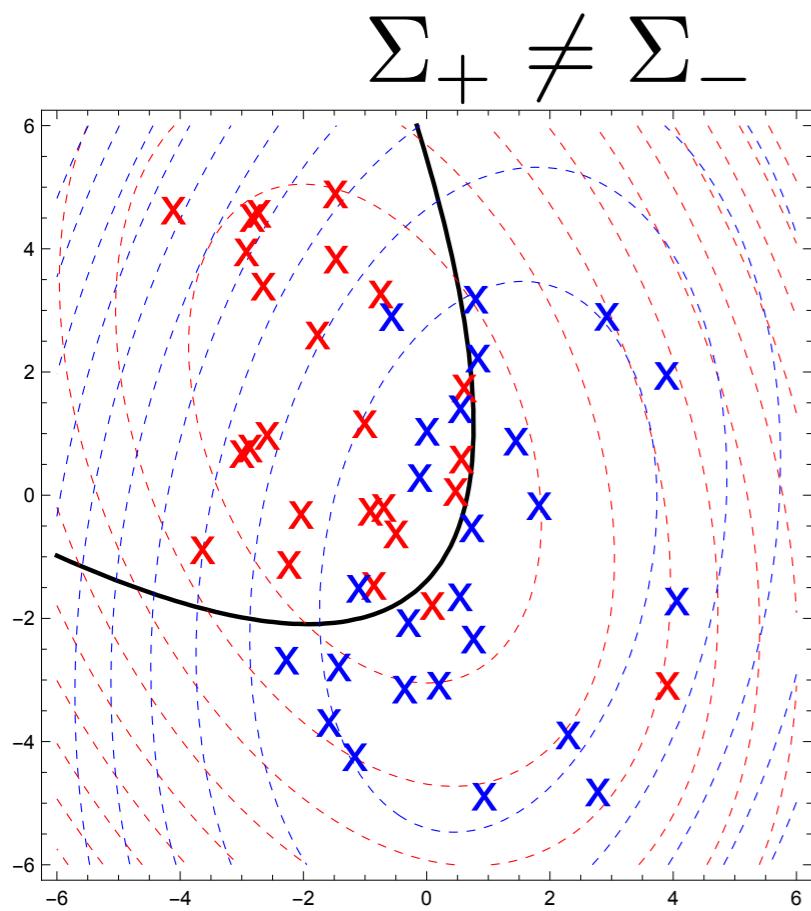
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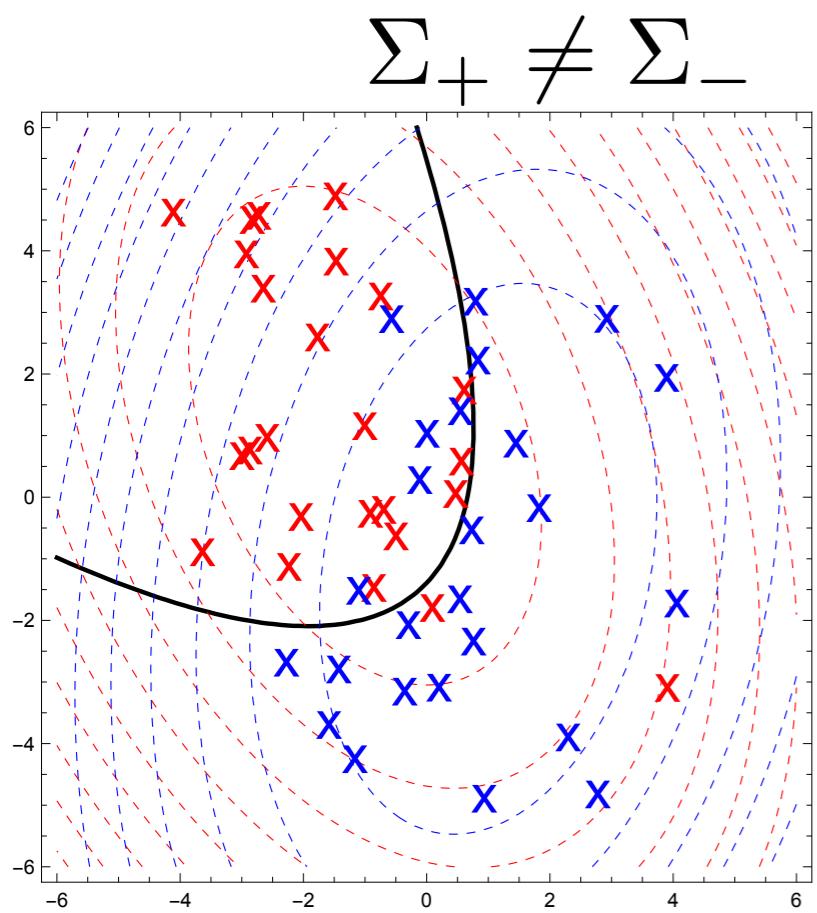


Discriminants (decision boundaries) for 2-class gaussians:
linear if covariances equal, quadratic if not

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If the covariance matrices are the same, the quadratic term is common to both classes, and drops out in the ratio

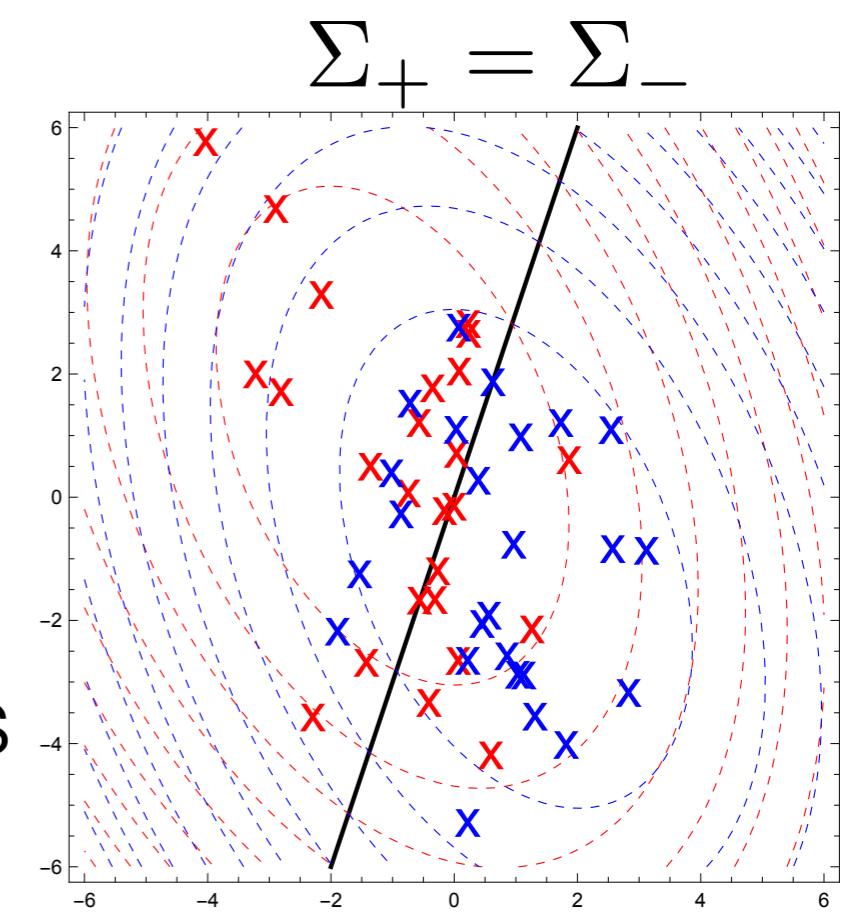
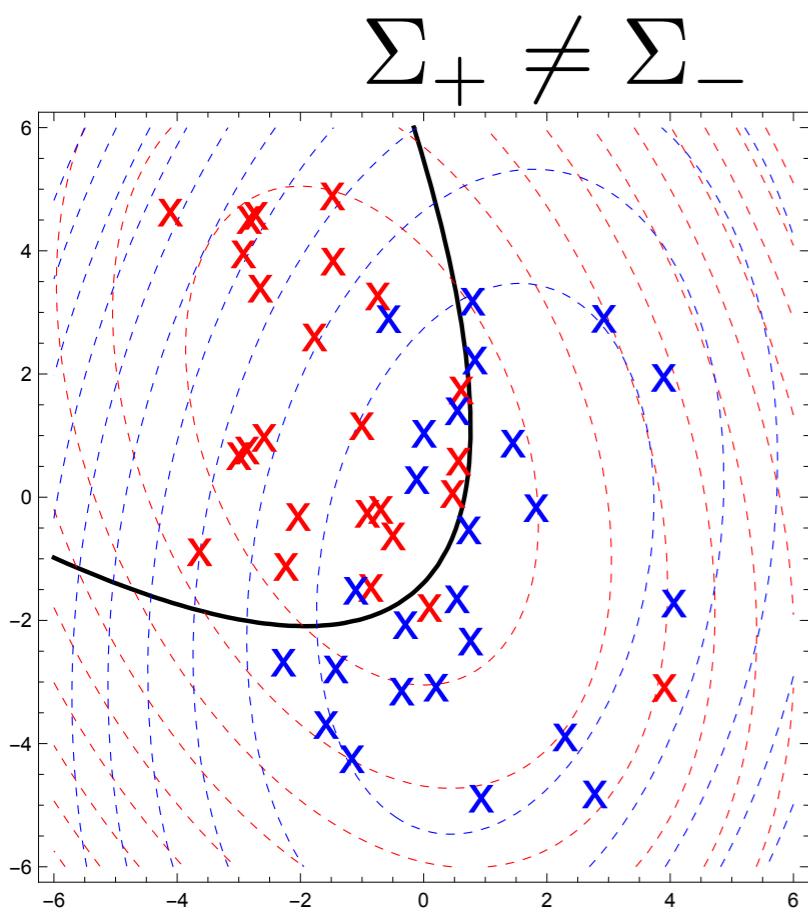


Discriminants (decision boundaries) for 2-class gaussians:
linear if covariances equal, quadratic if not

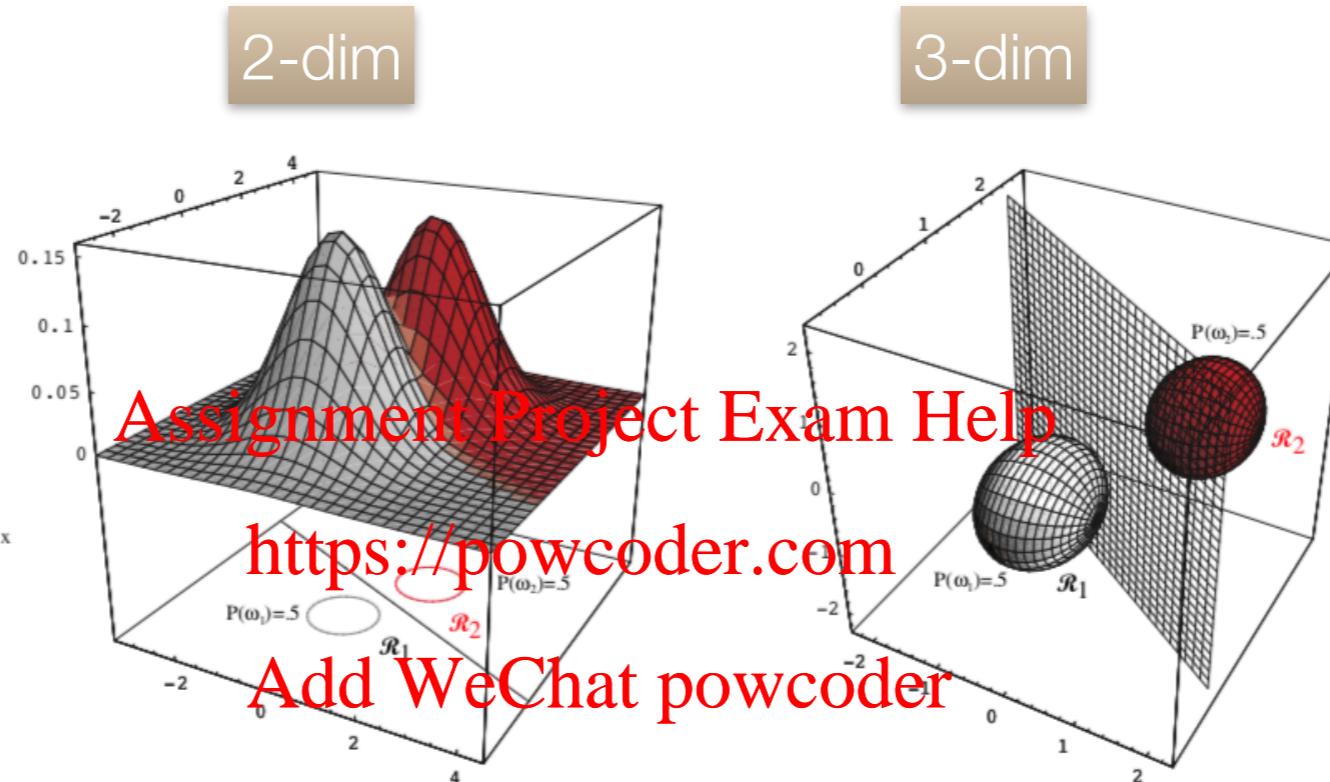
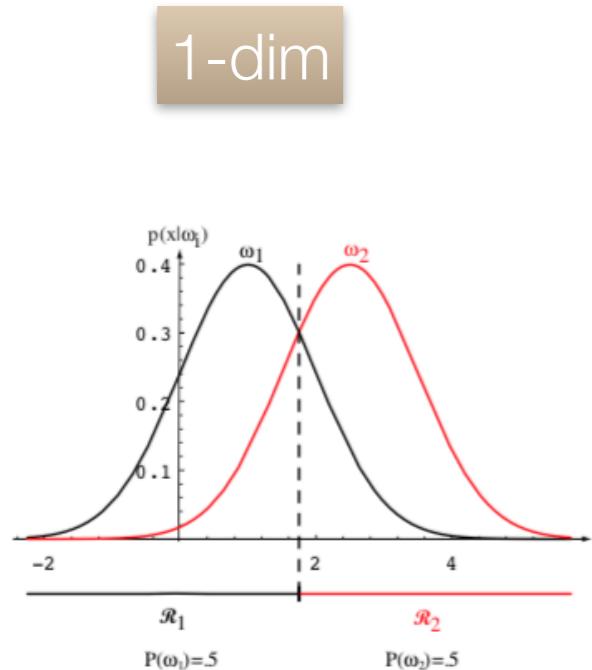
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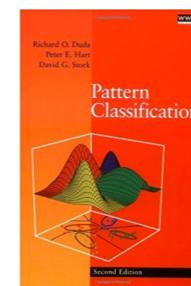
If the covariance matrices are the same, the quadratic term is common to both classes, and drops out in the ratio



