

Assignment Project Exam Help

Logistic Regression and MaxEnt

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- Generative models:

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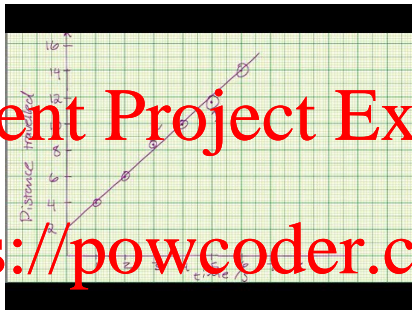
$$\Pr[y | \mathbf{x}] = \frac{\Pr[\mathbf{x} | y] \Pr[y]}{\Pr[\mathbf{x}]}$$

$$\propto \Pr[\mathbf{x} | y] \Pr[y] = \Pr[\mathbf{x}, y]$$

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- The key is to model the **generative** probability: $\Pr[\mathbf{x} | y]$.
- Example: Naive Bayes.
- Discriminative models:
 - Models $\Pr[y | \mathbf{x}]$ directly as $g(\mathbf{x}; \theta)$.
 - Example: Decision tree, Logistic Regression.
- Instance-based Learning.
 - Example: k NN classifier.

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Figure: Linear Regression

Task

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- Input: $(x^{(i)}, y^{(i)})$ pairs ($1 \leq i \leq n$)
- Preprocess: let $\mathbf{x}^{(i)} = [1 \quad x^{(i)}]^\top$
- Output: The best $\mathbf{w} = [w_0 \quad w_1]^\top$ such that $\hat{y} = \mathbf{w}^\top \mathbf{x}$ best explains the observations

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The criterion for “best”:

- Individual error: $\epsilon_i = \hat{y}^{(i)} - y^{(i)}$
- Sum of squared error: $\ell = \sum_{i=1}^n \epsilon_i^2$

Find \mathbf{w} such that ℓ is minimized.

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Taylor Series of $f(x)$ at point a

$$f(x) = \sum_{n=0}^{+\infty} \frac{f^{(n)}(a)}{n!} (x-a)^n$$

$$= f(a) + f'(a) \cdot (x-a) + \frac{f''(a)}{2} (x-a)^2 + o((x-a)^2) \quad (2)$$

- Intuitively, $f(x)$ is almost $f(a) + f'(a) \cdot (x-a)$ for all a if it is close to x .
- If $f(x)$ has local minimum x^* , then
 - $f'(x^*) = 0$, and
 - $f''(x^*) > 0$.

Minimum of the local minima is the global minimum if it is smaller than the function values at all the boundary points.

- Intuitively, $f(x)$ is almost $f(a) + \frac{f''(a)}{2} (x-a)^2$ if a is close to x^* .

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$$\begin{aligned}\frac{\partial \ell}{\partial w_j} &= \sum_{i=1}^n 2\epsilon_i \frac{\partial \epsilon_i}{\partial w_j} = \sum_{i=1}^n 2\epsilon_i \frac{\partial \mathbf{w}^\top \mathbf{x}^{(i)}}{\partial w_j} \\ &= \sum_{i=1}^n 2\epsilon_i x_j^{(i)} = 2 \sum_{i=1}^n (\hat{y}^{(i)} - y^{(i)}) x_j^{(i)}\end{aligned}$$

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By setting the above to 0, this essentially requires, for all j

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$$\sum_{i=1}^n \hat{y}^{(i)} x_j^{(i)} = \sum_{i=1}^n y^{(i)} x_j^{(i)}$$

what the model predicts

what the data says

Find the Least Square Fit for Linear Regression

In the simple 1D case, we have only two parameters in $\mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$

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Since $x_0^{(i)} = 1$, they are essentially

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$$\sum_{i=1}^n (w_0 + w_1 x_1^{(i)}) \cdot 1 = \sum_{i=1}^n y^{(i)} \cdot 1$$
$$\sum_{i=1}^n (w_0 + w_1 x_1^{(i)}) \cdot x_1^{(i)} = \sum_{i=1}^n y^{(i)} \cdot x_1^{(i)}$$

Example

Using the same example in [https://en.wikipedia.org/wiki/Linear_least_squares_\(mathematics\)](https://en.wikipedia.org/wiki/Linear_least_squares_(mathematics))

