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ECON 61001: *Hypothesis Testing: Power*

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- A statistical hypothesis is a conjecture about the distribution of one or more random variables.
- The classical theory of hypothesis testing provides a framework for deciding whether a particular hypothesis is correct.
- Within this framework, there are only two possible decisions: the hypothesis is true or it is not. A decision procedure for such a problem is called a test.

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Assume: our hypothesis involves θ , the parameter vector indexing distribution of V , and Θ denote the parameter space with $\Theta \subset \mathbb{R}^p$.

Divide Θ into two mutually exclusive and exhaustive parts:

$$\Theta_0 = \{\theta : \text{such that the hypothesis is true}\},$$

$$\Theta_1 = \{\theta : \text{such that the hypothesis is false}\}$$

Using this partition, we can state the object as being to test the null hypothesis,

$H_0: \theta \in \Theta_0$
against the alternative hypothesis,

$$H_1 : \theta \in \Theta_1.$$

Base inference on some **test statistic**; denoted by S_T .

Decision rule:

$$\begin{aligned} S_T \in R_0 &\Rightarrow H_0 \text{ is accepted or rather not rejected} \\ S_T \in R_1 &\Rightarrow H_0 \text{ is rejected in favour of } H_1 \end{aligned}$$

In the companion podcast discussed how R_0 and R_1 are chosen to control the probability of a Type I error.

Now consider the properties of the test under H_1 .

Power of a test

Let $P_\theta(\cdot)$ denote the probability of the event in parentheses if the parameter vector takes the value θ .

Define $\beta(\theta) = P_\theta(R_0) = 1 - P_\theta(R_1)$; that is, $\beta(\theta)$ describes the probability of a type II error for values of θ that satisfy H_1 .

The power function of the test is:

$$\pi(\cdot) = 1 - \beta(\cdot).$$

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→ for $\theta_* \in \Theta_1$, $\pi(\theta_*)$ is:

- the probability of correctly rejecting H_0 when $\theta = \theta_*$.
- the power of the test against the alternative $\theta = \theta_*$.

Suppose that:

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- $v_t \sim N(\theta, \sigma^2)$, $t = 1, 2, \dots, T$; assume σ^2 known.

- wish to test $H_0 : \theta = 0$ versus $H_1 : \theta \neq 0$.

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- Decision rule: reject $H_0 : \theta = 0$ in favour of $H_1 : \theta \neq 0$ at the 5% significance level if $|\tau_T| > 1.96$.

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So power of the test is given by

$$\pi(\theta) = P(|\tau_T| > 1.96 \mid \theta, \theta \in \Theta_1).$$

To evaluate $\pi(\theta)$ we need the distribution of τ_T if $\theta \neq 0$.

We have

$$\tau_T = \frac{\bar{v}_T}{\sqrt{\sigma^2/T}} = \frac{\bar{v}_T - \theta}{\sqrt{\sigma^2/T}} + \frac{\theta}{\sqrt{\sigma^2/T}} \sim N(\mu, 1),$$

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where $\mu = \theta/(\sqrt{\sigma^2/T})$.

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Power is a function of μ .

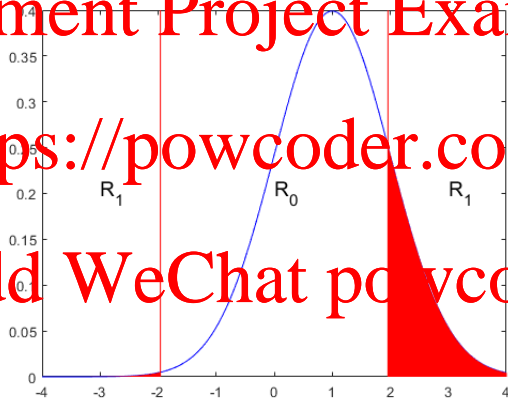
Next slide shows power for $\theta = \sigma/\sqrt{T}$, i.e. $\mu = 1$.

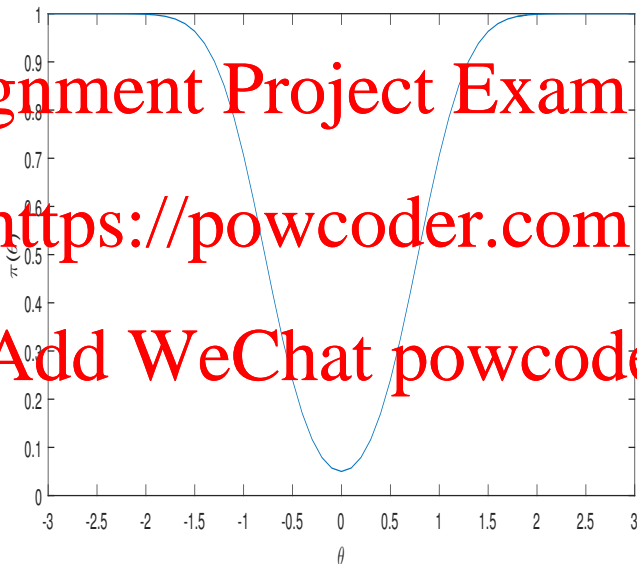
Example 2.8 in Lecture Notes: power of test when
 $\theta = \sigma/\sqrt{T}$

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From plot of power function for this two-sided test we see that:

- power depends on $|\theta|$ in this case <https://powcoder.com>
- $P_{\theta}(R_1) > \alpha$ for all $\theta \in \Theta_1 \rightarrow$ test is said to be unbiased.

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Further discussion of the topics in this podcast can be found in
Section 2.3.1 of the Lecture Notes.

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