Discriminants for Gaussians for equal priors: data distribution equally divided

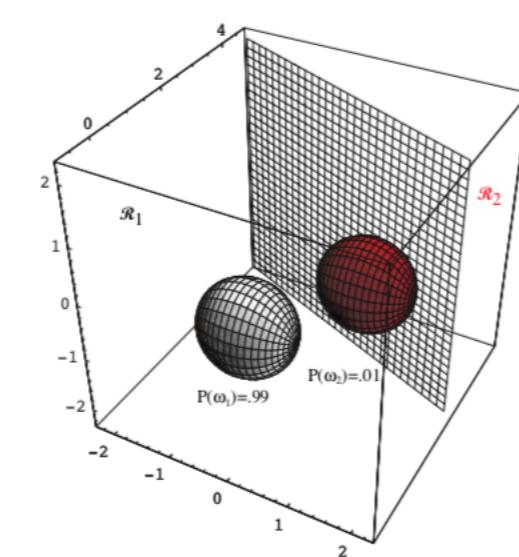
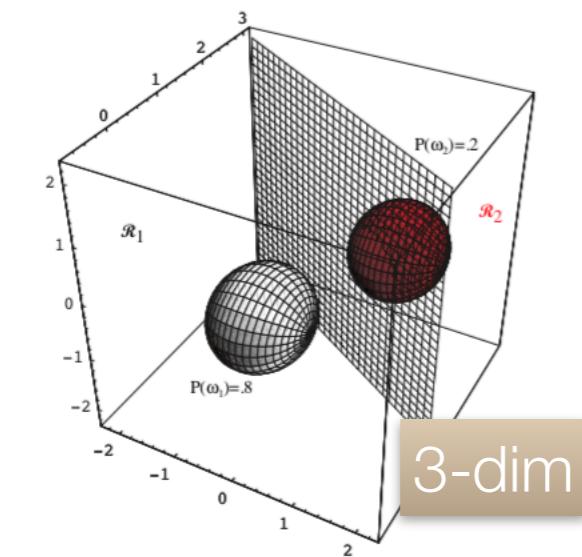
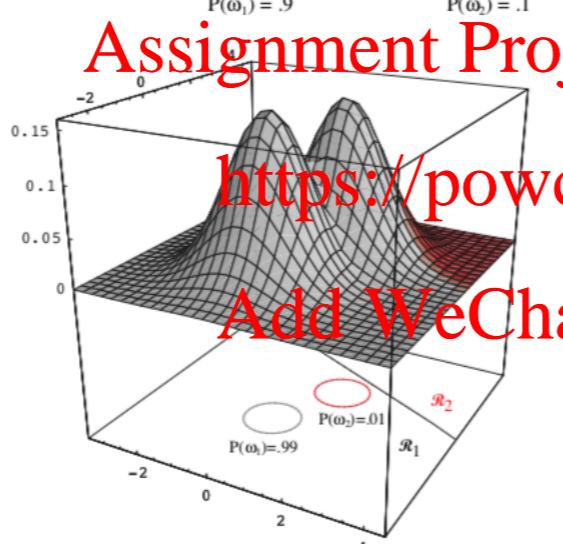
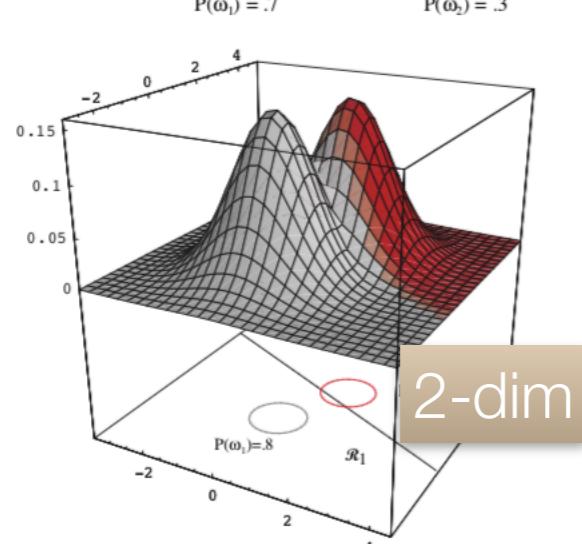
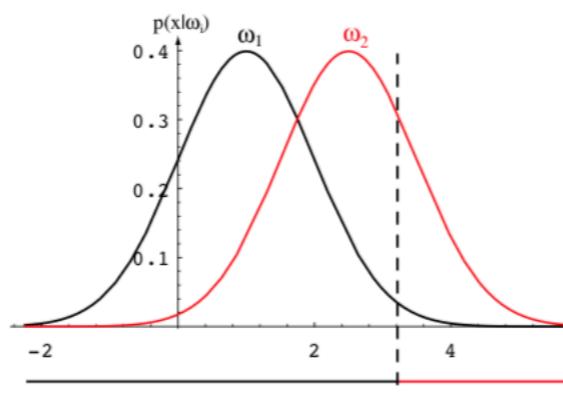
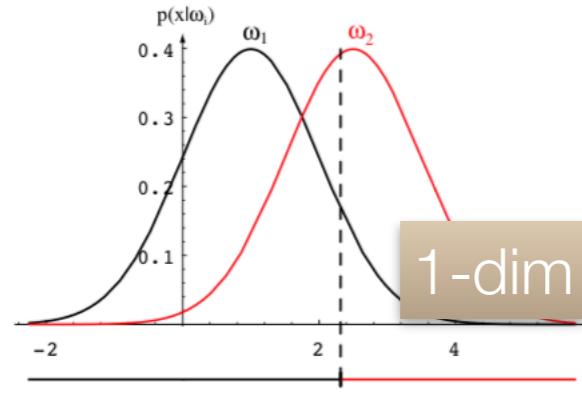


Equal priors: $\log(P_a/P_b) = 0$



From Duda and Hart

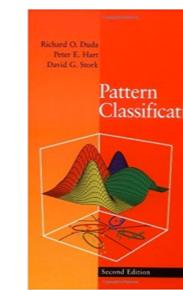
Discriminants for Gaussians: unequal priors shifts decision boundary by $\log(P_a/P_b)$



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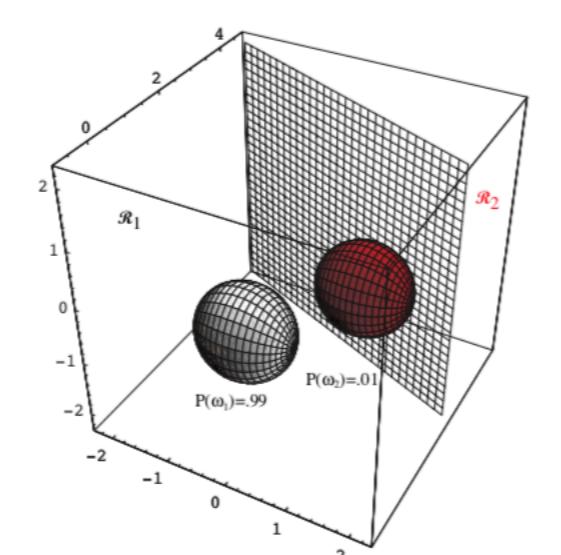
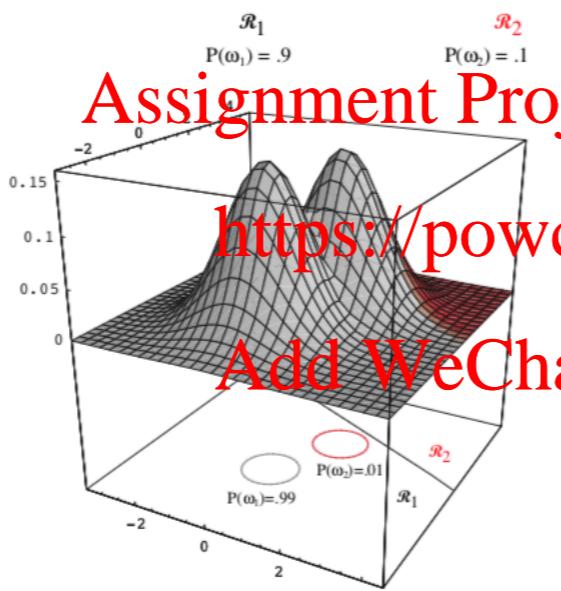
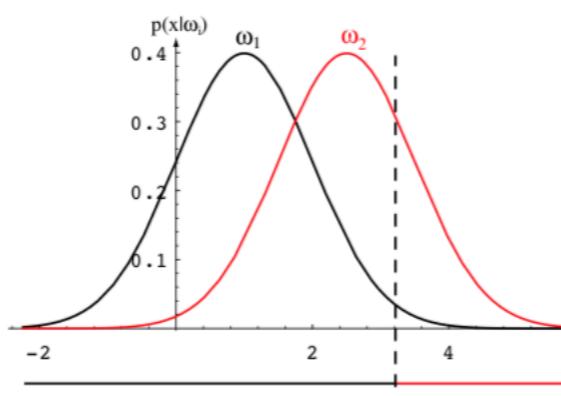
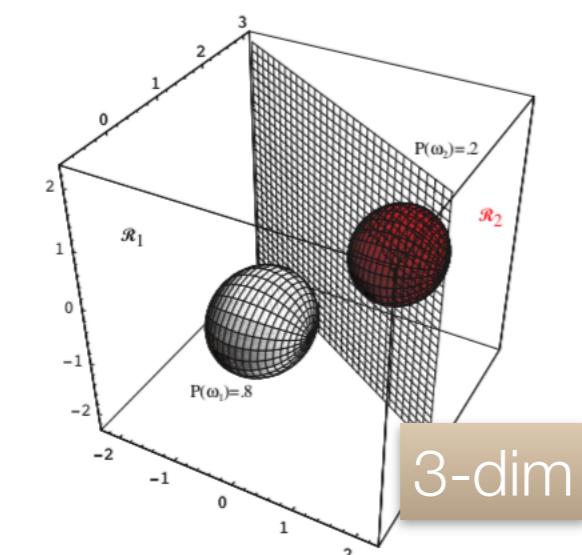
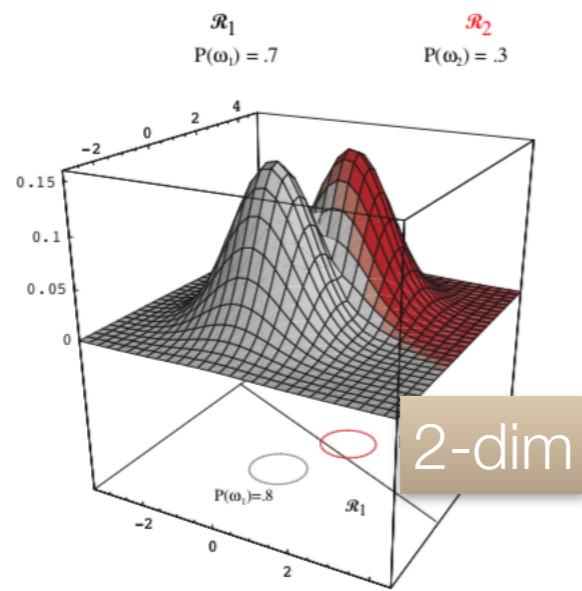
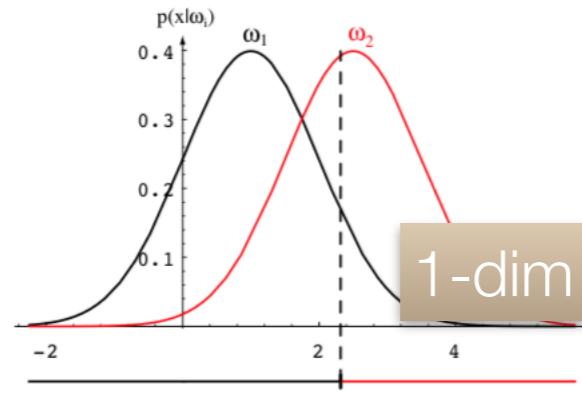
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From Duda and Hart