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$$\mathbf{X} = \begin{bmatrix} \text{---} & (x^{(1)})^\top & \text{---} \\ \text{---} & (x^{(2)})^\top & \text{---} \\ \text{---} & (x^{(3)})^\top & \text{---} \\ \text{---} & (x^{(4)})^\top & \text{---} \end{bmatrix} = \begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \end{bmatrix} \quad \mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix} \quad \mathbf{y} = \begin{bmatrix} 6 \\ 5 \\ 7 \\ 10 \end{bmatrix}$$

$$\begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \end{bmatrix} \begin{bmatrix} 6 \\ 5 \\ 7 \\ 10 \end{bmatrix} = \begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \end{bmatrix} \begin{bmatrix} \hat{y}_1 \\ \hat{y}_2 \\ \hat{y}_3 \\ \hat{y}_4 \end{bmatrix}$$

- Easily generalizes to more than 2-dim:

$$\mathbf{X} = \begin{bmatrix} 1 & x_1^{(1)} & \dots & x_m^{(1)} \\ 1 & \dots & \dots & \dots \\ 1 & x_1^{(i)} & \dots & x_m^{(i)} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_1^{(n)} & \dots & x_m^{(n)} \end{bmatrix} \quad \mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ \vdots \\ w_m \end{bmatrix} \quad \mathbf{y} = \begin{bmatrix} y^{(1)} \\ \vdots \\ y^{(i)} \\ \vdots \\ y^{(n)} \end{bmatrix}$$

- How to perform polynomial regression for one dimensional x ?

- $\hat{y} = w_0 + w_1 x + w_2 x^2 + \dots + w_m x^m$
 - Let $x_j^{(i)} = (x_1^{(i)})^j \implies$ Polynomial least square fitting
(<http://mathworld.wolfram.com/LeastSquaresFittingPolynomial.html>)

High-level idea:

- Observations (i.e., training data) are noisy
- $P(y^{(i)} | \hat{y}^{(i)}) = f_i(\mathbf{w})$
- Any \mathbf{w} is possible, but some \mathbf{w} is most likely.
 - Assuming independence of training examples, the likelihood of the training dataset is $\prod_i f_i(\mathbf{w})$.
 - We shall choose the \mathbf{w}^* that maximizes the likelihood.
 - Maximum likelihood estimation (MLE)
 - If we also incorporate some prior on \mathbf{w} , this becomes Maximum Posterior Estimation (MAP): If we assume some Gaussian prior on \mathbf{w} , this will add a ℓ_2 regularization term to the objective function.
- Many models and their variants can be deemed as different ways of estimating $P(y^{(i)} | \hat{y}^{(i)})$

Find \mathbf{w} such that $\|\mathbf{y} - \mathbf{X}\mathbf{w}\|_2$ is minimized.

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- What is $\mathbf{X}\mathbf{w}$ when \mathbf{X} is fixed?
 - It is the hyperplane spanned by the d column vectors of \mathbf{X} .
- \mathbf{y} in general is a vector outside the hyperplane. So the minimum distance is achieved when $\mathbf{X}\mathbf{w}^*$ is exactly the projection of \mathbf{y} on the hyperplane. This means (denote i -th column of \mathbf{X} as \mathbf{X}_i)

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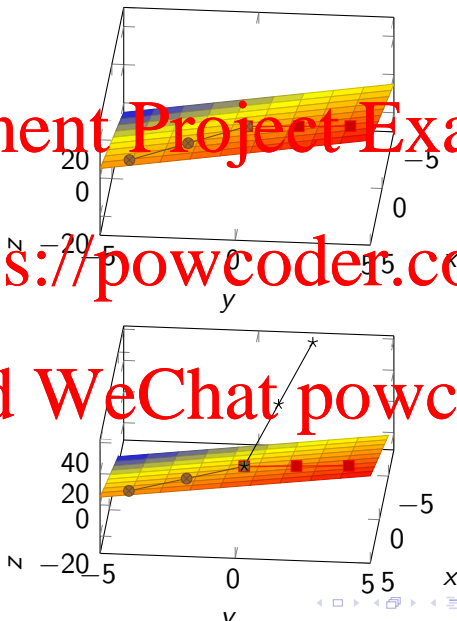
$$\left. \begin{array}{l} \mathbf{X}_1^\top (\mathbf{y} - \mathbf{X}\mathbf{w}) = 0 \\ \mathbf{X}_2^\top (\mathbf{y} - \mathbf{X}\mathbf{w}) = 0 \\ \dots\dots\dots \\ \mathbf{X}_d^\top (\mathbf{y} - \mathbf{X}\mathbf{w}) = 0 \end{array} \right\} \implies \mathbf{X}^\top (\mathbf{y} - \mathbf{X}\mathbf{w}) = 0$$

- $\mathbf{w} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} = \mathbf{X}^+ \mathbf{y}$ (\mathbf{X}^+ : pseudo inverse of \mathbf{X})

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Special case: $y^{(i)} \in \{0, 1\}$.