Discriminants for Gaussians: unequal priors shifts decision boundary by $\log(P_a/P_b)$



Prior overrides data distributions

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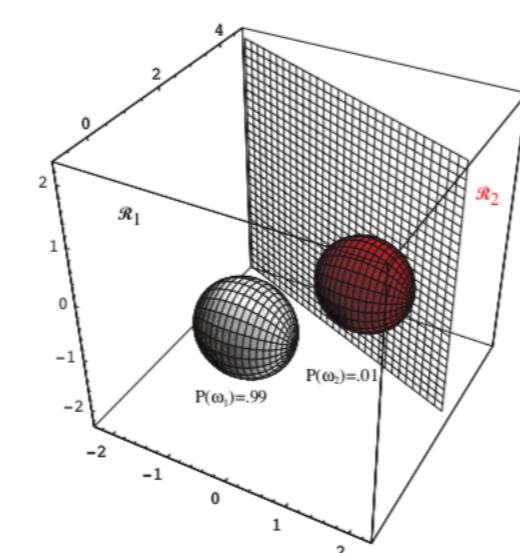
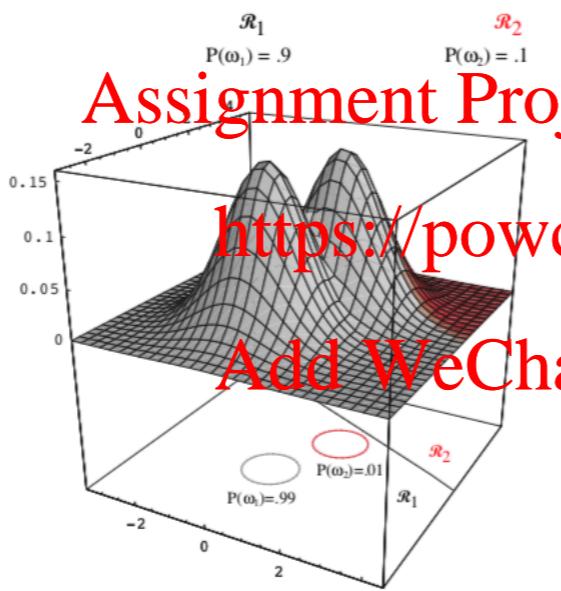
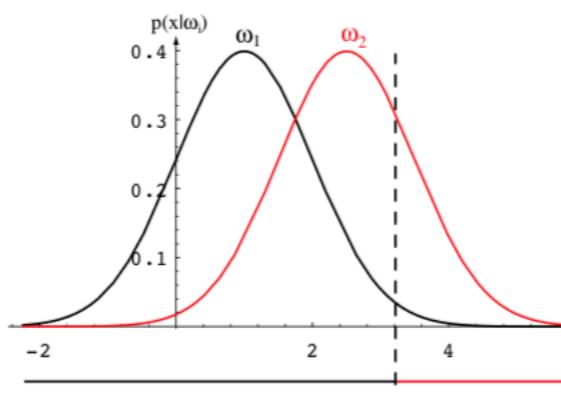
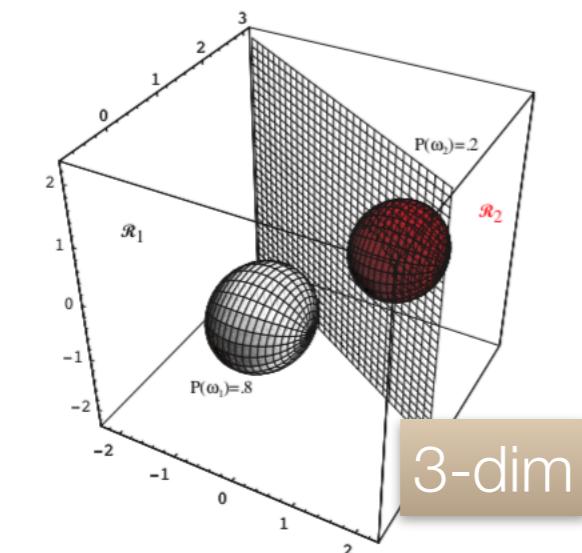
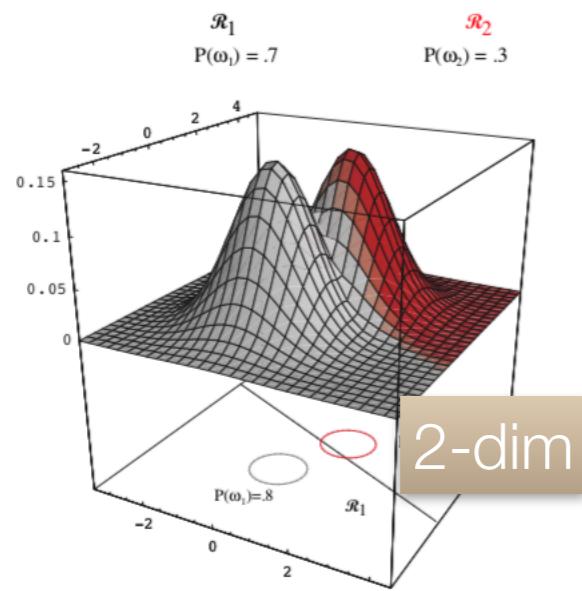
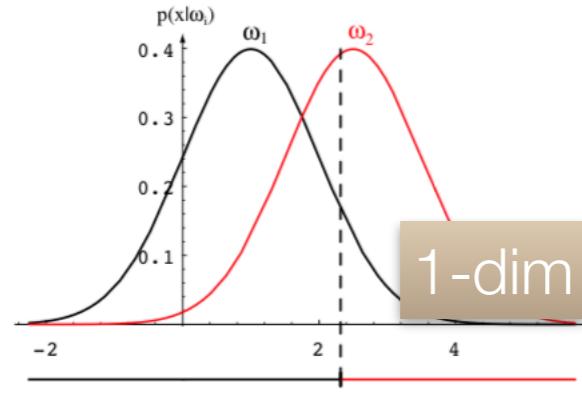
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From Duda and Hart

Discriminants for Gaussians: unequal priors shifts decision boundary by $\log(P_a/P_b)$



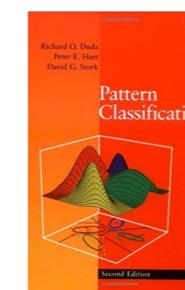
Prior overrides data distributions

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Notice how far boundary shifts

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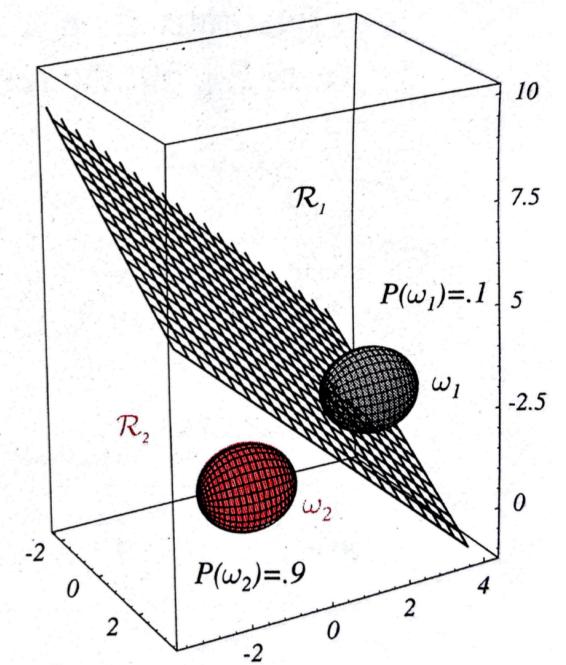
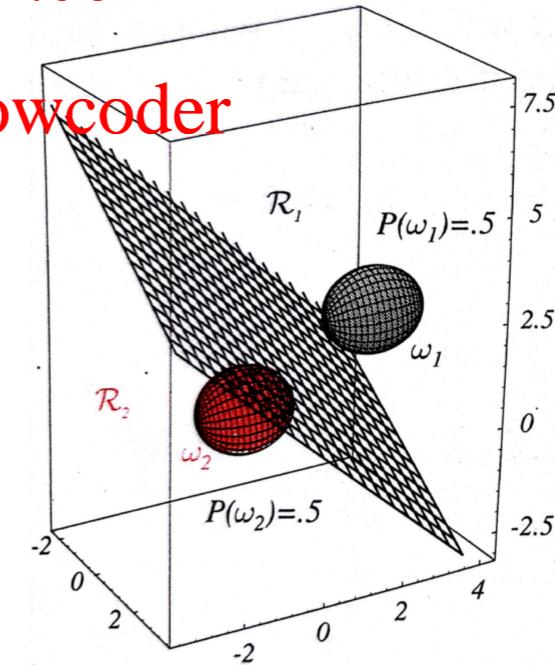
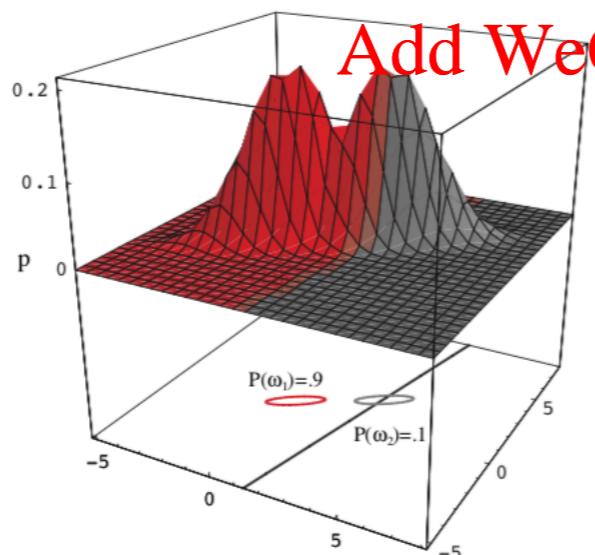
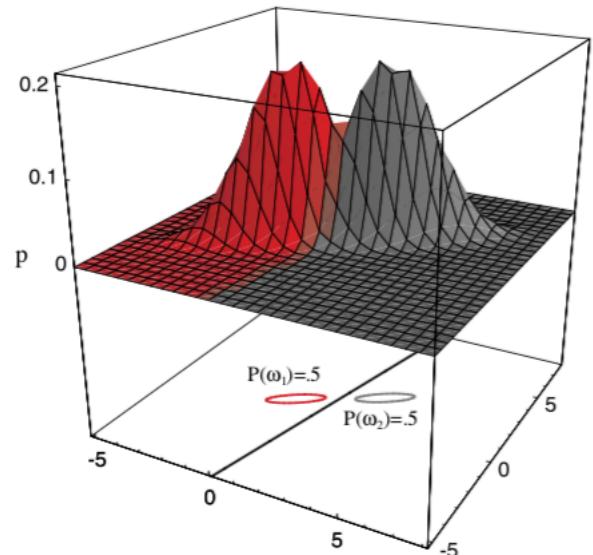


From Duda and Hart

Decision boundaries not perpendicular to line joining means
if covariance matrix not proportional to the identity matrix

$$\omega = \sum_{i=1}^k (\mu_+ - \mu_-)$$

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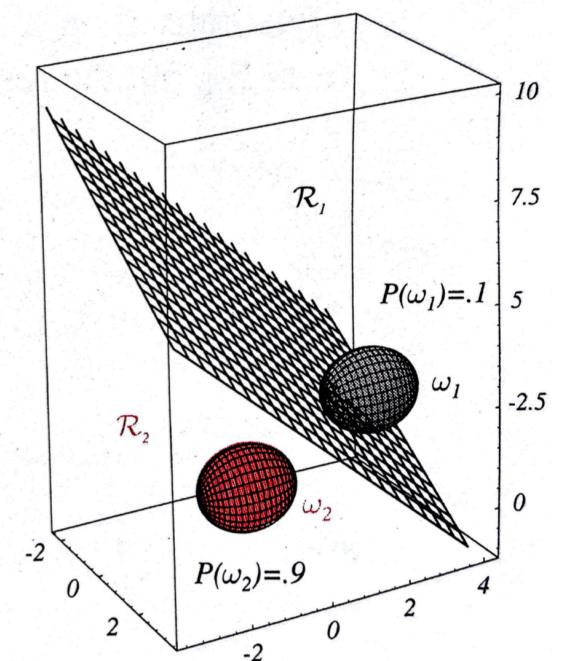
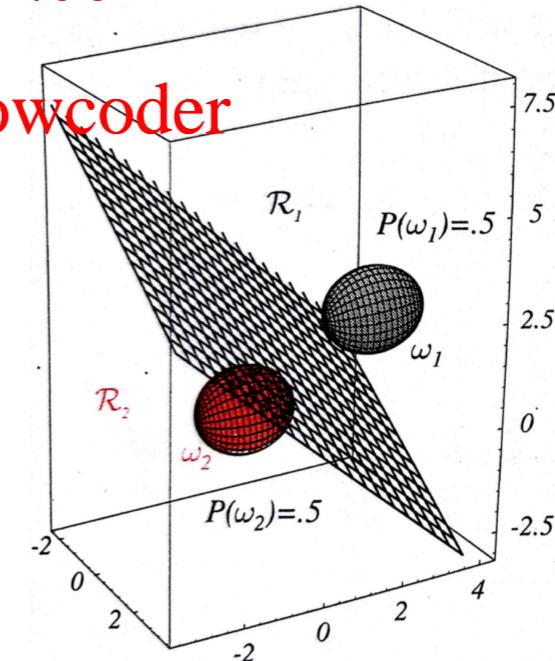
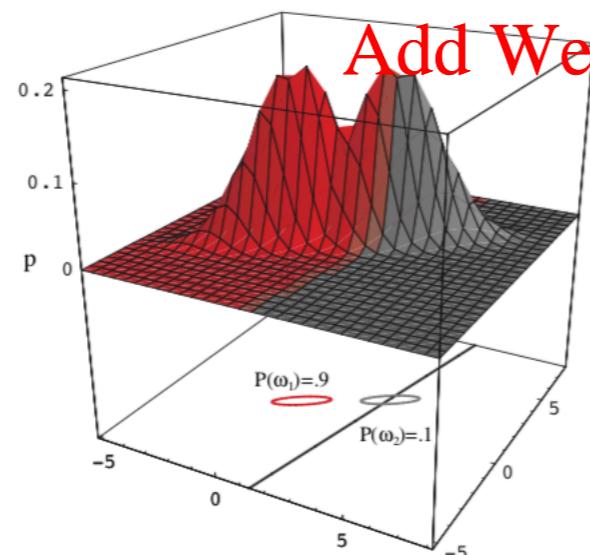
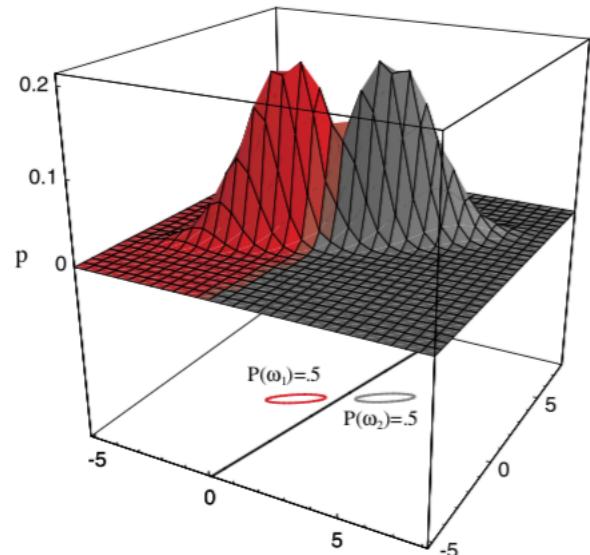
From Duda and Hart

Decision boundaries not perpendicular to line joining means if covariance matrix not proportional to the identity matrix

$$g_+(\mathbf{x}) - g_-(\mathbf{x}) = \ln \frac{P(C_+)}{P(C_-)} + (\boldsymbol{\mu}_+ - \boldsymbol{\mu}_-)^T \boldsymbol{\Lambda} \mathbf{x} + \mathbf{x}\text{-independent term}$$

$\boldsymbol{\omega} = \sum_{i=1}^k (\boldsymbol{\mu}_+ - \boldsymbol{\mu}_-)$

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From Duda and Hart

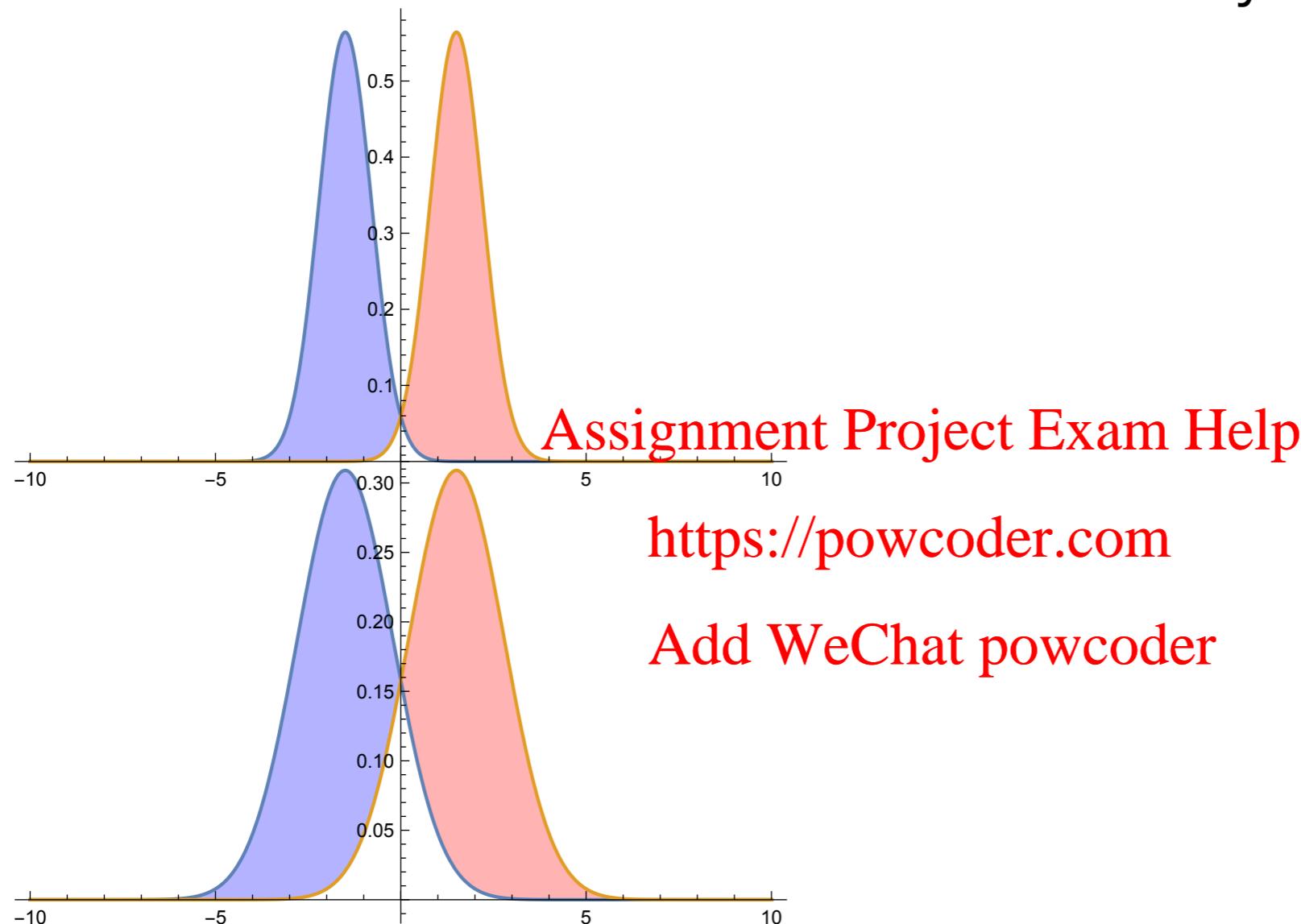
Distinctiveness relative to natural variation measured by distance between class means scaled by standard deviation:

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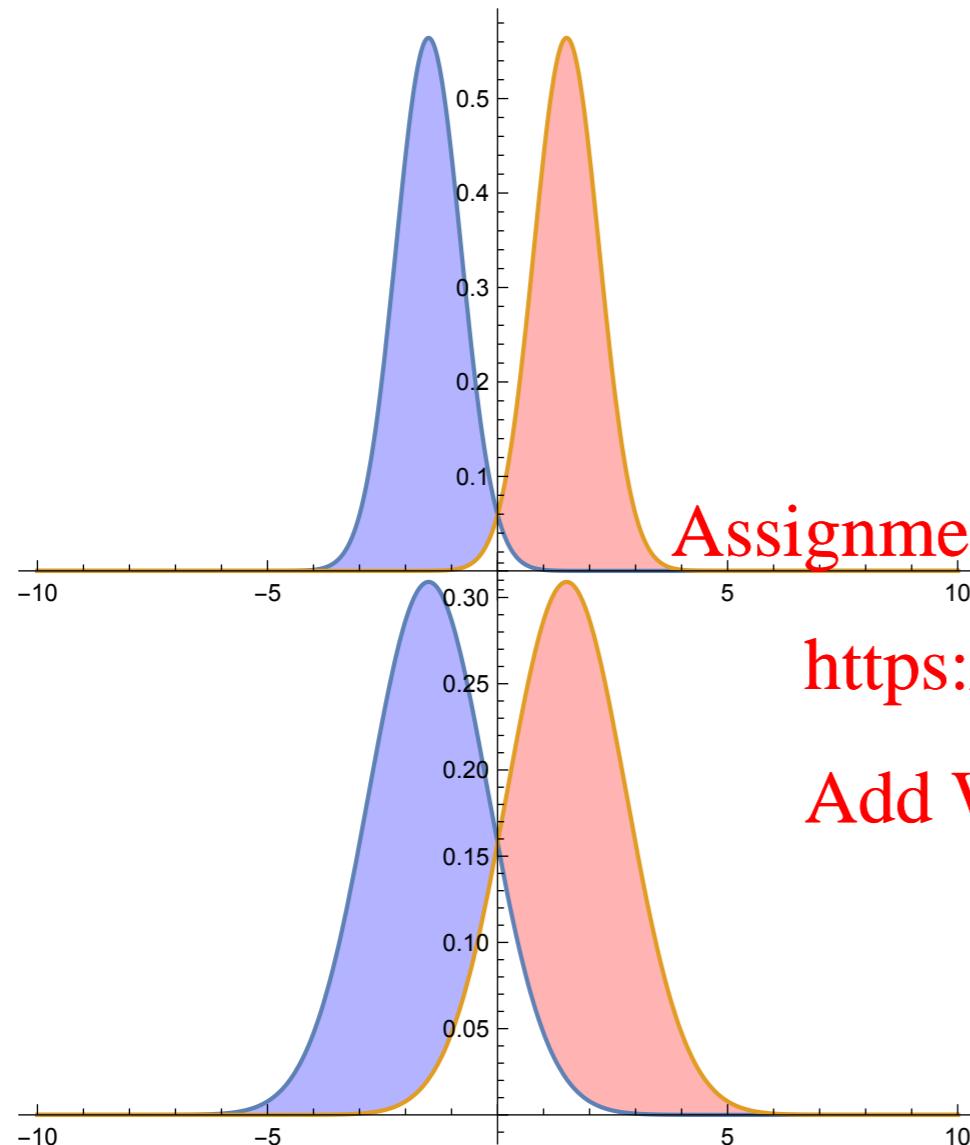
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Distinctiveness relative to natural variation measured by distance between class means scaled by standard deviation:



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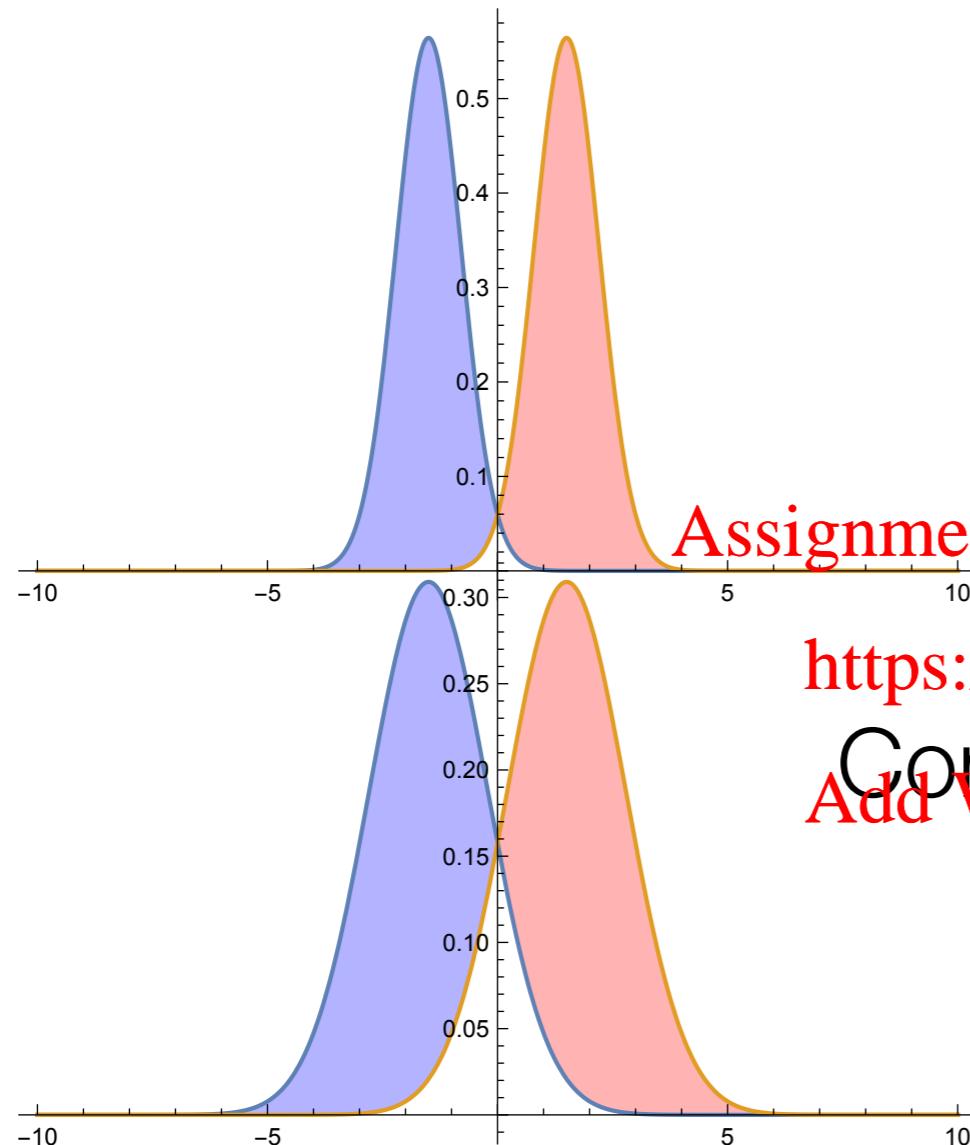
$$d = \frac{\mu_1 - \mu_2}{\sigma}$$