- Not appropriate to directly regress $y^{(i)}$.

- Rather, model $y^{(i)}$ as the observed outcome of a Bernoulli trial with an unknown parameter p_i

- How to model p_i

- We assume that p_i depends on $\mathbf{x} \triangleq \mathbf{X}_{i\cdot} \implies$ rename p_i to $p_{\mathbf{x}}$.
- Still had to estimate $p_{\mathbf{x}}$ reliably
 - MLE: $p_{\mathbf{x}} = \mathbf{E}[y = 1 \mid \mathbf{x}]$

- What can we say about $p_{\mathbf{x}+\epsilon}$ when $p_{\mathbf{x}}$ is given?

- Answer: we impose a linear relationship between $p_{\mathbf{x}}$ and \mathbf{x}

- What about a simple linear model $p_{\mathbf{x}} = \mathbf{w}^T \mathbf{x}$ for some \mathbf{w} ?

(Note: all points share the same parameter \mathbf{w})

- Problem: mismatch of the domains: vs
- Solution: mean function / inverse of link function:
 $g^{-1} : \mathbb{R} \rightarrow \text{params}$

- Solution: Link function $g(\text{parameters}) \rightarrow \mathbb{R}$

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- Equivalently, solve for p .

$$p = \frac{e^{\mathbf{w}^\top \mathbf{x}}}{1 + e^{\mathbf{w}^\top \mathbf{x}}} = \frac{1}{1 + e^{-\mathbf{w}^\top \mathbf{x}}} = \sigma(\mathbf{w}^\top \mathbf{x}) \quad (4)$$

Where $\sigma(z) = \frac{1}{1 + \exp(-z)}$.
Recall that $p_{\mathbf{x}} = \mathbf{E}[y = 1 \mid \mathbf{x}]$.

- Decision boundary is $p \geq 0.5$.
 - Equivalent to whether $\mathbf{w}^\top \mathbf{x} \geq 0$. Hence, LR is a linear classifier.

- Consider a training data point $\mathbf{x}^{(i)}$.
 - Recall that the conditional probability ($\Pr[y^{(i)} = 1 \mid \mathbf{x}^{(i)}]$) computed by the model is denoted by the shorthand notation p (which is a function of \mathbf{w} and $\mathbf{x}^{(i)}$).
 - The likelihood of $\mathbf{x}^{(i)}$ is $\begin{cases} p & , \text{ if } y^{(i)} = 1 \\ 1 - p & , \text{ otherwise} \end{cases}$, or equivalently,

- Hence, the likelihood of the whole training dataset is $p^{y^{(i)}}(1-p)^{1-y^{(i)}}$.

$$L(\mathbf{w}) = \prod_{i=1}^n p(\mathbf{x}^{(i)})^{y^{(i)}} (1 - p(\mathbf{x}^{(i)}))^{1-y^{(i)}}.$$

- Log-likelihood is (assume $\log \triangleq \ln$)

$$\ell(\mathbf{w}) = \sum_{i=1}^n y^{(i)} \log p(\mathbf{x}^{(i)}) + (1 - y^{(i)}) \log (1 - p(\mathbf{x}^{(i)})) \quad (5)$$

- To maximize ℓ , notice that it is concave. So take its partial derivatives

$$\frac{\partial \ell(\mathbf{w})}{\partial \mathbf{w}_j} = \sum_{i=1}^n \left(y^{(i)} \frac{1}{p(\mathbf{x}^{(i)})} \frac{\partial p(\mathbf{x}^{(i)})}{\partial \mathbf{w}_j} + (1 - y^{(i)}) \frac{1}{1 - p(\mathbf{x}^{(i)})} \frac{\partial (1 - p(\mathbf{x}^{(i)}))}{\partial \mathbf{w}_j} \right)$$
$$= \sum_{i=1}^n (\mathbf{x}^{(i)}_j y^{(i)} - \mathbf{x}^{(i)}_j p(\mathbf{x}^{(i)}))$$

- and set them to 0 essentially means, for all j

$$\sum_{i=1}^n \hat{y}^{(i)} \cdot \mathbf{x}^{(i)}_j = \sum_{i=1}^n p(\mathbf{x}^{(i)}) \mathbf{x}^{(i)}_j = \sum_{i=1}^n y^{(i)} \cdot \mathbf{x}^{(i)}_j$$

what the model predicts

what the data says

- Consider one dimensional x . The above condition is simplified to

$$\sum_{i=1}^n p^{(i)} x^{(i)} = \sum_{i=1}^n y^{(i)} x^{(i)}$$

- The RHS is essentially the sum of x values **only** for the training data in class $Y = 1$.