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Distinctiveness relative to natural variation measured by distance between class means scaled by standard deviation:



$$d = \frac{\mu_1 - \mu_2}{\sigma}$$

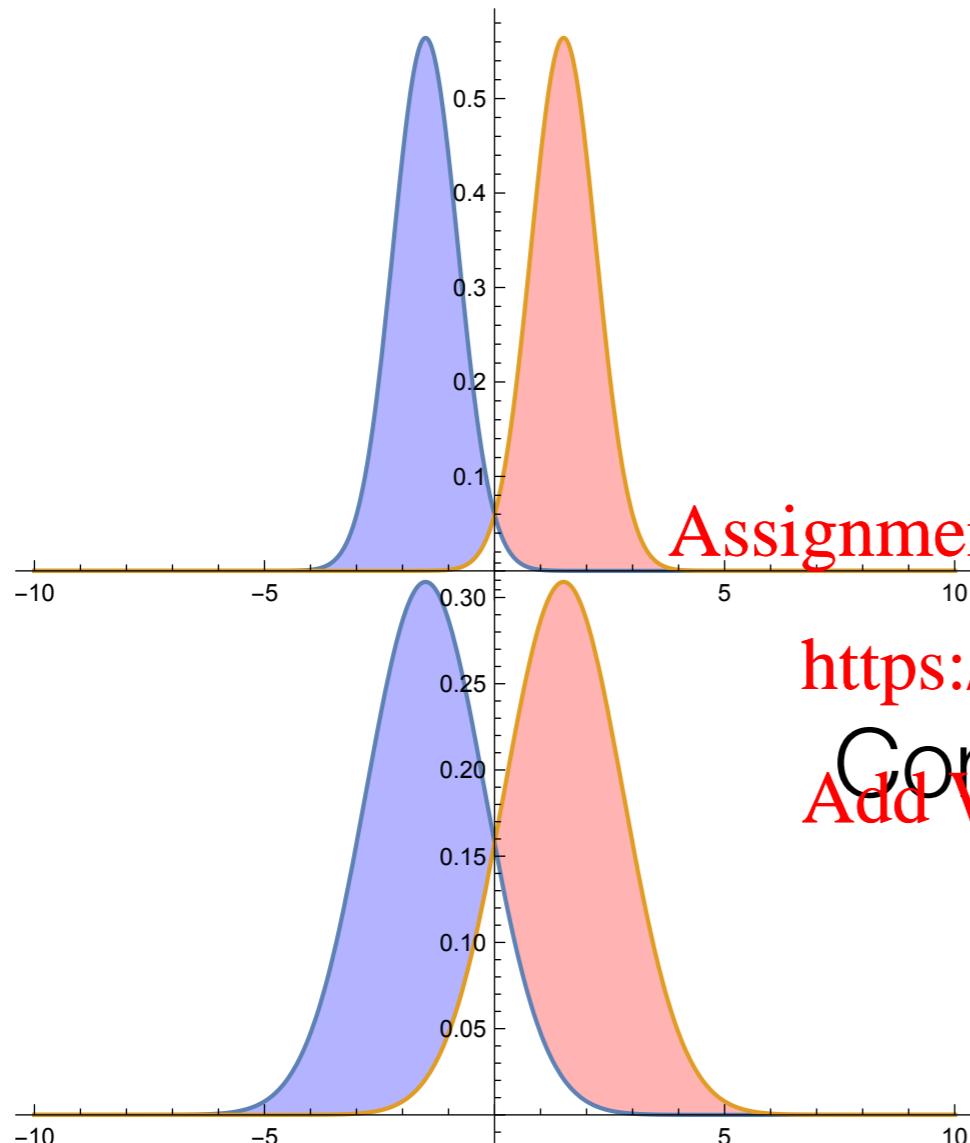
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Controls for overlap between classes

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Distinctiveness relative to natural variation measured by distance between class means scaled by standard deviation:



$$d = \frac{\mu_1 - \mu_2}{\sigma}$$

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Controls for overlap between classes
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Mahalanobis distance

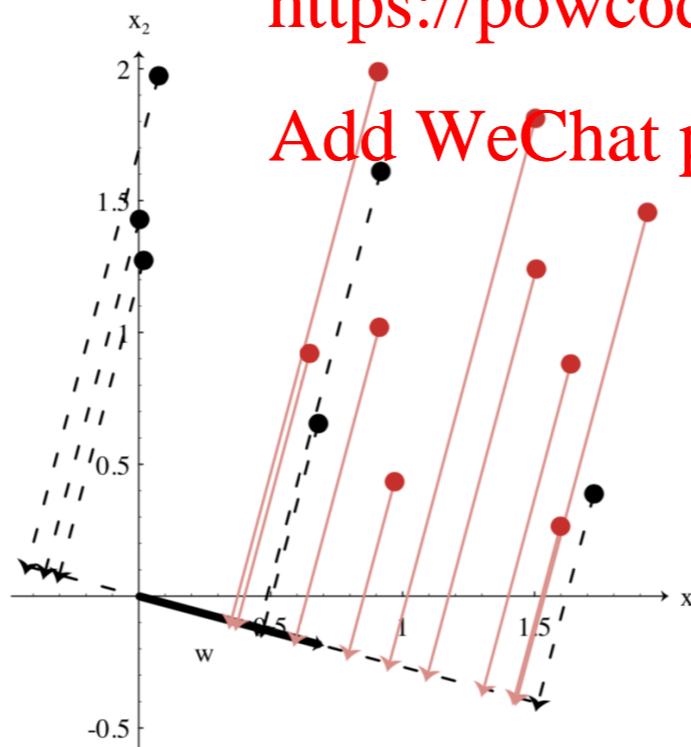
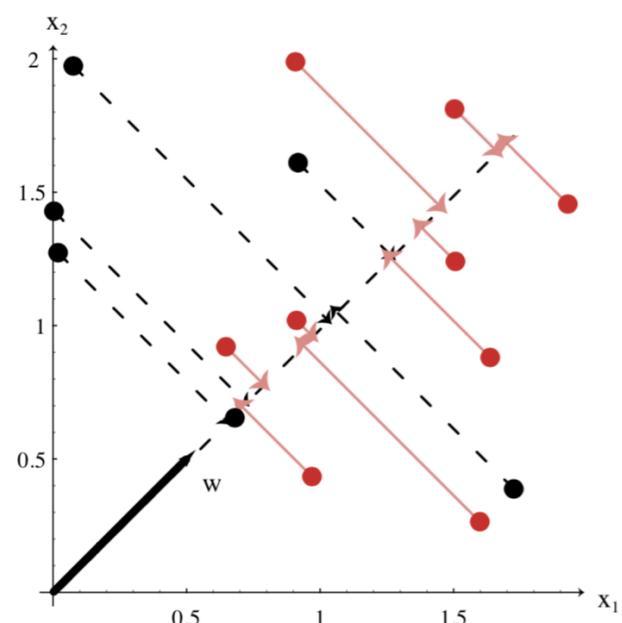
$$d(\mathbf{x}_1, \mathbf{x}_2) = \sqrt{(\mathbf{x}_1 - \mathbf{x}_2)^T \Sigma^{-1} (\mathbf{x}_1 - \mathbf{x}_2)}$$

Fisher Linear Discriminant Analysis: projections along discriminative directions

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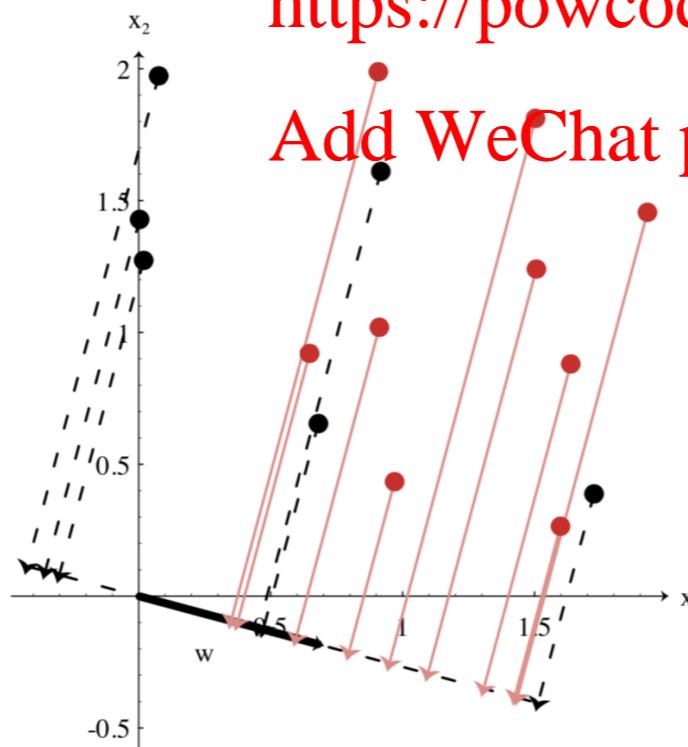
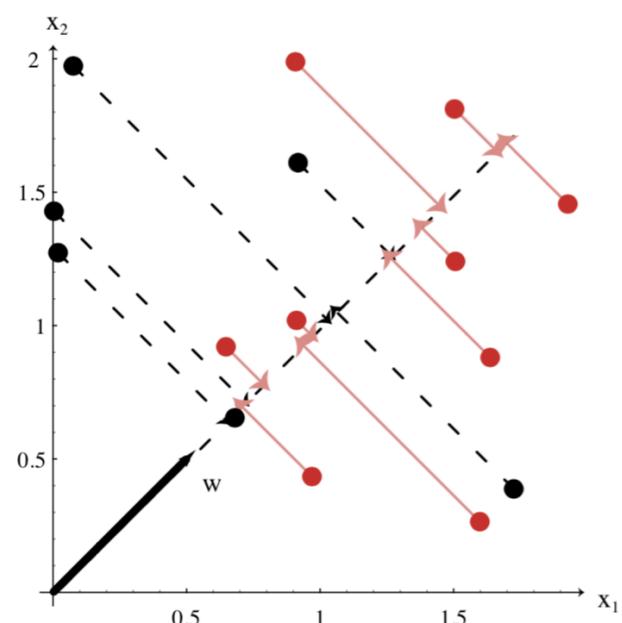


Fisher Linear Discriminant Analysis: projections along discriminative directions

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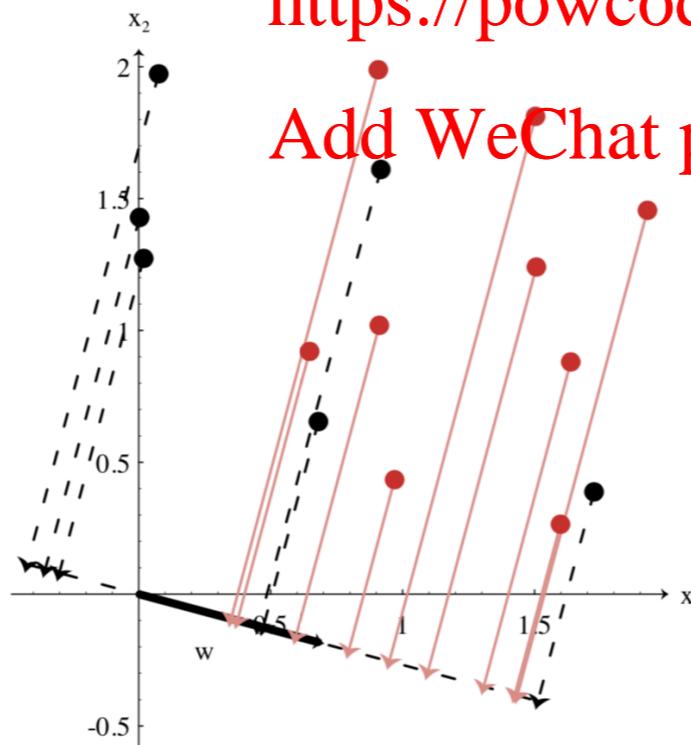
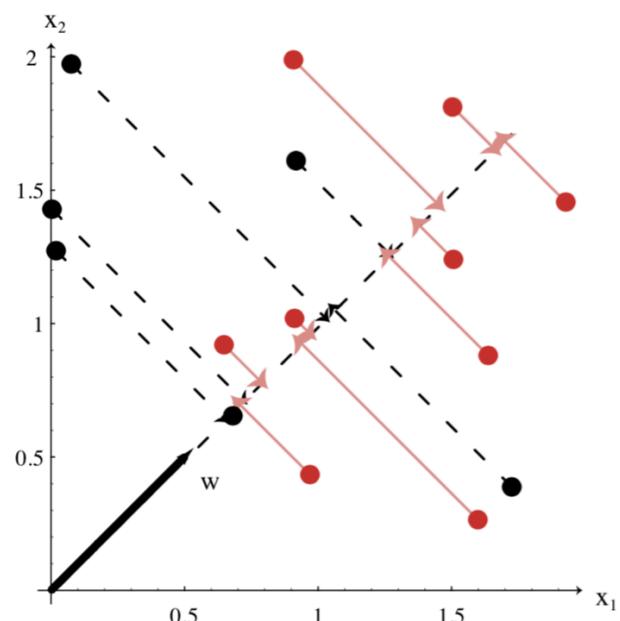
Fisher Linear Discriminant Analysis: projections along discriminative directions

$$\mu_c = \frac{1}{n_c} \sum_{\mathbf{x}^n \in \mathcal{D}_c} \mathbf{x}^n$$

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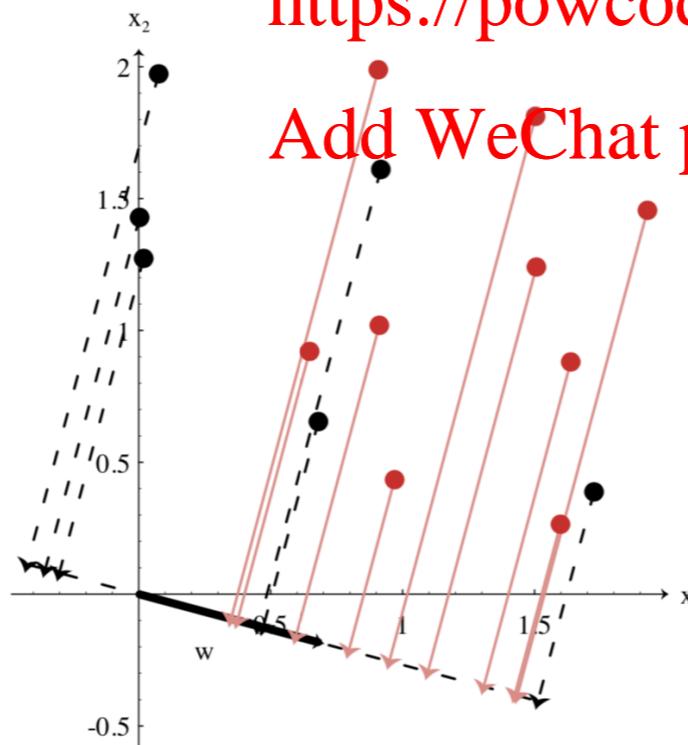
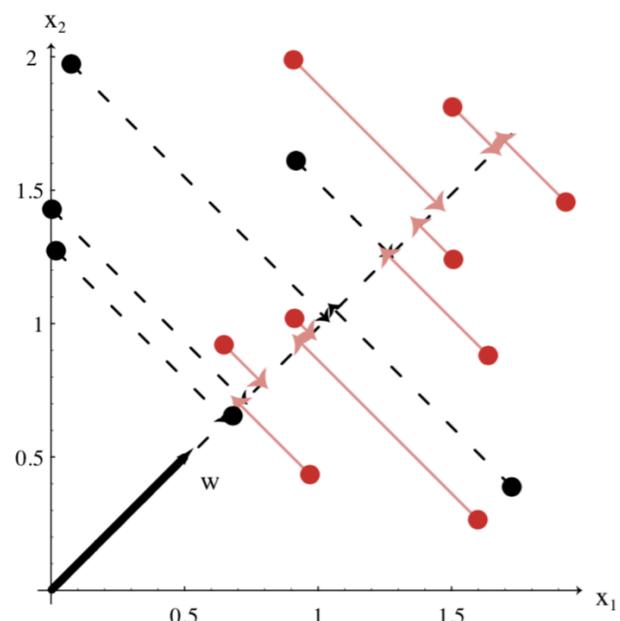
Fisher Linear Discriminant Analysis: projections along discriminative directions

$$\mu_c = \frac{1}{n_c} \sum_{\mathbf{x}^n \in \mathcal{D}_c} \mathbf{x}^n \quad m_c = \mathbf{w}^T \mu_c = \frac{1}{n_c} \sum_{\mathbf{x}^n \in \mathcal{D}_c} \mathbf{w}^T \mathbf{x}^n$$

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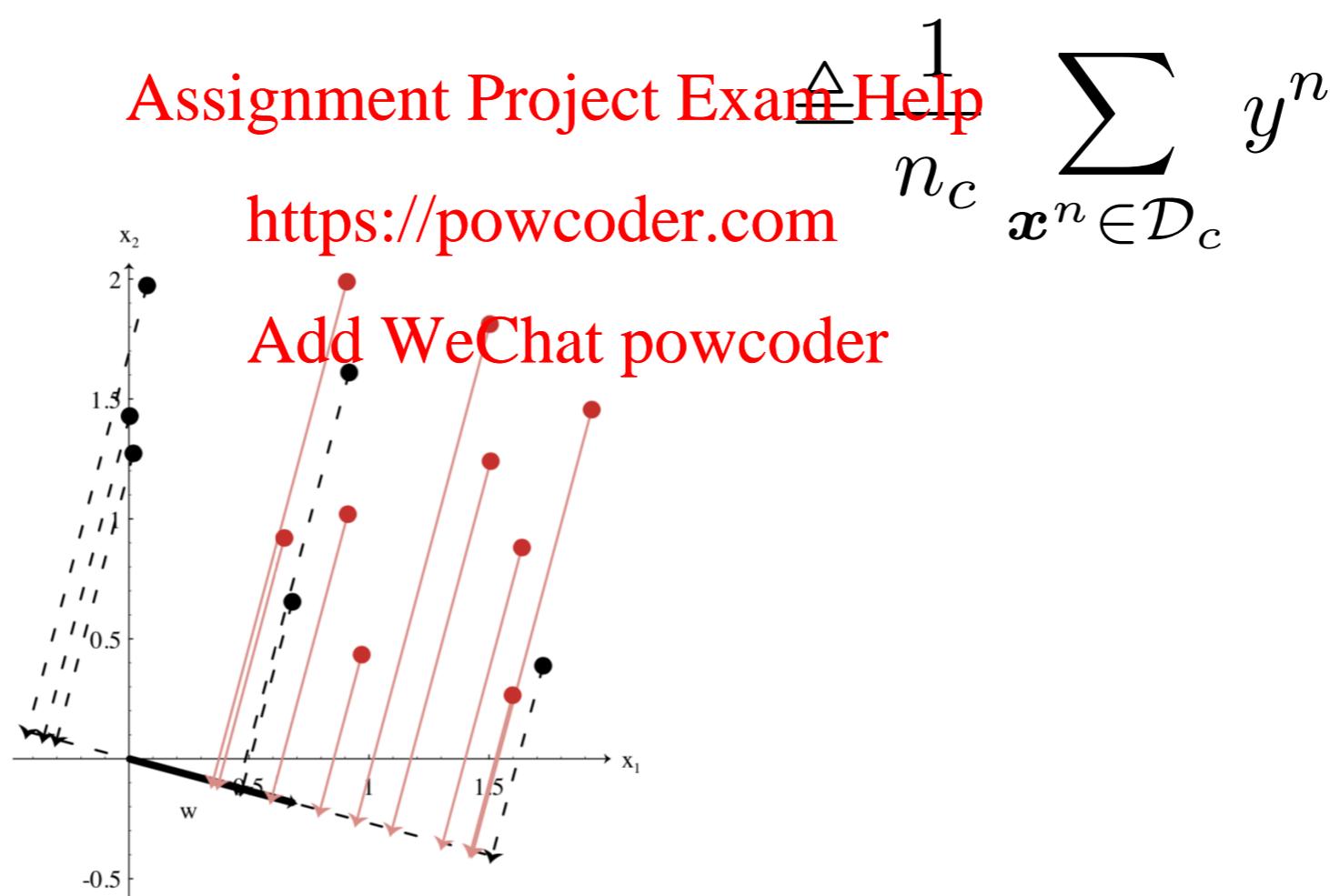
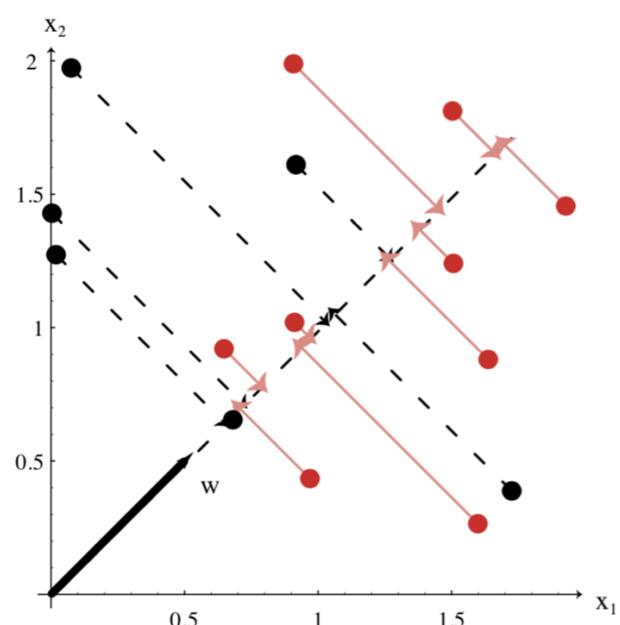
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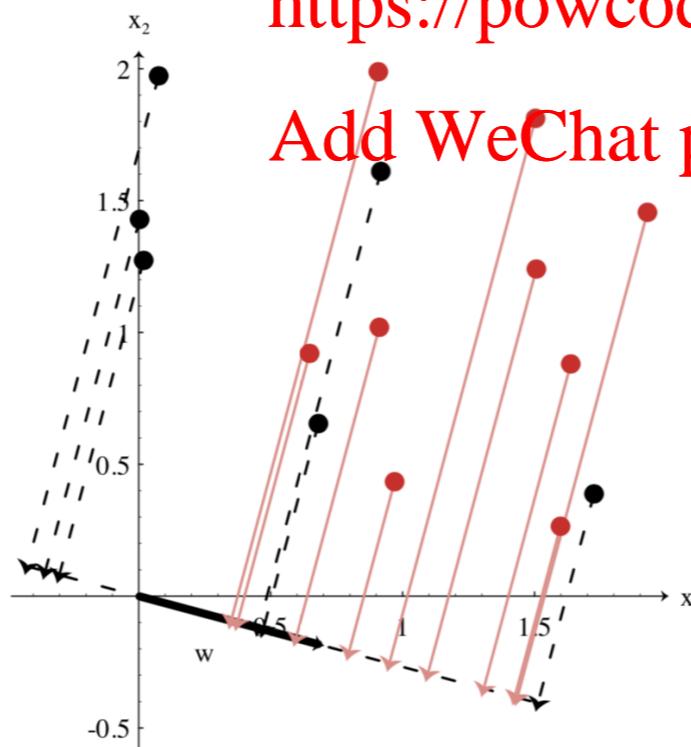
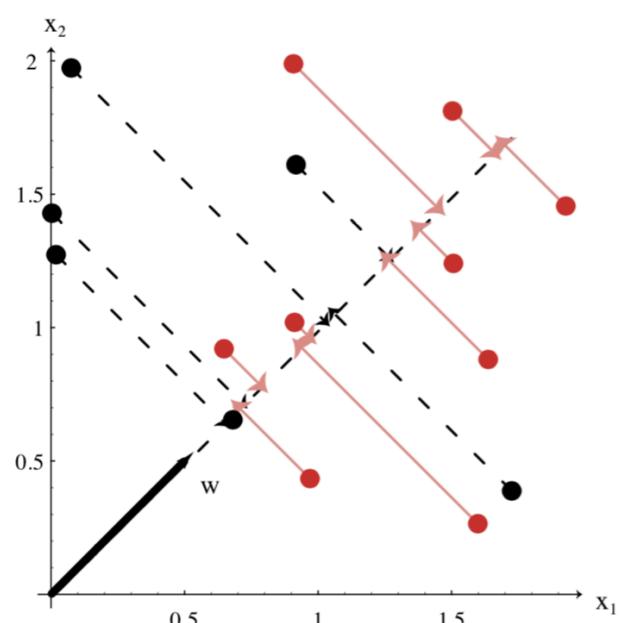
Fisher Linear Discriminant Analysis: projections along discriminative directions

$$\mu_c = \frac{1}{n_c} \sum_{x^n \in \mathcal{D}_c} x^n \quad m_c = w^T \mu_c = \frac{1}{n_c} \sum_{x^n \in \mathcal{D}_c} w^T x^n$$