- The LHS says: if we use our learned model to assign a probability (of belonging to the class $Y = 1$) for every training data, the LHS is the expected sum of x values.

- If this is still abstract, think of an example.

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- There is no closed-form solution to maximize ℓ .
- Use the Gradient Ascent algorithm to maximize ℓ .
- There are faster algorithms.

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- \mathbf{w} is initialized to some random value (e.g., $\mathbf{0}$).
- Since the gradient gives the *steepest* direction to increase a function's value, we move a small step towards that direction, i.e.,

$$w_j \leftarrow w_j + \alpha \frac{\partial \ell(\mathbf{w})}{\partial w_j}, \text{ or}$$

$$w_j \leftarrow w_j + \alpha \sum_{i=1} (y^{(i)} - p(\mathbf{x}^{(i)})) \mathbf{x}^{(i)}_j$$

where α (learning rate) is usually a small constant or decreasing over the epochs.

- Stochastic version: using the gradient on a **randomly** selected training instance, i.e.,

$$w_j \leftarrow w_j + \alpha (y^{(i)} - p(\mathbf{x}^{(i)})) \mathbf{x}^{(i)}_j$$

- Gradient Ascent moves to the “right” direction a tiny step a time. Can we find a good step size?

- Consider 1D case **minimize** $f(x)$ and the current point is a .

- $f(x) = f(a) + f'(a)(x - a) + \frac{f''(a)}{2}(x - a)^2$ for x near a .

- To minimize $f(x)$, take $\frac{\partial f(x)}{\partial x} = 0$, i.e.,

$$\frac{\partial f(x)}{\partial x} = 0$$

$$\Leftrightarrow f'(a) \cdot 1 + \frac{f''(a)}{2} \cdot 2(x - a) \cdot 1 = f'(a) + f''(a)(x - a) = 0$$

$$\Leftrightarrow x = a - \frac{f'(a)}{f''(a)}$$

- Can be applied to multiple dimension cases too \Rightarrow need to use ∇ (gradient) and Hess (Hessian).

- Regularization is another method to deal with **overfitting**.
 - It is designed to penalize large values of the model parameters.
 - Hence it *encourages* simpler models, which are less likely to overfit.
- Instead of optimizing for $\ell(\mathbf{w})$, we optimize $\ell(\mathbf{w}) + \lambda R(\mathbf{w})$.

- λ is a hyper-parameter that controls the strength of regularization.

- It is usually determined by cross validating with a list of possible values (e.g., 0.001, 0.01, 0.1, 1, 10, ...)

- Grid search: [http:](http://scikit-learn.org/stable/modules/grid_search.html)

- [/scikit-learn.org/stable/modules/grid_search.html](http://scikit-learn.org/stable/modules/grid_search.html)

- There are alternative methods.

- $R(\mathbf{w})$ quantifies the “size” of the model parameters. Popular choices are:

- L_2 regularization (**Ridge LR**) $R(\mathbf{w}) = \|\mathbf{w}\|_2^2$
 - L_1 regularization (**Lasso LR**) $R(\mathbf{w}) = \|\mathbf{w}\|_1$
 - L_1 regularization is more likely to result in sparse models.

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- LR can be generalized to multiple classes \Rightarrow MaxEnt.

$$\Pr[c | \mathbf{x}] \propto \exp(\mathbf{w}_c^\top \mathbf{x}) \implies \Pr[c | \mathbf{x}] = \frac{\exp(\mathbf{w}_c^\top \mathbf{x})}{Z}$$

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- Z is the normalization constant. = ?
- Let \mathbf{c}^* be the last class in C , then $\mathbf{w}_{\mathbf{c}^*} = \mathbf{0}$.
- Derive LR from MaxEnt. How?
- Both belong to *exponential* or *log-linear* classifiers.

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- Andrew Ng's note:
<http://cs229.stanford.edu/notes/cs229-notes1.pdf>
- Cosma Shalizi's note: <http://www.stat.cmu.edu/~cshalizi/uADA/12/lectures/ch12.pdf>
- Tom Mitchell's book chapter: <https://www.cs.cmu.edu/~tom/mlbook/NBayesLogReg.pdf>

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