Fisher Linear Discriminant Analysis: projections along discriminative directions

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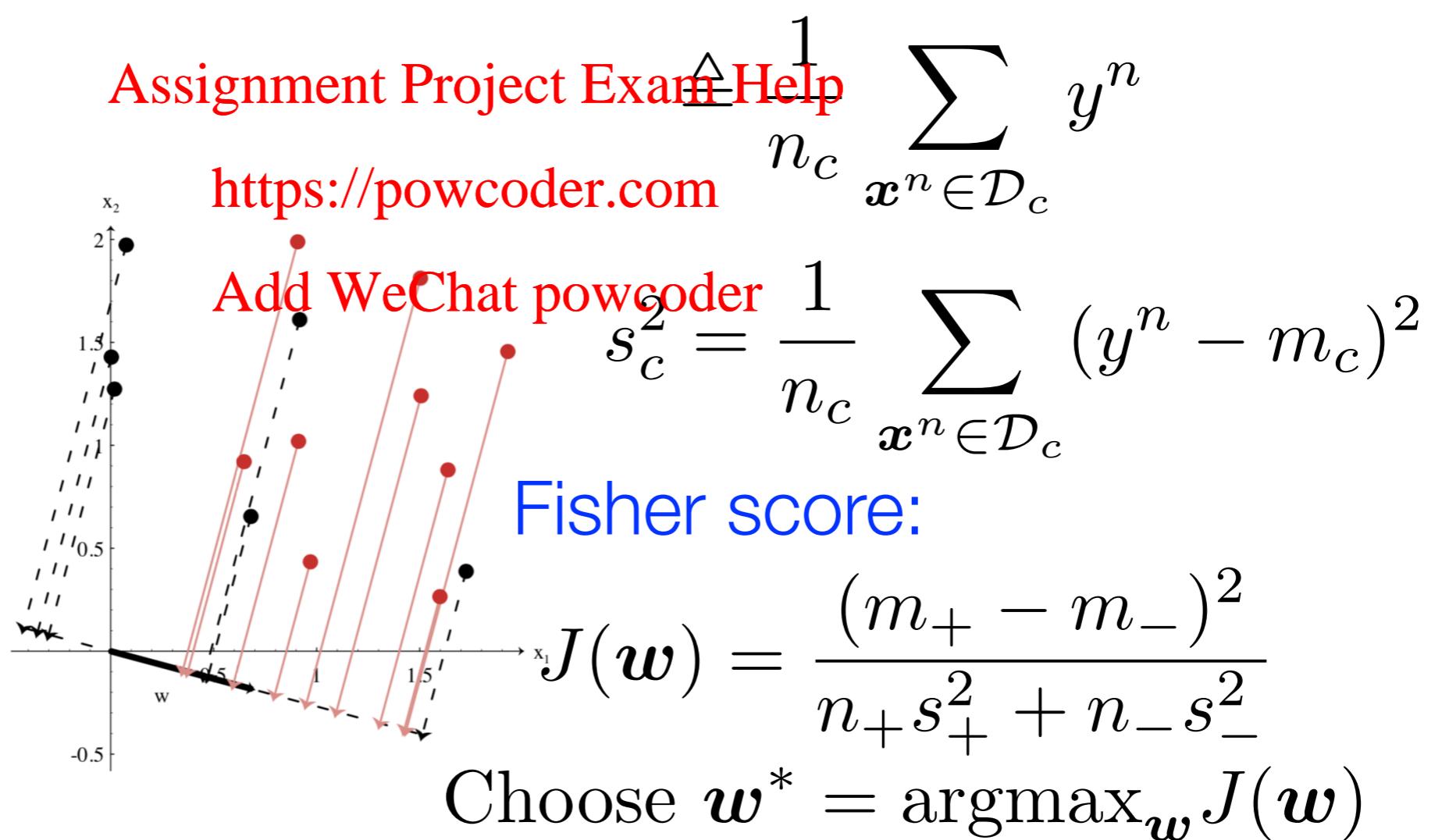
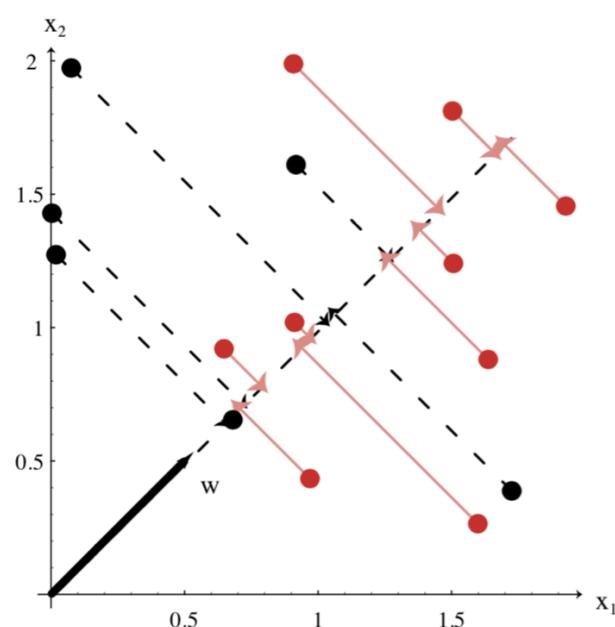


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$$s_c^2 = \frac{1}{n_c} \sum_{x^n \in \mathcal{D}_c} (y^n - m_c)^2$$

Fisher Linear Discriminant Analysis: projections along discriminative directions

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Rewrite Fisher score to make weight vector dependence explicit: between and within class scatter

$$s_c^2 = \frac{1}{n_c} \sum_{\mathbf{x}^n \in \mathcal{D}_c} (y^n - m_c)^2 = \frac{1}{n_c} \sum_{\mathbf{x}^n \in \mathcal{D}_c} (\mathbf{w}^T \mathbf{x}^n - \mathbf{w}^T \boldsymbol{\mu})^2$$

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denominator:

within class scatter

numerator:

between class scatter

$$J(\mathbf{w}) = \frac{(m_+ - m_-)^2}{n_+ s_+^2 + n_- s_-^2} \implies$$

$$J(\mathbf{w}) = \frac{\mathbf{w}^T \mathbf{S}_B \mathbf{w}}{\mathbf{w}^T \mathbf{S}_W \mathbf{w}}$$

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$$(\mathbf{a} \cdot \mathbf{b})^2 = (\mathbf{b}^\top \mathbf{a})(\mathbf{a}^\top \mathbf{b})$$

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numerator:

between class scatter

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 &= \frac{1}{n_c} \sum_{\mathbf{x}^n \in \mathcal{D}_c} \mathbf{w}^T (\mathbf{x}^n - \mu) (\mathbf{x}^n - \mu)^T \mathbf{w}
 \end{aligned}$$

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within class scatter

numerator:
between class scatter

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 &= \frac{1}{n_c} \sum_{\mathbf{x}^n \in \mathcal{D}_c} \mathbf{w}^T (\mathbf{x}^n - \mu) (\mathbf{x}^n - \mu)^T \mathbf{w} \\
 &\triangleq \mathbf{w}^T \mathbf{S} \mathbf{w}
 \end{aligned}$$

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denominator:
within class scatter

numerator:
between class scatter

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 &= \frac{1}{n_c} \sum_{\mathbf{x}^n \in \mathcal{D}_c} \mathbf{w}^T (\mathbf{x}^n - \mu) (\mathbf{x}^n - \mu)^T \mathbf{w} \\
 &\triangleq \mathbf{w}^T \mathbf{S} \mathbf{w} \quad \text{Assignment Project Exam Help}
 \end{aligned}$$

$$\begin{aligned}
 n_+ s_+^2 + n_- s_-^2 &= \mathbf{w}^T (n_+ \mathbf{S}_+ + n_- \mathbf{S}_-) \mathbf{w} \\
 &\triangleq \mathbf{w}^T \mathbf{S}_W \mathbf{w}
 \end{aligned}$$

denominator:
within class scatter

numerator:
between class scatter

$$J(\mathbf{w}) = \frac{(m_+ - m_-)^2}{n_+ s_+^2 + n_- s_-^2} \implies$$

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 \end{aligned}$$

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 n_+ s_+^2 + n_- s_-^2 &= \mathbf{w}^T (n_+ \mathbf{S}_+ + n_- \mathbf{S}_-) \mathbf{w} \\
 &\triangleq \mathbf{w}^T \mathbf{S}_W \mathbf{w}
 \end{aligned}$$

denominator:
within class scatter

$$\begin{aligned}
 (m_+ - m_-)^2 &= (\mathbf{w}^T (\mu_+ - \mu_-))^2 \\
 &= \mathbf{w}^T (\mu_+ - \mu_-) (\mu_+ - \mu_-)^T \mathbf{w}
 \end{aligned}$$

numerator:
between class scatter

$$J(\mathbf{w}) = \frac{(m_+ - m_-)^2}{n_+ s_+^2 + n_- s_-^2} \implies$$

$$J(\mathbf{w}) = \frac{\mathbf{w}^T \mathbf{S}_B \mathbf{w}}{\mathbf{w}^T \mathbf{S}_W \mathbf{w}}$$

Optimisation: find best direction, set gradients to 0

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Optimisation: find best direction, set gradients to 0

Partial derivatives (gradient) set to 0

$$\frac{\partial J(\mathbf{w})}{\partial \mathbf{w}} = \frac{\text{Assignment Project Exam Help}}{(\mathbf{w}^T \mathbf{S}_W \mathbf{w})^2} ((\mathbf{w}^T \mathbf{S}_W \mathbf{w}) \mathbf{S}_B \mathbf{w} - (\mathbf{w}^T \mathbf{S}_B \mathbf{w}) \mathbf{S}_W \mathbf{w})$$

reduces to a generalised eigenvalue problem:

$$\mathbf{S}_B \mathbf{w} = \lambda \mathbf{S}_W \mathbf{w}$$

Optimisation: find best direction, set gradients to 0

$$\left(\frac{u}{v}\right)' = \frac{1}{v^2} (u'v - uv')$$

Partial derivatives (gradient) set to 0

$$\frac{\partial J(\mathbf{w})}{\partial \mathbf{w}} = \frac{\text{Assignment Project Exam Help}}{(\mathbf{w}^T \mathbf{S}_W \mathbf{w})^2} ((\mathbf{w}^T \mathbf{S}_W \mathbf{w}) \mathbf{S}_B \mathbf{w} - (\mathbf{w}^T \mathbf{S}_B \mathbf{w}) \mathbf{S}_W \mathbf{w})$$

reduces to a ~~Add WeChat powcoder~~ generalised eigenvalue problem:

$$\mathbf{S}_B \mathbf{w} = \lambda \mathbf{S}_W \mathbf{w}$$

Optimisation: find best direction, set gradients to 0

$$\left(\frac{u}{v}\right)' = \frac{1}{v^2} (u'v - uv')$$

Partial derivatives (gradient) set to 0

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ratio

scalars

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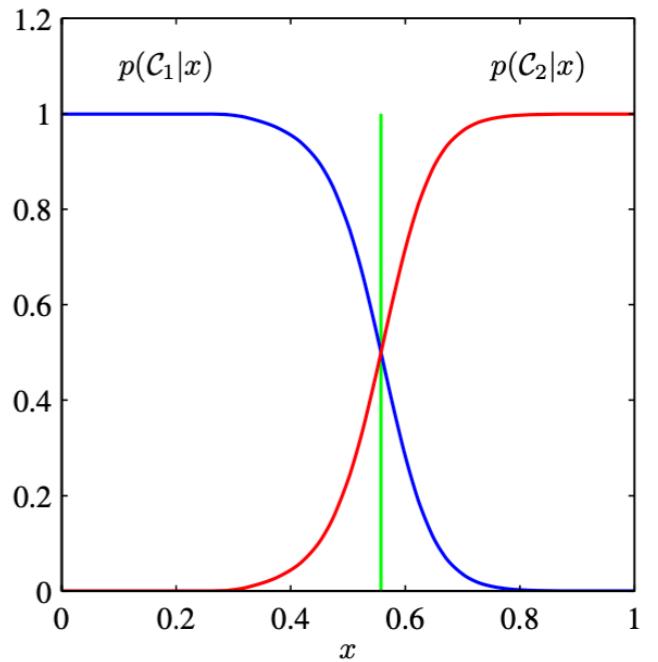
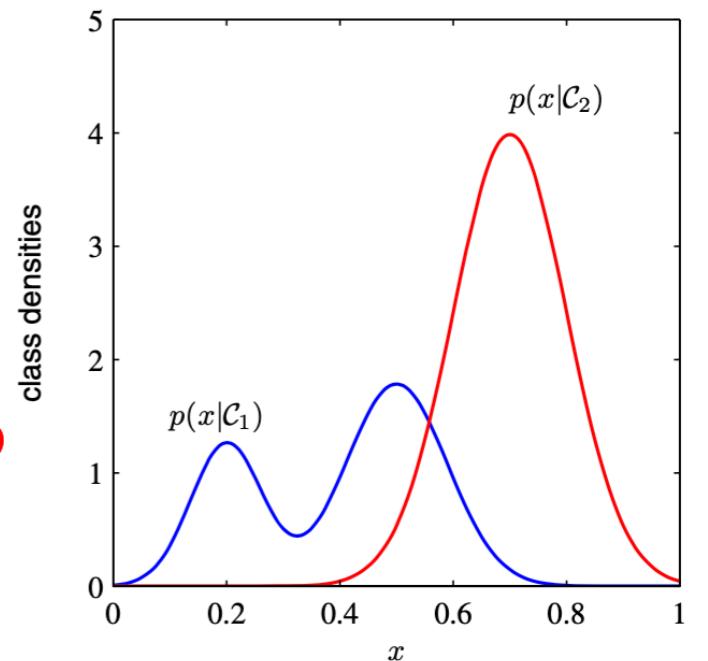
Extend to multiple classes, not just 2.

Classifiers can be generative or discriminative

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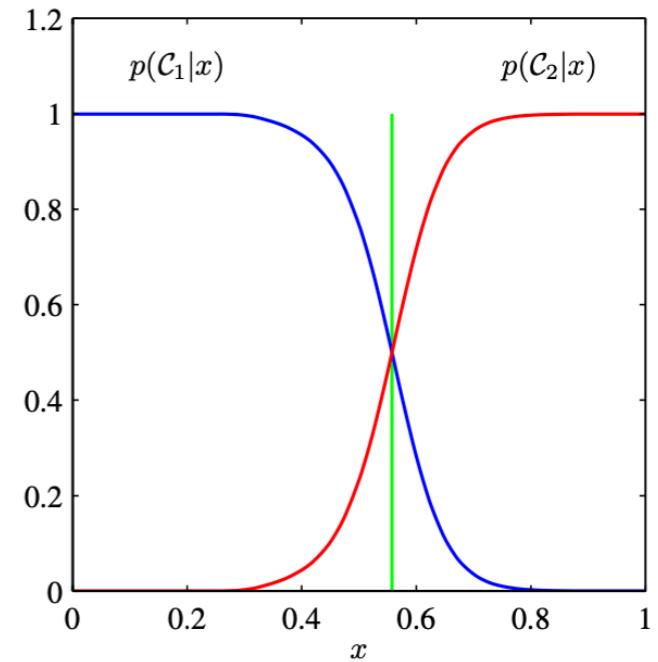
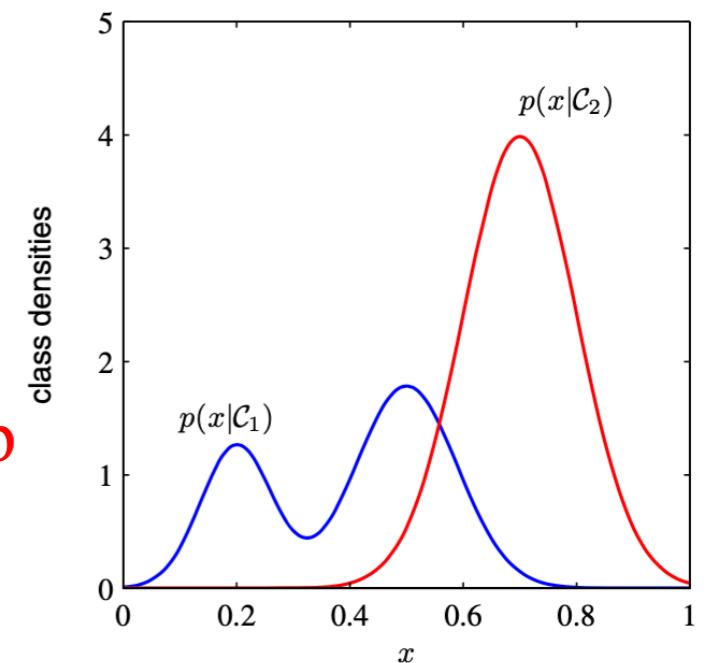
Classifiers can be generative or discriminative

- Generative – class conditional probabilities can be used to simulate (generate) data

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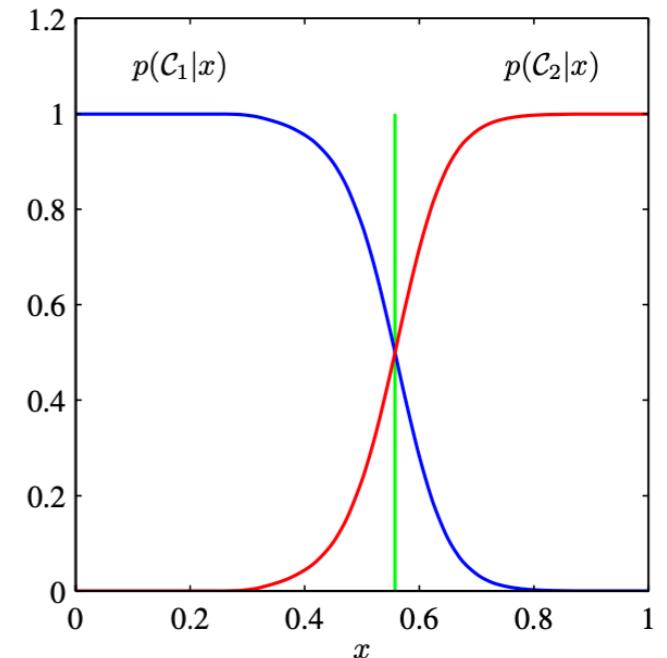
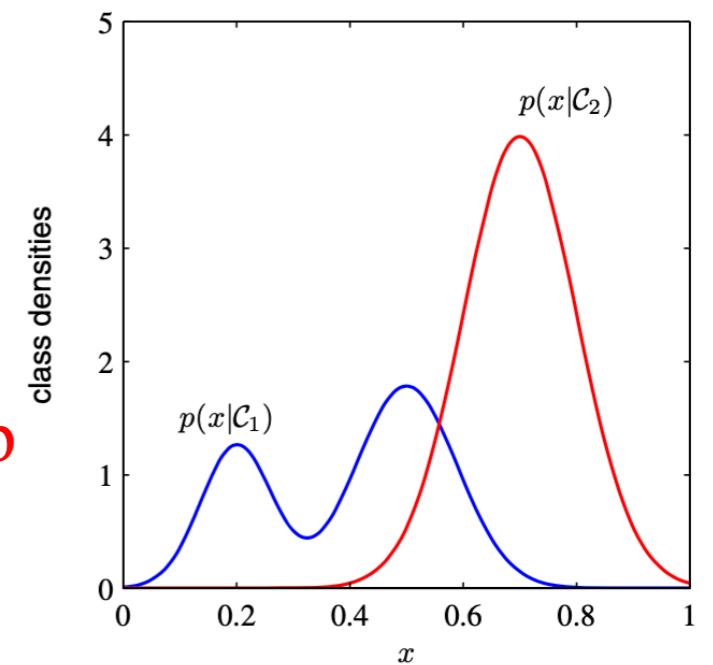
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Classifiers can be generative or discriminative

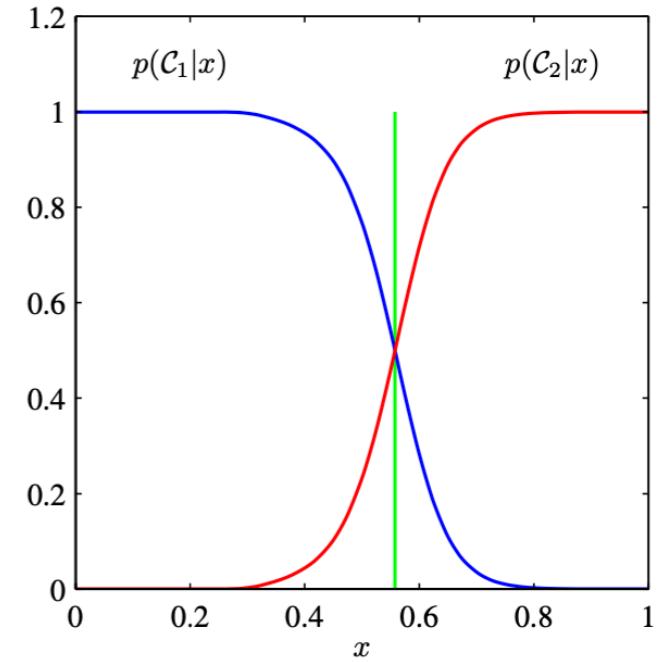
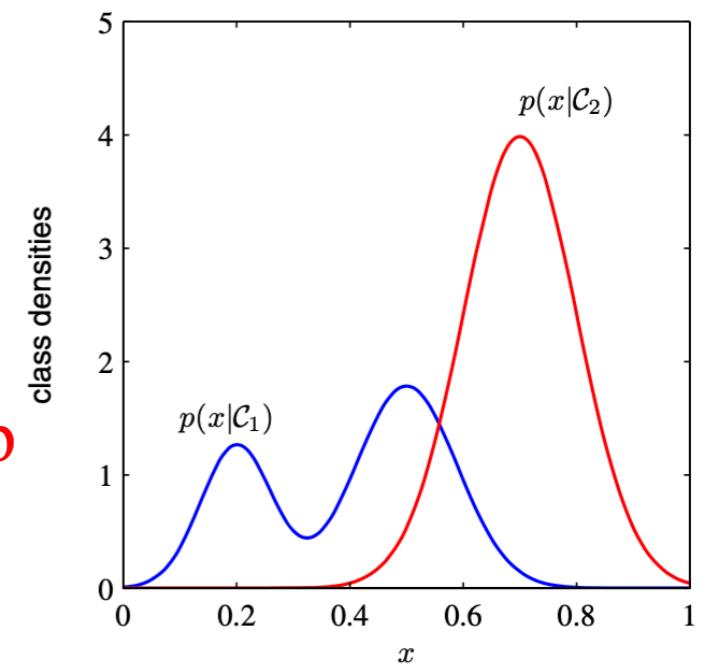
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- Generative – probabilities consistent with allocation of distance from decision boundaries

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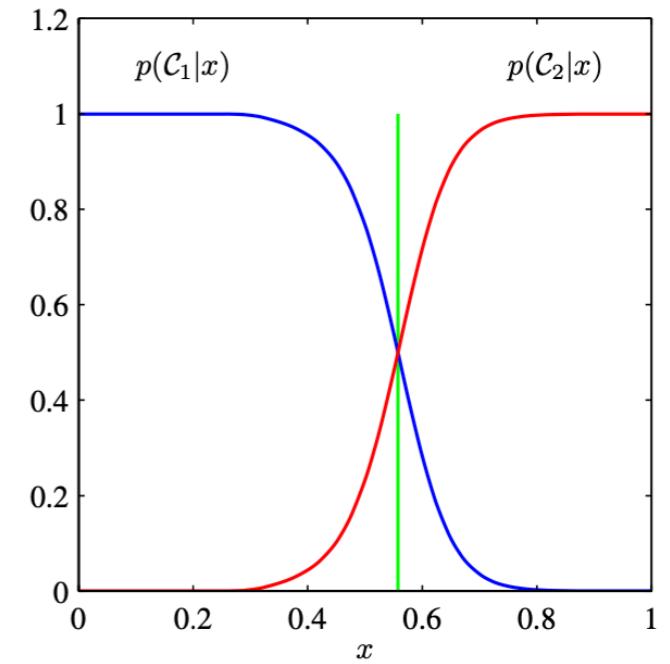
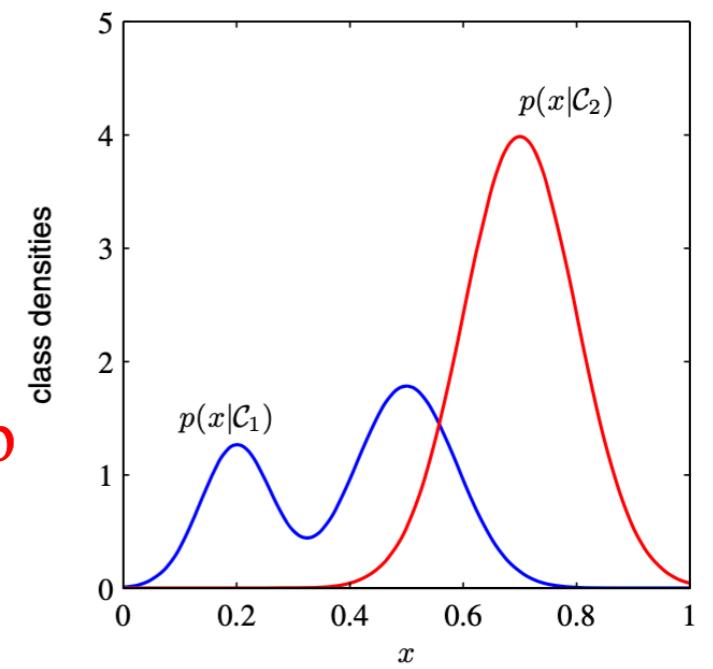
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- Discriminative – does not model data within classes, only what distinguishes a class from another
- Next time: maximum margin classifiers (Support Vector Machines)



References

- PRML, Bishop, Chapter 1, section 1.5
- PRML, Bishop, Chapter 4.
[Assignment Project Exam Help](#)
- FCML, Rogers-Girolami, Chapter 5.
<https://powcoder.com>.
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